Duration of Unemployment - Different Codings of Covariables

February 8, 2012

The unemployment data represent a contingency table with rows referring to gender and columns to duration of unemployment.

```
> unemployment <- matrix(c(403, 238, 167, 175), nrow=2, ncol=2)
> rownames(unemployment) <- c("male", "female")</pre>
> colnames(unemployment) <- c("<6 month",">6 month")
> unemployment
       <6 month >6 month
           403 167
male
female
             238
                      175
> rowSums(unemployment)
  male female
   570
          413
Calculation of odds and log-odds.
> ( odds_m <- 403/167 )</pre>
[1] 2.413174
> ( odds_w <- 238/175 )</pre>
[1] 1.36
> (log_odds_m < -log(403/167))
[1] 0.8809427
> ( log_odds_w <- log(238/175) )</pre>
[1] 0.3074847
For the fitting of a logit-model an alternative dataset is generated. First (0-1)-
coding is considered
> gender <- c(rep(1, 403+167), rep(0,238+175))
> unemp <- c(rep(1, 403), rep(0, 167), rep(1, 238), rep(0, 175))
```

```
For control, one can compute the crosstabulation of the generated data.
```

```
> table(gender, unemp)
      unemp
gender 0
             1
     0 175 238
     1 167 403
Fit of a logit model.
> bin <- glm(unemp ~ gender, family=binomial)</pre>
> summary(bin)
Call:
glm(formula = unemp ~ gender, family = binomial)
Deviance Residuals:
   Min
              1Q
                  Median
                                3Q
                                        Max
-1.5669 -1.3105
                 0.8327
                            0.8327
                                     1.0499
Coefficients:
            Estimate Std. Error z value Pr(>|z|)
(Intercept) 0.30748 0.09958 3.088 0.00202 **
gender
            0.57346
                        0.13559
                                 4.229 2.34e-05 ***
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
    Null deviance: 1270.3 on 982 degrees of freedom
Residual deviance: 1252.4 on 981 degrees of freedom
AIC: 1256.4
Number of Fisher Scoring iterations: 4
> bin$coef
(Intercept)
                 gender
  0.3074847
              0.5734580
> exp(bin$coef)
(Intercept)
                 gender
   1.360000
               1.774392
Now a dataset in effect-coding is created.
> gender_effect <- c(rep(1, 403+167), rep(-1,238+175))
For control, one can compute the crosstabulation of the generated data.
> table(gender_effect, unemp)
```

```
unemp
gender_effect 0
          -1 175 238
          1 167 403
Fit a logit model.
> bin_effect <- glm(unemp ~ gender_effect, family=binomial)</pre>
> summary(bin_effect)
Call:
glm(formula = unemp ~ gender_effect, family = binomial)
Deviance Residuals:
   Min
            1Q
                 Median
                               3Q
                                       Max
-1.5669 -1.3105
                0.8327
                          0.8327
                                    1.0499
Coefficients:
             Estimate Std. Error z value Pr(>|z|)
(Intercept)
             0.2867
                          0.0678 4.229 2.34e-05 ***
gender_effect
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
   Null deviance: 1270.3 on 982 degrees of freedom
Residual deviance: 1252.4 on 981 degrees of freedom
AIC: 1256.4
Number of Fisher Scoring iterations: 4
> bin_effect$coef
  (Intercept) gender_effect
   0.5942137
                0.2867290
> exp(bin_effect$coef)
  (Intercept) gender_effect
     1.811606
                  1.332063
  Now we consider education level as explanatory variable.
> unemp_level <- matrix(c(202, 307, 87, 45,
                         96, 162, 66, 18), nrow=4, ncol=2)
> colnames(unemp_level) <- c("Short term", "Long term")</pre>
> unemp_level
    Short term Long term
[1,]
           202
[2,]
           307
                     162
[3,]
            87
                      66
```

[4,]

45

18

```
> rowSums(unemp_level)
[1] 298 469 153 63
For the fitting of a logit-model a new dataset is generated. First (0-1)-coding is
considered.
> level <- factor(c(rep(1, 202+96), rep(2,307+162), rep(3,87+66), rep(4,45+18)))
> unemp_l <- c(rep(1, 202), rep(0, 96), rep(1, 307), rep(0, 162),
              rep(1, 87), rep(0, 66), rep(1, 45), rep(0, 18))
For control, one can compute the crosstabulation of the generated data.
> table(level, unemp_1)
     unemp_1
level 0 1
    1 96 202
    2 162 307
    3 66 87
    4 18 45
Fit a logit model on the data. Define the variable level as a factor with the
reference category 4.
> level <- relevel(level, ref=4)</pre>
> bin_l <- glm(unemp_l ~ level, family=binomial)</pre>
> summary(bin_1)
Call:
glm(formula = unemp_l ~ level, family = binomial)
Deviance Residuals:
   Min 1Q Median
                                3Q
                                        Max
-1.5829 -1.4581 0.8819 0.9206
                                     1.0626
Coefficients:
           Estimate Std. Error z value Pr(>|z|)
(Intercept) 0.9163 0.2789
                                 3.286 0.00102 **
            -0.1724
level1
                         0.3052 -0.565 0.57222
level2
             -0.2770
                         0.2953 -0.938 0.34818
level3
             -0.6400
                         0.3231 -1.981 0.04763 *
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
```

Null deviance: 1270.3 on 982 degrees of freedom Residual deviance: 1263.8 on 979 degrees of freedom

Number of Fisher Scoring iterations: 4

AIC: 1271.8

Now additionally quasi–variances can be computed. Therefore the function "qvcalc" from the "qvcalc"–library is used.

```
> library(qvcalc)
> qv<-qvcalc(bin_1,"level")
> summary(qv)
Model call: glm(formula = unemp_l ~ level, family = binomial)
Factor name: level
        estimate
                        SE
                              quasiSE
                                         quasiVar
    4 0.0000000 0.0000000 0.27888650 0.077777678
    1 -0.1723712 0.3051964 0.12396432 0.015367154
    2 -0.2770393 0.2953097 0.09710904 0.009430166
    3 -0.6400374 0.3231462 0.16323531 0.026645768
Worst relative errors in SEs of simple contrasts (%): 0 0 \,
Worst relative errors over *all* contrasts (%): 0 0
> plot(qv)
```

Intervals based on quasi standard errors

