A Handbook of Statistical Analyses Using ${\sf R}$

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CHAPTER 7

Density Estimation: Erupting Geysers and Star Clusters

- 7.1 Introduction
- 7.2 Density Estimation
- 7.3 Analysis Using R

```
7.3.1 A Parametric Density Estimate for the Old Faithful Data
```

```
R> logL <- function(param, x) {</pre>
      d1 \leftarrow dnorm(x, mean = param[2], sd = param[3])
      d2 \leftarrow dnorm(x, mean = param[4], sd = param[5])
      -sum(log(param[1] * d1 + (1 - param[1]) * d2))
+ }
R> startparam <- c(p = 0.5, mu1 = 50, sd1 = 3, mu2 = 80,
      sd2 = 3)
R> opp <- optim(startparam, logL, x = faithful$waiting,
      method = "L-BFGS-B", lower = c(0.01, rep(1,
          4)), upper = c(0.99, rep(200, 4)))
R> opp
$par
                   mu1
                               sd1
                                           mu2
 0.3608905 54.6120933 5.8723821 80.0934226
$value
[1] 1034.002
$counts
function gradient
      55
$convergence
```

[1] 0

Of course, optimising the appropriate likelihood 'by hand' is not very convenient. In fact, (at least) two packages offer high-level functionality for estimating mixture models. The first one is package mclust (Fraley et al., 2005) implementing the methodology described in Fraley and Raftery (2002). Here,

```
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```

```
R> data("faithful", package = "datasets")
   R> x <- faithful$waiting</pre>
   R> layout(matrix(1:3, ncol = 3))
   R> hist(x, xlab = "Waiting times (in min.)", ylab = "Frequency",
         probability = TRUE, main = "Gaussian kernel",
         border = "gray")
   R> lines(density(x, width = 12), lwd = 2)
   R> rug(x)
   R> hist(x, xlab = "Waiting times (in min.)", ylab = "Frequency",
         probability = TRUE, main = "Rectangular kernel",
11
         border = "gray")
   R> lines(density(x, width = 12, window = "rectangular"),
12
         lwd = 2)
13
   R> rug(x)
14
   R> hist(x, xlab = "Waiting times (in min.)", ylab = "Frequency",
15
         probability = TRUE, main = "Triangular kernel",
16
         border = "gray")
17
   R> lines(density(x, width = 12, window = "triangular"),
18
         lwd = 2)
19
   R> rug(x)
```

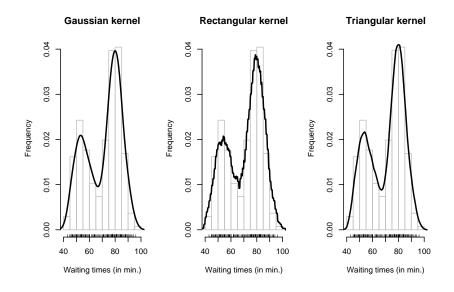


Figure 7.1 Density estimates of the geyser eruption data imposed on a histogram of the data.

```
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R> contour(x = CYGOB1d\$x1, y = CYGOB1d\$x2, z = CYGOB1d\$fhat,

+ xlab = "log surface temperature", ylab = "log light intensity")

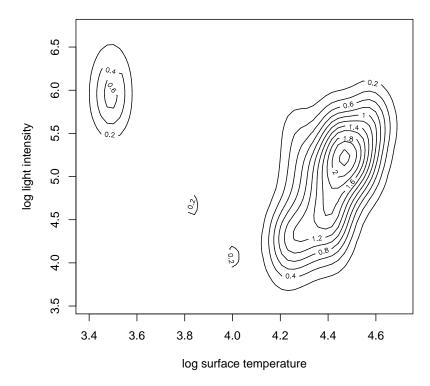


Figure 7.2 A contour plot of the bivariate density estimate of the CYGOB1 data, i.e., a two-dimensional graphical display for a three-dimensional problem.

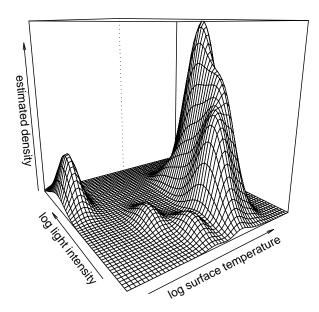


Figure 7.3 The bivariate density estimate of the CYGOB1 data, here shown in a three-dimensional fashion using the persp function.

a Bayesian information criterion (BIC) is applied to choose the form of the mixture model:

```
R> library("mclust")
R> mc <- Mclust(faithful$waiting)
R> mc
best model: equal variance with 2 groups
averge/median classification uncertainty: 0.015 / 0
and the estimated means are
```

```
R> mc$mu
        1
80.09624 54.62190
with estimated standard deviation (found to be equal within both groups)
R> sqrt(mc$sigmasq)
[1] 5.867345
The proportion is \hat{p} = 0.64. The second package is called flexmix whose func-
tionality is described by Leisch (2004). A mixture of two normals can be fitted
using
R> library("flexmix")
R> fl <- flexmix(waiting ~ 1, data = faithful, k = 2)
with \hat{p} = 0.36 and estimated parameters
R> parameters(fl, component = 1)
(Intercept)
    54.6287
$sigma
[1] 5.895234
R> parameters(fl, component = 2)
$coef
(Intercept)
   80.09858
$sigma
[1] 5.871749
```

We can get standard errors for the five parameter estimates by using a bootstrap approach (see Efron and Tibshirani, 1993). The original data are slightly perturbed by drawing n out of n observations with replacement and those artificial replications of the original data are called bootstrap samples. Now, we can fit the mixture for each bootstrap sample and assess the variability of the estimates, for example using confidence intervals. Some suitable R code based on the Mclust function follows. First, we define a function that, for a bootstrap sample indx, fits a two-component mixture model and returns \hat{p} and the estimated means (note that we need to make sure that we always get an estimate of p, not 1-p):

```
R> opar <- as.list(opp$par)
R> rx <- seq(from = 40, to = 110, by = 0.1)
R> d1 <- dnorm(rx, mean = opar$mu1, sd = opar$sd1)
R> d2 <- dnorm(rx, mean = opar$mu2, sd = opar$sd2)
R> f <- opar$p * d1 + (1 - opar$p) * d2
R> hist(x, probability = TRUE, xlab = "Waiting times (in min.)",
+ border = "gray", xlim = range(rx), ylim = c(0,
+ 0.06), main = "")
R> lines(rx, f, lwd = 2)
R> lines(rx, dnorm(rx, mean = mean(x), sd = sd(x)),
+ lty = 2, lwd = 2)
R> legend(50, 0.06, legend = c("Fitted two-component mixture density",
+ "Fitted single normal density"), lty = 1:2,
+ bty = "n")
```

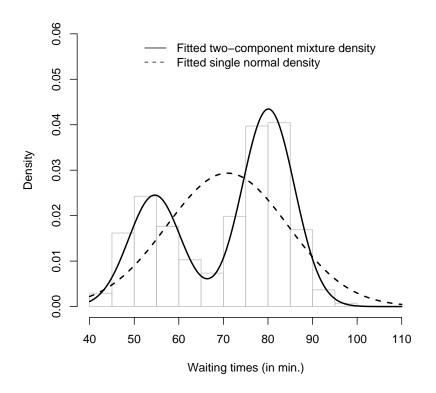


Figure 7.4 Fitted normal density and two-component normal mixture for geyser eruption data.

The function fit can now be fed into the boot function (Canty and Ripley, 2005) for bootstrapping (here 1000 bootstrap samples are drawn)

```
R> bootpara <- boot(faithful$waiting, fit, R = 1000)</pre>
```

We assess the variability of our estimates \hat{p} by means of adjusted bootstrap percentile (BCa) confidence intervals, which for \hat{p} can be obtained from

```
R> boot.ci(bootpara, type = "bca", index = 1)
BOOTSTRAP CONFIDENCE INTERVAL CALCULATIONS
Based on 1000 bootstrap replicates
boot.ci(boot.out = bootpara, type = "bca", index = 1)
Intervals :
            BCa
Level
     ( 0.3041,
                 0.4233 )
Calculations and Intervals on Original Scale
We see that there is a reasonable variability in the mixture model, however,
the means in the two components are rather stable, as can be seen from
R> boot.ci(bootpara, type = "bca", index = 2)
BOOTSTRAP CONFIDENCE INTERVAL CALCULATIONS
Based on 1000 bootstrap replicates
boot.ci(boot.out = bootpara, type = "bca", index = 2)
Intervals :
Level
            BCa
95% (53.42, 56.07)
Calculations and Intervals on Original Scale
for \hat{\mu}_1 and for \hat{\mu}_2 from
R> boot.ci(bootpara, type = "bca", index = 3)
BOOTSTRAP CONFIDENCE INTERVAL CALCULATIONS
Based on 1000 bootstrap replicates
CALL :
boot.ci(boot.out = bootpara, type = "bca", index = 3)
Intervals:
Level
            BCa
     (79.05, 81.01)
```

Finally, we show a graphical representation of both the bootstrap distribution of the mean estimates *and* the corresponding confidence intervals. For convenience, we define a function for plotting, namely

Calculations and Intervals on Original Scale

```
R> layout(matrix(1:2, ncol = 2))
R> bootplot(bootpara, 2, main = expression(mu[1]))
R> bootplot(bootpara, 3, main = expression(mu[2]))
```

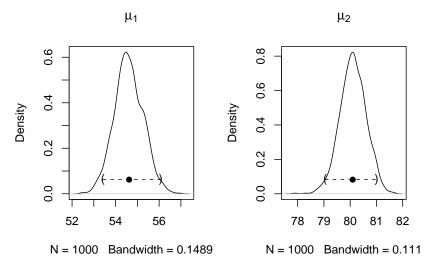


Figure 7.5 Bootstrap distribution and confidence intervals for the mean estimates of a two-component mixture for the geyser data.

```
R> bootplot <- function(b, index, main = "") {
        dens <- density(b$t[, index])
        ci <- boot.ci(b, type = "bca", index = index)$bca[4:5]
        est <- b$t0[index]
        plot(dens, main = main)
        y <- max(dens$y)/10
        segments(ci[1], y, ci[2], y, lty = 2)
        points(ci[1], y, pch = "(")
        points(ci[2], y, pch = ")")
        points(est, y, pch = 19)
        + }</pre>
```

The element t of an object created by boot contains the bootstrap replications of our estimates, i.e., the values computed by fit for each of the 1000 bootstrap samples of the geyser data. First, we plot a simple density estimate and then construct a line representing the confidence interval. We apply this function to the bootstrap distributions of our estimates $\hat{\mu}_1$ and $\hat{\mu}_2$ in Figure 7.5.

Bibliography

- Canty, A. and Ripley, B. D. (2005), boot: Bootstrap R (S-PLUS) Functions (Canty), URL http://CRAN.R-project.org, R package version 1.2-24. 9
- Efron, B. and Tibshirani, R. J. (1993), An Introduction to the Bootstrap, London, UK: Chapman & Hall/CRC. 7
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