glmpathcr: An **R** Package for Ordinal Response Prediction in High-dimensional Data Settings

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Abstract

This paper describes an R package, **glmpathcr**, that provides a function for fitting a penalized continuation ratio model when interest lies in predicting an ordinal response. The function, `glmpathcr` uses the coordinate descent fitting algorithm as implemented in `glmpath` and described by (Park and Hastie 2007a). Methods for extracting all estimated coefficients, extracting non-zero coefficient estimates, obtaining the predicted class, and obtaining the class-specific fitted probabilities have been implemented. Additionally, generic methods from **glmpath** including `summary`, `print`, and `plot` can be applied to a `glmpathcr` object.

Keywords: ordinal response, penalized models, LASSO, L$_1$ constraint, R.

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1. Introduction

High-throughput genomic experiments are frequently conducted for the purpose of examining whether genes are predictive of or significantly associated with phenotype. In many biomedical settings where histopathological or health status data are collected, phenotypic variables are recorded on an ordinal scale. Nevertheless, most often investigators neglect the ordinality of the phenotypic data and rather dichotomize the ordinal class than apply statistical methods suitable for two-class comparisons and predictions. This tendency to analyze ordinal data using dichotomous class methodologies may be due to the lack of available statistical methods and software for modeling an ordinal response in the presence of a high-dimensional covariate space. The approach of collapsing ordinal categories may neglect important information in the study (Armstrong and Sloan 1989).

A variety of statistical modeling procedures, namely, proportional odds, adjacent category, stereotype logit, and continuation ratio models can be used to predict an ordinal response. In this paper, we focus attention to the continuation ratio model because its likelihood can be easily re-expressed such that existing software can be readily adapted and used for model fitting. Suppose for each observation, $i = 1, \ldots, n$, the response Y_i belongs to one ordinal class $k = 1, \ldots, K$ and $\mathbf{x}_i$ represents a p-length vector of covariates. The backward formulation of the continuation ratio models the logit as

$$\texttt{logit}\left(P(Y = k | Y \leq k, \mathbf{X} = \mathbf{x})\right) = \alpha_k + \boldsymbol{\beta}_k^T \mathbf{x} \tag{1}$$

whereas the forward formulation models the logit as

$$\texttt{logit}\left(P(Y = k | Y \geq k, \mathbf{X} = \mathbf{x})\right) = \alpha_k + \boldsymbol{\beta}_k^T \mathbf{x}. \tag{2}$$

Rather than describe both formulations in detail, here we present the backward formulation, which is commonly used when progression through disease states from none, mild, moderate, severe is represented by increasing integer values, and interest lies in estimating the odds of more severe disease compared to less severe disease (Bender and Benner 2000). Therefore for $i = 1, \ldots, n$ we can construct a vector $\mathbf{y}_i$ from Y_i to represent ordinal class membership, such that $\mathbf{y}_i = (y_{i1}, y_{i2}, \ldots, y_{iK})^T$, where $y_{ik} = 1$ if the response is in category k and 0 otherwise, so that $n_i = \sum_{k=1}^{K} y_{ik} = 1$. Using the logit link, the equation representing the conditional probability for class k is

$$\delta_k(\mathbf{x}) = P(Y = k | Y \le k, \mathbf{X} = \mathbf{x}) = \frac{\exp(\alpha_k + \boldsymbol{\beta}^T \mathbf{x})}{1 + \exp(\alpha_k + \boldsymbol{\beta}^T \mathbf{x})}. \tag{3}$$

The likelihood for the continuation ratio model is then the product of conditionally independent binomial terms (Cox 1975), which is given by

$$L(\boldsymbol{\beta} | \mathbf{y}, \mathbf{x}) = \prod_{i=1}^{n} \delta_2^{y_{i2}} (1 - \delta_2)^{1 - \sum_{k=2}^{K} y_{ik}} \times \cdots \times \delta_K^{y_{iK}} (1 - \delta_K)^{1 - y_{iK}} \tag{4}$$

where here we have simplified our notation by not explicitly including the dependence of the conditional probability δ_k on $\mathbf{x}$. Further, simplifying our notation to let $\boldsymbol{\beta}$ represent the vector containing both the thresholds $(\alpha_2, \ldots, \alpha_K)$ and the log odds $(\beta_1, \ldots, \beta_p)$ for all $K - 1$ logits, the full parameter vector is

$$\boldsymbol{\beta} = (\alpha_2, \beta_{21}, \beta_{22}, \ldots, \beta_{2p}, \ldots, \alpha_K, \beta_{K,1}, \beta_{K,2}, \ldots, \beta_{K,p})^T \tag{5}$$

which is of length $(K-1)(p+1)$. As can be seen from equation 4, the likelihood can be factored into $K - 1$ independent likelihoods, so that maximization of the independent likelihoods will lead to an overall maximum likelihood estimate for all terms in the model (Bender and Benner 2000). A model consisting of $K - 1$ different $\boldsymbol{\beta}$ vectors may be overparameterized so to simplify, one commonly fits a constrained continuation model, which includes the $K - 1$ thresholds $(\alpha_2, \ldots, \alpha_K)$ and one common set of p slope parameters, $(\beta_1, \ldots, \beta_p)$. To fit a constrained continuation ratio model, the original dataset can be restructured by forming $K - 1$ subsets, where for classes $k = 2, \ldots, K$, the subset contains those observations in the original dataset up to class k. Additionally, for the kth subset, the outcome is dichotomized as $y = 1$ if the ordinal class is k and $y = 0$ otherwise. Furthermore, an indicator is constructed for each subset representing subset membership. Thereafter the $K - 1$ subsets are appended to form the restructured dataset, which represents the $K - 1$ conditionally independent datasets in equation 4. Applying a logistic regression model to this restructured dataset yields an L_1 constrained continuation ratio model.

2. Penalized Models

For datasets where the number of covariates p exceeds the sample size n, the backwards stepwise procedure cannot be undertaken. Furthermore, for any problem using a forward selection procedure the discrete variable inclusion process can exhibit high variance. Moreover, for high-dimensional covariate spaces, the best subset procedure is computationally prohibitive. Two penalized methods, ridge and L_1 penalization, places a penalty on a function of the coefficient estimates, thereby permitting a model fit even for high-dimensional

data Tibshirani (1996, 1997). A generalization of these penalized models can be expressed as,

$$\tilde{\beta} = \arg\min_{\beta}(\sum_{i=1}^{n}(y_i - \beta_0 - \sum_{j=1}^{p} x_{ij}\beta_j)^2 + \lambda \sum_{j=1}^{p} |\beta_j|^q) \tag{6}$$

for $q \geq 0$. When $q = 1$ we have the an L_1 penalized model, when $q = 2$ we have ridge regression. Values of $q \in (1, 2)$ provide a compromise between the L_1 and ridge penalized models. Because when $q > 1$ coefficients are no longer set exactly equal to 0, the elastic net penalty was introduced

$$\lambda \sum_{j=1}^{p}(\alpha\beta_j^2 + (1 - \alpha)|\beta_j|). \tag{7}$$

3. Implementation

The **glmpathcr** package was written in the R programming environment (R Development Core Team 2009) and depends on the **glmpath** package (Park and Hastie 2007b). Similar to the **Design** package which includes a function `cr.setup` for restructuring a dataset for fitting a forward continuation ratio model, in this package the model is fit by restructuring the dataset then passing the restructured dataset to a penalized logistic regression fitting function. However, unlike `cr.setup` which produces an object of class `list` from which the response and restructured independent variables are extracted and passed to a model fitting algorithm, in the **glmpathcr** package the restructuring functions are transparent to the user. Specifically, the **glmpathcr** package fits either a forward or backward (default) penalized constrained continuation ratio model by specification of `method="forward"` in the `glmpathcr` call. The glmpathcr function restructures the dataset to represent the $K - 1$ conditionally independent likelihoods and then fits the penalized continuation ratio model using the **glmpath** framework. Therefore, the predictor-corrector fitting procedure used by the `glmpath` function in the **glmpath** package is used in fitting the penalized continuation ratio model when invoking `glmpathcr`. This allows fitting a penalized model for situations where the number of covariates p exceed the sample size n. In addition, methods for extracting the best fitting model from the path using AIC and BIC criteria, obtaining predicted class and fitted class probabilities, and returning coefficient estimates were written in addition to adapting the `print`, `summary`, and `plot` methods from **glmpath** for a `glmpathcr` object.

4. Example

The **glmpathcr** package includes a filtered microarray dataset `diabetes` in which asymptomatic males not previously diagnosed with Type II diabetes were enrolled and subsequently were cross-classified as either normal controls (N=8), having impaired fasting glucose (N=7), or as Type II diabetics (N=9) based on a fasting glucose intolerance test. From the code below we can see that the classification variable is stored as `y` in the first column of the `diabetes` `data.frame`; all subsequent columns are the 11,066 Illumina probes having no negative expression values. In fitting the model we can extract the covariates into an object `x` and the ordinal outcome into the object `y`. The code for fitting a backward (default) continuation ratio model is given by

```
> library(glmpathcr)
> data(diabetes)
> dim(diabetes)

[1]    24 11067

> names(diabetes)[1:10]

 [1] "y"           "ILMN_1343291" "ILMN_1651228" "ILMN_1651229" "ILMN_1651236"
 [6] "ILMN_1651254" "ILMN_1651262" "ILMN_1651268" "ILMN_1651278" "ILMN_1651286"

> summary(diabetes$y)

                control impaired fasting glucose       type 2 diabetes
                      8                         7                     9

> x <- diabetes[, 2:dim(diabetes)[2]]
> y <- diabetes$y
> fit <- glmpathcr(x,y)
```

As with `glmpath` model objects, methods such as `summary` and `plot` can be applied to `glmpathcr` model objects, which are helpful for selecting the step at which to select the final model from the solution path.

```
> summary(fit)

        Df    Deviance       AIC       BIC
Step 1   3 5.248273e+01  58.48273  62.01690
Step 2   4 2.334393e+01  33.34393  39.23419
Step 5   5 2.047337e+01  30.47337  36.36364
Step 7   6 2.013650e+01  32.13650  39.20482
Step 10  7 1.587273e+01  29.87273  38.11911
Step 14  8 7.939954e+00  23.93995  33.36438
Step 17  9 7.530910e+00  25.53091  36.13339
Step 19 10 7.327910e+00  27.32791  39.10845
Step 21 11 7.055827e+00  29.05583  42.01442
Step 24 12 6.608876e+00  30.60888  44.74552
Step 28 13 3.583363e+00  29.58336  44.89806
Step 31 14 3.176650e+00  31.17665  47.66940
Step 34 15 2.491539e+00  32.49154  50.16235
Step 38 16 1.803196e+00  33.80320  52.65206
Step 40 17 1.718770e+00  35.71877  55.74569
Step 45 18 4.362169e-01  36.43622  57.64119
Step 47 18 9.654670e-02  36.09655  57.30152
Step 52 18 1.378866e-02  36.01379  57.21876
Step 53 19 1.013108e-02  40.01013  63.57121
```

```
Step 54   19 6.701491e-03 40.00670   63.56778
Step 55   20 6.639347e-03 40.00664   63.56772
Step 57   21 4.759662e-03 42.00476   66.74389
Step 58   22 4.746844e-03 44.00475   69.92193
Step 60   23 4.429498e-03 46.00443   73.09967
Step 62   24 4.097755e-03 48.00410   76.27739
Step 63   24 3.732833e-03 48.00373   76.27702
Step 64   23 3.581531e-03 46.00358   73.09882
Step 65   22 3.121086e-03 44.00312   69.92031
Step 66   21 3.064065e-03 42.00306   66.74219
Step 67   20 3.021735e-03 40.00302   63.56410
Step 68   19 2.718302e-03 38.00272   60.38574
Step 69   18 2.469391e-03 36.00247   57.20744
Step 70   17 2.348876e-03 34.00235   54.02926
Step 72   17 2.207919e-03 34.00221   54.02912
Step 74   18 2.103975e-03 38.00210   60.38513
Step 75   19 2.089284e-03 38.00209   60.38511
Step 76   20 2.057258e-03 40.00206   63.56313
Step 78   21 1.984285e-03 42.00198   66.74111
Step 80   22 1.952069e-03 44.00195   69.91914
Step 81   22 1.943316e-03 44.00194   69.91913
Step 87   22 1.663457e-03 46.00166   73.09690
Step 89   23 1.475852e-03 46.00148   73.09671
Step 91   24 1.419015e-03 48.00142   76.27471
Step 93   25 1.392416e-03 50.00139   79.45274
Step 94   25 1.372555e-03 50.00137   79.45272
Step 95   25 1.318877e-03 50.00132   79.45266
Step 97   26 1.248303e-03 52.00125   82.63065
Step 98   27 1.221304e-03 56.00122   88.98673
Step 99   28 1.167642e-03 58.00117   92.16473
Step 100  28 1.161706e-03 56.00116   88.98667
Step 101  27 1.135018e-03 54.00114   85.80859
Step 102  27 1.130148e-03 54.00113   85.80858
Step 103  27 1.121402e-03 54.00112   85.80857
Step 104  27 1.107786e-03 54.00111   85.80856
Step 105  28 1.102616e-03 56.00110   88.98661
Step 107  29 1.044026e-03 58.00104   92.16461
Step 108  30 1.008927e-03 60.00101   95.34262
Step 109  31 9.024878e-04 62.00090   98.52057
Step 110  32 8.992738e-04 64.00090  101.69862
Step 112  33 8.587383e-04 66.00086  104.87664
Step 113  34 8.524961e-04 68.00085  108.05468
Step 114  35 8.412025e-04 70.00084  111.23273
Step 115  35 8.398308e-04 70.00084  111.23272
Step 117  35 7.787257e-04 70.00078  111.23266
Step 118  36 7.567521e-04 72.00076  114.41069
Step 120  37 6.878064e-04 74.00069  117.58868
```

```
Step 121 38 6.849830e-04 76.00068 120.76673
```

```
> plot(fit, xvar = "step", type = "bic")
```
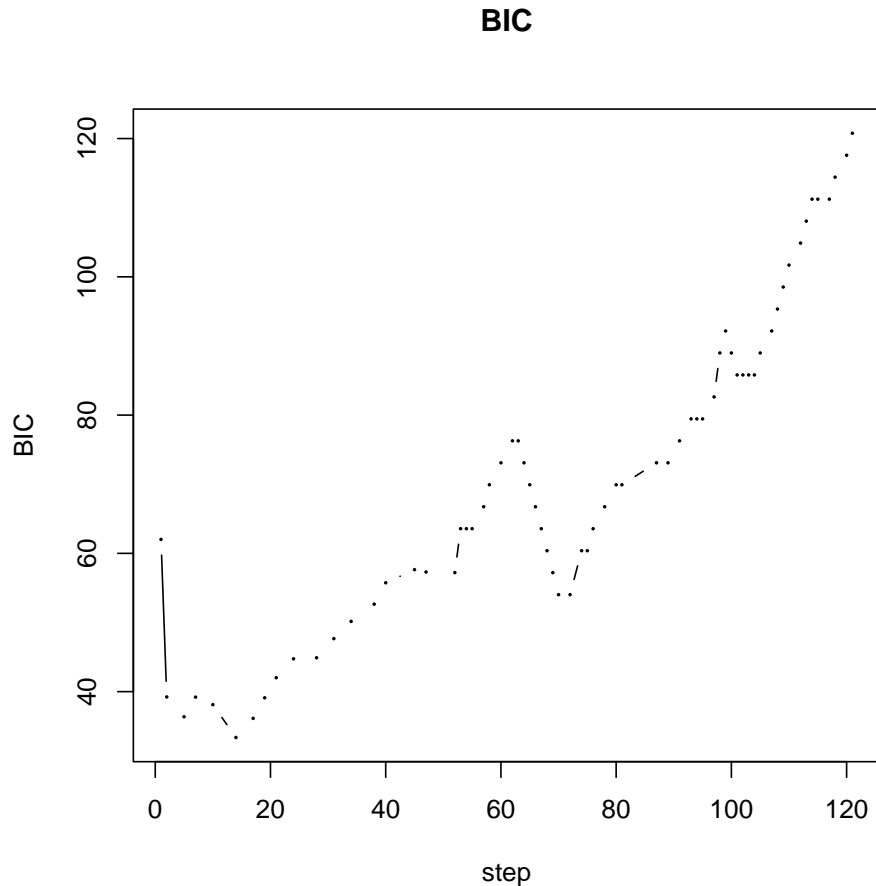
BIC

Figure 1: Plot of regularization path for `glmpathcr` object using simulated dataset, `data`.

Note that when plotting, the horizontal axis can be `norm`, `lambda`, or `step`, however extractor functions for `glmpathcr` generally require the step to be selected, so we have selected `xvar = "step"` in this example. The vertical axis can be coefficients, aic or bic. As one can see, there is a multitude of models fit from one call to `glmpathcr`. To faciliate extraction of best fitting models using commonly used criterion, the `model.select` function can be used. The `model.select` function extracts the best fitting model from the solution path, where the `which` parameter allows one to select either AIC or by default, BIC.

```
> BIC.step <- model.select(fit)
> BIC.step
```

```
[1] 14
```

```
> AIC.step <- model.select(fit, which = "AIC")
> AIC.step
```

```
[1] 14
```

In this example, the minimum BIC corresponds to a 8 degree of freedom model.

The `coef` function returns all estimated coefficients for a `glmpathcr` fitted model, where the model selected is indicated by step number, `s`. The `nonzero.coef` function returns only those non-zero coefficient estimates for a selected model.

```
> coefficients<-coef(fit, s=BIC.step)
> sum(coefficients!=0)
```

```
[1] 8
```

```
> nonzero.coef(fit, s=BIC.step)
```

```
    Intercept   ILMN_1705116   ILMN_1733757   ILMN_1758311   ILMN_1759232
18.3204861995   0.0334955981  -0.0003751318   0.0022607450  -0.0347775942
 ILMN_2100437            cp1            cp2
-0.0001679020   2.0164341494  -2.0164341493
```

Note that the `glmpathcr` function fits a penalized constrained continuation ratio model; therefore for K classes, there will be $K - 1$ intercepts representing the cutpoints between adjacent classes. In this package, the nomenclature for these cutpoints is to use "cp$k$" where $k = 1, \ldots, K - 1$. In this dataset, $K = 3$ so the intercepts are `cp1` and `cp2` with `Intercept` being an offset. The probe having the largest absolute coefficient estimate, `ILMN_1759232`, corresponds to the insulin receptor substrate 1 (IRS1) gene which is biologically meaningful.

Continuation ratio models predicts conditional probabilities so a new method to extract the fitted probabilities and predicted class was created. The `predict` and `fitted` functions are equivalent, and return either the predicted class or the fitted probabilities from the penalized continuation ratio model for a `glmpathcr` object. The user is required to supply the fitted `glmpathcr` model object, a data matrix `newx` that is either the same as the training data or an independent dataset having the same number and order of covariates as the training data, a vector `newy` that provides the class labels of the ordinal response. These functions extract the fitted values for the best fitting model using the BIC criteria by default, which can be changed to extracting the best fitting AIC model by supplying `which="AIC"`. By default, the predicted class is output. If one desired the fitted class-specific probabilities from the model, the `type="probs"` argument should be supplied.

```
> pred <- predict(fit)
> table(pred, y)
```

```
                          y
pred                       control impaired fasting glucose type 2 diabetes
  control                        8                        0               0
  impaired fasting glucose       0                        7               0
  type 2 diabetes                0                        0               9
```

```
> pred <- predict(fit, type="probs")
> pred

          control impaired fasting glucose type 2 diabetes
 [1,] 8.178419e-01              0.17830887      0.0038492509
 [2,] 9.239276e-01              0.07464256      0.0014297981
 [3,] 8.947141e-01              0.10324487      0.0020410068
 [4,] 9.530268e-01              0.04611633      0.0008568896
 [5,] 8.835483e-01              0.11416679      0.0022848810
 [6,] 9.070107e-01              0.09121017      0.0017791192
 [7,] 8.672968e-01              0.13005259      0.0026506120
 [8,] 8.784051e-01              0.11919565      0.0023992189
 [9,] 7.140319e-02              0.76841557      0.1601812423
[10,] 2.703789e-01              0.68655530      0.0430657811
[11,] 5.059698e-02              0.74290859      0.2064944306
[12,] 9.145857e-02              0.77758826      0.1309531684
[13,] 5.568355e-02              0.75132840      0.1929880462
[14,] 8.375911e-02              0.77530692      0.1409339747
[15,] 1.423227e-01              0.77016247      0.0875147915
[16,] 3.278482e-04              0.12689758      0.8727745672
[17,] 4.987756e-04              0.15400425      0.8454969782
[18,] 1.249965e-04              0.08046247      0.9194125351
[19,] 1.588154e-04              0.09020377      0.9096374108
[20,] 1.239401e-04              0.08013625      0.9197398137
[21,] 1.891746e-04              0.09801511      0.9017957105
[22,] 8.823833e-04              0.19912988      0.7999877353
[23,] 6.318903e-05              0.05792144      0.9420153759
[24,] 1.585017e-04              0.09011893      0.9097225673
```

For illustrative purposes, a forward continuation ratio model can be fit using the syntax

```
> fit <- glmpathcr(x, y, method="forward")
```

As before, the parameter estimates corresponding to the model attaining the minimum BIC can be extracted using the following code.

```
> coefficients<-coef(fit, s=BIC.step)
> nonzero.coef(fit, s=BIC.step)

    Intercept  ILMN_1701911  ILMN_1705116  ILMN_1733757  ILMN_1759232
-8.878742e+00  6.455187e-03 -2.238348e-02  2.963308e-04  2.964169e-02
  ILMN_2100437           cp1           cp2
 3.320597e-05 -1.601889e+00  1.601889e+00
```

and the predicted class can be obtained using

```
> pred <- predict(fit)
> table(pred, y)
```

```
                          y
pred                      control impaired fasting glucose type 2 diabetes
  control                       8                        1               0
  impaired fasting glucose      0                        6               0
  type 2 diabetes               0                        0               9
```

Summary

Herein we have described the **glmpathcr** package which works in conjunction with the **glmpath** package in the R programming environment. The package provides methods for fitting either a forward or backward penalized continuation ratio model. Moreover, the likelihood-based penalized continuation ratios models have been demonstrated to have good performance when applied to microarray gene expression datasets (Archer and Williams 2012) in comparison to corresponding penalized Bayesian continuation ratio models (Kiiveri 2008). A similar package, **glmnetcr**, which uses the **glmnet** fitting algorithm for fitting a penalized constrained continuation ratio model has also been developed and is available for download from the Comprehensive R Archive Network. Functions for extracting coefficients, extracting non-zero coefficients, and obtaining fitted probabilities and predicted class in the **glmnetcr** package follow those in **glmpathcr** and both packages have similar performance (Archer and Williams 2012). Therefore either the **glmpathcr** or **glmnetcr** package should be helpful when predicting an ordinal response for datasets where the number of covariates exceeds the number of available samples.

Acknowledgments

References

Archer KJ, Williams AA (2012). "L_1 penalized continuation ratio models for ordinal response prediction using high-dimensional datasets." *Statistics in Medicine*, **31**, 1464–1474.

Armstrong B, Sloan M (1989). "Ordinal regression models for epidemiologic data." *American Journal of Epidemiology*, **129**, 191–204.

Bender R, Benner A (2000). "Calculating ordinal regression models in SAS and S-Plus." *Biometrical Journal*, **42**, 677–699.

Cox D (1975). "Partial likelihood." *Biometrika*, **62**, 269–276.

Kiiveri HT (2008). "A general approach to simultaneous model fitting and variable elimination in response models for biological data with many more variables than observations." *BMC Bioinformatics*, **9**, 195.

Park M, Hastie T (2007a). "L1-regularization path algorithm for generalized linear models." *Journal of the Royal Statistical Society, B*, **64**, 659–677.

Park MY, Hastie T (2007b). *glmpath: L1 Regularization Path for Generalized Linear Models and Cox Proportional Hazards Model.* R package version 0.94.

R Development Core Team (2009). *R: A Language and Environment for Statistical Computing.* R Foundation for Statistical Computing, Vienna, Austria. ISBN 3-900051-07-0, URL http://www.R-project.org.

Tibshirani R (1996). "Regression shrinkage and selection via the Lasso." *Journal of the Royal Statistical Society, B*, **58**, 267–288.

Tibshirani R (1997). "The lasso method for variable selection in the Cox model." *Statistics in Medicine*, **16**, 385–395.

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