

Q1:

To prove Gauss-Markov Theorem, we need to prove unbiasedness and linearity

$$\text{First of all, } E(\hat{\beta}) = E((X^T X)^{-1} X^T y) = (X^T X)^{-1} X^T E(y) \\ = (X^T X)^{-1} X^T E(BX + \varepsilon)$$

$$E(s) > 0 \Rightarrow (X^T X)^{-1} X^T E(X\beta) \\ = (X^T X)^{-1} X^T X\beta$$

$$\text{since } (X^T X)^{-1} \cdot (X^T X) = I$$

$$\therefore (X^T X)^{-1} X^T XB = I \cdot \beta = \beta$$

$$\therefore E(\hat{\beta}) = \beta$$

In linear regression $y = \pi \beta + \varepsilon$

suppose $(x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)$

We have $y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \dots + \beta_k x_{ik} + \varepsilon_i$

from the form of linear model, we can see that OLS estimate is a linear function of the given data. The dependent variable y of each data can be expressed as a linear combination between $\beta_0, \beta_1, \beta_2, \dots, \beta_k$ and corresponding variables x_i . Linearity proved

$$\begin{cases} \hat{\alpha^T \beta} = \hat{\alpha^T \beta} + \varepsilon_1 \\ C^T \eta = \hat{\alpha^T \beta} + \varepsilon_2 \end{cases} \Rightarrow \begin{cases} \text{Var}(\hat{\alpha^T \beta} + \varepsilon_1) = \text{Var}(\hat{\alpha^T \beta}) \\ \text{Var}(\hat{\alpha^T \beta} + \varepsilon_2) = \text{Var}(C^T \eta) \end{cases}$$

$$\Rightarrow \begin{cases} \text{Var}(\hat{\alpha^T \beta}) = \text{Var}(\varepsilon_1) \\ \text{Var}(C^T \eta) = \text{Var}(\varepsilon_2) \end{cases}$$

Since $C^T \eta$ is an unbiased estimator for $\hat{\alpha^T \beta}$, $E(C^T \eta) = \hat{\alpha^T \beta}$
 $\therefore E(\varepsilon_2) = 0$

$$\begin{aligned} \text{Var}(\varepsilon_1) &\leq \text{Var}(\varepsilon_1 + \varepsilon_2) \\ \Rightarrow \text{Var}(\hat{\alpha^T \beta}) &\leq \text{Var}(C^T \eta) \end{aligned}$$

\therefore the least squares estimate in linear regression is BLUE.

Q₂:

Since columns x_0, \dots, x_p of X are orthogonal, they are pairwise orthogonal to each other. The dot product of different columns are '0'.

Linear regression: $y = \beta_0 x_0 + \beta_1 x_1 + \dots + \beta_p x_p + \epsilon$

$$\therefore \hat{\beta}_j = \frac{x_j \cdot y}{x_j^T x_j} = \frac{x_j \cdot y}{\|x_j\|^2} \quad (j \text{ from } 1 \text{ to } p)$$

Q₃, from the question: We assume the SVD of X is $U\Sigma V^T$.
 $U \in \mathbb{R}^{n \times (p+1)}$ satisfies $U^T U = I_r$, $V \in \mathbb{R}^{(p+1) \times r}$ satisfies $V^T V = I_r$ and $\Sigma = \text{diag}(\sigma_1, \dots, \sigma_r)$ is the diagonal matrix of positive singular values.

(a)

$$\begin{aligned} \text{SVD of } X: U\Sigma V^T &\Rightarrow (U\Sigma V^T)^T U\Sigma V^T = V\Sigma^T U^T U\Sigma V^T \\ \text{SVD of } X^T: (V\Sigma^T)^T &= V\Sigma^T \Sigma V^T \cdot I_r \\ &= V\Sigma^2 \cdot V^T \end{aligned}$$

$$\text{consider } X^T X \beta = X^T y$$

$$V\Sigma^2 V^T \beta = U\Sigma V^T y$$

$$V^T V \Sigma^2 V^T \beta = V^T U \Sigma V^T y$$

$$I_r \cdot \Sigma^2 V^T \beta = V^T U \Sigma V^T y$$

$$\Sigma^{-2} \cdot \Sigma^2 V^T \beta = \Sigma^2 V^T U \Sigma V^T y$$

$$I_r \cdot V^T \beta = \Sigma^2 V^T U \Sigma V^T y$$

$$V \cdot V^T \beta = V \Sigma^{-2} V^T U \Sigma V^T y$$

$$\Rightarrow \beta = V \Sigma^{-2} V^T U \Sigma V^T y$$

$$\beta = V \Sigma^{-1} \cdot \Sigma^{-1} V^T U \cdot \Sigma V^T y = V \cdot \Sigma^{-1} \cdot (\Sigma^{-1} \cdot V^T \cdot U \Sigma V^T y)$$

Since $\Sigma = \text{diag}(\sigma_1, \dots, \sigma_n) \Rightarrow \Sigma^{-1} = \Sigma^T$

Since $V^T V = I_r \Rightarrow V^T = V^{-1}$

$\therefore \Sigma^{-1} V^{-1}$ is the inverse of $U \Sigma V^T$

$$\therefore \beta = V \cdot \Sigma^{-1} \cdot (U^T y) \text{ which } \Sigma^{-1} \cdot V^T \cdot U \Sigma V^T y = U^T y$$

$$\therefore \beta_{\text{mns}} = V \Sigma^{-1} U^T y$$

(b) hint $\beta = \beta_{\text{mns}} + b$

$$X^T X \beta = X^T y \Rightarrow X^T X (\beta_{\text{mns}} + b) = X^T y$$

$$X^T X \beta_{\text{mns}} + X^T X b = X^T y$$

since β_{mns} is a solution to the normal equations:

$$X^T X \beta_{\text{mns}} = X^T y$$

$$\therefore X^T y + X^T X b = X^T y$$

$$\therefore X^T X b = 0$$

$X^T X$ is not invertible, b lies in the null space of $X^T X$ and can be expressed as a linear combination of the columns of V corresponding to 0 singular values.

Since $V^T V = I_r$, columns of V are orthogonal
 $\therefore b$ can be expressed as the form of $b = b_1 + b_2 \dots + b_n$,
 where b_i is a component in the null space of $X^T X$

$$\text{as for } \|b\| = \|f_{mn} + b\|$$

$$\therefore \|b\| \leq \|f_{mn}\| + \|b\|$$

$$\|b\| = \|b_1 + b_2 \dots + b_n\| \leq \|b_1\| + \|b_2\| \dots + \|b_n\|$$

However, $X^T X b = 0$ and $\|b\|$ is positive

$$\therefore \|b\| \geq \|f_{mn}\|$$

(c) the pseudo-inverse of X has properties of:

$$\left\{ \begin{array}{l} X^T X = I \\ X X^T X = X \\ (X^T X)^T = X^T X \\ (X X^T)^T = X X^T \end{array} \right.$$

When it comes to $V\Sigma^{-1}U^T$

$$V\Sigma^{-1}U^T X = V\Sigma^{-1}U^T \cdot U\Sigma V^T = I$$

$$X(V\Sigma^{-1}U^T)X = U\Sigma V^T$$

$$(V\Sigma^{-1}U^T X)^T = V\Sigma^{-1}U^T U\Sigma V^T$$

$$(XV\Sigma^{-1}U^T)^T = U\Sigma V^T \cdot V\Sigma^{-1}U^T$$

∴ Yes. $V\Sigma^{-1}X^T$ is the Pseudo-inverse of X .