

①

firstly use $\alpha_1, \alpha_2, \dots, \alpha_m$ and e_1, \dots, e_m to rebuild vector \hat{x}

$$\hat{x} = \sum_{i=1}^m \alpha_i e_i$$

$$\begin{aligned} J &= \|x - \hat{x}\|^2 = \left\| x - \sum_{i=1}^m \alpha_i e_i \right\|^2 = \left(x - \sum_{i=1}^m \alpha_i e_i \right) \cdot \left(x - \sum_{i=1}^m \alpha_i e_i \right)^T \\ &= x \cdot x - 2x \cdot \sum_{i=1}^m (\alpha_i e_i) + \left(\sum_{i=1}^m \alpha_i e_i \right) \cdot \left(\sum_{i=1}^m \alpha_i e_i \right) \end{aligned}$$

for eigenvectors are orthonormal, so $e_i \cdot e_j = 0$ ($i \neq j$) $e_i \cdot e_j = 1$

$$\therefore \sum \alpha_i e_i \cdot \sum \alpha_i e_i = \sum \alpha_i^2$$

$$\therefore J = x \cdot x - 2x \cdot \sum_{i=1}^m (\alpha_i e_i) + \sum_{i=1}^m (\alpha_i^2)$$

Since minimizing of squared error of J with respect of $\alpha_1, \dots, \alpha_m$.

We only need to consider the second term, the other terms are not depend on α_i .

$$\frac{\partial J}{\partial \alpha_i} = -2(x \cdot e_i) + 2\alpha_i = 0$$

$$\alpha_i = x \cdot e_i$$

$$\therefore \alpha_i = e_i^T x, i=1, 2, \dots, M$$

②

$$\left\{ \begin{array}{l} \eta = w^T x \\ u_i' = w^T u_i \\ \sigma_i'^2 = w^T \Sigma_i w \end{array} \right.$$

to Maximize the separation means , we want to maximize $(u_i' - u_j')^2$

to minimize the within class variance , we want to minimize $\sigma_i'^2 + \sigma_j'^2$

$$\frac{(u_i' - u_j')^2}{(\sigma_i'^2 + \sigma_j'^2)} = 1$$

take derivative of 'lag' with respect to w and set it to 0

$$2(u_i' - u_j')(\sigma_i'^2 + \sigma_j'^2)w - 2(u_i' - u_j')(\sigma_i'^2 + \sigma_j'^2)w + 2\lambda w = 0$$

$$\Rightarrow 2(u_i' - u_j')(\sigma_i'^2 + \sigma_j'^2 + \lambda)w = 0$$

as for $w \neq 0$

$$\text{we have } 2(u_i' - u_j')(\sigma_i'^2 + \sigma_j'^2 + \lambda) = 0$$

$$\text{since } (\sigma_i'^2 + \sigma_j'^2 + \lambda) > 0$$

$$\therefore u_i' - u_j' = 0 \quad u_i' = u_j'$$

\therefore When $u_i' = u_j'$, $J_1(w)$ is maximized, which is also satisfied

$$\text{in } J_1(w) = \frac{(u_i' - u_j')^2}{\sigma_i'^2 + \sigma_j'^2}$$

(b) incorporate the prior probabilities $P(c|w_i)$ into the criterion .

$$u_i' = w^T u_i$$

$$\text{variance of projected density} ; \quad \sigma_i'^2 = w^T \Sigma_i w$$

$$\|w\| = 1$$

We can get

$$2(u_1' - u_2') P(w_1) (\sigma_1^2 P(w_1) + \sigma_2^2 P(w_2)) w - 2(u_1' - u_2') P(w_1) (\sigma_1^2 P(w_1) \\ + \sigma_2^2 P(w_2)) \lambda w + 2\lambda w = 0$$

$$\Rightarrow 2(u_1' - u_2') P(w_1) (\sigma_1^2 P(w_1) + \sigma_2^2 P(w_2)) + \lambda w = 0$$

$w \neq 0$ and $(\sigma_1^2 P(w_1) + \sigma_2^2 P(w_2)) + \lambda$ is positive

$$\therefore (u_1' - u_2') P(w_1) = 0$$

$$u_1' = u_2'$$

this also stands when $J_2(w) = \frac{(u_1 - u_2)^2}{P(w_1)\sigma_1^2 + P(w_2)\sigma_2^2}$

(c) $J(w)$ maximize the separation between the mean of the projected densities while considering the variance of each class.

$T(w_2)$ incorporates the prior probabilities of the classes in addition to the w and σ .

Fisher's LDA usually maximize $J(w)$ to find the optimal answer.

Therefore, $J(w)$ is closer.