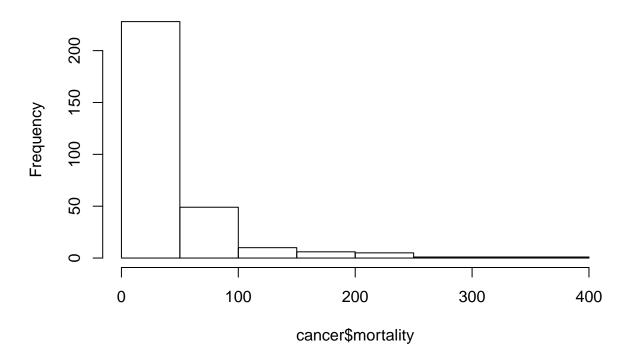
HW 1 #65

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Parts a, b, c, d, f

```
# a.
hist(cancer$mortality, freq = T)
```

Histogram of cancer\$mortality



```
# b.
mean(cancer$mortality)

## [1] 39.85714

sum(cancer$mortality)
```

[1] 11997

```
(length(cancer$mortality)-1)/length(cancer$mortality)*var(cancer$mortality) #population var

## [1] 2590.103

(length(cancer$mortality)-1)/length(cancer$mortality)*sd(cancer$mortality) #pop SD

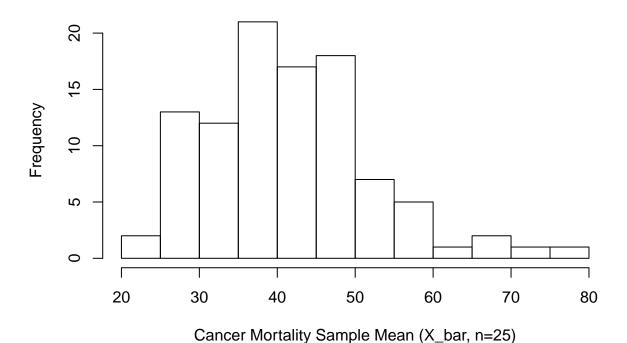
## [1] 50.80844

# Note: Since the variance function in R uses the n-1 in the denominator, I modify
#the result to reflect the actual population variance, not the estimator.

# c. Simulate sampling distribution of x_bar with n=25

xbars <- numeric(100)
for(i in 1:100){
    xbars[i] <- mean(sample(cancer$mortality, 25, replace = F))
}
hist(xbars, freq = T, xlab = "Cancer Mortality Sample Mean (X_bar, n=25)")</pre>
```

Histogram of xbars



```
# d.
x_bar <- mean(sample(cancer$mortality, 25, replace = F)) #n=25
x_bar #mean cancer mortality estimator</pre>
```

[1] 46.6

```
total <- x_bar*301 #total cancer mortality estimator</pre>
total
## [1] 14026.6
# e.
popvar <- (1-(1/301))*var(cancer$mortality)</pre>
sqrt(popvar) # pop SD
## [1] 50.89305
# f.
var_xbar \leftarrow (var(cancer\$mortality)/25)*(1-(25/301))
lower_lim <- x_bar - (1.96*sqrt(var_xbar))</pre>
upper_lim <- x_bar + (1.96*sqrt(var_xbar))</pre>
c(lower_lim, upper_lim) # 95% CI for X_bar
## [1] 27.46456 65.73544
#I am 95% confident that the average cancer mortality per county
#in the three states is contained within the values above
var_total <- (301^2)*var_xbar</pre>
lower_lim <- total - 1.96*sqrt(var_total)</pre>
upper_lim <- total + 1.96*sqrt(var_total)</pre>
c(lower_lim, upper_lim) # 95% CI for T
## [1] 8266.832 19786.368
#I am 95% confident that the total cancer for all counties
#in the three states is contained within the values above
#Conclusion: Yes, the intervals cover the population values.
Part g. (repeat for n=100)
# d.
x_bar <- mean(sample(cancer$mortality, 100, replace = F)) #n=100</pre>
x_bar #mean cancer mortality estimator
## [1] 34.18
total <- x_bar*301 #total cancer mortality estimator</pre>
total
```

[1] 10288.18

```
popvar <- (1-(1/301))*var(cancer$mortality)</pre>
sqrt(popvar) # pop SD
## [1] 50.89305
# f.
var_xbar \leftarrow (var(cancer\$mortality)/100)*(1-(100/301))
lower_lim <- x_bar - (1.96*sqrt(var_xbar))</pre>
upper_lim <- x_bar + (1.96*sqrt(var_xbar))</pre>
c(lower_lim, upper_lim) # 95% CI for X_bar
## [1] 26.01508 42.34492
#I am 95% confident that the average cancer mortality per county
#in the three states is contained within the values above
var_total <- (301^2)*var_xbar</pre>
lower_lim <- total - 1.96*sqrt(var_total)</pre>
upper_lim <- total + 1.96*sqrt(var_total)</pre>
c(lower_lim, upper_lim) # 95% CI for T
## [1] 7830.539 12745.821
#I am 95% confident that the total cancer mortality for all counties
#in the three states is contained within the values above
#Conclusion: Yes, the intervals cover the population values.
```

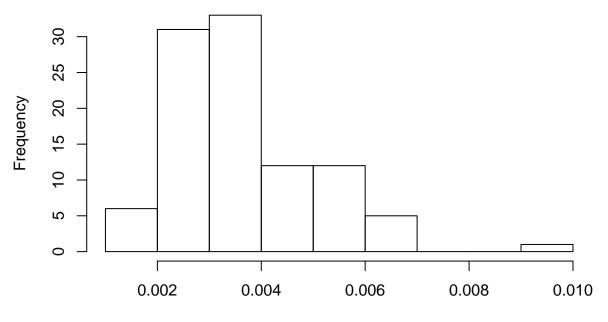
Part h.

This is effective because the ratio estimates achieve comparable precision with a much smaller

Parts i, j, k

```
# i.
cancer$prop <- cancer$mortality/cancer$pop
ratios <- numeric(100)
for(i in 1:100){
    y_bar <- mean(sample(cancer$mortality, 25, replace = F))
    x_bar <- mean(sample(cancer$pop, 25, replace = F))
    ratios[i] <- y_bar/x_bar
}
hist(ratios, freq = T, xlab = "Sample Proportion Mean (p_hats, n=25)")</pre>
```

Histogram of ratios



#Compared to part c, the range of the distribution is more constrained.

Sample Proportion Mean (p_hats, n=25)

```
# j.
y_sample <- sample(cancer$mortality, 25, replace = F)
x_sample <- sample(cancer$pop, 25, replace = F)
y_bar <- mean(y_sample)
x_bar <- mean(x_sample)
R <- y_bar/x_bar #ratio estimator

y_bar_R <- R*mean(cancer$pop)
y_bar_R

## [1] 62.83246

T_R <- R*sum(cancer$pop)
T_R

## [1] 18912.57

# These estimates are comparable to the values formulated
# in part d.

# k. confidence intervals</pre>
```

```
- 2*R*cov(x_sample,y_sample))
lower <- y_bar_R - (1.96*sqrt(y_bar_R))</pre>
upper <- y_bar_R + (1.96*sqrt(y_bar_R))</pre>
c(lower,upper) # 95% CI for Ybar_R
## [1] 47.29614 78.36877
# I am 95% confident that the average number of cancer mortalities in each county
# falls within the above values.
var_T_R <- (301^2)*var_ybar_R</pre>
lower \leftarrow T_R - (1.96*sqrt(var_T_R))
upper \leftarrow T_R + (1.96*sqrt(var_T_R))
c(lower,upper) # 95% CI for T_R
## [1] 10949.25 26875.89
#I am 95% confident that the total cancer mortality for all counties
#in the three states is between the above values.
Parts l, m, n
# l. four strata
cancer$group <- ntile(cancer$pop, n=4)</pre>
n1 <- mean(sample(cancer$mortality[cancer$group==1], size = 6))</pre>
n2 <- mean(sample(cancer$mortality[cancer$group==2], size = 6))</pre>
n3 <- mean(sample(cancer$mortality[cancer$group==3], size = 6))</pre>
n4 <- mean(sample(cancer$mortality[cancer$group==4], size = 6))</pre>
W < -75/301
xbar s \leftarrow W*(n1+n2+n3+n4)
xbar_s #pop mean mortality estimate
## [1] 42.9402
T s \leftarrow 301*xbar s
T_s #pop total mortality estimate
## [1] 12925
# m.
#sampling fraction for proportional allocation is simply 1/4th for each strata,
#since each strata is equally sized.
#sampling fraction for optimal allocation:
denom <-
```

 $var_ybar_R \leftarrow (1/25)*(1-((24)/300))*(((R^2)*var(x_sample)) + var(y_sample))$

W*var(cancer\$mortality[cancer\$group==1]) + W*var(cancer\$mortality[cancer\$group==2]) +

```
W*var(cancer$mortality[cancer$group==3]) + W*var(cancer$mortality[cancer$group==4])

n_1 <- var(cancer$mortality[cancer$group==1])/denom

n_2 <- var(cancer$mortality[cancer$group==2])/denom

n_3 <- var(cancer$mortality[cancer$group==3])/denom

n_4 <- var(cancer$mortality[cancer$group==4])/denom

c(W*n_1, W*n_2, W*n_3, W*n_4) #optimal allocation sampling fraction</pre>
```

[1] 0.003474048 0.009457559 0.032221686 0.954846707

[1] 2.325144

Here we conclude that the variance of the mean proportional allocation is 232.5144143 percent of the variance under optimal allocation. Therefore, optimal allocation is clearly preferred.