Quantifying nitrogen oxides and ammonia via frequency modulation in gas sensors

- DRAFT

Kvantifiering av kväveoxider och ammoniak via frekvensmodulering i gassensorer

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Abstract

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Acknowledgments

Thank you for reading my draft! :)

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3.1 Data acquisition details	
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List of acronyms and abbreviations

AC Alternating Current. 2 ${\bf GBCO}$ Gate Bias Cycled Operation. 2 Hz Hertz. 11 mA miliamperes. 10 NIPALS Nonlinear Iterative Partial Least Squares. 6, 7, 8 **OLS** Ordinary Least Squares. 4 PC Principal Component. 5, 6, 7 PCA Principal Components Analysis. 5, 7 PCR Principal Components Regression. 5, 7, 8 PLS Partial Least Squares. 8 PLSR Partial Least Squares Regression. 2, 7, 8 **ppm** parts per million. 10 \mathbf{RSS} Residual Sum of Squares. 5 SAS Sensor and Actuator Systems. 10 SCR Selective Catalytic Reduction. 1, 2 SiC-FET Silicon Carbide Field Effect Transistor. 2, 10

 ${f TCO}$ Temperature Cycled Operation. 2



Introduction

1.1 Motivation

Nitric Oxide (NO) and Nitrogen Dioxide (NO₂), commonly referred together as NO_x , are hazardous gases to the environment and to humans. Its main sources are combustion processes in transportation, and industrial processes such as (but not limited to) auto mobiles, trucks, boats, industrial boilers, turbines, etc. (USEPA 2019).

 $\mathrm{NO_x}$ exposure to humans can cause respiratory illnesses such bronchitis, emphysema and can worsen heart disease (Boningari and Smirniotis 2016). Environmentally, $\mathrm{NO_x}$ are deemed precursors of adverse phenomena such as smog, acid rain, and the depletion of ozone (O₃) (Alberto Bernabeo, Webster, and Onofri n.d.). It is of high interest, therefore, to reduce $\mathrm{NO_x}$ emissions.

One well studied and successful method of reducing emissions is Selective Catalytic Reduction (SCR), which consists in the reduction of NO_x by ammonia (NH₃) into nitrogen gas (N₂) and water (H₂O) (Forzatti 2001), both harmless components. The process is based in the following reactions (Forzatti 2001):

•
$$4 \text{ NH}_3 + 4 \text{ NO} + \text{O}_2 \longrightarrow 4 \text{ N}_2 + 6 \text{ H}_2 \text{O}$$

•
$$2 \text{ NH}_3 + \text{NO} + \text{NO}_2 \longrightarrow 2 \text{ N}_2 + 3 \text{ H}_2 \text{O}$$

$$\bullet \ 8\,\mathrm{NH_3} + 6\,\mathrm{NO_2} \longrightarrow 7\,\mathrm{N_2} + 12\,\mathrm{H_2O}$$

One key element in these reactions, however, is the amount of ammonia dosed into the SCR systems. Ammonia itself is hazardous to humans, causing skin and respiratory irritation, among other illnesses (ASTDR 2004). More importantly, ammonia is one of the main sources of nitrogen pollution and it has direct negative impact on biodiversity via nitrogen deposition in soil and water (Guthrie, Giles, Dunkerley, Tabaqchali, Harshfield, Ioppolo, and Manville 2018).

Hence it is also desired to keep ammonia emissions to a minimum. Too much ammonia in the SCR catalyst will guarantee NO_x reduction at the expense of undesired ammonia emissions. Concurrently, too little ammonia will impede SCR to occur properly, beating the purpose of the catalyst and as a consequence, undesired NO_x emissions.

To monitor gasses concentrations, chemical sensors are deployed, one of which is the Silicon Carbide Field Effect Transistor (SiC-FET). The identification and quantification of gasses is normally achieved through multiple sensor in so called sensor arrays. Ideally each sensor in the array needs to have different responses to different compounds (Bastuck 2019). The deployment of multiple sensors, on the other hand, proves itself cumbersome due to the increased chances of failure, and decalibration of the system should one or multiple sensors be replaced (Bastuck 2019).

One solution to this problem is the cycled operation of one single sensor, referred as virtual multi-sensor (Bastuck 2019). By cycling the working point parameters of the sensor, different substances react differently in the sensor surface, which in turn produces different responses. Temperature Cycled Operation (TCO), Gate Bias Cycled Operation (GBCO), and the combination of the two have been proven to increase selectivity of SiC-FET sensors (Bastuck 2019).

TCO, in contrast with a constant temperature evaluation, produces unique transient sensor responses, i.e. each gas mixture yields a slightly different sensor output. This unique gas signature increases selectivity (Bur, Bastuck, Lloyd Spetz, Andersson, and Schütze 2014). Additionally, the high temperatures reached in these cycles help in the cleansing of the sensor surface, preparing it for the new mixtures to come.

Frequency modulation tries to achieve the same goal: avoid steady state responses in exchange of unique signatures that could help identify/quantify the gasses at hand. It consists on operating the sensor in Alternating Current (AC). One then can regulate the frequency of this operation and create cycles of different frequencies, similar to what is done in TCO. This is equivalent to GBCO, but with more frequency changes and achieving overall higher frequencies.

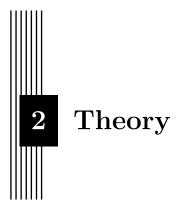
The main question is: given these set of unique sensor responses, how one can quantify the gasses that produced them? The answer lies in multivariate regression techniques. Partial Least Squares Regression (PLSR) has been used in chemometrics extensively and it has been proven to be good at this task (Bastuck 2019) (Wold, Sjöström, and Eriksson 2001). Other multivariate regression methods, naturally, can also be used. This is the aim of this thesis work, which is shown in the following section.

1.2 Aim

The aim of this thesis is to investigate different regression methods, namely: PLSR, Ridge Regression and (neural nets XXXX - TENTATIVE), and their fit to correctly quantify gas mixtures such NO_x and Ammonia subjected to sensor frequency modulation.

1.3 Research questions

- 1. Is it possible to achieve acceptable prediction levels for NO_{x} and Ammonia using frequency modulation?
- 2. Which method yields best predictions of gas concentrations?



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The quantification of gases based on the sensor response can be viewed as a multivariate multiple regression problem where the predictors, i.e. features derived from the sensor signal, are used to predict multiple responses, i.e. the concentrations of pertinent gases. This chapter briefly exposes the theory behind some of these models. Their implementation, on the other hand, will be discussed in Chapter 4 - Methods.

2.1 Ordinary Least Squares Regression

A simple, first approach would be to tackle the problem with a Ordinary Least Squares (OLS) regression model. As Friedman, Hastie, Tibshirani, et al. 2001 explains, each output $\mathbf{Y} = [Y_1, Y_2, ..., Y_K]^{\mathsf{T}}$ has its own linear model. Now, given a set of n observations $\mathbf{X} = [X_1, X_2, ..., X_n]^{\mathsf{T}}$ and each observation having p+1 features, e.g. $X_i = [1, x_{i1}, x_{i2}, ...x_{ip}]$, i = 1, 2, ...n, the concatenation of all linear models can be written in matrix form as in Equation 2.1.

$$Y = XB + E \tag{2.1}$$

Where:

- **B**: $(p+1\times K)$ matrix of regression coefficients (with the +1 referring to the intercept term);
- **E**: $(N \times K)$ matrix of residuals.

The objective is then to find the coefficients $\hat{\mathbf{B}}$ which minimizes the Residual Sum of Squares (RSS), which is summarized by Equation 2.3 (Friedman, Hastie, Tibshirani, et al. 2001):

$$\hat{\mathbf{B}} = \underset{\mathbf{B}}{\operatorname{arg\,min}} \operatorname{RSS}(\mathbf{B}) \tag{2.2}$$

In turn, the RSS, as the name suggests, is defined as the difference between real and predicted values, squared, which in matrix form is written as (Friedman, Hastie, Tibshirani, et al. 2001):

$$RSS(\mathbf{B}) = Tr[(\mathbf{Y} - \mathbf{X}\mathbf{B})^{\mathsf{T}}(\mathbf{Y} - \mathbf{X}\mathbf{B})]$$
(2.3)

Finally, solving for $\hat{\mathbf{B}}$ yields (Friedman, Hastie, Tibshirani, et al. 2001):

$$\hat{\mathbf{B}} = (\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1}\mathbf{X}^{\mathsf{T}}\mathbf{Y} \tag{2.4}$$

For the problem at hand, in addition to the high number of features, it is often the case that sensor data points are acquired in quick succession, which in turn leads to highly correlated features (Bastuck 2019), which can result in high variance in a least squares model (Friedman, Hastie, Tibshirani, et al. 2001). It is natural, therefore, to progress towards methods that incorporate dimensionality reduction such as Principal Components Regression (PCR).

2.2 Principal Component Analysis

One way to define Principal Components Analysis (PCA) is to view it as a orthogonal projection of the data into a principal space of lower dimension such that the variance of this projection is maximized (Bishop 2006).

Just as before, consider the collection of n observations is $\mathbf{X} = [X_1, X_2, ..., X_n]^{\mathsf{T}}$ with covariance matrix $\mathbf{\Sigma}$. Additionally, consider a matrix $\mathbf{P} = [\mathbf{p_1}, \mathbf{p_2}, ..., \mathbf{p_n}]^{\mathsf{T}}$ where $\mathbf{p_i}$ is a row vector of coefficients referring to the i-th linear combination (Johnson and Wichern 2013):

$$T_i = \mathbf{X} \mathbf{p}_i^{\mathsf{T}} \qquad i = 1, 2, ..., n \tag{2.5}$$

The variance and covariance of these new variables T_i can be written as follows:

$$Var(T_i) = \mathbf{p}_i^{\mathsf{T}} \mathbf{\Sigma} \mathbf{p}_i \qquad i = 1, 2, ..., n$$
 (2.6)

$$Cov(T_i, T_k) = \mathbf{p}_i^{\mathsf{T}} \mathbf{\Sigma} \mathbf{p}_k \qquad i, k = 1, 2, ..., n$$
(2.7)

The first Principal Component (PC) is then the linear combination with maximum variance, i.e. the linear combinations that maximizes $Var(T_1)$, with the constraint that the coefficient

vector $\mathbf{p_1}$ has unit length. In summary, the first PC is computed as (Johnson and Wichern 2013):

$$T_1 = \mathbf{X} \mathbf{p}_1^{\mathsf{T}}$$

that maximizes $Var(\mathbf{X} \mathbf{p}_1^{\mathsf{T}})$ (2.8)
subject to $\mathbf{p}_1^{\mathsf{T}} \mathbf{p}_1 = 1$

The second PC, similarly to the first, is the linear combination with maximum variance, but with an added extra constraint: this new linear combination must be orthogonal to the previous one, i.e. they must be linearly independent:

$$T_2 = \mathbf{X}\mathbf{p}_2^{\mathsf{T}}$$

that maximizes $Var(\mathbf{X}\mathbf{p}_2^{\mathsf{T}})$
subject to $\mathbf{p}_2^{\mathsf{T}}\mathbf{p}_2 = 1$
and $Cov(T_1, T_2) = 0$ (2.9)

The k-th PC is then:

$$T_k = \mathbf{X} \mathbf{p}_k^{\mathsf{T}}$$

that maximizes $\operatorname{Var}(\mathbf{X} \mathbf{p}_k^{\mathsf{T}})$
subject to $\mathbf{p}_k^{\mathsf{T}} \mathbf{p}_k = 1$
and $\operatorname{Cov}(T_j, T_k) = 0$ for $k > j$

It can be shown that these desired linear combinations can be written in terms of the eigenvalues (λ) and eigenvectors (\mathbf{e}) of Σ , the covariance matrix of \mathbf{X} (Johnson and Wichern 2013). Namely, for the k-th PC:

$$T_k = \mathbf{X} \mathbf{e}_{\mathbf{k}}^{\mathsf{T}}$$

$$\operatorname{Var}(T_k) = \mathbf{e}_{\mathbf{k}}^{\mathsf{T}} \mathbf{\Sigma} \mathbf{e}_{\mathbf{k}} = \lambda_k$$

$$\operatorname{Cov}(T_j, T_k) = \mathbf{e}_{\mathbf{k}}^{\mathsf{T}} \mathbf{\Sigma} \mathbf{e}_{\mathbf{j}} = 0 \quad \text{for } k \neq j$$
(2.11)

There are several ways of computing PCs. Many of which involving finding aforementioned eigenvalues and eigenvectors. These calculations can be computationally expensive, depending on the desired number of extracted PCs (Bishop 2006). One option is the Nonlinear Iterative Partial Least Squares (NIPALS) algorithm, also called Power Method. It has two clear advantages: "it can handle missing data and computes the components sequentially" (Dunn 2021).

$$T = XP \tag{2.12}$$

2.3 Principal Component Regression

The objective of PCA is find a matrix \mathbf{P} such that the linear transformation

$$T = XP \tag{2.13}$$

yields new variables $(t_1, t_2, ..., t_m)$ that are uncorrelated and arranged in decreasing order of variance. **T** is named scores and **P** Principal Component (PC) of **X** (Ng 2013). Since the matrix **T** is ordered, it follows that most of the variance on the data **X** is captured by the first k-th PC (Ng 2013). This approximation of **X** is defined in Equation 2.14:

$$\mathbf{T}_{|\mathbf{k}} = \mathbf{X} \mathbf{P}_{|\mathbf{k}} \tag{2.14}$$

Finally, PCR is simply performing linear regression on $\mathbf{T}_{|\mathbf{k}}$ instead of \mathbf{X} :

$$\mathbf{y} = \mathbf{T}_{|\mathbf{k}}\beta + \epsilon \tag{2.15}$$

And the regression coefficients just as in linear regression:

$$\hat{\beta}^{PCR} = (\mathbf{T}_{|\mathbf{k}}^{\mathbf{T}} \mathbf{T}_{|\mathbf{k}})^{-1} \mathbf{T}_{|\mathbf{k}}^{\mathbf{T}} \mathbf{y}$$
 (2.16)

The Nonlinear Iterative Partial Least Squares (NIPALS) algorithm can also compute Principal Component and its scores (Ng 2013) (Wright n.d.).

Algorithm 1: Nonlinear Iterative Partial Least Squares (NIPALS)

```
Result: First k Principal Components i = 1;
\mathbf{X_i} = \mathbf{X};
while i < k do

repeat

Choose \mathbf{t_i} as any column of \mathbf{X_i};
Compute loadings \mathbf{p_i} = \frac{\mathbf{X_i^T t_i}}{\mathbf{t_i^T t_i}};
Let \mathbf{p_i} = \frac{\mathbf{p_i}}{\sqrt{\mathbf{p_i^T p_i}}};
Compute scores \mathbf{t_i} = \frac{\mathbf{X_i p_i}}{\mathbf{p_i^T p_i}};
until Until\ \mathbf{t_i}\ converges;
\mathbf{X_{i+1}} = \mathbf{X_i} - \mathbf{t_i p_i^T};
i += 1;
```

2.4 Partial Least Squares Regression

PLSR, much like PCR, also tries to reduce dimensionality via linear combinations of the inputs. In this technique, however, also takes into account the dependent variables y. One key

advantage of PLSR is that it seeks axes with most variance (like PCR) and high correlation with the dependent variables (Friedman, Hastie, Tibshirani, et al. 2001).

The main idea can be described as decomposing both the design matrix \mathbf{X} and response matrix \mathbf{Y} as follows (Ng 2013), similarly to what was done in Section 2.3:

$$\mathbf{X} = \mathbf{T}\mathbf{P}^{\mathbf{T}} \tag{2.17}$$

$$Y = UQ^{T}$$
 (2.18)

Instead of simply running NIPALS on \mathbf{X} and \mathbf{Y} separately. PLSR uses information from \mathbf{Y} to decompose \mathbf{X} and vice-versa (Ng 2013).

Algorithm 2: Partial Least Squares (PLS)

```
\begin{split} \textbf{Result:} & \text{First k Partial Least Squares directions} \\ \textbf{i} &= 1; \\ \textbf{X}_{\textbf{i}} &= \textbf{X}; \\ \textbf{Y}_{\textbf{i}} &= \textbf{Y}; \\ \textbf{while } i < k \text{ do} \\ & | \textbf{repeat} \\ & | \textbf{Compute loading of } \textbf{X}_{\textbf{i}} \text{ based on score of } \textbf{Y}_{\textbf{i}} \text{: } \textbf{p}_{\textbf{i}} = \frac{\textbf{X}_{\textbf{i}}^{\textbf{T}} \textbf{u}_{\textbf{i}}}{\|\textbf{X}_{\textbf{i}}^{\textbf{T}} \textbf{u}_{\textbf{i}}\|} \; ; \\ \textbf{Compute score of } \textbf{X}_{\textbf{i}} \text{: } \textbf{t}_{\textbf{i}} &= \textbf{X} \textbf{p}_{\textbf{i}}; \\ \textbf{Compute loading of } \textbf{y}_{\textbf{i}} \text{ based on score of } \textbf{X}_{\textbf{i}} \text{: } \textbf{q}_{\textbf{i}} &= \frac{\textbf{Y}_{\textbf{i}}^{\textbf{T}} \textbf{u}_{\textbf{i}}}{\|\textbf{Y}_{\textbf{i}}^{\textbf{T}} \textbf{t}_{\textbf{i}}\|} \; ; \\ \textbf{Compute score of } \textbf{Y}_{\textbf{i}} \text{: } \textbf{u}_{\textbf{i}} &= \textbf{Y} \textbf{q}_{\textbf{i}}; \\ \textbf{until } \textit{Until } \textbf{t}_{\textbf{i}} \; \textit{converges}; \\ \textbf{X}_{\textbf{i+1}} &= \textbf{X}_{\textbf{i}} - \textbf{t}_{\textbf{i}} \textbf{p}_{\textbf{i}}^{\textbf{T}} \; ; \\ \textbf{Y}_{\textbf{i+1}} &= \textbf{Y}_{\textbf{i}} - \textbf{u}_{\textbf{i}} \textbf{q}_{\textbf{i}}^{\textbf{T}} \; ; \\ \textbf{i} &= 1; \\ \end{split}
```

After finding the k partial least squares directions from the Algorithm 2 above, the score matrices **T** and **U** are found. The regression coefficients β^{PLS} are found by the relation (Ng 2013):

$$\mathbf{U} = \mathbf{T}\beta^{\text{PLS}} \tag{2.19}$$

Finally, by the substitution of Equation 2.19 on 2.18 (Ng 2013):

$$\mathbf{Y} = \mathbf{X} \mathbf{P} \beta^{\text{PLS}} \mathbf{Q}^{\mathbf{T}} \tag{2.20}$$

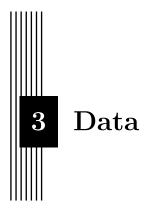
2.5 Ridge Regression

Another option is shrink regression coefficients via a penalty term. As stated on (Friedman, Hastie, Tibshirani, et al. 2001), "ridge coefficients minimize a penalized sum of squares", as shown on Equation 2.21 and Equation 2.22.

$$\hat{\beta}^{\text{ridge}} = \operatorname{argmin}_{\beta} \left\{ \sum_{i=1}^{N} (y_i - \beta_0 - \sum_{j=1}^{p} x_{ij} \beta_j)^2 + \lambda \sum_{j=1}^{p} \beta_j^2 \right\}$$
 (2.21)

Where $\lambda \geq 0$ is a parameter that controls strength of the penalization. This could also be written in matrix form:

$$\hat{\beta}^{\text{ridge}} = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y}$$
(2.22)



3.1 Data acquisition

The data was acquired at the Sensor and Actuator Systems (SAS) laboratory at Linköping University. The experiment — as shown on Figure 3.1 — consisted of exposing different gas combinations to the SiC-FET sensor under a certain frequency cycle and recording its response, measured in miliamperes (mA). The is then used to extract secondary features, namely average and slope values from certain regions of the frequency cycle.

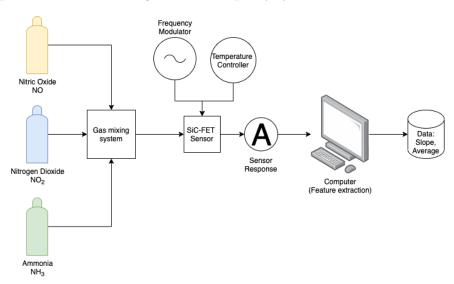


Figure 3.1: Schema of the data acquisition process.

In more detail, NO, NO₂ and NH₃ had five possible concentration values each: 10, 20, 40, 80 and 160 parts per million (ppm). The experiment was designed to encompass all possible combinations of these gasses, which totals to 125 different gas mixtures. Each feature was

submitted to the same frequency cycle five times. The cycle consists of 16 unique frequencies: 0.05, 0.1, 0.25, 0.5, 1, 2, 5, 10, 25, 50, 100, 200, 500, 1000, 2500 and 5000 Hertz (Hz). A typical raw sensor response for frequency modulation experiments is shown on Figure 3.1.

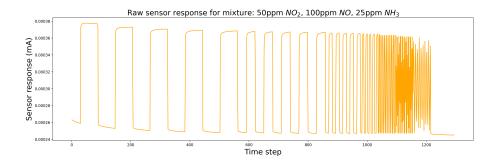


Figure 3.2: An example of row sensor response

For each frequency in each cycle, two slope and two average features were extracted. These measurements were taken during a 0.4 seconds window of time, alternating between slope and average as shown on Figure 3.1.

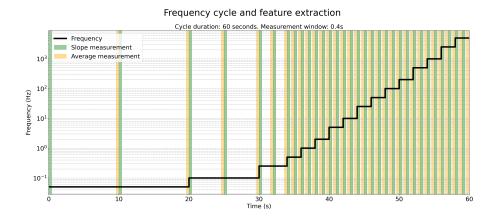


Figure 3.3: Feature measurements times per cycle.

Finally, all 125 gas mixtures were subjected to the experiment three times, each time at a different temperature. Table 3.1 summarizes the data acquisition details.

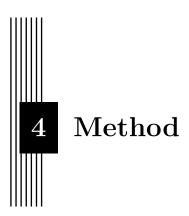
For specific timestamps and measurement durations, the reader is referred to Appendix 7.

3.2 Raw data

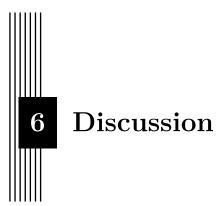
3.3 Secondary data

Table 3.1: Data acquisition details

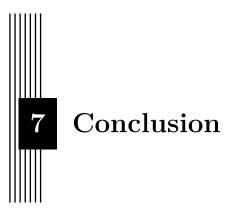
Parameter	Value
Factors (gases)	3
Levels (concentrations)	5
Frequencies	16
Features per frequency	4 (2 slopes and 2 averages)
Features per cycle	64
Number of cycles	5
Data points per mixture	320
Number of mixtures	125
Datapoints per experiment	40.000
Number of experiments	3
Total data points	120.000

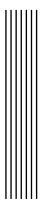






- 6.1 Results
- 6.2 Method
- 6.3 The work in a wider context

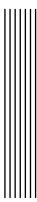




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Appendix

Appendix A: Data acquisition time stamps

Frequency (Hz)	Duration (s)	Feature	Start time (s)	End time (s)
		Slope	0,0	0,4
0.05	20	Average	9,6	10,0
0.05	20	Slope	10,0	10,4
		Average	19,6	20,0
		Slope	20,0	20,4
0.1	10	Average	24,6	25,0
5.2		Slope	25,0	25,4
		Average	29,6	30,0
		Slope	30,0	30,4
0.25	4	Average	31,6	32,0
		Slope	32,0	32,4
		Average	33,6	34,0
		Slope	34,0	34,4
0.5	2	Average	34,6	35,0
		Slope Average	35,0 35,6	35,4 36,0
		Slope	36,0	36,4
		Average	36,6	37,0
1.0	2	Slope	37,0	37,4
		Average	37,6	38,0
		Slope	38,0	38,4
		Average	38,6	39,0
2.0	2	Slope	39,0	39,4
		Average	39,6	40,0
		Slope	40,0	40,4
5.0	_	Average	40,6	41,0
5.0	2	Slope	41,0	41,4
		Average	41,6	42,0
		Slope	42,0	42,4
10.0	2	Average	42,6	43,0
10.0	2	Slope	43,0	43,4
		Average	43,6	44,0
		Slope	44,0	44,4
25.0	2	Average	44,6	45,0
	_	Slope	45,0	45,4
		Average	45,6	46,0
		Slope	46,0	46,4
50.0	2	Average	46,6	47,0
		Slope	47,0	47,4
		Average	47,6	48,0
		Slope	48,0	48,4
100.0	2	Average Slope	48,6 49,0	49,0 49,4
		Average	49,6	50,0
		Slope	50,0	50,4
	2	Average	50,6	51,0
200.0		Slope	51,0	51,4
		Average	51,6	52,0
		Slope	52,0	52,4
E00.0	,	Average	52,6	53,0
500.0	2	Slope	53,0	53,4
		Average	53,6	54,0
		Slope	54,0	54,4
1000.0	2	Average	54,6	55,0
1000.0		Slope	55,0	55,4
		Average	55,6	56,0
	2	Slope	56,0	56,4
2500.0		Average	56,6	57,0
		Slope	57,0	57,4
		Average	57,6	58,0
	2	Slope	58,0	58,4
5000.0		Average	58,6	59,0
		Slope	59,0	59,4
		Average	59,6	60,0

Figure 7.1: Data acquisition timestamps.