

# Predictive Bias and Calibration

DS 6030 | Fall 2024

calibration.pdf

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# 1 Probability Modeling

## 1.1 Data

```
#: Load Data, Create binary column (y)
data(Default, package="ISLR")
Default = Default %>% as_tibble() %>%
  mutate(y = if_else(default == "Yes", 1L, 0L), .after=default)
```

default	y	student	balance	income
No	0	No	1384.9	40131
Yes	1	Yes	1889.3	22652
Yes	1	Yes	1740.8	18161
Yes	1	Yes	2123.4	23836
No	0	Yes	856.7	15523
No	0	No	310.1	31446
No	0	No	1248.9	31960
Yes	1	No	1823.6	44260

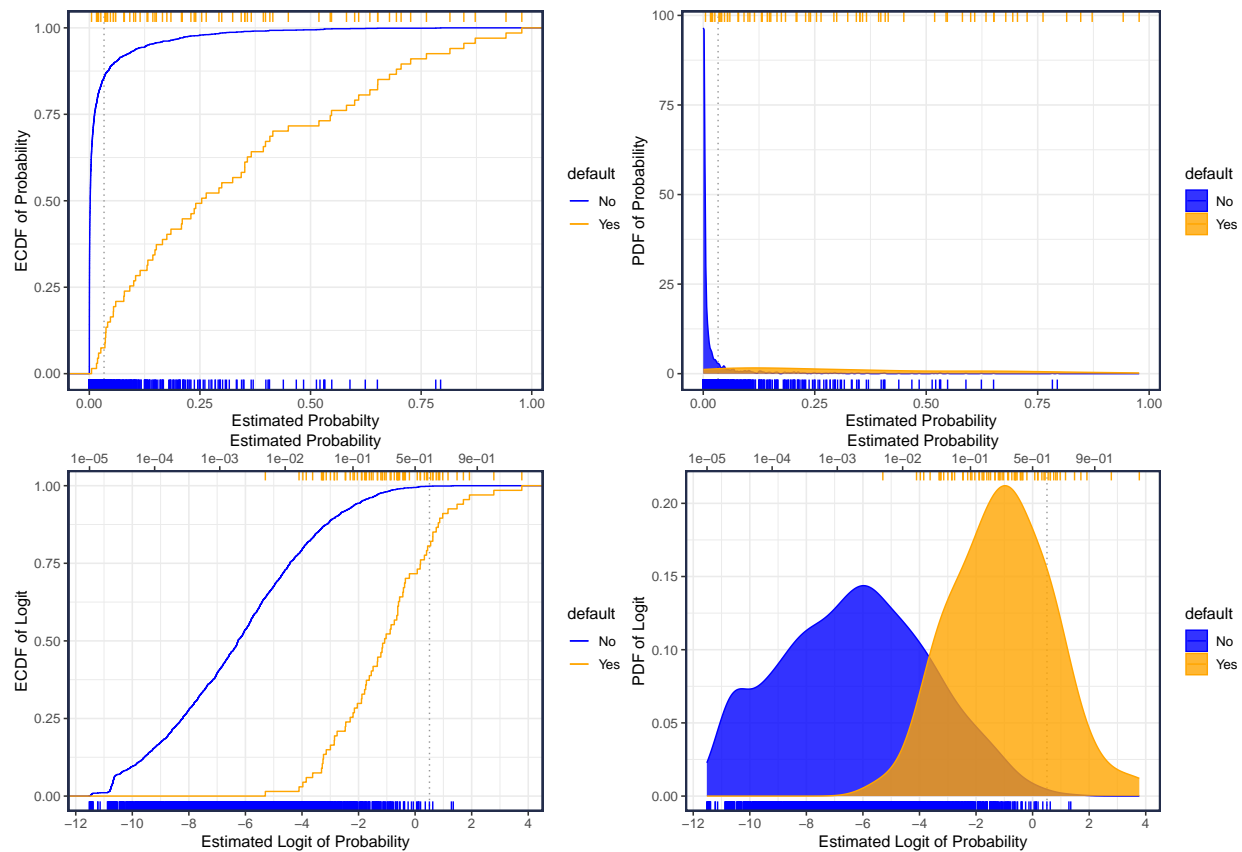
```
#: train/test split
set.seed(2019)
test = sample(nrow(Default), size=2000)
train = -test
```

## 1.2 Logistic Regression

```
#: fit logistic regression model on training data
fit_lr = glm(default~student + balance + income, family='binomial',
  data = Default[train, ])
```

## 1.3 Predictive Performance

```
#: table of predictions and true values
tbl_lr = Default[test,] %>%
  mutate(
    g = factor(y, c(0,1)),
    p_hat = predict(fit_lr, ., type="response"),
    gamma_hat = predict(fit_lr, ., type="link"),
  )
```



```

#: Threshold of 1/2
tbl_lr %>%
  mutate(Y_hat = ifelse(p_hat > 1/2, 1L, 0L)) %>%
  count(y, Y_hat)
#> # A tibble: 4 x 3
#>   y Y_hat     n
#>   <int> <int> <int>
#> 1     0     0 1922
#> 2     0     1   11
#> 3     1     0   48
#> 4     1     1   19

```

```

#: Threshold of 1/4
tbl_lr %>%
  mutate(Y_hat = ifelse(p_hat > 1/4, 1L, 0L)) %>%
  count(y, Y_hat)
#> # A tibble: 4 x 3
#>   y Y_hat     n
#>   <int> <int> <int>
#> 1     0     0 1891
#> 2     0     1   42
#> 3     1     0   33
#> 4     1     1   34

```

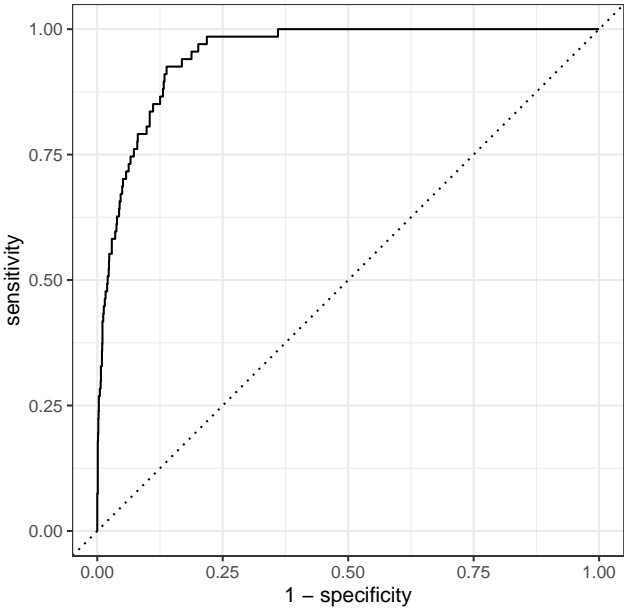
```

library(yardstick)
roc_curve(tbl_lr, truth = g, p_hat, event_level="second") %>%
  autoplot()

```

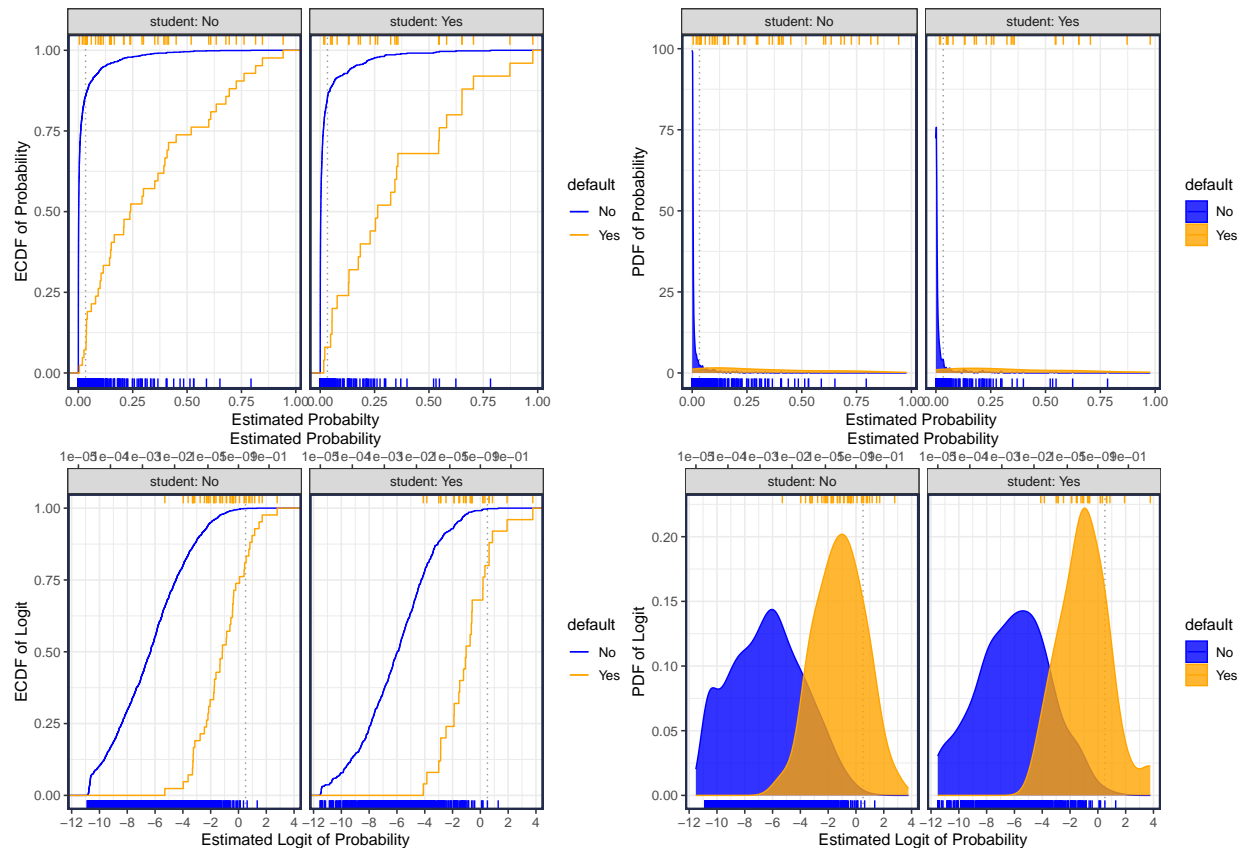
thres	$\hat{P}$	$\hat{N}$	FP	FN	Cost(FP=1, FN=2)	Cost(FP=2, FN=1)	misclass
0.1	178	1822	128	17	162	273	145
0.2	101	1899	62	28	118	152	90
0.3	59	1941	29	37	103	95	66
0.4	39	1961	17	45	107	79	62
0.5	30	1970	11	48	107	70	59
0.6	18	1982	4	53	110	61	57
0.7	10	1990	2	59	120	63	61
0.8	5	1995	0	62	124	62	62
0.9	2	1998	0	65	130	65	65

n	log_loss	brier	mae	auroc
2000	0.082	0.023	0.046	0.951



## 1.4 Performance by Group

All performance scores can also be assessed at the group level (i.e., over subsets of the features). Here we explore the predicted output by Student status.



```
#: Threshold of 1/2
tbl_lr %>%
  mutate(Y_hat = ifelse(p_hat > 1/2, 1L, 0L)) %>%
  count(student, y, Y_hat)
#> # A tibble: 8 x 4
#>   student y Y_hat n
#>   <fct>   <int> <int> <int>
#> 1 No      0      0 1356
#> 2 No      0      1    6
#> 3 No      1      0   31
#> 4 No      1      1   11
#> 5 Yes     0      0  566
#> 6 Yes     0      1    5
#> # i 2 more rows
```

```
#: Threshold of 1/4
tbl_lr %>%
  mutate(Y_hat = ifelse(p_hat > 1/4, 1L, 0L)) %>%
  count(student, y, Y_hat)
#> # A tibble: 8 x 4
#>   student y Y_hat n
#>   <fct>   <int> <int> <int>
#> 1 No      0      0 1334
#> 2 No      0      1   28
#> 3 No      1      0   22
```

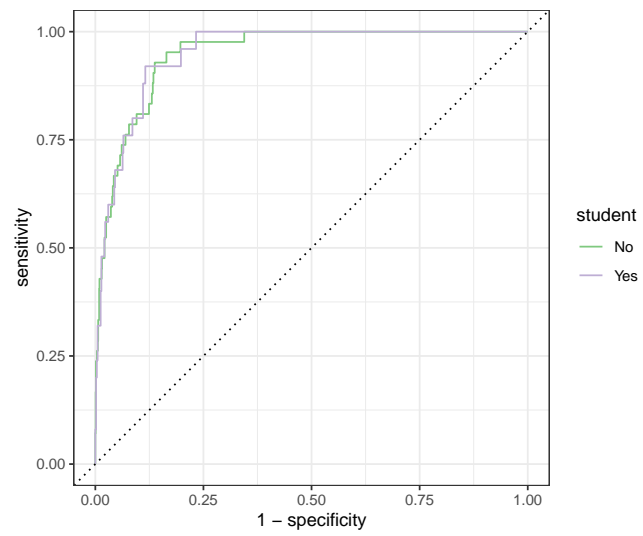
```
#> 4 No      1      1      20
#> 5 Yes     0      0     557
#> 6 Yes     0      1      14
#> # i 2 more rows
```

thres	student	$\hat{P}$	$\hat{N}$	FP	FN	Cost(FP=1, FN=2)	Cost(FP=2, FN=1)	misclass
0.1	No	114	1290	83	11	105	177	94
0.2	No	63	1341	39	18	75	96	57
0.3	No	38	1366	20	24	68	64	44
0.4	No	26	1378	12	28	68	52	40
0.5	No	17	1387	6	31	68	43	37
0.6	No	11	1393	2	33	68	37	35
0.7	No	6	1398	1	37	75	39	38
0.8	No	3	1401	0	39	78	39	39
0.9	No	1	1403	0	41	82	41	41

thres	student	$\hat{P}$	$\hat{N}$	FP	FN	Cost(FP=1, FN=2)	Cost(FP=2, FN=1)	misclass
0.1	Yes	64	532	45	6	57	96	51
0.2	Yes	38	558	23	10	43	56	33
0.3	Yes	21	575	9	13	35	31	22
0.4	Yes	13	583	5	17	39	27	22
0.5	Yes	13	583	5	17	39	27	22
0.6	Yes	7	589	2	20	42	24	22
0.7	Yes	4	592	1	22	45	24	23
0.8	Yes	2	594	0	23	46	23	23
0.9	Yes	1	595	0	24	48	24	24

student	n	log_loss	brier	mae	auroc
No	1404	0.076	0.021	0.042	0.949
Yes	596	0.096	0.028	0.054	0.951

```
library(yardstick)
tbl_lr %>% group_by(student) %>%
  roc_curve(truth = g, p_hat, event_level="second") %>%
  autoplot() + scale_color_brewer(type = "qual")
```



## 2 Predictive Bias and Calibration

A risk model is said to be *calibrated* if the predicted probabilities are equal to the true risk (probabilities).

→ binary classification model

$$\Pr(Y = 1 \mid \hat{p}(x) = p) = p \quad \text{for all } p$$

To evaluate the calibration of a model's predictions, we need to estimate the proportion of the observations with  $Y = 1$  and  $\hat{p} \approx p$ .

Calibration *plots* can be used to measure drift, fairness, and model/algorithmic bias. We could for example use binning (regressograms/histograms), kNN, smoothing, or isotonic regression.

### 2.1 Binning

Here is an example of binning. I'll partition the predictions such that there are 10 groups of equal width.

```
tbl_grp = tbl_lr %>%
  mutate(grp = cut_width(p_hat, width = .1, boundary = 1)) %>%
  group_by(grp) %>%
  summarize(
    n = n(),
    lower = min(p_hat),
    upper = max(p_hat),
    p_hat = mean(p_hat),
    n_1 = sum(y),
    # bayesian (uniform prior for each bin)
    p_1 = (n_1 + 1) / (n + 2), # posterior mean
    beta_lower = qbeta(.025, n_1 + 1, n - n_1 + 1),
    beta_upper = qbeta(.975, n_1 + 1, n - n_1 + 1),
    # frequentist
    p_1_bar = mean(y),
    moe = 1.96 * sqrt(p_1_bar * (1 - p_1_bar) / n)
  )
```

$$\frac{n_1}{n_1 + n_0}$$

$$P_1 = \frac{n_1}{n}$$

bin 1

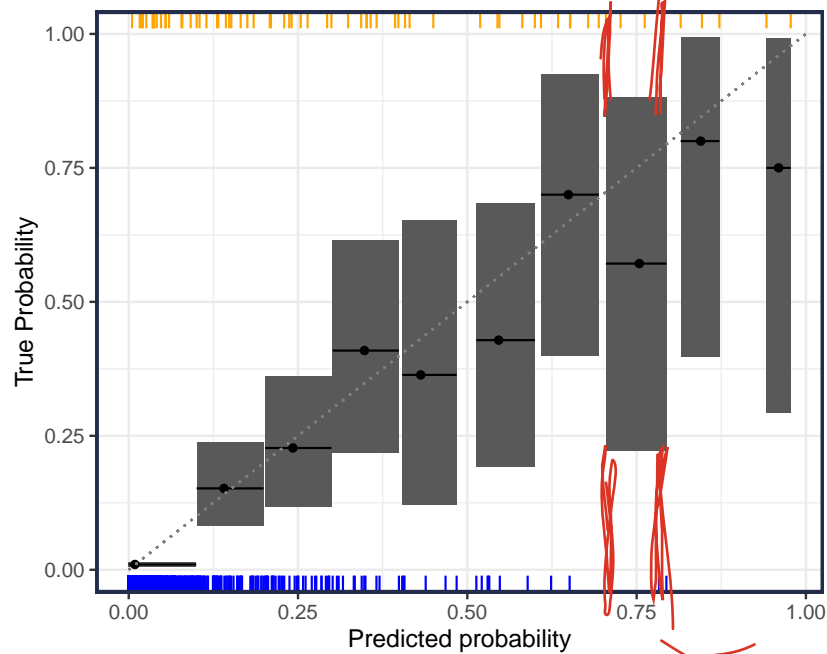
grp	n	lower	upper	p_hat	n_1	p_1	beta_lower	beta_upper	p_1_bar	moe
[0,0.1]	1822	0.000	0.100	0.009	17	0.010	0.006	0.015	0.009	0.004
(0.1,0.2]	77	0.100	0.199	0.141	11	0.152	0.082	0.238	0.143	0.078
(0.2,0.3]	42	0.202	0.299	0.242	9	0.227	0.118	0.360	0.214	0.124
(0.3,0.4]	20	0.301	0.399	0.348	8	0.409	0.218	0.616	0.400	0.215
(0.4,0.5]	9	0.404	0.484	0.431	3	0.364	0.122	0.652	0.333	0.308
(0.5,0.6]	12	0.513	0.600	0.546	5	0.429	0.192	0.684	0.417	0.279
(0.6,0.7]	8	0.609	0.694	0.649	6	0.700	0.400	0.925	0.750	0.300
(0.7,0.8]	5	0.705	0.794	0.754	3	0.571	0.223	0.882	0.600	0.429
(0.8,0.9]	3	0.815	0.872	0.845	3	0.800	0.398	0.994	1.000	0.000
(0.9,1]	2	0.942	0.978	0.960	2	0.750	0.292	0.992	1.000	0.000

bin 10

```
tbl_grp %>%
  ggplot(aes(p_hat, p_1)) +
    geom_rect(aes(xmin=lower, xmax=upper,
                  ymin=beta_lower,
                  ymax=beta_upper))
    ) +
  geom_point()
```



```
geom_linerange(aes(xmin=lower, xmax=upper)) +
geom_segment(x=0,xend=1, y=0, yend=1, linetype = 3, color = "grey50") +
labs(x = "Predicted probability", y = "True Probability") +
geom_rug(data = tbl_lr %>% filter(y==0),
  aes(x=p_hat, color=default), sides="b",
  inherit.aes = FALSE) +
geom_rug(data = tbl_lr %>% filter(y==1),
  aes(x=p_hat,color=default), sides="t",
  inherit.aes = FALSE) +
scale_color_manual(values=c(Yes="orange", No="blue"))
```

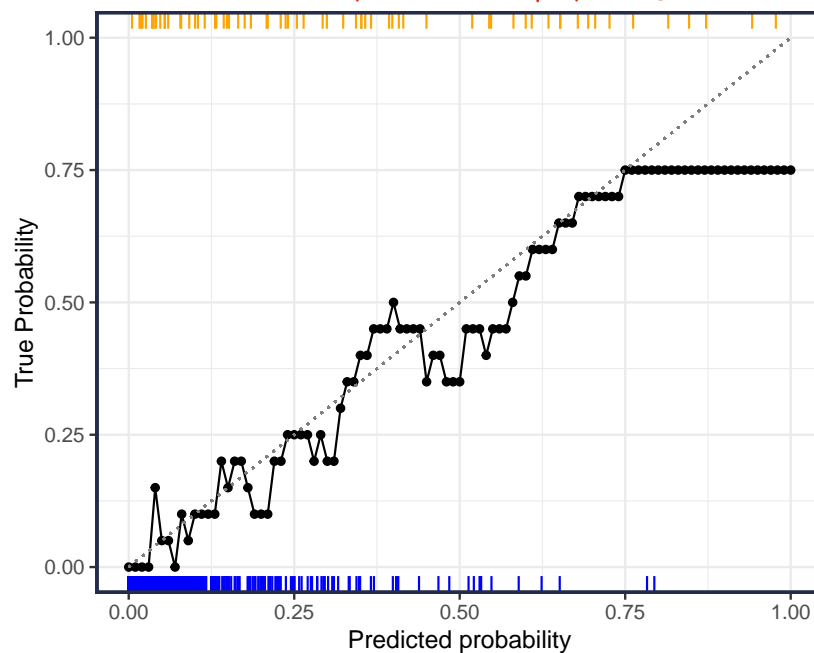


$P = \frac{3}{5} = 60$   
 $CI(P) = \text{point est} \pm MOE$   
 $= 60 \pm$   
 $196 \sqrt{\frac{6 \times 4}{5}}$   
 $95\%$   
 $SE$

## 2.2 Nearest Neighbor

K-nearest neighbor ( $k = 20$ ):

$[0.1, 0.8]$



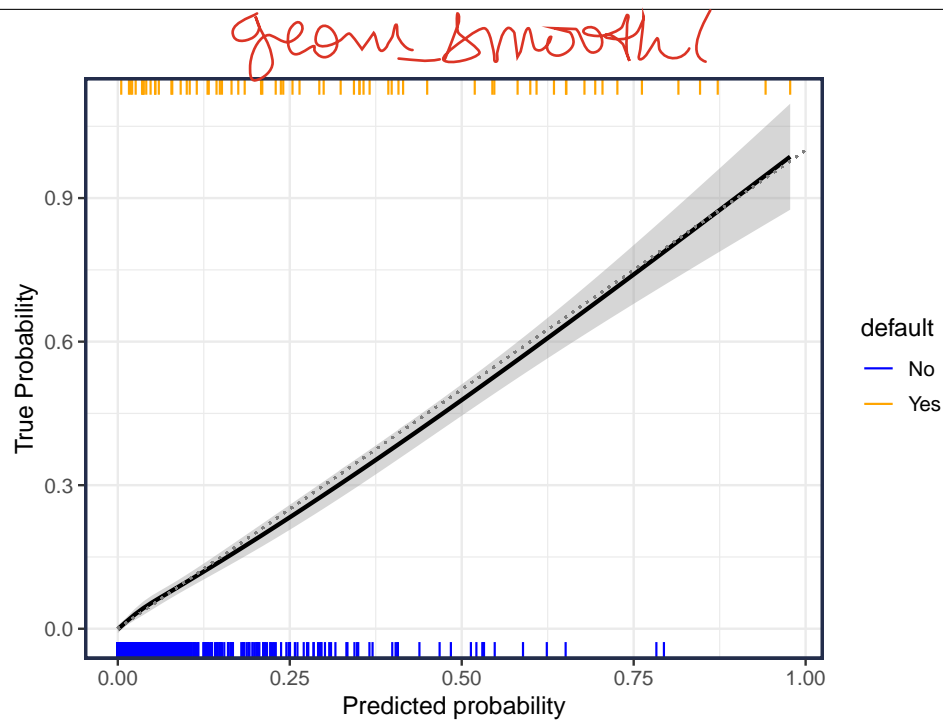
test data  
 predictor =  $\hat{p}$   
 outcome =  $y \in \{0, 1\}$

But knn is not expected to work well at the edges ( $\hat{p}$  close to 0 or 1).

## 2.3 Smoothing splines

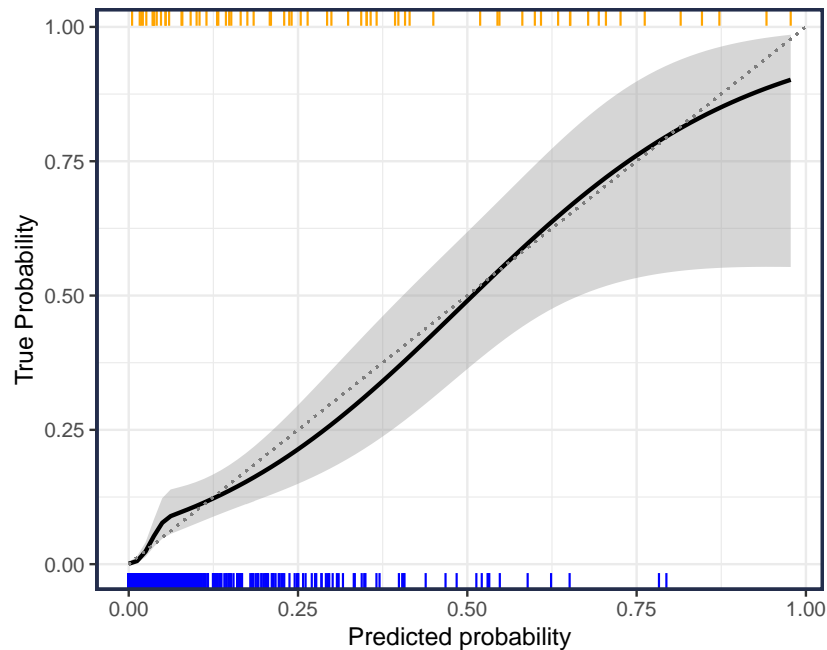
Using smoothing splines (brier score loss):

```
#: The Regular geom_smooth() uses an MSE (brier score) loss
tbl_lr %>%
  ggplot(aes(p_hat)) +
  geom_smooth(aes(y = y), color="black") +
  geom_segment(x=0, xend=1, y=0, yend=1, linetype = 3, color = "grey50") +
  geom_rug(data = . %>% filter(y==0), aes(color=default), sides="b") +
  geom_rug(data = . %>% filter(y==1), aes(color=default), sides="t") +
  scale_color_manual(values=c(Yes="orange", No="blue")) +
  labs(x = "Predicted probability", y = "True Probability")
```



Using smoothing splines (log-loss):

```
#: Here we use family argument in geom_smooth() to force log-loss
tbl_lr %>%
  ggplot(aes(p_hat)) +
  geom_smooth(aes(y = y),
              method = "gam",
              method.args = list(family = binomial()),
              color = "black") +
  geom_segment(x = 0, xend = 1, y = 0, yend = 1, linetype = 3, color = "grey50") +
  geom_rug(data = . %>% filter(y==0), aes(color=default), sides="b") +
  geom_rug(data = . %>% filter(y==1), aes(color=default), sides="t") +
  scale_color_manual(values=c(Yes="orange", No="blue")) +
  labs(x = "Predicted probability", y = "True Probability")
```



if you do linear regression on binary data, then you are violating homoskedasticity principal

default  
— No  
— Yes

## 2.4 Summary

It looks like the logistic regression model is decently calibrated for *overall* calibration (i.e., aggregated). The predicted probabilities are suitably close to the observed proportions in a test set. That is, with 2000 test observations there is not enough evidence to suggest clear predictive bias.

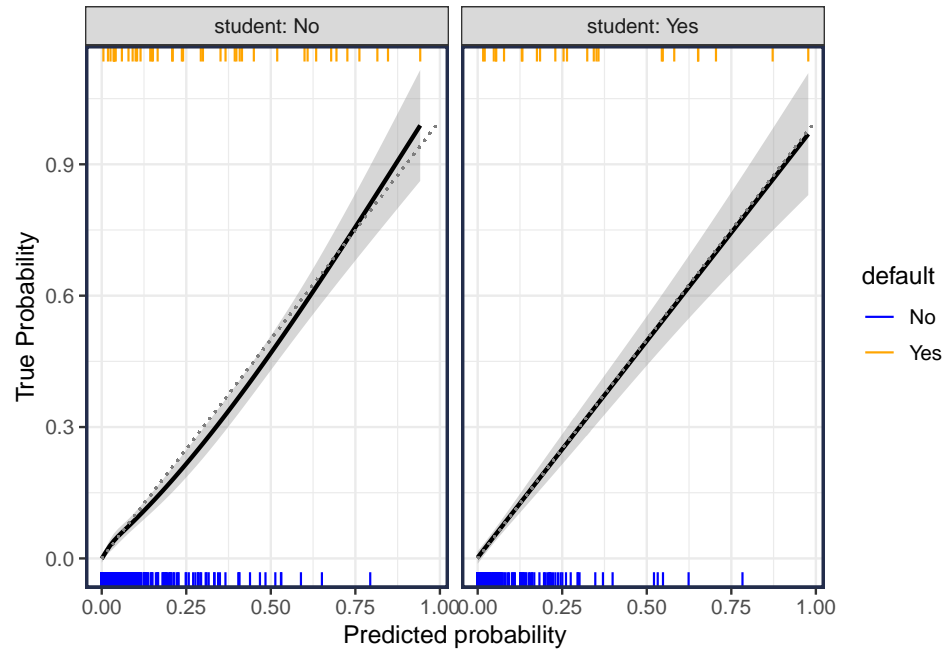
But, does predictive performance and calibration hold for all groups?

## 2.5 Calibration by group

Consider comparing the predictive performance of our models for Students and Non-Students.

$$\Pr(Y = 1 \mid \hat{p}(x) = p, X = x) = p \quad \text{for all } p \text{ and } x$$

```
tbl_lr %>%
  ggplot(aes(p_hat)) +
  geom_smooth(aes(y = y), color="black") +
  geom_segment(x=0,xend=1, y=0, yend=1, linetype = 3, color = "grey50") +
  labs(x = "predicted probability", y = "true - predicted") +
  geom_rug(data = . %>% filter(y==0), aes(color=default), sides="b") +
  geom_rug(data = . %>% filter(y==1), aes(color=default), sides="t") +
  scale_color_manual(values=c(Yes="orange", No="blue")) +
  labs(x = "Predicted probability", y = "True Probability") +
  facet_wrap(~student, labeller = label_both) # <--- Faceting
```



### 3 Testing for Calibration

The main idea is to test the (null) hypothesis that

$$\Pr(Y = 1 \mid \hat{p}(x) = p) = p \quad \text{for all } p$$

or equivalently,

$$\Pr(Y = 1 \mid \hat{p}(x)) - \hat{p}(x) = 0 \quad \text{for all } \hat{p}$$

The plots shown above visually explore the hypothesis. We can also take a model-based approach.

#### 3.1 Logistic Regression

For a calibrated model the estimated  $\hat{p}(x)$  should be close to the true  $p(x)$ ,

$$p(x) \approx \hat{p}(x)$$

$$\text{logit } p(x) \approx \text{logit } \hat{p}(x)$$

Logistic Regr

$$\text{logit } p = \beta_0 + \beta_1 X$$

To test this, we introduce a bias term  $b(x)$  and test for  $b(x) = 0 \forall x$ .

$$\text{logit } p(x) = b(x) + \text{logit } \hat{p}(x)$$

Notice the above expression is the same form as logistic regression.

This means we can use logistic regression to test for mis-calibration (predictive bias). To do this, use  $\text{logit } \hat{p}(x)$  as an *offset* in the logistic regression model. An *offset* is a term that has a fixed weight/coefficient of 1.

We also need to specify the form of the bias term before we can test it. We give a few examples below.

#### 3.2 Linear bias

To check for linear deviation, we specify the bias term as  $b(x) = \beta_0 + \beta_1 \text{logit } \hat{p}(x)$ .

$$\text{logit } p(x) = \beta_0 + \beta_1 \text{logit } \hat{p}(x) + (\text{logit } \hat{p}(x)) : \text{offset}$$

Fit on a hold-out set, and check how far  $\beta_0$  and  $\beta_1$  are from 0.

```
calibrated = glm(y ~ gamma_hat + offset(gamma_hat),
  family = binomial,
  data = tbl_lr)
# Note that gamma_hat = logit(p_hat)

calibrated %>% broom::tidy()
#> # A tibble: 2 x 5
#>   term                estimate std.error statistic p.value
#>   <chr>              <dbl>    <dbl>    <dbl>    <dbl>
#> 1 (Intercept)    -0.0224      0.220    -0.102     0.919
#> 2 gamma_hat       0.00589     0.0923     0.0638     0.949
```

No significance suggests not enough evidence to reject null of no bias against linear bias.

### 3.3 Non-linear bias

We can introduce splines to detect non-linear deviations:

```
glm(y~splines::bs(gamma_hat, df = 3) + offset(gamma_hat),
    family = binomial,
    data = tbl_lr) %>%
  broom::tidy()
#> Warning: glm.fit: fitted probabilities numerically 0 or 1 occurred
#> # A tibble: 4 x 5
#>   term                                estimate std.error statistic p.value
#>   <chr>                                <dbl>      <dbl>      <dbl>  <dbl>
#> 1 (Intercept)                        -27.0        22.5       -1.20   0.230
#> 2 splines::bs(gamma_hat, df = 3)1      43.7        36.7        1.19   0.233
#> 3 splines::bs(gamma_hat, df = 3)2      18.4        15.3        1.21   0.228
#> 4 splines::bs(gamma_hat, df = 3)3      30.5        26.1        1.17   0.243
```

Again, no significance suggests not enough evidence to reject null of no bias against non-linear bias (although we could try other smoothing).

### 3.4 Bias in predictors

We can also see if the predictions should be adjusted for regions in features space.

Including an interaction term too.

```
glm(y ~ student + student:gamma_hat + offset(gamma_hat),
    family = binomial,
    data = tbl_lr) %>%
  broom::tidy()
#> # A tibble: 4 x 5
#>   term                                estimate std.error statistic p.value
#>   <chr>                                <dbl>      <dbl>      <dbl>  <dbl>
#> 1 (Intercept)                        -0.0734     0.276     -0.266   0.791
#> 2 studentYes                          0.141       0.458      0.307   0.759
#> 3 studentNo:gamma_hat                  0.00509     0.115      0.0443   0.965
#> 4 studentYes:gamma_hat                 0.00368     0.155      0.0237   0.981
```

```
glm(y ~ splines::bs(balance, 3) + offset(gamma_hat),
    family = binomial,
    data = tbl_lr) %>%
  broom::tidy()
#> # A tibble: 4 x 5
#>   term                                estimate std.error statistic p.value
#>   <chr>                                <dbl>      <dbl>      <dbl>  <dbl>
#> 1 (Intercept)                        -15.5        15.6     -0.997   0.319
#> 2 splines::bs(balance, 3)1            26.3        27.0      0.974   0.330
#> 3 splines::bs(balance, 3)2             9.37         9.26      1.01   0.312
#> 4 splines::bs(balance, 3)3            18.4        19.5      0.946   0.344
```

```
glm(y ~ splines::bs(income, 3) + offset(gamma_hat),
    family = binomial,
    data = tbl_lr) %>%
  broom::tidy()
#> # A tibble: 4 x 5
#>   term                                estimate std.error statistic p.value
#>   <chr>                                <dbl>      <dbl>      <dbl>  <dbl>
#> 1 (Intercept)                         0.435        1.04      0.419   0.675
#> 2 splines::bs(income, 3)1            -0.756        2.83     -0.267   0.789
#> 3 splines::bs(income, 3)2            -0.299        1.65     -0.182   0.856
```

```
#> 4 splines::bs(income, 3)3 -1.10 2.56 -0.431 0.667
```

The logistic regression model appears well-calibrated. This isn't surprising as the log-loss metric encourages good calibration.

However, even using logistic regression we could get poor calibration due to over-fitting, drift, or mis-specification.



## 4 SVM Calibration

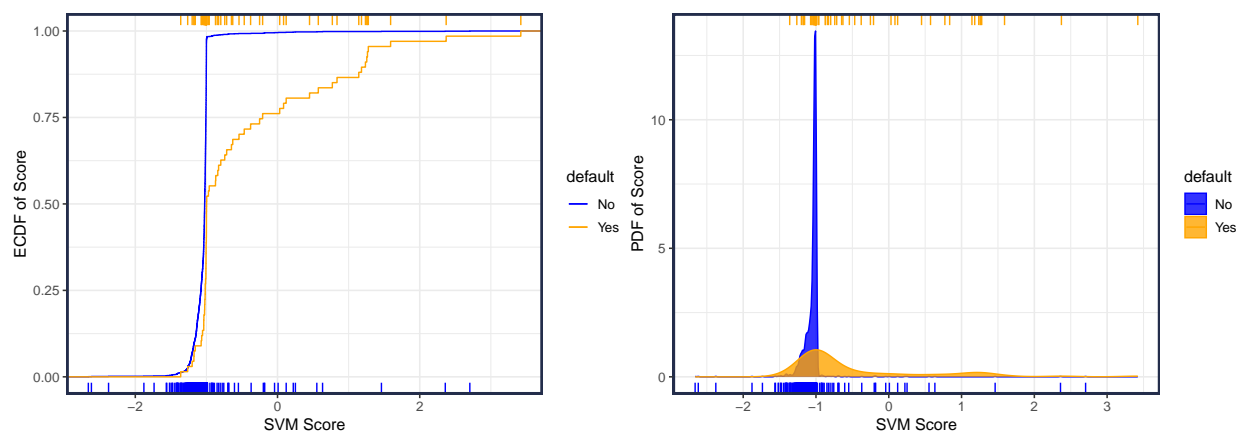
### 4.1 SVM Modeling

```
library(kernlab)
# Radial basis kernel
fit_svm = ksvm(default~student + balance + income,
               data = Default[train, ],
               scaled = TRUE, # be sure to scale the predictors
               kernel = "rbfdot", # kernel
               sigma = .01, C = 100, # tuning parameters
               )

# table of predictions and true values
tbl_svm = Default[test,] %>%
  mutate(
    g = factor(y, c(0,1)),
    score = predict(fit_svm, ., type = "decision")[,1],
    class = predict(fit_svm, ., "response"),
  )
```

SVMs don't naturally output a probability, but rather a score (or decision) value that indicates the observations' distance to the decision boundary. We called the SVM score output  $\hat{f}(x)$  in the class notes.

default	y	student	balance	income	g	score	class
No	0	Yes	1026.1	16431	0	-1.023	No
No	0	Yes	1378.2	20120	0	-1.002	No
No	0	No	487.1	55459	0	-1.156	No
No	0	Yes	1252.1	13861	0	-0.999	No
No	0	No	0.0	30208	0	-1.004	No
No	0	No	921.3	25350	0	-1.046	No



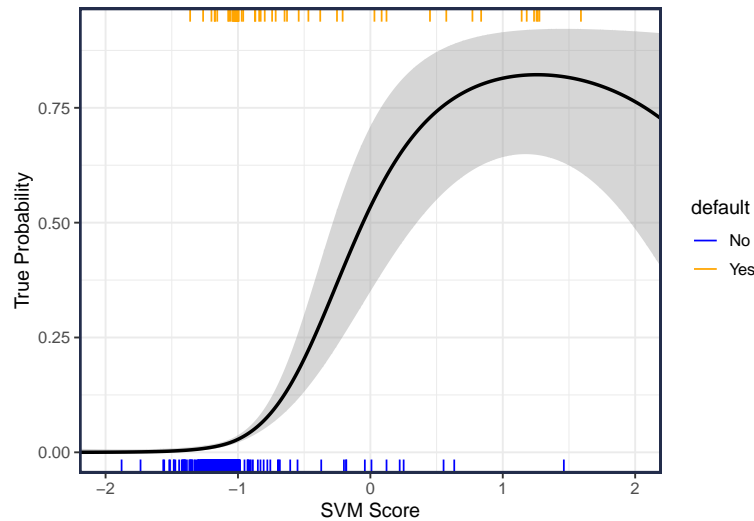
Let's check out how well the SVM scores map to a probability. Here I'll go with the smoothing spline approach (log-loss):

```
# Here we use family argument in geom_smooth() to force log-loss
tbl_svm %>%
  ggplot(aes(score)) +
  geom_smooth(aes(y = y), n=1000,
             method = "gam",
             method.args = list(family = binomial()),
```

```

color="black") +
geom_rug(data = . %>% filter(y==0), aes(color=default), sides="b") +
geom_rug(data = . %>% filter(y==1), aes(color=default), sides="t") +
scale_color_manual(values=c(Yes="orange", No="blue")) +
labs(x = "SVM Score", y = "True Probability") +
scale_x_continuous(breaks = seq(-5, 5, by=1)) +
coord_cartesian(xlim = c(-2, 2))

```



## 4.2 SVM Calibration

Let  $\hat{f}_i = \hat{f}(x_i)$  be the score output from an SVM model. Platt's idea was to fit a logistic regression model using  $\hat{f}(x)$  as the predictor variable. The *calibrated probabilities* are the predictions from this model. That is:

$$\text{logit}(p(x)) = \beta_0 + \beta_1 \hat{f}(x)$$

or

$$p(x) = 1 / \left( 1 + e^{-(\beta_0 + \beta_1 \hat{f}(x))} \right)$$

```

calibrated_svm = glm(y~score, family="binomial", data = tbl_svm)
broom::tidy(calibrated_svm)
#> # A tibble: 2 x 5
#>   term      estimate std.error statistic  p.value
#>   <chr>      <dbl>    <dbl>    <dbl>    <dbl>
#> 1 (Intercept) -1.25    0.294    -4.24 2.24e- 5
#> 2 score        2.30    0.289     7.95 1.93e-15

```

```

tbl_svm = tbl_svm %>%
  mutate(
    p_hat = predict(calibrated_svm, ., type="response"),
    gamma_hat = log(p_hat) - log(1-p_hat)
  )

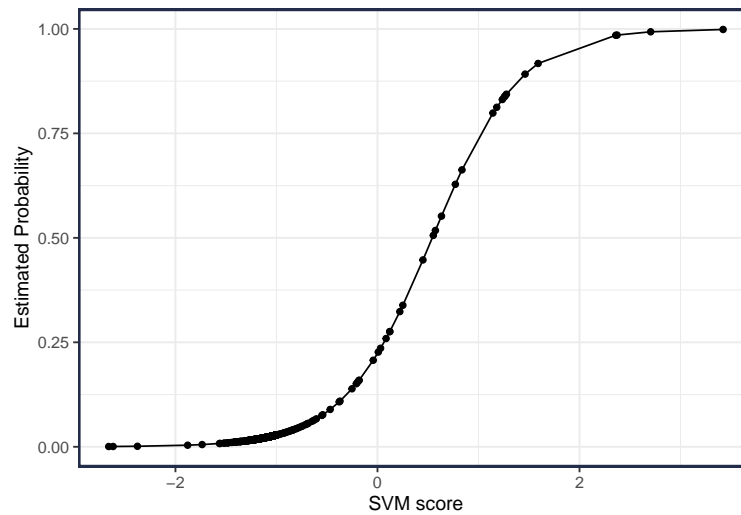
```

```

tbl_svm %>%
  ggplot(aes(score, p_hat)) +
  geom_point() +

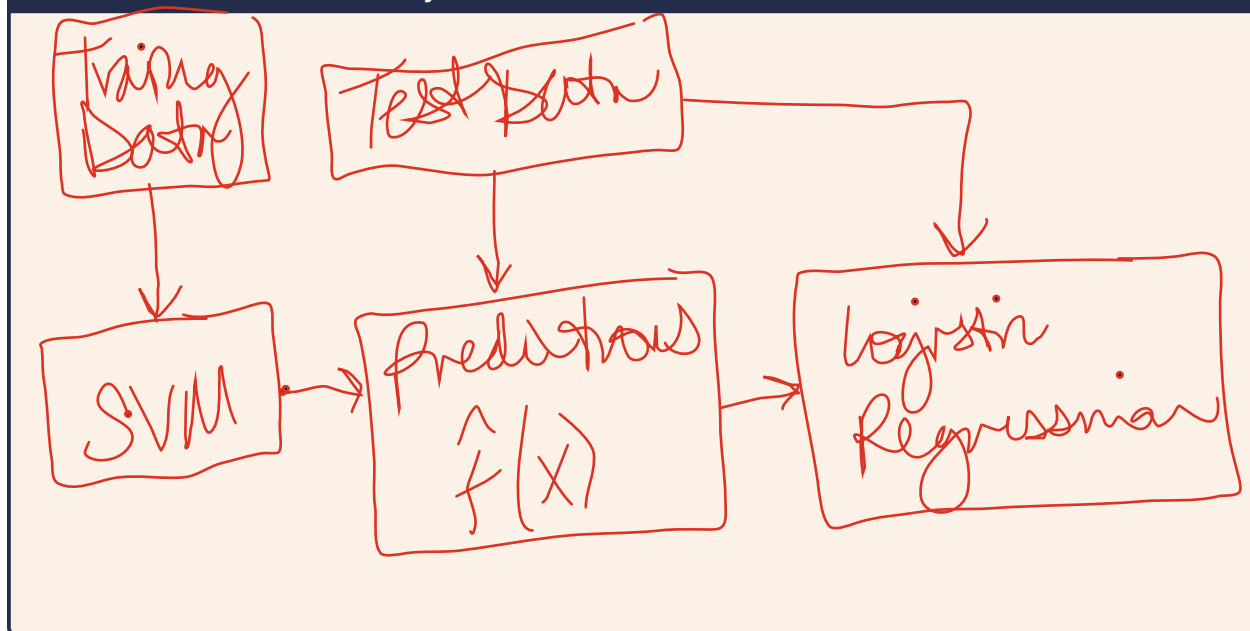
```

```
geom_line() +
labs(x = "SVM score", y = "Estimated Probability")
```



Notice that when the SVM score is zero ( $\hat{f}(x) = 0$ ), the estimated probability is 0.25; this is pretty far from the expected 0.50.

#### Sketch of SVM Calibration System



#### 4.2.1 Testing Calibration

How well did Platt's method work? Is the 2 parameter logistic transformation sufficiently complex?

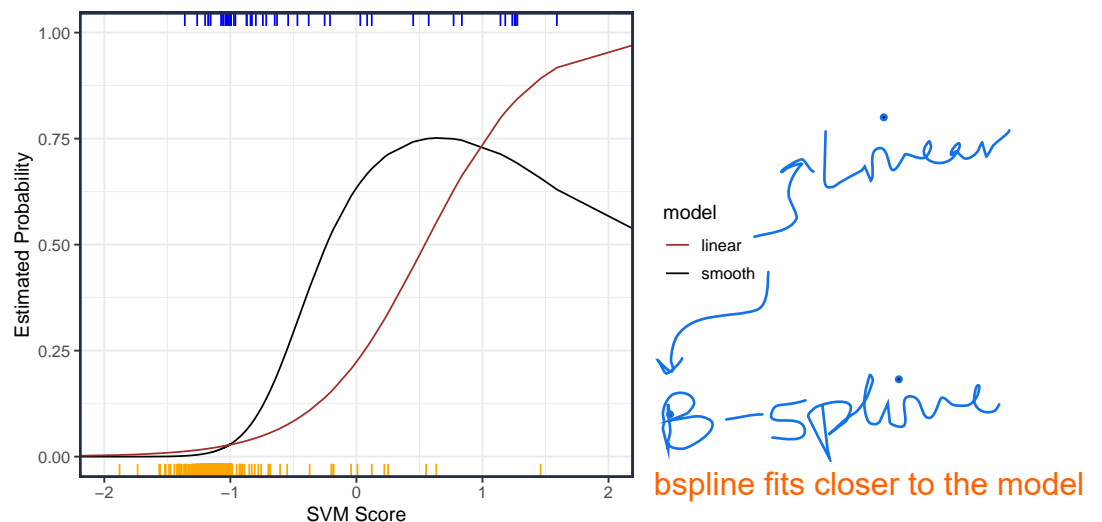
We can for test this, just like we did earlier. However, we are going to see a problem with using splines!

```
calibrated_svm_2 = glm(y~splines::bs(score, df=3),
                      family="binomial", data = tbl_svm)
broom::tidy(calibrated_svm_2)
#> # A tibble: 4 x 5
```

```
#>   term                                estimate std.error statistic    p.value
#>   <chr>                                <dbl>      <dbl>      <dbl>    <dbl>
#> 1 (Intercept)                        -24.4        4.51       -5.41 6.39e- 8
#> 2 splines::bs(score, df = 3)1         39.8        9.89        4.03 5.68e- 5
#> 3 splines::bs(score, df = 3)2         19.4        2.33        8.31 9.19e-17
#> 4 splines::bs(score, df = 3)3         25.9        5.87        4.41 1.05e- 5
```

Notice that the bspline with 3 df (4 edf with intercept) has statistically significant terms. But what does this non-linear calibration look like?

```
tbl_svm %>%
  ggplot(aes(score)) +
  geom_line(aes(y = p_smooth, color = "smooth")) +
  geom_line(aes(y = p_hat, color = "linear")) +
  geom_rug(data = . %>% filter(y==0), color = "orange", sides="b") +
  geom_rug(data = . %>% filter(y==1), color = "blue", sides="t") +
  scale_color_manual(name = "model", values=c(smooth="black", linear="brown")) +
  labs(x = "SVM Score", y = "Estimated Probability") +
  scale_x_continuous(breaks = seq(-5, 5, by=1)) +
  coord_cartesian(xlim = c(-2, 2))
```



That dip doesn't look good! It suggests a non-monotonic transformation which is not appealing. It would mean an individual that SVM classifies as a default ( $\hat{f} < 0$ ) could actually be given a higher probability than someone that SVM scored higher!

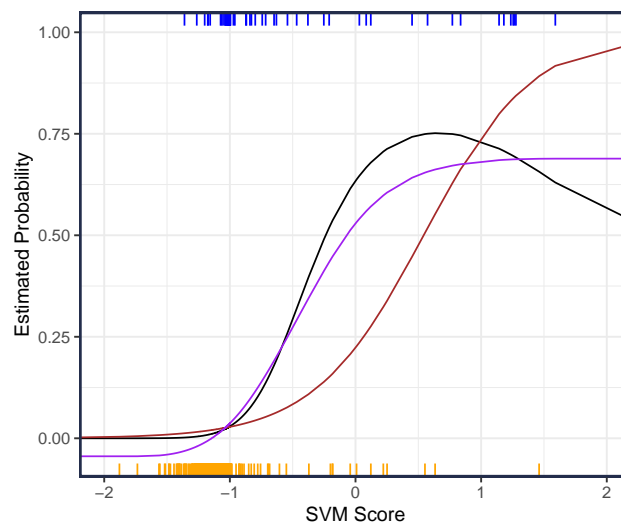
#### 4.2.2 Isotonic Calibration forces the model to go up or down

Enter isotonic (or monotonic) splines. These are special splines that produce a monotonic prediction; something perfect for our application.

```
library(scam)
iso = scam(y~s(score, bs="mpi"), data=tbl_svm)
summary(iso)
#>
#> Family: gaussian
#> Link function: identity
#>
#> Formula:
#> y ~ s(score, bs = "mpi")
```

```
#> <environment: 0x5633dcd96ac0>
#>
#> Parametric coefficients:
#>               Estimate Std. Error t value Pr(>|t|)
#> (Intercept)  0.03350    0.00362    9.25  <2e-16 ***
#> ---
#> Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
#>
#> Approximate significance of smooth terms:
#>               edf Ref.df    F p-value
#> s(score)      3      3 158  <2e-16 ***
#> ---
#> Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
#>
#> R-sq.(adj) =  0.1905   Deviance explained = 19.2%
#> GCV score = 0.026275   Scale est. = 0.026222   n = 2000
```

```
tbl_svm %>%
  ggplot(aes(score)) +
  geom_line(aes(y = p_smooth, color = "smooth")) +
  geom_line(aes(y = p_hat, color = "linear")) +
  geom_line(aes(y = p_iso, color = "isotonic")) +
  geom_rug(data = . %>% filter(y==0), color = "orange", sides="b") +
  geom_rug(data = . %>% filter(y==1), color = "blue", sides="t") +
  scale_color_manual(name = "model",
                     values=c(smooth="black", linear="brown",
                              isotonic = "purple")) +
  labs(x = "SVM Score", y = "Estimated Probability") +
  scale_x_continuous(breaks = seq(-5, 5, by=1)) +
  coord_cartesian(xlim = c(-2, 2))
```



stays constant /  
increases  
here

It does appear that the more complex isotonic smoother deviates from the logistic, but with only 3 edf in the smoothing portion, it isn't too different.

### 4.2.3 Hold-out data

#### Calibration with Validation Data

### 4.2.4 Built-in Calibration

Set `prob.model=TRUE` in the `ksvm()` function to get probability estimates. Note that this uses 3-fold cross-validation to predict out-of-sample scores. Then uses weighted logistic regression to estimate the calibration function.

```
library(kernlab)
# Radial basis kernel
fit_svm = ksvm(default~student + balance + income,
               data = Default[train, ],
               prob.model = TRUE, # ** Need to set to TRUE to get probabilities
               scaled = TRUE,    # be sure to scale the predictors
               kernel = "rbfdot", # kernel
               sigma = .01, C = 100, # tuning parameters
               )
```

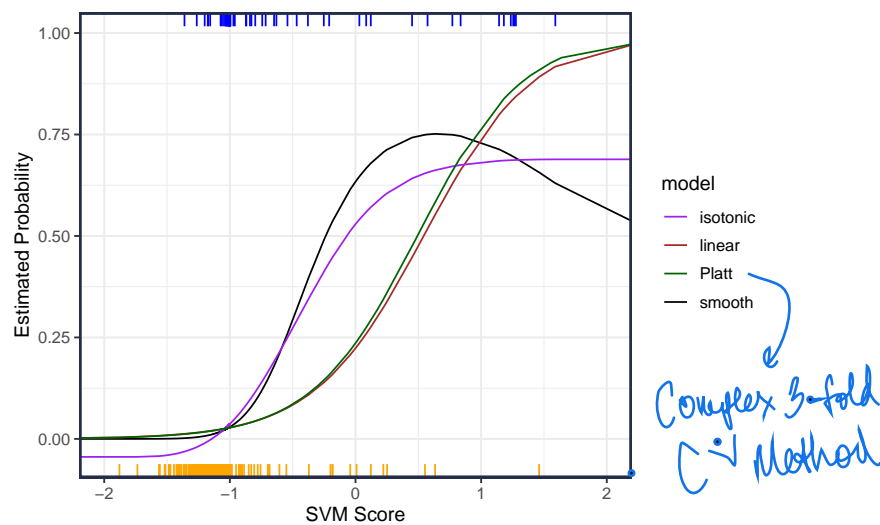
```
# table of predictions and true values
tbl_svm = tbl_svm %>%
  mutate(
    score_ksvm = predict(fit_svm, ., type = "decision")[,1],
```

```

p_ksvm = predict(fit_svm, ., type = "probabilities")[, "Yes"]
)

tbl_svm %>%
  ggplot(aes(score)) +
  geom_line(aes(y = p_smooth, color = "smooth")) +
  geom_line(aes(y = p_hat, color = "linear")) +
  geom_line(aes(y = p_iso, color = "isotonic")) +
  geom_line(aes(x=score_ksvm, y = p_ksvm, color = "Platt")) +
  geom_rug(data = . %>% filter(y==0), color = "orange", sides="b") +
  geom_rug(data = . %>% filter(y==1), color = "blue", sides="t") +
  scale_color_manual(name = "model",
    values=c(smooth="black", linear="brown",
             isotonic = "purple", Platt="darkgreen")) +
  labs(x = "SVM Score", y = "Estimated Probability") +
  scale_x_continuous(breaks = seq(-5, 5, by=1)) +
  coord_cartesian(xlim = c(-2, 2))

```



The built-in cross-validation calibration matches pretty closely to just implementing logistic regression on the in-sample data. I recommend using out-of-sample data to “calibrate”, but if using a simple (e.g., 2 parameter logistic regression) model, using the same data to estimated model parameters and calibration may not be too bad.