

Simulations midterm



Before we Sim Prob Review

- we need to review some basic probability terms...
- **Random Variable** \rightarrow a random phenomenon where function X provides a real # to every outcome $P(X=x)$
- we can define random variables w/ **Cumulative Distribution Function (CDF)** ...
$$F(x) = P(X \leq x) = \begin{cases} \sum_{x \leq x} P(X=x), & \text{if discrete} \\ \int_{-\infty}^x f(x) dx, & \text{if continuous} \end{cases}$$
- we also like to know the **Expected Value** of $g(x)$...
$$E[g(x)] = \sum_{x \in X} g(x)P(X=x) \text{ if discrete } E[g(x)] = \int_{-\infty}^{\infty} g(x)f(x)dx \text{ if continuous}$$
- The **covariance** of X and Y is ...
$$\text{Cov}(X,Y) = E[(X-\mu_X)(Y-\mu_Y)] = E[XY] - E[X]E[Y] \text{ and } \dots \text{Cov}(X,X) = \text{Var}(X)$$

and if X and Y are independent \Rightarrow The correlation between X and Y ...
$$\text{Cov}(X,Y) = 0 \quad \rho(X,Y) = \text{Cov}(X,Y) / \sqrt{\text{Var}(X)\text{Var}(Y)}$$
- **Central Limit Theorem** \rightarrow w/ a sequence of iid r.v.s $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$ w/ $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$
 $\bar{Z}_n \approx N(0,1)$ as $n \rightarrow \infty$, and $s_n^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2 \rightarrow \sigma^2$ as $n \rightarrow \infty$
- We can use this to create **Confidence Intervals** which give us a range that our true μ is w/ some% confidence. Given by ...
$$\bar{X}_n \pm 1.96 \sqrt{\frac{s_n^2}{n}}$$

Random # Gen.

- Our simulation packages need random # generation to operate. But computers can't actually generate truly random outputs like say a coin toss would. The create **Pseudo Random Numbers**
- **Random Number** \rightarrow a variable from a $\text{Unif}(0,1)$
- **Random Variable** \rightarrow random outcomes of all other distributions. Can be obtained by transforming random #.
- main method computers use ... **Linear Congruential Generators** ... $a(\text{multiplier}) c(\text{increment}) m(\text{modulus}) z(\text{seed})$
generates sequence of integers between 0 and $m-1$... use that to get U_i in $[0,1]$...
$$Z_n = (aZ_{n-1} + c) \bmod m$$

$$U_n = \frac{Z_n}{m}$$
- cycle lengths of Z_n 's will be $\leq m$. Obviously not really random. We pick $m \geq 10^9$ to make look more random.
- we use LCG's cause they're fast, low storage, and reproducible, and have good statistical properties based on period.
- We measure an LCG's period, or how long (n) it takes for $Z_n = Z_0$. Largest period \equiv Full Period $\equiv m-1$
- Two types of LCG's ... **multiplicative** $c=0$ or **mixed** $c>0$
$$Z_n = (aZ_{n-1}) \bmod m$$
 or
$$Z_n = (aZ_{n-1} + c) \bmod m$$

has full period (desirable) when m is a prime # and a is a primitive root of m .
easier thm ... $a > 1$ and $(a^{m-1} - 1) \bmod m = 0$ and for all $i < m-1$, $(a^i - 1) \bmod m \neq 0$
* a multiplicative LCG w/ $m = 2^k$ has a max period of 2^{k-2} achieved if ...
 $\square Z_0$ is odd
 $\square a$ has form $8k+3$ or $8k+5$ for some integer k
- has full period if c and m are relatively prime (only pos. integer divides both is 1) and a is prime and divides m .
easier thm ...
* a mixed LCG w/ $m = 2^k$, has full period if ...
 $\square c$ is odd
 $\square a$ has form $4j+1$ for some integer j

Test Rand. #'s

- for our random #'s we can look at empirical tests to test their actual randomness.
- **Chi-squared Test** \rightarrow look at histogram of observations vs fitted density function. Look at how many obsv. out of n are in each bin. null hypothesis H_0 : X_i 's are iid w/ dist. given. F . k bins. N_j = # of obsv. in bin p_j = prob of outcome according to F . test stat. $\chi^2 = \sum_{j=1}^k (N_j - np_j)^2 / np_j$, if H_0 true, dist should match χ^2 . for Uniform $[0,1]$ $\chi^2 = \sum_{j=1}^k (N_j - \frac{n}{k})^2$
- **Serial Test** \rightarrow a 1 dimensional χ^2 test to test uniformity and indep. between successive obsv. H_0 : U_i 's are iid $U(0,1)$
- for both χ^2 and Serial test the prob. of rejecting a false H_0 is low.
- **Autocorrelation Test** \rightarrow test autocorrelation of a sequence at various lag j 's. We can't get the true autocor. we need to estimate ...
$$\hat{\rho}_j = \frac{12}{n+1} \sum_{k=0}^n (U_{1+kj} U_{1+(k+1)j}) - 3 \quad \hat{\gamma}_j = \frac{13k+7}{(n+1)^2} \quad h = \lfloor \frac{n-1}{j} \rfloor - 1$$

w/ H_0 : $\rho_j = 0$
for large j test stat ... $Z_j = \hat{\rho}_j / \sqrt{\hat{\gamma}_j} \sim N(0,1)$

Sims

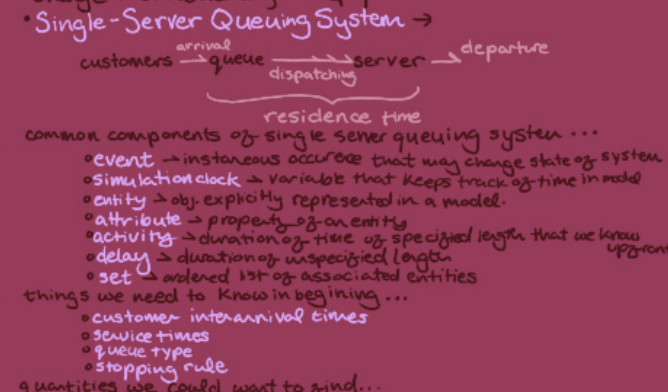
- We categorize simulations in 3 ways. Static vs Dynamic, Deterministic vs Stochastic, and Continuous vs Discrete

Static Sim

- A static simulation is one where time plays no role. It relies on random # or random variat generation for things to happen.
- When solving problems we need to remember we're not getting optimal solutions. We're getting output.
- examples \rightarrow coin flip, news vendor problem, proj. management

Discrete Event Sim

- DES \rightarrow when a system evolves over time where state var. change instantaneously at sep. points in time.
- **Single-Server Queuing System** \rightarrow



- quantities we could want to find ...
- **Expected Avg. Delay in Queue** \rightarrow wait time for cust.
$$d(n) = E\left[\frac{\sum_{i=1}^n D_i}{n}\right] \rightarrow D_i$$
's delays for each cust.
$$\hat{d}(n) = \frac{\sum_{i=1}^n d_i}{n} \rightarrow$$
 estimator

- **Expected Avg. # customers in Queue** $\rightarrow Q(t)$ = # in queue at time t . $T(n)$ = time n th customer exits queue
$$\hat{q}(n) = \frac{\int_0^{T(n)} Q(t) dt}{T(n) - 0} \rightarrow$$
 estimator

- **Expected Server Utilization** \rightarrow
$$B(t) = \begin{cases} 1 & \text{if busy at } t \\ 0 & \text{if not} \end{cases}$$

$$u(n) = \frac{\sum_{i=1}^n B(t_i)}{T(n) - 0} \rightarrow$$
 estimator

- If we are running simulations by hand, there are 2 strategies to do so: **Process-Interaction Approach** & **Event Scheduling Approach**.

Output of Sims

- What kind of output analysis we do depends on the simulation ...
- we want to use CI to account for sampling error, but simulations will never produce iid Normal output, so we can't apply classical statistical techniques.
- So what can we do with output ...

Terminating Sim

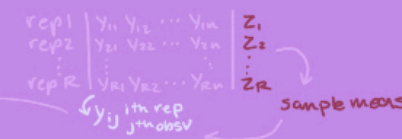
- means we know when the system will end and we want to observe the system at an end.
- we could be working w/ discrete Y_n or continuous $Y(t)$ data and our goal is to find unbiased estimators \bar{Y}_n for some property (ie avg wait time).
- we'd also like an estimator for $\text{Var}(\bar{Y}_n)$, but because our Y_i 's are correlated w get extremely biased output. So we use ...

Method of Independent Replications

- \rightarrow conduct R indep. replications, each w/ n observations, to ensure indep.

$$\bar{Z}_i = \frac{1}{n} \sum_{j=1}^n Y_{i,j}, i=1 \dots R$$

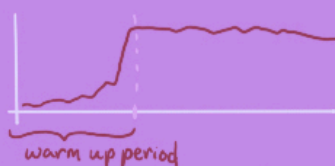
- \rightarrow each initialized under same conditions



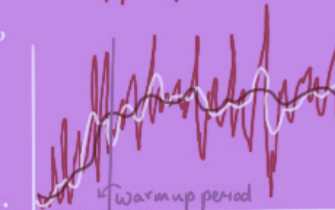
- The Z_i 's from the independent replications are iid and can use them to find a grand sample mean, sample variance ($\hat{V} = \frac{1}{R-1} \sum_{i=1}^R (Z_i - \bar{Z}_R)^2$) ($\text{Var}(\bar{Z}_R) = \frac{\hat{V}}{R}$), and a confidence interval $\bar{Z}_R \pm 1.96 \sqrt{\hat{V}/R}$

Steady State Sim

- we want to know about a system during its run, while it's in its normal state. we want its steady state mean μ .
- a problem \rightarrow **initialization bias** \rightarrow we don't care about the time it took the system to warm up to steady state.
- easiest way to deal is to truncate the warm up period. so ...
$$\bar{X}_{n,\ell} = \frac{1}{n-\ell} \sum_{i=\ell+1}^n X_i$$
- but how do we choose truncate point ℓ ...



- **Welch's Method** \rightarrow we start w/ the avg. of k rep's \bar{X}_j 's. Then for chosen window w , moving averages
$$\bar{X}_j(w) = \begin{cases} \frac{1}{w} \sum_{i=j-w+1}^j \bar{X}_i, & w+1 \leq j \leq n-w \\ \frac{1}{j} \sum_{i=1}^j \bar{X}_i, & 1 \leq j \leq w \end{cases}$$



- this finds the average of a window sized w , rather than an avg for each \bar{X}_j .
- Find averages for many w 's, plot, and visually find warm up period.

- we prefer to err on the side of the truncation point being too large, even if it makes slightly bigger CI's ...
- we make CI's of steady state μ in 2 ways ...
- **Multiple Replications Method** \rightarrow determine truncation ℓ , run k indep. reps of length n , cutting off each. Can find CI for μ ... $\bar{Y} \pm 1.96 \sqrt{\hat{V}/k}$
 \rightarrow note ... $\ell \uparrow$ for fixed $n \Rightarrow$ systematic + sampling error \downarrow
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grand mean of estimates \downarrow var estimator \downarrow

- **Batch Means Method** \rightarrow take a very long sim and cut off b once at the beginning, then split into b nonoverlapping batches of size m . ($n = bm$)
we find a CI for μ ... $\bar{Y} \pm 1.96 \sqrt{\hat{V}_b/b}$
grand mean of estimates \downarrow var estimator \downarrow # batches
- but how to choose m ? rec. $10 \leq b \leq 30$