

# Øksendal: Stochastic Differential Equations

Solutions Manual  
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# **Introduction**

This is a solutions manual for Stochastic Differential Equations, 6<sup>th</sup> edition, by Bernt Øksendal.

It was last updated in January 2026. Errors in the solutions to §3.1 and §3.2 have been corrected.  
The following problems have been solved to date:

- Chapter 2: Problems #1-17
- Chapter 3: Problems #1-17
- Chapter 4: Problems #1-15
- Chapter 5: Problems #1-17
- Chapter 7: Problems #1-11.

## Some Mathematical Preliminaries (§2)

2.1 Suppose  $X : \Omega \rightarrow \mathbb{R}$  is a function that assumes countably many values  $\{a_j\}$  in  $\mathbb{R}$ .

- (a) Note that  $X$  is a random variable if and only if it is measurable. If  $X : \Omega \rightarrow \mathbb{R}$  is measurable, then  $U = X^{-1}(\mathbb{R} \setminus a_k) \in \mathcal{F}$  and thus  $X^{-1}(a_k) = \Omega \setminus U \in \mathcal{F}, \forall k$ . On the other hand, if  $X^{-1}(a_k) \in \mathcal{F}, \forall k$ , then Borel set  $V \subseteq \mathbb{R}$ ,  $X^{-1}(V) = \bigcup_{a_k \in V} X^{-1}(a_k) \in \mathcal{F}$  and thus  $X$  is measurable.
- (b) Compute  $\mathbb{E}(|X|) = \int_{\mathbb{R}} |x| d\mathbb{P}_X = \int_{\bigcup_{k=1}^{\infty} \{a_k\}} |x| d\mathbb{P}_X = \sum_{k=1}^{\infty} |a_k| \mathbb{P}(X = a_k)$ .
- (c) If  $\mathbb{E}(|X|) < \infty$ , then the series

$$\mathbb{E}(X) = \int_{\mathbb{R}} x d\mathbb{P}_X = \int_{\bigcup_{k=1}^{\infty} \{a_k\}} x d\mathbb{P}_X = \sum_{k=1}^{\infty} a_k \mathbb{P}(X = a_k)$$

is absolutely convergent and therefore converges.

- (d) If  $f$  is measurable and  $|f|$  is bounded by  $M$ , then

$$\mathbb{E}(|f(X)|) = \int_{\mathbb{R}} |f(x)| d\mathbb{P}_X \leq \int_{\mathbb{R}} M d\mathbb{P}_X = M \int_{\mathbb{R}} d\mathbb{P}_X = M < \infty.$$

Hence,

$$\mathbb{E}(f(X)) = \int_{\mathbb{R}} f(x) d\mathbb{P}_X = \int_{\bigcup_{k=1}^{\infty} \{a_k\}} f(x) d\mathbb{P}_X = \sum_{k=1}^{\infty} f(a_k) \mathbb{P}(X = a_k)$$

is absolutely convergent and therefore converges.

2.2 Let  $F(x) = \mathbb{P}(X \leq x)$  be the distribution function of  $X$ .

- (a) By monotonicity of  $\mathbb{P}$ ,  $0 = \mathbb{P}(\emptyset) \leq \mathbb{P}(X \leq x) \leq P(\mathbb{R}) = 1$ . Now, by the Monotone Convergence Theorem,

$$\lim_{n \rightarrow \infty} F(n) = \lim_{n \rightarrow \infty} \int_{\mathbb{R}} \chi_{(-\infty, n]} d\mathbb{P}(x) = \int_{\mathbb{R}} d\mathbb{P}(x) = 1.$$

Similarly, for  $G(n) := 1 - F(-n)$ , we have

$$\lim_{n \rightarrow \infty} G(n) = \lim_{n \rightarrow \infty} \int_{\mathbb{R}} (1 - \chi_{(-\infty, -n]}) d\mathbb{P}(x) = 1.$$

Moreover,  $F$  is increasing by monotonicity of  $P$  and finally, again by Monotone Convergence,

$$\lim_{h \rightarrow 0^+} 1 - F(x + h) + F(x) = \lim_{h \rightarrow 0^+} \int_{\mathbb{R}} (1 - \chi_{(x, x+h]}) d\mathbb{P}(x) = \int_{\mathbb{R}} d\mathbb{P}(x) = 1$$

and so  $\lim_{h \rightarrow 0^+} F(x + h) = F(x)$ , i.e.  $F$  is right-continuous.

(b) Compute the expectation

$$\mathbb{E}(g(X)) = \int_{\mathbb{R}} g(x) d\mathbb{P}(x) = \int_{\mathbb{R}} g(x) \chi_{(-\infty, x]} d\mathbb{P}(x) = \int_{\mathbb{R}} g(x) dF(x).$$

(c) Compute the density of  $B_t^2$

$$\begin{aligned} F(u) &:= \mathbb{P}(B_t^2 \leq u) = \mathbb{P}(-\sqrt{u} \leq B_t \leq \sqrt{u}) \\ &= 2 \int_{[0, \sqrt{u}]} p(y) dy \\ &= 2 \int_{[0, u]} \frac{p(\sqrt{u})}{2\sqrt{u}} du \\ &= \int_{(-\infty, u]} \chi_{[0, \infty)} \frac{p(\sqrt{u})}{\sqrt{u}} du. \end{aligned}$$

and so  $p(u) = \chi_{[0, \infty)} \frac{p(\sqrt{u})}{\sqrt{u}}$  where  $p(u)$  is the density of  $B_t$ .

2.3 Since  $\mathcal{H}_i$  is a  $\sigma$ -algebra,  $\emptyset \in \mathcal{H}_i$ ,  $\forall i \in I$ . So  $\emptyset \in \mathcal{H} = \cap_{i \in I} \mathcal{H}_i$ . If  $\{U_j\}_{j \in \mathbb{N}} \subseteq \mathcal{H}$ , then  $\{U_j\}_{j \in \mathbb{N}} \subseteq \mathcal{H}_i$  for each  $i \in I$  and so  $\Omega \setminus U_j \in \mathcal{H}_i$  and  $\cup_{j \in \mathbb{N}} U_j \in \mathcal{H}_i$ ,  $\forall i \in I$ . Conclude that  $\Omega \setminus \cup_{j \in \mathbb{N}} U_j \in \mathcal{H}$  and  $\cup_{j \in \mathbb{N}} U_j \in \mathcal{H}$  and  $\mathcal{H} = \cap_{i \in I} \mathcal{H}_i$  is also a  $\sigma$ -algebra.

2.4 Let  $X : \Omega \mapsto \mathbb{R}$  be a random variable with  $\mathbb{E}(|X|^p) < \infty$ .

(a) Let  $A = \{\omega \in \Omega \mid |X| \geq \lambda > 0\}$  and compute

$$\mathbb{E}(|X|^p) = \int_{\Omega} |X|^p d\mathbb{P} \geq \int_A |X|^p d\mathbb{P} \geq \lambda^p \int_A d\mathbb{P} = \lambda^p \mathbb{P}(|X| \geq \lambda).$$

(b) By Chebychev,  $\mathbb{P}(|X| \geq \lambda) = \mathbb{P}(e^{|X|} \geq e^\lambda) \leq \frac{1}{e^{k\lambda}} \mathbb{E}(e^{|X|}) = M e^{-k\lambda}$ .

2.5 Since the measures are  $\sigma$ -finite,  $f(x, y) = xy$  is  $\mathbb{P}_X \otimes \mathbb{P}_Y$  measurable and  $\mathbb{E}(|XY|) < \infty$ , apply Fubini-Tonelli and compute

$$\begin{aligned} \mathbb{E}(XY) &= \int_{\mathbb{R}^2} xy d\mathbb{P}_{XY}(x, y) \\ &= \int_{\mathbb{R}^2} xy d\mathbb{P}_X(x) \otimes d\mathbb{P}_Y(y) \\ &= \int_{\mathbb{R}} y \left( \int_{\mathbb{R}} x d\mathbb{P}_X(x) \right) d\mathbb{P}_Y(y) \\ &= \mathbb{E}(X) \int_{\mathbb{R}} y d\mathbb{P}_Y(y) \\ &= \mathbb{E}(X)\mathbb{E}(Y). \end{aligned}$$

2.6 (Borel-Cantelli) Let  $\{A_k\}_{k=1}^{\infty} \subseteq \mathcal{F}$  and suppose  $\sum_{k=1}^{\infty} \mathbb{P}(A_k) < \infty$ . Then

$$\mathbb{P}(\bigcap_{m=1}^{\infty} \bigcup_{k=m}^{\infty} A_k) \leq \lim_{m \rightarrow \infty} \sup_{k \geq m} \mathbb{P}(A_k) = 0$$

by dominated convergence.

2.7 Let  $\Omega = \bigsqcup_{i=1}^n G_i$ .

- (a) Note  $\emptyset \in \mathcal{G}$  and  $\mathcal{G}$  is closed under unions by construction. It is also closed under complements as  $\Omega \setminus G_i = \bigcup_{j \neq i} G_j \in \mathcal{G}$ .
- (b) Write a new sequence defined by  $F_i = G_i \setminus \bigcup_{j \leq i} F_j$  and  $\{F_i\}$  will satisfy (a).
- (c) Note that  $\{X^{-1}(x \in \mathbb{R})\} \subseteq \mathcal{F}$  is disjoint. So, by (a) and (b),  $\mathcal{F}$  is finite if and only if all but finitely many  $X^{-1}(x \in \mathbb{R})$  are empty.

2.8 Let  $B_t$  be a 1-dimensional Wiener process.

- (a) By Equation 2.2.3, since  $B_t \sim N(0, t)$ ,

$$\mathbb{E}(e^{iuB_t}) = \exp\left(-\frac{u^2}{2}\mathbb{V}(B_t) + iu\mathbb{E}(B_t)\right) = e^{-\frac{u^2}{2}}.$$

- (b) Comparing power series coefficients, we deduce that

$$\frac{(iu)^{2n}}{(2n)!} \mathbb{E}(B_t^{2n}) = \frac{1}{n!} \left(-\frac{u^2 t}{2}\right)^n,$$

and so  $\mathbb{E}(B_t^{2n}) = \frac{(2n)!}{2^n n!} t^n$ .

- (c) Integrating by parts, compute the  $n^{\text{th}}$  moment of  $B_t$

$$\begin{aligned} \mathbb{E}(B_t^{2k}) &= \frac{1}{\sqrt{2\pi t}} \int_{\mathbb{R}} x^{2k} e^{-\frac{x^2}{2t}} dx \\ &= x^{2k-1} \sqrt{\frac{2t}{\pi}} \int_{\mathbb{R}} ue^{-u^2} du \Big|_{x=-\infty}^{x=\infty} - \int_{\mathbb{R}} (2k-1)x^{2k-2} \sqrt{\frac{2t}{\pi}} \int_{\mathbb{R}} ue^{-u^2} du \\ &= -(2k-1) \sqrt{\frac{2t}{\pi}} \int_{\mathbb{R}} x^{2k-2} \left(\frac{-1}{2} e^{-\frac{x^2}{2t}}\right) dx \\ &= (2k-1)t \cdot \frac{1}{\sqrt{2\pi t}} \int_{\mathbb{R}} x^{2k-2} e^{-\frac{x^2}{2t}} dx \\ &= (2k-1)t \mathbb{E}(B_t^{2k-2}). \end{aligned}$$

As  $\mathbb{E}(B_t^2) = t$ , we have that  $\mathbb{E}(B_t^{2k}) = \frac{(2k)!t^{k-1}}{2^k k!} \cdot t = \frac{(2k)!t^k}{2^k k!}$ .

- (d) Check the base case,  $n = 2k = 2$ , where  $\mathbb{E}(B_t^2)] = \frac{2! \cdot t}{2 \cdot 1!} = t$ . If the claim is true for  $n = 2k$ , then

$$\mathbb{E}(B_t^{2k+2}) = (2k-1)t\mathbb{E}(B_t^{2k}) = (2k+1)t \cdot \frac{(2k)!t^k}{2^k k!} = \frac{(2k+2)!t^{k+1}}{2^{k+1}(k+1)!},$$

and so it is also true for  $n = 2(k+1) = 2k+2$ , thus completing the induction step.

2.9 Note that  $\{X_t\}$  and  $\{Y_t\}$  have the same distributions since neither distribution has any atoms and they agree except on a zero set  $\forall t \geq 0$ . Yet  $t \mapsto X_t$  is discontinuous while  $t \mapsto Y_t$  is continuous.

2.10 As  $B_t$  is Brownian,  $B_{t+h} - B_t \sim N(0, h)$ . Since  $h$  is fixed,  $\{B_{t+h} - B_t\}_{h \geq 0}$  have the same distributions  $\forall t \geq 0$ .

2.11 As  $B_0 = (B_0^{(1)}, B_0^{(2)}, \dots, B_0^{(n)}) = 0$ ,  $B_0^{(j)} = 0$  for all  $j \in \{1, \dots, n\}$ .  $B_t$  is almost surely continuous only if its components are almost surely continuous. Each component is normally distributed with  $\mathbb{E}(B_t^{(j)}) = 0$  as  $\mathbb{E}(B_t) = \vec{0}$  and  $\text{Cov}(B_t^{(i)}, B_t^{(j)}) = t\delta_{ij}$  as  $\text{Cov}(B_t) = tI$ .

2.12 Let  $W_t := B_{s+t} - B_s$  where  $s \geq 0$  is fixed. Then  $W_0 = B_s - B_s = 0$  and  $W_t$  is almost surely continuous as the sum of two almost surely continuous stochastic processes. Noting  $W_{t_2} - W_{t_1} = B_{s+t_2} - B_{s+t_1}$  is independent of both  $B_{s+t_1}$  and  $B_s$ , deduce that  $W_{t_2} - W_{t_1}$  is independent of  $W_{t_1} = B_{s+t_1} - B_s$ . The expected value is  $\mathbb{E}(W_t) = \mathbb{E}(B_{s+t}) - \mathbb{E}(B_s) = 0$  and the variance is

$$\begin{aligned}\mathbb{V}(W_t) &= \mathbb{E}((B_{s+t} - B_s)^2) \\ &= \mathbb{E}(B_{s+t}^2) - 2\mathbb{E}(B_s B_{s+t}) + \mathbb{E}(B_s^2) \\ &= \mathbb{E}(B_{s+t}^2) - 2\mathbb{E}(B_s(B_{s+t} - B_s)) - \mathbb{E}(B_s^2) \\ &= \mathbb{E}(B_{s+t}^2) - 2\mathbb{E}(B_s)\mathbb{E}(B_{s+t} - B_s) - \mathbb{E}(B_s^2) \\ &= (s+t) - 0 - s \\ &= t.\end{aligned}$$

Since  $W_t$  is the sum of two normal distributions, it is also normal and  $W_t \sim N(0, t)$ .

2.13 Compute

$$\mathbb{P}_0(B_t \in D_\rho) = \int_{|\vec{x}|<\rho} \frac{1}{2\pi t} e^{-\frac{|\vec{x}|^2}{2t}} d^2 \vec{x} = \frac{2\pi}{2\pi t} \int_0^\rho r e^{-\frac{r^2}{2t}} dr = \int_0^{\frac{\rho^2}{2t}} e^{-u} du = 1 - e^{-\frac{\rho^2}{2t}}.$$

## 2.14 Compute

$$\begin{aligned}
\mathbb{E}_x \left( \int_{[0,\infty]} \chi_K(B_t) dt \right) &= \int_{[0,\infty]} \mathbb{P}(B_t \in K) dt \\
&= \int_{[0,\infty]} \left( \int_K \frac{1}{(2\pi t)^{n/2}} e^{-\frac{|\vec{x}-\vec{y}|^2}{2t}} d^n \vec{x} \right) dt \\
&\leq \int_{[0,\infty]} \left\| \frac{1}{(2\pi t)^{n/2}} e^{-\frac{|\vec{x}-\vec{y}|^2}{2t}} \right\|_\infty \mu(K) dt \\
&= 0
\end{aligned}$$

and deduce that the expected total time spent in  $K$  is 0.

2.15 Note that  $UU^T = I$ , whence  $|\det U| = 1$  and the probability measures are identical by change of variables. It follows that both are Brownian.

2.16 Let  $W_t = \frac{1}{c} B_{c^2 t}$ . We have  $W_0 = B_0 = 0$  and that  $W_t$  is absolutely continuous as a scaling of absolutely continuous  $B_t$ . Finally,

$$\begin{aligned}
\mathbb{P}_0(W_t \in U) &= \mathbb{P}_0(B_{c^2 t} \in cU) \\
&= \int_{cU} p(c^2 t, 0, y) dy \\
&= \int_{cU} \frac{1}{c} p(t, 0, y/c) dy \\
&= \int_U \frac{1}{c} p(t, 0, y') (cdy') \\
&= \mathbb{P}_0(B_t \in U),
\end{aligned}$$

and so  $W_t$  is also a Brownian motion.

2.17 Let  $X_t(\cdot)$  be a continuous stochastic process.

(a) Recall that  $\mathbb{E}(B_t) = 0$ ,  $\mathbb{E}(B_t^2) = t$  and  $\mathbb{E}(B_t^4) = 3t^2$ . Then

$$\begin{aligned}
\mathbb{E} \left( \left( \sum_k (\Delta B_k^2 - \Delta t_k) \right)^2 \right) &= \mathbb{E} \left( \left( \sum_k (\Delta B_k^2 - \Delta t_k)^2 \right) \right) \\
&= \sum_k \left( \mathbb{E}(\Delta B_k^4) - 2\Delta t_k \mathbb{E}(\Delta B_k^2) + \Delta t_k^2 \right) \\
&= \sum_k \left( 3\Delta t_k^2 - 2\Delta t_k^2 + \Delta t_k^2 \right) \\
&= 2 \sum_k \Delta t_k^2.
\end{aligned}$$

So  $\langle B, B \rangle_t^{(2)}(w) = t$ .

(b) Note that the Brownian motion has positive quadratic variation  $t$  on  $[0, t]$ . So

$$\langle B, B \rangle_t^{(1)}(w) \geq \lim_{\|\Delta B_k\| \rightarrow 0^+} \frac{\langle B, B \rangle_t^{(2)}(w)}{\|\Delta B_k\|} = \infty.$$

## Itô Integrals (§3)

3.1 Using a dyadic mesh, compute

$$\begin{aligned}
\int_0^t s dB_s &= \lim_{n \rightarrow \infty} \left( \frac{\lfloor 2^n t \rfloor}{2^n} \left( B_t - B_{\lfloor 2^n t \rfloor} \right) + \sum_{j=0}^{\lfloor 2^n t \rfloor - 1} \frac{j}{2^n} \left( B_{\frac{j+1}{2^n}} - B_{\frac{j}{2^n}} \right) \right) \\
&= \lim_{n \rightarrow \infty} \frac{\lfloor 2^n t \rfloor}{2^n} (B_t - B_{\lfloor 2^n t \rfloor}) + \lim_{n \rightarrow \infty} \frac{\lfloor 2^n t \rfloor - 1}{2^n} B_{\lfloor 2^n t \rfloor} - \lim_{n \rightarrow \infty} \frac{1}{2^n} \sum_{j=0}^{\lfloor 2^n t \rfloor - 1} B_{\frac{j}{2^n}} \\
&= t B_t - \int_0^t B_s ds,
\end{aligned}$$

where the first term vanishes as  $B_t - B_{\lfloor 2^n t \rfloor} \stackrel{d}{\sim} \mathcal{N}(0, t - 2^{-n} \lfloor 2^n t \rfloor) \xrightarrow{d} 0$ .

3.2 Using the same mesh from §3.1, compute

$$\begin{aligned}
\int_0^t B_s^2 dB_s &= \lim_{n \rightarrow \infty} \sum_{j=0}^{\lfloor 2^n t \rfloor - 1} B_{j \cdot 2^{-n}}^2 (B_{(j+1) \cdot 2^{-n}} - B_{j \cdot 2^{-n}}) \\
&= \lim_{n \rightarrow \infty} \sum_{j=0}^{\lfloor 2^n t \rfloor - 1} \left( \frac{1}{3} B_{(j+1) \cdot 2^{-n}}^3 - \frac{1}{3} B_{j \cdot 2^{-n}}^3 - B_{j \cdot 2^{-n}} (B_{(j+1) \cdot 2^{-n}} - B_{j \cdot 2^{-n}})^2 - \frac{1}{3} (B_{(j+1) \cdot 2^{-n}} - B_{j \cdot 2^{-n}})^3 \right) \\
&= \lim_{n \rightarrow \infty} \left( \frac{1}{3} B_{\lfloor 2^n t \rfloor \cdot 2^{-n}} - \left[ \sum_{j=0}^{\lfloor 2^n t \rfloor - 1} 2^{-n} B_{j \cdot 2^{-n}} \right] + \mathcal{O}(2^{-n/2}) \right) \\
&= \frac{1}{3} B_t^3 - \int_0^t B_s ds.
\end{aligned}$$

3.3 Let  $\{\mathcal{N}_t\}$  be some filtration and let  $\{\mathcal{H}_t^{(X)}\}$  be the filtration of process  $X_t$ .

(a) Compute

$$\mathbb{E}(X_t | \mathcal{H}_s^{(X)}) = \mathbb{E}(\mathbb{E}(X_t | \mathcal{N}_s) | \mathcal{H}_s^{(X)}) = \mathbb{E}(H_s | \mathcal{H}_s^{(X)}) = H_s.$$

(b) Compute

$$\mathbb{E}(X_t) = \mathbb{E}(\mathbb{E}(X_t | H_0^{(X)})) = \mathbb{E}(X_0).$$

(c) Let  $Y \sim \text{Bernoulli}(0.5)$  and fix  $X_0 = 2Y - 1$ . Then  $X_t = t \cdot \text{sgn}(X_0)$  satisfies  $\mathbb{E}(X_t) = \mathbb{E}(X_0) = 0$ , but  $\mathbb{E}(X_t | \mathcal{F}_s) = t \cdot \text{sgn}(X_0) \neq s \cdot \text{sgn}(X_0)$ .

3.4 Compute

$$\begin{aligned}\mathbb{E}(B_t + 4t | \mathcal{F}_s) &= B_s + 4t \neq B_s + 4s \\ \mathbb{E}(B_t^2 | \mathcal{F}_s) &= \mathbb{E}((B_t - B_s)^2 + 2B_s(B_t - B_s) + B_s^2 | \mathcal{F}_s) = B_s^2 + t - s \neq B_s^2 \\ \mathbb{E}\left(t^2 B_t - 2 \int_0^t u B_u du | \mathcal{F}_s\right) &= t^2 B_s - 2 \int_0^s u B_u du - 2 \int_s^t u B_s du = s^2 B_s - 2 \int_0^s u B_u du \\ \mathbb{E}(B_t^{(1)} B_t^{(2)} | \mathcal{F}_s) &= \mathbb{E}(B_t^{(1)} | \mathcal{F}_s) \mathbb{E}(B_t^{(2)} | \mathcal{F}_s) = B_s^{(1)} B_s^{(2)},\end{aligned}$$

and deduce that only the last two are martingales.

3.5 Verify  $\mathbb{E}(|B_t^2 - t|) \leq \mathbb{E}(B_t^2) + t = 2t < \infty$  and compute

$$\mathbb{E}(B_t^2 - t | \mathcal{F}_s) = \mathbb{E}((B_t - B_s)^2 + 2B_s(B_t - B_s) + B_s^2 - t | \mathcal{F}_s) = B_s^2 + t - s - t = B_s^2 - s.$$

to deduce that  $X_t := B_t^2 - t$  is a martingale.

3.6 Verify  $\mathbb{E}(|B_t^3 - 3tB_t|) \leq \sqrt{\mathbb{E}(B_t^2)}(\sqrt{\mathbb{E}(B_t^4)} + 3t) = (3 + \sqrt{3})t^{3/2} < \infty$  and compute

$$\begin{aligned}\mathbb{E}(B_t^3 - 3tB_t | \mathcal{F}_s) &= \mathbb{E}((B_t - B_s)^3 + 3B_s(B_t - B_s)^2 + 3B_s^2(B_t - B_s) + B_s^3 - 3tB_s | \mathcal{F}_s) \\ &= 3B_s(t - s) + B_s^3 - 3tB_s \\ &= B_s^3 - 3sB_s\end{aligned}$$

to deduce that  $Y_t := B_t^3 - 3tB_t$  is a martingale.

3.7 In this question, the formula for Itô iterated integrals is derived.

(a) Note that  $\{0 \leq u_1 \dots \leq u_n\}$  is Borel measurable and  $\chi_{0 \leq u_1 \dots \leq u_n}$  is  $\mathcal{F}_t$ -adapted. Finally  $\mathbb{E}\left(\int_0^T f(t_1, \dots, t_n, \omega)^2 dt_1 \dots dt_n\right) \leq T^n < \infty$ .

(b) For  $n \in \{1, 2, 3\}$

$$\begin{aligned}1! \int_0^t dB_u &= B_t = t^{1/2} H_1\left(\frac{B_t}{\sqrt{t}}\right) \\ 2! \int_0^t \int_0^v dB_u dB_v &= 2 \int_0^t B_v dB_v = B_t^2 - t = t H_2\left(\frac{B_t}{\sqrt{t}}\right) \\ 3! \int_0^t \int_0^w \int_0^v dB_u dB_v dB_w &= 3 \int_0^t (B_w^2 - w) dB_w = B_t^3 - 3tB_t = t^{3/2} H_3\left(\frac{B_t}{\sqrt{t}}\right).\end{aligned}$$

(c) Deduce that  $d(B_t^3 - 3tB_t) = 3(B_t^2 - t) dB_t$  and so  $Y_t := B_t^3 - 3tB_t$  is a martingale.

3.8 There exists continuous martingale  $M_t$  iff there exists  $Y \in L^1$  such that  $M_t = \mathbb{E}(Y | \mathcal{F}_t)$ .

(a) Verify that  $\mathbb{E}(\mathbb{E}(Y | \mathcal{F}_t)) \leq \mathbb{E}(\mathbb{E}(|Y| | \mathcal{F}_t)) = \mathbb{E}(|Y|) < \infty$  and

$$\mathbb{E}(M_t | \mathcal{F}_s) = \mathbb{E}(\mathbb{E}(Y | \mathcal{F}_t) | \mathcal{F}_s) = \mathbb{E}(Y | \mathcal{F}_s) = M_s.$$

(b) If  $M_t$  is a continuous martingale such that  $\sup_{t>0} \mathbb{E}(|X|^p) < \infty$  for  $p \in (1, \infty)$ , then  $\exists M$  such that  $\|M_t - M\|_{L^1} \rightarrow 0$  as  $t \rightarrow \infty$ . So let  $Y = M$  and

$$\begin{aligned} \lim_{s \rightarrow \infty} \int_{\Omega_s} |M_s - \mathbb{E}(M | \mathcal{F}_s)| d\mathbb{P} &= \lim_{s \rightarrow \infty} \int_{\Omega_s} |\mathbb{E}(M_s - M | \mathcal{F}_s)| d\mathbb{P} \\ &\leq \lim_{s \rightarrow \infty} \int_{\Omega_s} \mathbb{E}(|M_s - M| | \mathcal{F}_s) d\mathbb{P} \\ &= \lim_{s \rightarrow \infty} \int_{\Omega_s} |M_s - M| d\mathbb{P} \\ &= 0. \end{aligned}$$

3.9 Compute

$$\begin{aligned} \int_0^T B_t \circ dB_t &= \lim_{n \rightarrow \infty} \sum_{j=0}^{\lfloor nt \rfloor - 1} \frac{1}{2} (B_{\frac{jt}{n}} + B_{\frac{(j+1)t}{n}})(B_{\frac{(j+1)t}{n}} - B_{\frac{jt}{n}}) \\ &= \lim_{n \rightarrow \infty} \sum_{j=0}^{\lfloor nt \rfloor - 1} B_{\frac{jt}{n}} (B_{\frac{(j+1)t}{n}} - B_{\frac{jt}{n}}) + \lim_{n \rightarrow \infty} \sum_{j=0}^{\lfloor nt \rfloor - 1} \frac{1}{2} (B_{\frac{(j+1)t}{n}} - B_{\frac{jt}{n}})^2 \\ &= \frac{1}{2} B_t^2 - \frac{t}{2} + \frac{t}{2} \\ &= \frac{1}{2} B_t^2. \end{aligned}$$

3.10 If  $f(t, \omega)$  varies smoothly in  $t$ , then the Itô and Stratonovich integrals coincide. Compute

$$\int_0^T f(t, \omega) \circ dB_t = \int_0^T f(t, \omega) dB_t + \frac{1}{2} \langle f(t, \omega), B_t \rangle^{(2)}$$

and

$$\begin{aligned} \mathbb{E}(\langle f(t, \omega), B_t \rangle^{(2)})^2 &\leq \mathbb{E}(\langle B_t, B_t \rangle^{(2)}) \mathbb{E}(\langle f(t, \omega), f(t, \omega) \rangle^{(2)}) \\ &\leq T \lim_{\|\Delta t_k\| \rightarrow 0^+} \sup_{|\Delta t_k|} \frac{T}{|\Delta t_k|} (K |\Delta t_k|^{1+\varepsilon}) \\ &= KT^2 \lim_{\|\Delta t_k\| \rightarrow 0^+} \|\Delta t_k\|^\varepsilon \\ &= 0. \end{aligned}$$

3.11 Define white noise  $W_t^{(N)} = \max\{-N, \min\{W_t, N\}\}$ . Since  $W_t$  and  $W_s$  are independent and identically distributed, it follows that  $W_t^{(N)}$  and  $W_s^{(N)}$  are as well. If  $W_t$  is continuous, then since  $|W_t^{(N)}| \leq N$  and by bounded convergence

$$\lim_{t \rightarrow s} 2\mathbb{E}(W_t^{(N)})^2 = \lim_{t \rightarrow s} \mathbb{E}(|W_t^{(N)} - W_s^{(N)}|^2) = 0.$$

But then  $W_t \xrightarrow{\text{a.s.}} 0$ , which is a contradiction.

3.12 Let  $\circ dB_t$  denote the Stratonovich differential.

(i) Since  $\alpha X_t \circ dB_t = \frac{\alpha^2}{2} X_t dt + \alpha X_t dB_t$ ,

$$dX_t = (\gamma + \frac{\alpha^2}{2}) X_t dt + \alpha X_t dB_t.$$

Since  $(t^2 + \cos(X_t)) \circ dB_t = -\frac{\sin(X_t)}{2} (t^2 + \cos(X_t)) dt + (t^2 + \cos(X_t)) dB_t$ ,

$$dX_t = \frac{\sin(X_t)}{2} (\cos(X_t) - t^2) dt + (t^2 + \cos(X_t)) dB_t.$$

(ii) Since  $\alpha X_t dB_t = \alpha X_t \circ dB_t - \frac{\alpha^2}{2} X_t dt$ ,

$$dX_t = (r - \frac{\alpha^2}{2}) X_t dt + \alpha X_t \circ dB_t.$$

Since  $X_t^2 dB_t = X_t^2 \circ dB_t - X_t^3 dt$ ,

$$dX_t = (2e^{-X_t} - X_t^3) dt + X_t^2 \circ dB_t.$$

3.13 Let  $X_t$  be continuous in mean square. Calculate

$$(a) \lim_{s \rightarrow t} \mathbb{E}[(B_t - B_s)^2] = \lim_{s \rightarrow t} \mathbb{E}[(B_{t-s})^2] = \lim_{s \rightarrow t} (t-s) = 0$$

$$(b) \lim_{s \rightarrow t} \mathbb{E}[(f(B_t) - f(B_s))^2] \leq \lim_{s \rightarrow t} C^2 \mathbb{E}[(B_t - B_s)^2] = 0$$

(c) and finally by Itô isometry,

$$\begin{aligned} \lim_{n \rightarrow \infty} \mathbb{E} \left[ \left( \int_S^T (X_s - \phi_n(s)) dB_s \right)^2 \right] &= \lim_{n \rightarrow \infty} \mathbb{E} \left[ \int_S^T (X_s - \phi_n(s))^2 ds \right] \\ &= \lim_{n \rightarrow \infty} \mathbb{E} \left[ \sum_j \int_{t_j^{(n)}}^{t_{j+1}^{(n)}} (X_t - X_{t_j^{(n)}})^2 dt \right] \\ &\leq (T-S) \lim_{n \rightarrow \infty} \sup_{1 \leq j \leq n} \mathbb{E}[(X_t - X_{t_j^{(n)}})^2] \\ &= 0. \end{aligned}$$

3.14 Show that  $h(\omega)$  is  $\mathcal{F}_t$  measurable if and only if it is the pointwise limit of a sum-product of bounded continuous functions  $g(B_{t_j})$ .

- (a) Assume that  $h$  is bounded since  $\{h_n(\omega) := h(\omega)\mathbb{1}_{\{|h(\omega)| < n\}}\}$  converges pointwise to  $h$ .
- (b) Let  $\mathcal{H}_n$  be the  $\sigma$ -algebra generated by  $B(t_j)$  for  $t_j = \frac{j}{2^n} \leq t$ . Then  $\mathcal{F}_t = \sigma(\cup_n \mathcal{H}_n)$  and so by Corollary (C.9),  $h = \mathbb{E}[h|\mathcal{F}_n] = \lim_{n \rightarrow \infty} \mathbb{E}[h|\mathcal{H}_n]$ .
- (c) By Doob-Dynkin,  $\mathbb{E}[h|\mathcal{H}_n](\omega) = g(B_{t_1}, \dots, B(t_{\lfloor 2^n t \rfloor}))$ . Since  $C(\mathbb{R}^k)$  is dense in  $L^1(\mathbb{R}^k)$  and by Stone-Weierstrass  $P(\mathbb{R}^k)$  is dense in  $C(\mathbb{R}^k)$ , a limiting sequence must exist.

3.15 Suppose  $C + \int_S^T f(t, \omega) dB_t(\omega) = D + \int_S^T g(t, \omega) dB_t(\omega)$ . Then we have that

$$C - D = \mathbb{E}[C - D] = \mathbb{E}\left[\int_S^T g(t, \omega) dB_t(\omega) - \int_S^T f(t, \omega) dB_t(\omega)\right] = 0 \implies C = D,$$

and by Itô isometry,

$$0 = \mathbb{E}\left[\left(\int_S^T g(t, \omega) dB_t(\omega) - \int_S^T f(t, \omega) dB_t(\omega)\right)^2\right] = \int_S^T \mathbb{E}[(g(t, \omega) - f(t, \omega))^2] ds,$$

whence  $g(t, \omega) = f(t, \omega)$  almost surely for  $(t, \omega) \in [S, T] \times \Omega$ .

3.16 By Jensen's inequality,  $\mathbb{E}[\mathbb{E}[X|\mathcal{H}]^2] \leq \mathbb{E}[\mathbb{E}[X^2|\mathcal{H}]] = \mathbb{E}[X^2]$ .

3.17 Let  $\mathcal{G}$  be a finite  $\sigma$ -algebra with partition  $\Omega = \bigsqcup_{i=1}^m G_i$ .

- (a) Note that  $\mathbb{E}[X|\mathcal{G}](\omega) = \sum_{i=1}^m c_i \mathbb{1}_{G_i}(\omega) = c_i$  on  $G_i$ .
- (b) Show that

$$\int_{G_i} \left( \frac{\int_{G_i} X d\mathbb{P}}{\mathbb{P}(G_i)} d\mathbb{P} \right) = \frac{\int_{G_i} X d\mathbb{P}}{\mathbb{P}(G_i)} \int_{G_i} 1 d\mathbb{P} = \int_{G_i} X d\mathbb{P}, \forall i \in \{1, \dots, m\}.$$

- (c) By part (b),  $c_i = \frac{\int_{G_i} X d\mathbb{P}}{\mathbb{P}(G_i)}$ . Show for  $\omega \in G_i$  that

$$\begin{aligned} \mathbb{E}[X|\mathcal{G}](\omega) &= \sum_{i=1}^m \frac{\int_{G_i} X d\mathbb{P}}{\mathbb{P}(G_i)} \mathbb{1}_{G_i}(\omega) \\ &= \frac{\int_{G_i} X d\mathbb{P}}{\mathbb{P}(G_i)} \\ &= \frac{\sum_{k=1}^m a_k \mathbb{P}(X = a_k, \omega \in G_i)}{\mathbb{P}(G_i)} \\ &= \sum_{k=1}^m a_k \mathbb{P}(X = a_k | G_i). \end{aligned}$$

## The Itô Formula (§4)

### 4.1 Compute

- (a)  $dX_t = d(B_t^2) = 2B_t dB_t + d[B, B]_t = 2B_t dB_t + dt$
- (b)  $dX_t = d(2 + t + e^{B_t}) = (1 + \frac{1}{2}e^{B_t}) dt + e^{B_t} dB_t$
- (c)  $dX_t = d((B_t^{(1)})^2 + (B_t^{(2)})^2) = 2B_t^{(1)} dB_t^{(1)} + 2B_t^{(2)} dB_t^{(2)} + 2 dt$
- (d)  $dX_t = d((t_0 + t, B_t)) = (dt, dB_t)$
- (e) and finally

$$\begin{aligned} dX_t &= d((B_t^{(1)} + B_t^{(2)} + B_t^{(3)}, (B_t^{(2)})^2 - B_t^{(1)} B_t^{(3)})) \\ &= (dB_t^{(1)} + dB_t^{(2)} + dB_t^{(3)}, 2B_t^{(2)} dB_t^{(2)} + dt - B_t^{(3)} dB_t^{(1)} - B_t^{(1)} dB_t^{(3)}). \end{aligned}$$

### 4.2 Using Itô's Lemma, differentiate

$$d\left(\frac{1}{3}B_t^3 - \int_0^t B_s ds\right) = B_t^2 dB_t + B_t d[B, B]_t - B_t dt = B_t^2 dB_t$$

and deduce that

$$\int_0^t B_s^2 dB_s = \frac{1}{3}B_t^3 - \int_0^t B_s ds.$$

### 4.3 Let $X_t$ and $Y_t$ be Itô processes. Then, letting $f(t, x, y) = xy$ and by Itô's formula

$$\begin{aligned} d(X_t Y_t) &= f_t(t, X_t, Y_t) dt + f_x(t, X_t, Y_t) dX_t + f_y(t, X_t, Y_t) dY_t \\ &\quad + \frac{1}{2}f_{xx}(t, X_t, Y_t) d[X, X]_t + f_{xy}(t, X_t, Y_t) d[X, Y]_t + \frac{1}{2}f_{yy}(t, X_t, Y_t) d[Y, Y]_t \\ &= Y_t dX_t + X_t dY_t + d[X, Y]_t \end{aligned}$$

and deduce the integration of parts formula

$$\begin{aligned} \int_0^t X_s dY_s &= \int_0^t (d(X_s Y_s) - Y_s dX_s - d[X, Y]_s) \\ &= X_t Y_t - X_0 Y_0 - \int_0^t Y_s dX_s - \int_0^t d[X, Y]_s. \end{aligned}$$

4.4 Let  $Z_t = \exp \left( \int_0^t \langle \theta(s, \omega), dB_s \rangle - \frac{1}{2} |\theta(s, \omega)|^2 ds \right)$ .

(a) Then, letting  $Z_t = e^{Y_t}$  and by Itô's formula,

$$\begin{aligned} dZ_t &= e^{Y_t} dY_t + \frac{1}{2} e^{Y_t} d[Y, Y]_t \\ &= Z_t \left( \langle \theta(t, \omega), dB_t \rangle - \frac{1}{2} |\theta(t, \omega)|^2 dt + \frac{1}{2} \sum_{i,j=1}^n [\theta_i(s, \omega) dB^{(i)}, \theta_j(s, \omega) dB^{(j)}]_s \right) \\ &= Z_t \langle \theta(t, \omega), dB_t \rangle. \end{aligned}$$

(b) It suffices to check that

$$\begin{aligned} [\mathbb{E}(|Z_t|)]^2 &= \left[ \mathbb{E} \left( \left| \int_0^t dZ_s \right| \right) \right]^2 \\ &= \left[ \mathbb{E} \left( \left| \int_0^t Z_s \langle \theta(s, \omega), dB_s \rangle \right| \right) \right]^2 \\ &\leq \mathbb{E} \left( \int_0^t \sum_{i=1}^n |Z_s \theta_i(s, \omega)| dB_s^{(i)} \right)^2 \\ &= \mathbb{E} \left( \sum_{i,j=1}^n \int_0^t |Z_s \theta_i(s, \omega)| |Z_s \theta_j(s, \omega)| d[B^{(i)}, B^{(j)}]_s \right) \\ &= \sum_{i=1}^n \mathbb{E} \left( \int_0^t |Z_s \theta_i(s, \omega)|^2 ds \right) \\ &< \infty. \end{aligned}$$

4.5 Let  $\beta_k(t) = \mathbb{E}(B_t^k)$ . Then, by Itô's lemma,

$$dB_t^k = kB_t^{k-1} dB_t + \frac{1}{2} k(k-1) B_t^{k-2} dt$$

and so

$$\beta_k(t) = \mathbb{E}(B_t^k) = \mathbb{E} \left( \int_0^t dB_s^k \right) = \int_0^t \mathbb{E} \left( \frac{1}{2} k(k-1) B_s^{k-2} \right) ds = \frac{1}{2} k(k-1) \int_0^t \beta_{k-2}(s) ds.$$

Deduce that  $\beta_4(t) = 6 \int_0^t \beta_2(s) ds = 6 \cdot \frac{t^2}{2} = 3t^2$  and  $\beta_6(t) = 15 \int_0^t 3s^2 ds = 15t^3$ .

4.6 Define geometric Brownian motions  $X_t = e^{ct+\alpha B_t}$  and  $Y_t = e^{ct+\sum_{j=1}^n \alpha_j B_t^{(j)}}$ .

(a) Calculate

$$\begin{aligned} dX_t &= ce^{ct+\alpha B_t} dt + \alpha e^{ct+\alpha B_t} dB_t + \frac{1}{2} \alpha^2 e^{ct+\alpha B_t} d[B, B]_t \\ &= X_t \left( \left( c + \frac{\alpha^2}{2} \right) dt + \alpha dB_t \right). \end{aligned}$$

(b) Calculate

$$\begin{aligned} dY_t &= Y_t \left( c dt + \sum_{j=1}^n \alpha_j dB_t^{(j)} + \frac{1}{2} \sum_{i,j=1}^n \alpha_i \alpha_j d[B^{(i)}, B^{(j)}]_t \right) \\ &= Y_t \left( (c + \frac{1}{2} \sum_{j=1}^n \alpha_j^2) dt + \sum_{j=1}^n \alpha_j dB_t^{(j)} \right). \end{aligned}$$

4.7 Let  $X_t$  solve  $dX_t = v(t, \omega) dB_t$ .

(a) Note that  $B_t$  is a martingale while  $B_t^2$  is not.

(b) Define  $M_t = X_t^2 - \int_0^t v(s, \omega)^2 ds$ . Then

$$\begin{aligned} dM_t &= 2X_t dX_t + [dX, dX]_t - v(t, \omega)^2 dt \\ &= 2X_t v(t, \omega) dB_t + (v(t, \omega)^2 - v(t, \omega)^2) dt \\ &= 2X_t v(t, \omega) dB_t. \end{aligned}$$

Moreover,

$$\begin{aligned} \mathbb{E}(|M_t|) &\leq \mathbb{E}(X_t^2) + \mathbb{E}\left(\int_0^t v(s, \omega)^2 ds\right) \\ &= \mathbb{E}\left(\int_0^t v(s, \omega) dB_s\right)^2 + \mathbb{E}\left(\int_0^t v(s, \omega)^2 ds\right) \\ &= 2\mathbb{E}\left(\int_0^t v(s, \omega)^2 ds\right) \\ &< \infty. \end{aligned}$$

4.8 Let  $f(x^{(1)}, \dots, x^{(n)})$  be a function of class  $C^2$ .

(a) By Itô's lemma,

$$\begin{aligned} d(f(B_t)) &= \sum_{i=1}^n \partial_i f(B_t) dB_t^{(i)} + \frac{1}{2} \sum_{i,j=1}^n \partial_{ij}^2 f(B_t) d[B^{(i)}, B^{(j)}]_t \\ &= \langle \nabla f(B_t), dB_t \rangle + \frac{1}{2} \Delta f(B_t) dt \end{aligned}$$

and so

$$f(B_t) - f(B_0) = \int_0^t d(f(B_s)) = \int_0^t \langle \nabla f(B_s), dB_s \rangle + \frac{1}{2} \int_0^t \Delta f(B_s) ds.$$

(b) Assume that  $g$  is of class  $C^1$  everywhere, as well as  $C^2$  and uniformly bounded outside of finitely many points with  $|g''(z)| \leq M$  for  $z \notin \{z_1, \dots, z_k\}$ . Then the set of functions  $\{f\}$  of class  $C^2$  uniformly bounded with  $|f''(z)| \leq M$  are  $C^k$ -dense. So we can extract

a sequence  $\{f_k\}$  such that  $f_k \Rightarrow g$ ,  $f'_k \Rightarrow g'$  as well as  $f''_k \rightarrow g''$  and  $|f''_k| \leq M$  on  $\mathbb{R} \setminus \{z_1, \dots, z_k\}$ . So

$$\begin{aligned} & \lim_{k \rightarrow \infty} \left| (f_k - g)(B_t) + (f_k - g)(0) + \int_0^t (f'_k - g') dB_s + \frac{1}{2} \int_0^t (f''_k - g'') ds \right| \\ & \leq \lim_{k \rightarrow \infty} |(f_k - g)(B_t)| + |(f_k - g)(0)| + t \|f'_k - g'\|_\infty + \frac{1}{2} \int_0^t |f''_k - g''| ds \\ & = 0, \end{aligned}$$

where the last term vanishes by bounded convergence.

4.9 Clearly

$$\int_0^t v \frac{\partial g_n}{\partial x}(s, X_s) \chi_{s \leq \tau_n} dB_s = \int_0^{t \wedge \tau_n} v \frac{\partial g}{\partial x}(s, X_s) dB_s$$

and the result follows by Itô's lemma where  $dX_t = u dt + v dB_t$ . Since  $\mathbb{E}(|X_t|) < \infty$ , it follows that  $\lim_{n \rightarrow \infty} \mathbb{P}(\tau_n > t) = \lim_{n \rightarrow \infty} \mathbb{P}(X_t < n) = 1$  and so the identity holds almost surely.

4.10 (*Tanaka*) In this problem, Tanaka's formula for Brownian motion is derived.

(a) Substitute  $u \equiv 0$  and  $v \equiv 1$  here. Then as  $g''_\varepsilon(x) = \frac{1}{\varepsilon} \chi_{|x|<\varepsilon}(x)$

$$\frac{1}{2} \int_0^t \frac{d^2 g_\varepsilon}{dx^2}(B_s) ds = \frac{1}{2\varepsilon} \int_0^t \chi_{|B_s|<\varepsilon} ds = \frac{1}{2\varepsilon} |\{s \in [0, t] \mid |B_s| < \varepsilon\}|.$$

(b) Differentiate to get

$$\int_0^t g'_\varepsilon(B_s) \chi_{|B_s|<\varepsilon} dB_s = \int_0^t \frac{B_s}{\varepsilon} \chi_{|B_s|<\varepsilon} dB_s,$$

and apply Itô isometry to get

$$\lim_{\varepsilon \rightarrow 0^+} \mathbb{E} \left( \int_0^t \frac{B_s}{\varepsilon} \chi_{|B_s|<\varepsilon} dB_s \right)^2 = \lim_{\varepsilon \rightarrow 0^+} \mathbb{E} \left( \int_0^t \frac{B_s^2}{\varepsilon^2} \chi_{|B_s|<\varepsilon} ds \right) \leq \lim_{\varepsilon \rightarrow 0^+} \int_0^t \mathbb{P}(|B_s| < \varepsilon) ds = 0.$$

(c) As  $\varepsilon \rightarrow 0$  for  $g(x) = x$ ,

$$\begin{aligned} |B_t| &= |B_0| + \lim_{\varepsilon \rightarrow 0^+} \int_0^t \text{sgn}(B_s) \chi_{|B_s| \geq \varepsilon} ds + \lim_{\varepsilon \rightarrow 0^+} \frac{1}{2\varepsilon} |\{s \in [0, t] \mid |B_s| < \varepsilon\}| \\ &= |B_0| + \int_0^t \text{sgn}(B_s) ds + L_t. \end{aligned}$$

4.11 Let  $X_t = e^{t/2} \cos(B_t)$ ,  $Y_t = e^{t/2} \sin(B_t)$  and  $Z_t = (B_t + t)e^{-B_t-t/2}$ . Compute

- (a)  $dX_t = \frac{1}{2} e^{t/2} \cos(B_t) dt - e^{t/2} \sin(B_t) dB_t + \frac{1}{2} (-e^{t/2} \cos(B_t)) d[B, B]_t = -e^{t/2} \sin(B_t) dB_t$
- (b)  $dY_t = \frac{1}{2} e^{t/2} \sin(B_t) dt + e^{t/2} \cos(B_t) dB_t + \frac{1}{2} (-e^{t/2} \sin(B_t)) d[B, B]_t = e^{t/2} \cos(B_t) dB_t$

(c) and finally

$$\begin{aligned} dZ_t &= e^{-B_t-t/2}d(B_t+t) + (B_t+t)d(e^{-B_t-t/2}) + d[B_t+t, e^{-B_t-t/2}] \\ &= e^{-B_t-t/2}(dt + dB_t) - \frac{1}{2}X_t dt - X_t dB_t - e^{-B_t-t/2} dt + \frac{1}{2}(B_t+t)e^{-B_t-t/2} dt \\ &= e^{-B_t-t/2}(1-t-B_t) dB_t. \end{aligned}$$

4.12 The given condition implies  $\mathbb{E}(|X_t|) < \infty$ . So  $X_t$  is a martingale if and only if  $\mathbb{E}(X_t | \mathcal{F}_s) = X_s$ . Then

$$\mathbb{E}\left(\int_s^t u(r, \omega) dr | \mathcal{F}_s\right) = \mathbb{E}(X_t - X_s | \mathcal{F}_s) = 0.$$

Moreover by dominated convergence

$$\mathbb{E}(u(t, \omega) dr | \mathcal{F}_s) = \mathbb{E}\left(\frac{d}{ds} \int_s^t u(r, \omega) dr | \mathcal{F}_s\right) = 0.$$

Then

$$u(t, \omega) = \mathbb{E}(u(t, \omega) | \mathcal{F}_t) = \lim_{s \rightarrow t^-} \mathbb{E}(u(t, \omega) | \mathcal{F}_s) = 0.$$

4.13 Let  $dX_t = u(t, \omega) dt + dB_t$  where  $u(t, \omega) \in \mathcal{V}([0, T])$ . Then  $Y_t = X_t M_t$  is a martingale, where

$$M_t = \exp\left(-\int_0^t u(r, \omega) dB_r - \frac{1}{2} \int_0^t u^2(r, \omega) dr\right)$$

since  $\mathbb{E}(|M_t|) < \infty$  (see question 4b),  $\mathbb{E}(|X_t|) \leq \sqrt{t} \left( \sqrt{\int_0^t u^2(r, \omega) dr} + 1 \right) < \infty$  and

$$\begin{aligned} d(X_t M_t) &= M_t dX_t + X_t dM_t + d[X, M]_t \\ &= M_t(u(t, \omega) dt + dB_t) + M_t X_t (-u(t, \omega) dB_t - \frac{1}{2}u^2(t, \omega) dt) \\ &\quad - M_t u(t, \omega) dt + \frac{1}{2} M_t X_t u^2(t, \omega) dt \\ &= M_t(1 - u(t, \omega) X_t) dB_t. \end{aligned}$$

4.14 In this problem, the martingale representation of stochastic processes is explicitly shown.

(a) Compute  $dF_t = dB_t$ ,  $\mathbb{E}(F_T) = 0$  and

$$dF_t - d\mathbb{E}(F_t) = 1 dB_t \implies f(t, \omega) = 1.$$

(b) Compute  $dF_t = B_t dt$ ,  $\mathbb{E}(F_T) = 0$  and

$$dF_t - d\mathbb{E}(F_t) = B_t dt = d(TB_T) - t dB_t = (T - t) dB_t \implies f(t, \omega) = T - t.$$

(c) Compute  $dF_t = 2B_t dB_t + dt$ ,  $\mathbb{E}(F_T) = T$  and

$$dF_t - d\mathbb{E}(F_t) = 2B_t dB_t + 1 dt - 1 dt = 2B_t dB_t \implies f(t, \omega) = 2B_t.$$

(d) Compute  $dF_t = 3B_t^2 dB_t + 3B_t dt$ ,  $\mathbb{E}(F_T) = 0$  and

$$\begin{aligned} dF_t - d\mathbb{E}(F_t) &= 3B_t^2 dB_t + 3B_t dt \\ &= 3B_t^2 + 3(T-t)) dB_s \implies f(t, \omega) = 3B_t^2 + 3T - 3t. \end{aligned}$$

(e) Recall that  $e^{B_t-t/2}$  is a martingale and compute

$$d(e^{B_t-t/2}) = e^{B_t-t/2} dB_t.$$

Deduce that

$$e^{B_T} = e^{T/2} \left( 1 + \int_0^T e^{B_t-t/2} dB_t \right) \implies f(t, \omega) = e^{B_t+(T-t)/2}.$$

(f) Find martingale  $e^{t/2} \sin(B_t)$  and compute

$$d(e^{t/2} \sin(B_t)) = e^{t/2} \cos(B_t) dB_t$$

Deduce that

$$\sin(B_T) = e^{-T/2} \int_0^T e^{t/2} \cos(B_t) dB_t \implies f(t, \omega) = e^{-(T-t)/2} \cos(B_t).$$

4.15 Define  $X_t = (x^{1/3} + \frac{1}{3}B_t)^3$ . Then

$$\begin{aligned} dX_t &= 3X_t^{2/3} d(x^{1/3} + \frac{1}{3}B_t) + 3X_t^{1/3} d \left[ x^{1/3} + \frac{1}{3}B_t, x^{1/3} + \frac{1}{3}B_t \right] \\ &= X_t^{2/3} dB_t + \frac{1}{3}X_t^{1/3} dt. \end{aligned}$$

# Stochastic Differential Equations (§5)

## 5.1 Compute

- (a)  $dX_t = d(e^{B_t}) = e^{B_t} dB_t + \frac{1}{2} e^{B_t} d[B, B]_t = \frac{1}{2} X_t dt + X_t dB_t$
- (b)  $dX_t = d\left(\frac{B_t}{1+t}\right) = \frac{1}{1+t} dB_t - \frac{B_t}{(1+t)^2} dt = \frac{1}{1+t} dB_t - \frac{1}{1+t} X_t dt$
- (c)  $dX_t = d(\sin(B_t)) = \cos(B_t) dB_t - \frac{1}{2} \sin(B_t) dt = \cos(B_t) dB_t - \frac{1}{2} X_t dt$
- (d)  $dX_t^{(1)} = dt$  and

$$dX_t^{(2)} = d(e^t B_t) = e^t dB_t + e^t B_t dt = e^t dB_t + X_t^{(2)} dt.$$

- (e) and finally differentials

$$d(\cosh(B_t)) = \sinh(B_t) dB_t + \frac{1}{2} \cosh(B_t) dt$$

and

$$d(\sinh(B_t)) = \cosh(B_t) dB_t + \frac{1}{2} \sinh(B_t) dt$$

to deduce that

$$\begin{pmatrix} dX_t^{(1)} \\ dX_t^{(2)} \end{pmatrix} = \frac{1}{2} \begin{pmatrix} X_t^{(1)} \\ X_t^{(2)} \end{pmatrix} dt + \begin{pmatrix} X_t^{(2)} \\ X_t^{(1)} \end{pmatrix} dB_t.$$

5.2 Let  $X_t^{(1)} = a \cos(B_t)$  and  $X_t^{(2)} = b \sin(B_t)$ . Then

$$dX_t^{(1)} = -a \sin(B_t) dB_t - \frac{a}{2} \cos(B_t) dt = -\frac{1}{2} X_t^{(1)} dt - \frac{a}{b} X_t^{(2)} dB_t$$

and

$$dX_t^{(2)} = b \cos(B_t) dB_t - \frac{b}{2} \sin(B_t) dt = -\frac{1}{2} X_t^{(2)} dt + \frac{b}{a} X_t^{(1)} dB_t.$$

5.3 The solution is given by

$$X_t = X_0 \exp \left( \left( r - \frac{1}{2} \sum_{k=1}^n \alpha_k^2 \right) t + \sum_{k=1}^n \alpha_k dB_k \right).$$

5.4 In this problem, solutions to stochastic differential equations are found.

(a) The solution to  $dX_t^{(1)} = dt + dB_t^{(1)}$  is  $X_t^{(1)} = X_0^{(1)} + t + B_t^{(1)}$  and

$$dX_t^{(2)} = X_t^{(1)} dB_t^{(2)} = (X_0^{(1)} + t + B_t^{(1)}) dB_t^{(2)}$$

is

$$X_t^{(2)} = X_0^{(2)} + X_0^{(1)} B_t^{(2)} + \int_0^t (s + B_s^{(1)}) dB_s^{(2)}.$$

(b) Using integrating factors, solve  $dX_t = X_t dt + dB_t$  for

$$e^{-t} X_t - X_0 = \int_0^t e^{-s} dB_s$$

and deduce that the solution  $X_t$  is

$$X_t = e^t X_0 + \int_0^t e^{t-s} dB_s.$$

(c) Using integrating factors, solve  $dX_t = -X_t dt + e^{-t} dB_t$  for

$$e^t X_t - X_0 = \int_0^t dB_s$$

and deduce that the solution  $X_t$  is

$$X_t = e^{-t} (X_0 + B_t).$$

5.5 The Langevin equation is given by

$$dX_t - \mu X_t dt = \sigma dB_t.$$

(a) Using integrating factors, solve for

$$e^{-\mu t} X_t - X_0 = \int_0^t e^{-\mu s} \sigma dB_s$$

and deduce that the solution  $X_t$  is

$$X_t = e^{\mu t} X_0 + \sigma \int_0^t e^{\mu(t-s)} dB_s.$$

(b) The expected value of  $X_t$  is

$$\mathbb{E}(X_t) = e^{\mu t} X_0$$

and, by Itô isometry, the variance of  $X_t$  is

$$\mathbb{V}(X_t) = \mathbb{E} \left( \sigma^2 \left( \int_0^t e^{\mu(t-s)} dB_s \right)^2 \right) = \mathbb{E} \left( \sigma^2 \int_0^t e^{2\mu(t-s)} ds \right) = \frac{\sigma^2}{2\mu} (e^{2\mu t} - 1).$$

5.6 Suppose  $Y_t$  is given by

$$dY_t = r dt + \alpha Y_t dB_t.$$

Using integrating factors, solve for

$$d(e^{-\alpha B_t} Y_t) = e^{-\alpha B_t} Y_t \left( r - \frac{\alpha^2}{2} \right) dt$$

and

$$e^{-\alpha B_t + \frac{\alpha^2}{2} t} Y_t - Y_0 = \int_0^t r e^{-\alpha B_s + \frac{\alpha^2}{2} s} ds.$$

Deduce that

$$Y_t = e^{\alpha B_t - \frac{\alpha^2}{2} t} Y_0 + r \int_0^t e^{\alpha(B_t - B_s) - \frac{\alpha^2}{2}(t-s)} ds.$$

5.7 The Ornstein-Uhlenbeck process is given by

$$dX_t = (m - X_t) dt + \sigma dB_t.$$

(a) Using integrating factors, solve for

$$e^t X_t - X_0 = \int_0^t e^s m ds + \int_0^t e^s \sigma dB_s$$

and deduce that the solution  $X_t$  is

$$X_t = e^{-t} X_0 + m(1 - e^{-t}) + \sigma \int_0^t e^{s-t} dB_s.$$

(b) The expected value of  $X_t$  is

$$\mathbb{E}(X_t) = m + e^{-t}(X_0 - m)$$

and the variance of  $X_t$  is

$$\mathbb{V}(X_t) = \mathbb{E} \left( \sigma^2 \left( \int_0^t e^{s-t} dB_s \right)^2 \right) = \mathbb{E} \left( \sigma^2 \int_0^t e^{2s-2t} ds \right) = \frac{\sigma^2}{2}(1 - e^{-2t}).$$

5.8 Consider the stochastic differential equation

$$\begin{pmatrix} dX_t^{(1)} \\ dX_t^{(2)} \end{pmatrix} = \begin{pmatrix} 0 & 1 \\ -1 & 0 \end{pmatrix} \begin{pmatrix} X_t^{(1)} \\ X_t^{(2)} \end{pmatrix} dt + \begin{pmatrix} \alpha dB_t^{(1)} \\ \beta dB_t^{(2)} \end{pmatrix}.$$

By d'Alembert's formula, it has a solution of the form

$$X_t = e^{At} X_0 + \int_0^t e^{A(t-s)} g(s) ds,$$

where

$$e^{At} = \begin{pmatrix} 1 & 1 \\ i & -1 \end{pmatrix} \begin{pmatrix} e^{it} & 0 \\ 0 & e^{-it} \end{pmatrix} \begin{pmatrix} 1 & 1 \\ i & -1 \end{pmatrix}^{-1} = \begin{pmatrix} \cos(t) & \sin(t) \\ -\sin(t) & \cos(t) \end{pmatrix}.$$

Conclude that the solutions are

$$X_t^{(1)} = X_0^{(1)} \cos(t) + X_0^{(2)} \sin(t) + \alpha \int_0^t \cos(t-s) dB_s^{(1)} + \beta \int_0^t \sin(t-s) dB_s^{(2)}$$

and

$$X_t^{(2)} = -X_0^{(1)} \sin(t) + X_0^{(2)} \cos(t) - \alpha \int_0^t \sin(t-s) dB_s^{(1)} + \beta \int_0^t \cos(t-s) dB_s^{(2)}.$$

5.9 Let  $dX_t = \ln(1 + X_t^2) dt + \chi_{\{X_t > 0\}} X_t dB_t$ . It suffices to check that

$$|b(t, x)| + |\sigma(t, x)| = \ln(1 + x^2) + \chi_{\{x > 0\}} |x| \leq \frac{2}{e}(|x| + 1) + |x| \leq 2(|x| + 1),$$

$\mathbb{E}(|X_0|^2) = \alpha^2 < \infty$ , and

$$|b(t, x) - b(t, y)| + |\sigma(t, x) - \sigma(t, y)| \leq |\ln(x^2) - \ln(y^2)| + |x - y| \leq 3|x - y|.$$

Hence, by Theorem 5.2.1, there is a unique strong solution to the stochastic differential equation.

5.10 Calculate

$$\begin{aligned} \mathbb{E}(X_t^2) &= \mathbb{E} \left( Z + \int_0^t b(s, X_s) ds + \int_0^t \sigma(s, X_s) dB_s \right)^2 \\ &\leq 3 \left( \mathbb{E}(Z^2) + \mathbb{E} \left( \int_0^t b(s, X_s) ds \right)^2 + \mathbb{E} \left( \int_0^t \sigma(s, X_s) dB_s \right)^2 \right) \\ &\leq 3 \left( \mathbb{E}(Z^2) + T \mathbb{E} \left( \int_0^t b(s, X_s)^2 ds \right) + \mathbb{E} \left( \int_0^t \sigma(s, X_s)^2 ds \right) \right) \\ &\leq 3\mathbb{E}(Z^2) + 6C^2 \left( T + \int_0^t \mathbb{E}(|X_s|^2) ds \right) (T + 1) \\ &= (3\mathbb{E}(Z^2) + 6C^2 T(T + 1)) + 6C^2 (T + 1) \int_0^t \mathbb{E}(|X_s|^2) ds. \end{aligned}$$

and apply Gronwall to derive the result.

5.11 Consider the stochastic process

$$Y_t = a(1-t) + bt + (1-t) \int_0^t \frac{dB_s}{1-s}.$$

Then  $Y_0 = a$  and, for  $t \in [0, 1)$ ,  $Y_t$  solves

$$\begin{aligned} dY_t &= (b-a)dt - \int_0^t \frac{dB_s}{1-s} dt + (1-t) \frac{dB_t}{1-t} \\ &= \frac{1}{1-t} \left( (b-a)(1-t) - (1-t) \int_0^t \frac{dB_s}{1-s} \right) dt + dB_t \\ &= \frac{1}{1-t} \left( b - a(1-t) - bt - (1-t) \int_0^t \frac{dB_s}{1-s} \right) dt + dB_t \\ &= \frac{b - Y_t}{1-t} dt + dB_t. \end{aligned}$$

Finally by Itô isometry  $\mathbb{E} \left( (1-t)^2 \int_0^t \frac{dB_s}{1-s} \right)^2 = (1-t)^2 \int_0^t \frac{1}{(1-s)^2} ds = (1-t)t \rightarrow 0$  as  $t \rightarrow 1^-$  and so limit  $\lim_{t \rightarrow 1^-} Y_t \stackrel{\text{a.s.}}{\equiv} b$ .

5.12 Let  $y''(t) + (1 + \varepsilon W_t)y(t) = 0$  where  $W_t = \frac{dB_t}{dt}$  is 1-dimensional white noise.

(a) Rewrite

$$\begin{pmatrix} dy_t \\ d\dot{y}_t \end{pmatrix} = \begin{pmatrix} 0 & 1 \\ -1 & 0 \end{pmatrix} \begin{pmatrix} y_t \\ \dot{y}_t \end{pmatrix} dt + \begin{pmatrix} 0 & 0 \\ -\varepsilon & 0 \end{pmatrix} \begin{pmatrix} y_t \\ \dot{y}_t \end{pmatrix} dB_t.$$

(b) Check that, if  $y(t) = y(0) + y'(0)t + \int_0^t (r-t)y(r) dr + \int_0^t \varepsilon(r-t)y(r) dB_r$ , then

$$y'(t) = y'(0) - \int_0^t y(r) dr - \int_0^t \varepsilon y(r) dB_r = y'(0) - \int_0^t y(r)(1 + \varepsilon W_r) dr$$

and  $y''(t) = -(1 + \varepsilon W_r) dr$ .

5.13 Let  $x''_t + a_0 x'_t + w^2 x_t = (T_0 - \alpha_0 x'_t) \eta W_t$  where  $W_t$  is 1-dimensional white noise. Then

$$\begin{pmatrix} dx_t \\ d\dot{x}_t \end{pmatrix} = \begin{pmatrix} 0 & 1 \\ -w^2 & -a_0 \end{pmatrix} \begin{pmatrix} x_t \\ \dot{x}_t \end{pmatrix} dt + \begin{pmatrix} 0 & 0 \\ 0 & -\alpha_0 \eta \end{pmatrix} \begin{pmatrix} x_t \\ \dot{x}_t \end{pmatrix} dB_t + \begin{pmatrix} 0 \\ T_0 \eta \end{pmatrix} dB_t$$

and by d'Alembert's formula the solution is

$$X_t = e^{At} X_0 + \int_0^t e^{A(t-s)} K X_s dB_s + \int_0^t e^{A(t-s)} M dB_s.$$

The eigenvalues of  $A$  satisfy  $\lambda^2 + a_0\lambda + w^2 = 0$  and are  $\lambda_{\pm} = -\frac{a_0}{2} \pm \sqrt{w^2 - \frac{a_0^2}{4}}i =: -\lambda \pm \xi i$ . Then take the exponential of matrix  $A$

$$\begin{aligned} e^{At} &= \begin{pmatrix} 1 & 1 \\ \lambda_+ & \lambda_- \end{pmatrix} \begin{pmatrix} e^{\lambda_+ t} & 0 \\ 0 & e^{\lambda_- t} \end{pmatrix} \begin{pmatrix} 1 & 1 \\ \lambda_+ & \lambda_- \end{pmatrix}^{-1} \\ &= \frac{1}{\lambda_- - \lambda_+} \begin{pmatrix} \lambda_- e^{\lambda_+ t} - \lambda_+ e^{\lambda_- t} & e^{\lambda_- t} - e^{\lambda_+ t} \\ -\lambda_- \lambda_+ (e^{\lambda_- t} - e^{\lambda_+ t}) & \lambda_- e^{\lambda_- t} - \lambda_+ e^{\lambda_+ t} \end{pmatrix} \\ &= -\frac{1}{2\xi i} \begin{pmatrix} e^{-\lambda t} (-\lambda \cdot 2i \sin(\xi t) - \xi i \cdot 2 \cos(\xi t)) & e^{-\lambda t} (-2i \sin(\xi t)) \\ -w^2 e^{-\lambda t} (-2i \sin(\xi t)) & e^{-\lambda t} (-\lambda \cdot 2i \sin(\xi t) - \xi i \cdot 2 \cos(\xi t) + 2\lambda \cdot 2i \sin(\xi t)) \end{pmatrix} \\ &= \frac{e^{-\lambda t}}{\xi} \begin{pmatrix} \lambda \sin(\xi t) + \xi \cos(\xi t) & \sin(\xi t) \\ -w^2 \sin(\xi t) & \lambda \sin(\xi t) + \xi \cos(\xi t) - 2\lambda \sin(\xi t) \end{pmatrix} \\ &= \frac{e^{-\lambda t}}{\xi} ((\lambda \sin(\xi t) + \xi \cos(\xi t))I + A \sin(\xi t)). \end{aligned}$$

Next, letting  $y_s = \dot{x}_s$ ,  $g_t = e^{-\lambda t} \frac{\sin(\xi t)}{\xi}$  and  $h_t = e^{-\lambda t} \frac{\xi \cos(\xi t) - \lambda \sin(\xi t)}{\xi}$ , compute

$$e^{A(t-s)} K X_s = -\frac{\alpha_0 \eta e^{-\lambda(t-s)}}{\xi} \begin{pmatrix} 0 & \sin(\xi(t-s)) \\ 0 & \xi \cos(\xi(t-s)) - \lambda \sin(\xi(t-s)) \end{pmatrix} \begin{pmatrix} x_s \\ \dot{x}_s \end{pmatrix} = \begin{pmatrix} -\alpha_0 \eta y_s g_{t-s} \\ -\alpha_0 \eta y_s h_{t-s} \end{pmatrix}$$

and

$$e^{A(t-s)} M = \frac{T_0 \eta e^{-\lambda(t-s)}}{\xi} \begin{pmatrix} \sin(\xi(t-s)) \\ \xi \cos(\xi(t-s)) - \lambda \sin(\xi(t-s)) \end{pmatrix} = \begin{pmatrix} \eta T_0 g_{t-s} \\ \eta T_0 h_{t-s} \end{pmatrix}.$$

It follows that

$$x_t = \eta \int_0^t (T_0 - \alpha_0 y_s) g_{t-s} dB_s$$

and

$$y_t = \eta \int_0^t (T_0 - \alpha_0 y_s) h_{t-s} dB_s.$$

5.14 Letting  $Z_t = F(\mathbf{B}_t)$ , where  $\mathbf{B}_t = B_t^{(1)} + iB_t^{(2)}$ , calculate

$$\begin{aligned} dZ_t &= F_x(\mathbf{B}_t) dB_t^{(1)} + F_y(\mathbf{B}_t) dB_t^{(2)} \\ &\quad + \frac{1}{2} F_{xx}(\mathbf{B}_t) d[B^{(1)}, B^{(1)}]_t + F_{xy}(\mathbf{B}_t) d[B^{(1)}, B^{(2)}]_t + F_{yy}(\mathbf{B}_t) d[B^{(2)}, B^{(2)}]_t \\ &= (u_x + iv_x) dB_t^{(1)} + (u_y + iv_y) dB_t^{(2)} + \frac{1}{2}(u_{xx} + iv_{xx} + u_{yy} + iv_{yy}) dt \\ &= \langle F'(\mathbf{B}_t), dB_t \rangle + \frac{1}{2}(v_{xy} - iu_{xy} + u_{yy} + iv_{yy}) dt \\ &= \langle F'(\mathbf{B}_t), dB_t \rangle + \frac{1}{2}(-u_{yy} - iv_{yy} + u_{yy} + iv_{yy}) dt \\ &= \langle F'(\mathbf{B}_t), dB_t \rangle. \end{aligned}$$

5.15 Consider the non-linear stochastic differential equation

$$dX_t = rX_t(K - X_t)dt + \beta X_t dB_t, \quad X_0 = x > 0.$$

Comparing to the deterministic Bernoulli equation, do a substitution  $Y_t = X_t^{-1}$ , then

$$\begin{aligned} dY_t &= -rY_t(K - X_t)dt - \beta Y_t dB_t + \beta^2 Y_t dt \\ &= (-rK + \beta^2)Y_t dt - \beta Y_t dB_t + r dt. \end{aligned}$$

Next do a new change of variables

$$Z_t = Y_t e^{(rK - \beta^2)t}$$

and calculate

$$\begin{aligned} dZ_t &= -\beta Z_t dB_t + re^{(rk - \beta^2)t} dt \\ \implies Z_t &= e^{-\beta B_t} \left( x^{-1} + r \int_0^t e^{(rk - \beta^2)s + \beta B_s} ds \right). \end{aligned}$$

Conclude that

$$X_t = \frac{e^{(rk - \beta^2)t}}{Z_t} = \frac{e^{(rk - \beta^2)t + \beta B_t}}{x^{-1} + r \int_0^t e^{(rk - \beta^2)s + \beta B_s} ds}.$$

5.16 Consider the non-linear stochastic differential equation

$$dX_t = f(t, X_t)dt + c(t)X_t dB_t, \quad X_0 = x.$$

(a) Let  $F_t(\omega) = \exp\left(-\int_0^t c(s) dB_s + \frac{1}{2} \int_0^t c(s)^2 ds\right)$ . Then calculate

$$\begin{aligned} d(F_t X_t) &= X_t dF_t + F_t dX_t + d[F_t, X_t] \\ &= X_t \left[ F_t \left( -c(t) dB_t - \frac{1}{2} c(t)^2 dt - \frac{1}{2} c(t)^2 dt \right) \right] \\ &\quad + [f(t, X_t)F_t dt + c(t)X_t F_t dB_t] - c(t)^2 F_t X_t dt \\ &= f(t, X_t)F_t dt. \end{aligned}$$

(b) Defining  $Y_t = F_t X_t$ , deduce that

$$\frac{dY_t}{dt} = F_t(\omega)f(t, F_t^{-1}(\omega)Y_t(\omega)).$$

(c) Consider  $dX_t = X_t^{-1} + \alpha X_t dB_t$ ,  $X_0 = x > 0$ . Then

$$\frac{dY_t}{dt} = e^{-2\alpha B_t + \alpha^2 t} Y_t^{-1},$$

which implies

$$Y_t = \sqrt{Y_0^2 + 2 \int_0^t e^{-2\alpha B_s + \alpha^2 s} ds}$$

and

$$X_t = e^{\alpha B_t - \frac{\alpha^2}{2} t} \sqrt{x^2 + 2 \int_0^t e^{-2\alpha B_s + \alpha^2 s} ds}.$$

(d) Consider  $dX_t = X_t^\gamma dt + \alpha X_t dB_t$ ,  $X_0 = x > 0$ . Then

$$\frac{dY_t}{dt} = e^{-(1-\gamma)B_t + (1-\gamma)\frac{\alpha^2}{2}t} Y_t^\gamma,$$

which implies

$$Y_t = \left( Y_0^{1-\gamma} + (1-\gamma) \int_0^t e^{-(1-\gamma)B_s + (1-\gamma)\frac{\alpha^2}{2}s} ds \right)^{\frac{1}{1-\gamma}}$$

and

$$X_t = e^{\alpha B_t - \frac{\alpha^2}{2} t} \left( x^{1-\gamma} + (1-\gamma) \int_0^t e^{-(1-\gamma)B_s + (1-\gamma)\frac{\alpha^2}{2}s} ds \right)^{\frac{1}{1-\gamma}}.$$

5.17 Let  $v \geq 0$  satisfy  $v(t) \leq C + A \int_0^t v(s) ds$  and consider quantity  $w(t) = \int_0^t v(s) ds$ . Then

$$w'(t) = v(t) \leq C + A \int_0^t v(s) ds = C + Aw(t).$$

Then for  $f(t) = w(t)e^{-At}$ , calculate

$$f'(t) = e^{-At} (w'(t) - Aw(t)) \leq Ce^{-At}$$

and

$$\begin{aligned} w(t)e^{-At} &\leq \int_0^t Ce^{-As} ds = \frac{C}{A}(1 - e^{-At}) \\ \implies w(t) &\leq \frac{C}{A}(e^{At} - 1). \end{aligned}$$

Deduce that

$$v(t) \leq C + Aw(t) \leq Ce^{At}.$$

## Diffusions: Basic Properties (§7)

7.1 Using Theorem 7.3.3, the generators of the Itô diffusions are

- (a)  $(Af)(x) = \mu x f'(x) + \frac{1}{2} \sigma^2 f''(x)$  for  $f \in C_0^2(\mathbb{R})$
- (b)  $(Af)(x) = r x f'(x) + \frac{1}{2} \alpha^2 x^2 f''(x)$  for  $f \in C_0^2(\mathbb{R})$
- (c)  $(Af)(y) = r f'(y) + \frac{1}{2} \alpha^2 y^2 f''(y)$  for  $f \in C_0^2(\mathbb{R})$
- (d)  $\mathbf{b} = (1, \mu X)^T$  and  $\sigma = (0, \sigma)^T$  to deduce for  $f \in C_0^2(\mathbb{R}^2)$

$$(Af)(y_1, y_2) = \frac{\partial f}{\partial y_1} + \mu y_2 \frac{\partial f}{\partial y_2} + \frac{1}{2} \sigma^2 \frac{\partial^2 f}{\partial y_2^2}$$

- (e)  $(Af)(x_1, x_2) = \frac{\partial f}{\partial x_1} + x_2 \frac{\partial f}{\partial x_2} + \frac{1}{2} e^{2x_1} \frac{\partial^2 f}{\partial x_2^2}$  for  $f \in C_0^2(\mathbb{R}^2)$ .

(f) Since  $d[B_i, B_j] = \delta_{ij} dt$ , for  $f \in C_0^2(\mathbb{R}^2)$ , deduce

$$(Af)(x_1, x_2) = \frac{\partial f}{\partial x_1} + \frac{1}{2} \frac{\partial^2 f}{\partial x_1^2} + \frac{1}{2} x_1^2 \frac{\partial^2 f}{\partial x_2^2}$$

- (g) and finally for  $\mathbf{X}$  and  $f \in C_0^2(\mathbb{R}^n)$

$$\begin{aligned} (Af(\mathbf{x})) &= (rx)_k \cdot \nabla f(\mathbf{x}) + \frac{1}{2} \sum_{j=1}^n \sum_{k,\ell=1}^n (\alpha_{kj} x_k) (\alpha_{\ell j} x_\ell) \frac{\partial^2 f}{\partial x_k \partial x_\ell} \\ &= (rx)_k \cdot \nabla f(\mathbf{x}) + \frac{1}{2} \sum_{k,\ell=1}^n x_k x_\ell \left( \sum_{j=1}^n \alpha_{kj} \alpha_{\ell j} \right) \frac{\partial^2 f}{\partial x_k \partial x_\ell}. \end{aligned}$$

7.2 In this problem, the Itô diffusion processes are found.

- (a) Since  $b = 1$ ,  $\frac{1}{2}\sigma^2 = 2$ , it follows  $dX_t = b dt + \sigma dB_t = dt \pm \sqrt{2} dB_t$ .
- (b) Rewrite

$$(Af)(t, x) = (1, cx) \cdot (Df)(t, x) + \frac{1}{2} \begin{pmatrix} 0 & \alpha x \end{pmatrix} (D^2 f)(t, x) \begin{pmatrix} 0 \\ \alpha x \end{pmatrix}$$

and deduce that

$$dY_t = \begin{pmatrix} dY_1 \\ dY_2 \end{pmatrix} = \begin{pmatrix} 1 \\ cY_2 \end{pmatrix} dt \pm \begin{pmatrix} 0 \\ \alpha Y_2 \end{pmatrix} dB_t.$$

(c) The equation reduces to

$$(Af)(x_1, x_2) = (2x_2, \ln(1 + x_1^2 + x_2^2)) \cdot (Df)(x_1, x_2) + \frac{1}{2} \begin{pmatrix} 1 & 0 \end{pmatrix} (D^2f)(x_1, x_2) \begin{pmatrix} 1 \\ 0 \end{pmatrix} \\ + \frac{1}{2} \begin{pmatrix} x_1 & 1 \end{pmatrix} (D^2f)(x_1, x_2) \begin{pmatrix} x_1 \\ 1 \end{pmatrix}$$

and thus

$$\begin{pmatrix} dX_1 \\ dX_2 \end{pmatrix} = \begin{pmatrix} 2X_2 \\ \ln(1 + X_1^2 + X_2^2) \end{pmatrix} dt \pm \begin{pmatrix} 1 \\ 0 \end{pmatrix} dB_t^{(1)} \pm \begin{pmatrix} X_1 \\ 1 \end{pmatrix} dB_t^{(2)}.$$

7.3 Let  $B_t$  be standard Brownian motion and define  $X_t = x \cdot e^{ct+\alpha B_t}$ . We prove it for  $\mathcal{P}(\mathbb{R})$ , which is dense by Stone-Weierstrass in the set of bounded Borel measurable functions. Fix  $f(x) = x^n$  and since  $\sigma(X_t) = \sigma(B_t)$ , we have

$$\begin{aligned} \mathbb{E}[f(X_{t+h}) | \mathcal{F}_t] &= \mathbb{E}[x^n \cdot e^{cn(t+h)+\alpha n(B_t+(B_{t+h}-B_t))} | \mathcal{F}_t] \\ &= x^n e^{cn(t+h)} e^{\alpha n B_t} \mathbb{E}[e^{\alpha n(B_{t+h}-B_t)} | \mathcal{F}_t] \\ &= x^n e^{cn(t+h)} e^{\alpha n B_t} \mathbb{E}[e^{\alpha n(B_{t+h}-B_t)}] \\ &= x^n e^{cn(t+h)} \mathbb{E}[e^{\alpha n(B_t+(B_{t+h}-B_t))} | B_t] \\ &= x^n e^{cn(t+h)} \mathbb{E}[e^{\alpha n(B_t+(B_{t+h}-B_t))} | X_t] \\ &= \mathbb{E}[f(X_{t+h}) | X_t]. \end{aligned}$$

7.4 Let  $B_t$  be Brownian motion starting at  $x > 0$ . The generator is  $(Af)(x) = \frac{1}{2}f''(x)$ . Let  $\tau = \inf\{t | B_t = 0\}$ .

(a) Consider exit time  $\sigma_k$  from interval  $A_k = \{0 < y < k\}$  where  $X_0 = x \in (0, k)$ . Applying Dynkin's formula for  $f(x) = x$ , calculate

$$k\mathbb{P}(X_{\sigma_k} = k) = x \implies \mathbb{P}(X_{\sigma_k} = 0) = 1 - \frac{x}{k} \xrightarrow{k \rightarrow \infty} 1.$$

(b) Applying Dynkin's formula again for  $f(x) = x^2$ , derive

$$\begin{aligned} kx &= k^2\mathbb{P}(X_{\sigma_k} = k) = x^2 + \mathbb{E}[X_{\sigma_k}] \\ &= x^2 + \mathbb{E}[\sigma_k | X_{\sigma_k} = 0]\mathbb{P}(X_{\sigma_k} = 0) + \mathbb{E}[\sigma_k | X_{\sigma_k} = k]\mathbb{P}(X_{\sigma_k} = k) \\ &\xrightarrow{k \rightarrow \infty} x^2 + \mathbb{E}[\tau], \end{aligned}$$

and so  $\mathbb{E}[\tau] = \infty$ .

7.5 By Theorem 5.2.1, there exists a unique  $t$ -continuous solution  $X_t(\omega)$ . Applying Dynkin's formula with  $f(\mathbf{x}) = |\mathbf{x}|^2$  and  $\tau = \min\{t, \tau_R\}$ , calculate

$$\begin{aligned}\mathbb{E}[|X_\tau|^2] &= \mathbb{E}[X_t^2]\mathbb{P}(t \leq \tau_R) + R^2\mathbb{P}(\tau_R < t) \\ &= |X_0|^2 + \mathbb{E}\left[\int_0^\tau (Af)(X_s) ds\right] \\ &= |X_0|^2 + \mathbb{E}\left[\int_0^\tau \mathbf{b}(s, X_s) \cdot 2X_s + \frac{1}{2}\sigma(s, X_s)(2\mathbf{I}_n)\sigma^T(s, X_s) ds\right] \\ &\leq |X_0|^2 + \mathbb{E}\left[\int_0^\tau |\mathbf{b}|^2 + |X_s|^2 + |\sigma|^2 ds\right] \\ &\leq |X_0|^2 + \mathbb{E}\left[\int_0^\tau C^2(1 + |X_s|)^2 + |X_s|^2 ds\right] \\ &\leq |X_0|^2 + (2C^2 + 1)\mathbb{E}\left[\int_0^\tau (1 + |X_s|^2) ds\right].\end{aligned}$$

Passing to the limit as  $R \rightarrow \infty$ ,

$$\mathbb{E}[X_t^2] \leq |X_0|^2 + (2C^2 + 1)\mathbb{E}\left[\int_0^t (1 + |X_s|^2) ds\right].$$

Letting  $f(t) := \mathbb{E}[X_t^2] + 1$ ,  $f(t) - f(0) \leq (2C^2 + 1) \int_0^t f(s) ds$ . By applying Gronwall's Lemma (or checking that  $g(t) := f(t)e^{-(2C^2+1)t}$  is decreasing), the proof is complete.

7.6 Let  $g(x, \omega) = (f \circ F)(x, t, t+h, \omega)$ .

- (a) It suffices to check that  $x \mapsto F(x, \dots)$  is continuous. Consider two paths starting at  $X_t(\omega)$ . Then check

$$\begin{aligned}\mathbb{E}[|X_{t+h}^{(1)} - X_{t+h}^{(2)}|^2] &\leq |X_t^{(1)} - X_t^{(2)}|^2 + (2D^2 + 1)\mathbb{E}\left[\int_t^{t+h} |X_u^{(1)} - X_u^{(2)}|^2 du\right]\end{aligned}$$

and apply the previous problem.

- (b) Follow the hint.

7.7 Let  $B_t$  be  $n$ -dimensional Brownian motion starting at  $x$  and  $D$  be a ball centred at  $x$ .

- (a) Recall  $\tilde{B}_t = UB_t$  where  $U$  is orthogonal is also Brownian. Then for  $S \subset \partial D$

$$\mu_D^x(S) = \mathbb{P}(B_{\tau_D} \in S) = \mathbb{P}(\tilde{B}_{\tau_D} \in U \cdot S) = \mu_D^x(U \cdot S).$$

Hence, the harmonic measure is rotation-invariant.

(b) Calculate

$$\begin{aligned}
u(x) &= \mathbb{E}^x[\phi(B_{\tau_W})] \\
&= \mathbb{E}^x[\mathbb{E}^x[\phi(B_{\tau_W}) \mid B_{\tau_D}]] \\
&= \mathbb{E}^x[\mathbb{E}^{B_{\tau_D}}[\phi(B_{\tau_W})]] \\
&= \mathbb{E}^x[u(B_{\tau_D})] \\
&= \int_{\partial D} u(y) d\mu_D^x(y) \\
&= \int_{\partial D} u(y) d\sigma(y),
\end{aligned}$$

which completes the proof.

7.8 Let  $\mathcal{N}_t$  be a right-continuous family of  $\sigma$ -algebras of subsets of  $\Omega$ , containing all sets of measure zero.

- (a) Since  $\mathcal{N}_t$  is closed under finite unions and intersections, it follows that  $\{\min\{\tau_1, \tau_2\} \leq t\} = \cup_{j=1}^2 \{\tau_j \leq t\} \in \mathcal{N}_t$  and  $\{\max\{\tau_1, \tau_2\} \leq t\} = \cap_{j=1}^2 \{\tau_j \leq t\} \in \mathcal{N}_t$ .
- (b) Check  $\{\tau \leq t\} = \lim_{n \rightarrow \infty} \{\tau_n \leq t\} = \cup_{j=1}^{\infty} \{\tau_n \leq t\} \in \mathcal{N}_t$ .
- (c)  $F$  is a  $G_{1/n}$  set and  $\{\tau_n \leq t\} \in \mathcal{M}_t$  for  $\tau_n = \{\inf t > 0 \mid X_t \notin G_{1/n}\}$ . Since  $G_{1/n} \downarrow F$  and given that  $\tau_n$  is a decreasing family of stopping times, it follows by part (b) that  $\tau_F = \cup_{j=1}^n \{\tau_n \leq t\} \in \mathcal{M}_t$ .

7.9 Let  $X_t$  be geometric Brownian motion

$$dX_t = rX_t dt + \alpha X_t dB_t, \quad X_0 = x > 0.$$

- (a) Compute  $(Af)(x) = rx f'(x) + \frac{1}{2}\alpha^2 x^2 f''(x)$  for  $f \in C_0^2(\mathbb{R})$ , whence for  $f(x) = x^\gamma$

$$(Af)(x) = \left(r + \frac{1}{2}\alpha^2(\gamma - 1)\right) \gamma x^\gamma.$$

- (b) Applying Dynkin's formula with  $f(x) = x^{\gamma_1}$ ,  $\gamma_1 = 1 - \frac{2r}{\alpha^2}$  and  $\sigma_k = \min\{k, \tau\}$ , calculate

$$\begin{aligned}
f(x) &= x^{\gamma_1} = x^{\gamma_1} + \mathbb{E}\left[\int_0^{\sigma_k} Af(X_s) ds\right] \\
&= \mathbb{E}[f(X_{\sigma_k})] \\
&= R^{\gamma_1} \mathbb{P}(\tau \leq k) + \mathbb{E}[f(X_k) \mid k < \tau] \mathbb{P}(k < \tau) \\
&\stackrel{k \rightarrow \infty}{\rightarrow} R^{\gamma_1} p,
\end{aligned}$$

and deduce  $p = \left(\frac{x}{R}\right)^{\gamma_1}$ .

- (c) Consider exit time  $\sigma_\rho$  from annulus  $A_\rho = \{\rho < y < R\}$  where  $X_0 = x \in (\rho, R)$ . Using parts (a) and (b), note that

$$\mathbb{P}(X_{\sigma_\rho} = R) = \frac{x^\gamma - \rho^\gamma}{R^\gamma - \rho^\gamma},$$

and thus for exit time  $\sigma_\rho$  from  $A_\rho$ , calculate

$$\begin{aligned} & \mathbb{E} \left[ \int_0^{\sigma_\rho} \left( r - \frac{1}{2}\alpha^2 \right) ds \right] \\ &= \left( r - \frac{1}{2}\alpha^2 \right) \mathbb{E}[\sigma_\rho] \\ &= \left( r - \frac{1}{2}\alpha^2 \right) \mathbb{E}[\sigma_\rho | X_{\sigma_\rho} = R] \mathbb{P}(X_{\sigma_\rho} = R) + \left( r - \frac{1}{2}\alpha^2 \right) \mathbb{E}[\sigma_\rho | X_{\sigma_\rho} = \rho] \mathbb{P}(X_{\sigma_\rho} = \rho) \\ &\xrightarrow{\rho \rightarrow 0^+} \left( r - \frac{1}{2}\alpha^2 \right) \mathbb{E}[\tau] \end{aligned}$$

and

$$\mathbb{E}[\ln(X_{\sigma_\rho})] = \ln(R)\mathbb{P}(X_{\sigma_\rho} = R) + \ln(\rho)\mathbb{P}(X_{\sigma_\rho} = \rho) \xrightarrow{\rho \rightarrow 0^+} \ln(R).$$

Applying Dynkin's formula with  $f(x) = \ln(x)$ , deduce

$$\ln(R) = \ln(x) + \left( r - \frac{1}{2}\alpha^2 \right) \mathbb{E}[\tau],$$

which completes the proof.

7.10 Let  $X_t$  be geometric Brownian motion.

- (a) Using the Markov property,

$$\begin{aligned} \mathbb{E}^x[X_T | \mathcal{F}_t] &= \mathbb{E}^{X_t}[X_{T-t}] \\ &= X_t + \mathbb{E}^{X_t} \left[ \int_0^{T-t} r X_u du + \int_0^{T-t} \alpha X_u dB_u \right] \\ &= X_t + r \int_0^{T-t} \mathbb{E}^{X_t}[X_u] du. \end{aligned}$$

Letting  $f(s) = \mathbb{E}^{X_t}[X_{s-t}]$ , we have  $f'(s) = rf(s)$  by differentiation and so

$$\mathbb{E}^x[X_T | \mathcal{F}_t] = f(T-t) = f(0)e^{r(T-t)} = X_t e^{r(T-t)}.$$

- (b) The solution is given by  $X_t = xe^{rt}M_t$ , where  $M_t = e^{\alpha B_t - \frac{1}{2}\sigma^2 t}$  is a martingale. So

$$\mathbb{E}^x[X_T | \mathcal{F}_t] = \mathbb{E}[X_T | X_t] = xe^{rT}\mathbb{E}[M_T | X_t] = xe^{rT}M_t = X_t e^{r(T-t)}.$$

7.11 Let  $\tau$  be a stopping time. By the strong Markov property,

$$\begin{aligned}\mathbb{E}^x[g(X_\tau)] &= \mathbb{E}^x\left[\mathbb{E}^{X_\tau}\left[\int_0^\infty f(X_u) du\right]\right] \\ &= \mathbb{E}^x\left[\int_0^\infty \mathbb{E}^{X_\tau}[f(X_u)] du\right] \\ &= \mathbb{E}^x\left[\int_0^\infty \mathbb{E}[f(X_{u+\tau}) | \mathcal{F}_\tau] du\right] \\ &= \mathbb{E}^x\left[\int_0^\infty f(X_{u+\tau}) du\right] \\ &= \mathbb{E}^x\left[\int_\tau^\infty f(X_t) dt\right],\end{aligned}$$

which completes the proof.

## **Other Chapters (§6, §8 –12)**

**The Filtering Problem (§6)**

**Other Topics in Diffusion Theory (§8)**

**Applications to Boundary Value Problems (§9)**

**Applications to Optimal Stopping (§10)**

**Applications to Stochastic Control (§11)**

**Applications to Mathematical Finance (§12)**