Multicollinearity, identification, and estimable functions

Simen Gaure

ABSTRACT. Since there is quite a lot of confusion here and there about what happens when factors are collinear; here is a walkthrough of the identification problems which may arise in models with many dummies, and how lfe handles them. (Or, at the very least, attempts to handle them).

1. Context

The **lfe** package is used for ordinary least squares estimation, i.e. models which conceptually may be estimated by 1m as

```
lm(y ~ x1 + x2 + ... + f1 + f2 + ... + fn)
```

where f1,f2,...,fn are factors. The standard method is to introduce a dummy variable for each level of each factor. This is too much as it introduces multicollinearities in the system. Conceptually, the system may still be solved, but there are many different solutions. In all of them, the difference between the coefficients for each factor will be the same.

The ambiguity is typically solved by removing a single dummy variable for each factor, this is termed a reference. This is like forcing the coefficient for this dummy variable to zero, and the other levels are then seen as relative to this zero. Other ways to solve the problem is to force the sum of the coefficients to be zero, or one may enforce some other constraint, typically via the contrasts argument to lm. The default in lm is to have a reference level in each factor, and a common intercept term.

In **lfe** the same estimation can be performed by

$$felm(y ~ x1 + x2 + ... + G(f1) + G(f2) + ... + G(fn))$$

Since felm conceptually does exactly the same as lm, the contrasts approach may work there too. Or rather, it is actually not necessary that felm handles it at all, it is only necessary if one needs to fetch the coefficients for the factor levels with getfe.

Ife is intended for very large datasets, with factors with many levels. Then the approach with a single constraint for each factor may sometimes not be sufficient. The standard example in the econometrics literature (see e.g. [2]) is the case with two factors, one for individuals, and one for firms these individuals work for, changing jobs now and then. What happens in practice is that the labour market may be disconnected, so that one set of individuals move between one set of firms, and

1

another (disjoint) set of individuals move between some other firms. This happens for no obvious reason, and is data dependent, not intrinsic to the model. There may be several such components. I.e. there are more multicollinearities in the system than the obvious ones. In such a case, there is no way to compare coefficients from different connected components, it is not sufficient with a single individual reference. The problem may be phrased in graph theoretic terms (see e.g. [1, 3, 4]), and it can be shown that it is sufficient with one reference level in each of the connected components. This is what **lfe** does, in the case with two factors it identifies these components, and force one level to zero in one of the factors.

In the examples below, rather small randomly generated datasets are used. **Ife** is hardly the best solution for these problems, they are solely used to illustrate some concepts. I can assure the reader that no CPUs, sleeping patterns, romantic relationships, trees or cats, nor animals in general, were harmed during data collection and analysis.

2. Identification with two factors

In the case with two factors, i.e. two G() terms in the model, identification is well-known. getfe will partition the dataset into connected components, and introduce a reference level in each component:

```
library(lfe)
## Loading required package:
                               Matrix
## Loading required package:
                               Rcqmin
## Loading required package:
                               numDeriv
set.seed(42)
x1 \leftarrow rnorm(20)
f1 <- sample(8, length(x1), replace = TRUE)/10
f2 <- sample(8, length(x1), replace = TRUE)/10
e1 \leftarrow sin(f1) + 0.02 * f2^2 + rnorm(length(x1))
y \leftarrow 2.5 * x1 + (e1 - mean(e1))
summary(est \leftarrow felm(y \sim x1 + G(f1) + G(f2)))
##
## Call:
      felm(formula = y ~ x1 + G(f1) + G(f2))
##
##
## Residuals:
##
     Min
              1Q Median
                             30
                                   Max
## -1.399 -0.279 0.000 0.436
                                 0.981
##
## Coefficients:
      Estimate Std. Error t value Pr(>|t|)
##
## x1
                    0.377
                              6.71
                                   0.0011 **
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.13 on 5 degrees of freedom
## Multiple R-squared: 0.973 Adjusted R-squared: 0.894
## F-statistic: 13.1 on 14 and 5 DF, p-value: 0.0051
```

We examine the estimable function produced by efactory.

```
ef <- efactory(est)</pre>
is.estimable(ef, est$fe)
## [1] TRUE
getfe(est)
##
            effect obs comp fe idx
## f1.0.1 0.376275
                   2
                         1 f1 0.1
## f1.0.2 -0.081100
                         2 f1 0.2
                   3
## f1.0.3 -0.686880
                         1 f1 0.3
## f1.0.4 0.573177 4
                        1 f1 0.4
                   2
## f1.0.5 0.479142
                       1 f1 0.5
## f1.0.6 1.413020
                   3
                        1 f1 0.6
## f1.0.7 0.844956
                   1
                         2 f1 0.7
## f1.0.8 0.926434
                   4 1 f1 0.8
## f2.0.1 -0.004011
                         1 f2 0.1
                    3
## f2.0.2 0.000000
                   5
                         1 f2 0.2
## f2.0.3 -1.518667
                         1 f2 0.3
                   1
## f2.0.4 0.000000
                   2
                         2 f2 0.4
## f2.0.5 -1.894524
                    2
                         1 f2 0.5
## f2.0.6 -0.884319
                    3
                         1 f2 0.6
## f2.0.7 -0.609110
                    3
                         1 f2 0.7
## f2.0.8 -0.968652
                   1 1 f2 0.8
```

As we can see from the comp entry, there are two components, with f1=0.2, f1=0.7 and f2=0.4. A reference is introduced in each of the components, i.e. f2.0.2=0 and f2.0.4=0. If we look at the dataset, the component structure becomes clearer:

```
data.frame(f1, f2, comp = est$cfactor)
##
       f1 f2 comp
## 1 0.4 0.6
                1
## 2 0.4 0.8
                 1
## 3 0.1 0.7
                 1
## 4 0.8 0.5
                 1
## 5 0.4 0.7
                 1
## 6 0.8 0.2
                 1
## 7
     0.8 0.3
                 1
## 8 0.6 0.7
                 1
## 9 0.8 0.6
                 1
## 10 0.5 0.2
## 11 0.3 0.1
                 1
## 12 0.3 0.2
                 1
## 13 0.4 0.2
                 1
## 14 0.7 0.4
                 2
## 15 0.1 0.2
                 1
## 16 0.6 0.6
                 1
## 17 0.6 0.1
                 1
## 18 0.2 0.4
                 2
```

```
## 19 0.3 0.5 1
## 20 0.5 0.1 1
```

Observation 14 and 18 belong to component 2; no other observation has f1=0.7, f1=0.2 or f2=0.4, thus it is clear that coefficients for these can not be compared to other coefficients. 1m is silent about this component structure, hence coefficients are hard to interpret. Though, predictive properties and residuals are the same:

```
f1 <- factor(f1)</pre>
f2 <- factor(f2)
summary(lm(y ~ x1 + f1 + f2))
##
## Call:
## lm(formula = y ~ x1 + f1 + f2)
##
## Residuals:
##
    1
                              3
                                       4
                                                 5
##
   4.18e-01 4.72e-16 6.36e-01 4.91e-01 -1.40e+00 -2.26e-01 1.94e-16
         8 9
                            10
                                  11 12
                                                    13
   7.64e-01 -2.64e-01 2.05e-01 8.15e-01 -3.24e-01 9.81e-01
##
                                                             1.53e-16
##
        15
             16
                             17
                                      18
                                                19
## -6.36e-01 -1.54e-01 -6.10e-01 -1.25e-16 -4.91e-01 -2.05e-01
## Coefficients: (1 not defined because of singularities)
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 0.37226
                         1.21296
                                    0.31
                                           0.7713
               2.53052
                          0.37710
                                     6.71
                                           0.0011 **
## x1
## f10.2
              -0.45336
                          2.02866
                                   -0.22
                                           0.8320
## f10.3
              -1.06316
                          1.24867
                                   -0.85
                                           0.4334
## f10.4
               0.19690
                                  0.18
                                           0.8619
                         1.07494
## f10.5
               0.10287
                          1.27009
                                    0.08
                                           0.9386
## f10.6
                                     0.90
                                           0.4101
               1.03674
                          1.15377
## f10.7
               0.47269
                         1.67007
                                    0.28
                                           0.7885
## f10.8
               0.55016
                       1.25945
                                  0.44
                                           0.6805
## f20.2
               0.00401
                          0.96913
                                    0.00
                                           0.9969
## f20.3
              -1.51466
                          1.56345
                                   -0.97
                                           0.3771
## f20.4
                    NA
                              NA
                                     NA
                                               NA
## f20.5
              -1.89051
                          1.50465
                                   -1.26
                                           0.2645
## f20.6
              -0.88031
                          1.11170
                                   -0.79
                                           0.4643
## f20.7
              -0.60510
                          1.09728
                                   -0.55
                                           0.6051
## f20.8
              -0.96464
                          1.58571
                                   -0.61
                                           0.5695
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 1.13 on 5 degrees of freedom
## Multiple R-squared: 0.973, Adjusted R-squared: 0.899
## F-statistic: 13.1 on 14 and 5 DF, p-value: 0.0051
```

3. Identification with three or more factors

In the case with three or more factors, there is no general intuitive theory (yet) for handling identification problems. **Ife** resorts to the simple-minded approach that non-obvious multicollinearities arise among the first two factors, and assumes it is sufficient with a single reference level for each of the remaining factors, i.e. that they in principle could be specified as ordinary dummies. In other words, the order of the factors in the model specification is important. A typical example would be 3 factors; individuals, firms and education:

```
est <- felm(logwage ~ x1 + x2 + G(id) + G(firm) + G(edu))
getfe(est)</pre>
```

This will result in the same number of references as if using the model

```
logwage \sim x1 + x2 + G(id) + G(firm) + edu
```

though it may run faster (or slower). Alternatively, one could specify the model as

```
logwage \sim x1 + x2 + G(firm) + G(edu) + G(id)
```

This would not account for a partioning of the labour market along individual/firm, but along firm/education, using a single reference level for the individuals. In this example, there is some reason to suspect that it is not sufficient, depending on how edu is specified. There exists no general scheme that sets up suitable reference groups when there are more than two factors. It may happen that the default is sufficient. The function getfe will check whether this is so, and it will yield a warning about 'non-estimable function' if not. With some luck it may be possible to rearrange the order of the factors to avoid this situation.

There is nothing special with lfe in this respect. You will meet the same problem with lm, it will remove a reference level (or dummy-variable) in each factor, but the system will still contain multicollinearities. You may remove reference levels until all the multicollinearities are gone, but there is no obvious way to interpret the resulting coefficients.

To illustrate, the classical example is when you include a factor for age (in years), a factor for observation year, and a factor for year of birth. You pick a reference individual, e.g. age=50, year=2013 and birth=1963, but this is not sufficient to remove all the multicollinearities. If you analyze this problem (see e.g. [6]) you will find that the coefficients are only identified up to linear trends. You may force the linear trend between birth=1963 and birth=1990 to zero, by removing the reference level birth=1990, and the system will be free of multicollinearities. In this case the birth coefficients have the interpretation as being deviations from a linear trend between 1963 and 1990, though you do not know which linear trend. The age and year coefficients are also relative to this same unknown trend.

In the above case, the multicollinearity is obviously built into the model, and it is possible to remove it and find some intuitive interpretation of the coefficients. In the general case, when either lm or getfe reports a handful of non-obvious spurious multicollinearites between factors with many levels, you probably will not be able to find any reasonable way to interpret coefficients. Of course, certain linear

combinations of coefficients will be unique, i.e. estimable, and these may be found by e.g. the procedures in [5, 8], but the general picture is muddy.

Ife does not provide a solution to this problem, however, getfe will still provide a vector of coefficients which results from finding a non-unique solution to a certain set of equations. To get any sense from this, an estimable function must be applied. The simplest one is to pick a reference for each factor and subtract this coefficient from each of the other coefficients in the same factor, and add it to a common intercept, however in the case this does not result in an estimable function, you are out of luck. If you for some reason believe that you know of an estimable function, you may provide this to getfe via the ef-argument. There is an example in the getfe documentation. You may also test it for estimability with the function is.estimable, this is a probabilistic test which almost never fails (see [4, Remark 6.2]).

4. Specifying an estimable function

A model of the type

```
y ~ x1 + x2 + f1 + f2 + f3
```

may be written in matrix notation as

$$(1) y = X\beta + D\alpha + \epsilon,$$

where X is a matrix with columns x1 and x2 and D is matrix of dummies constructed from the levels of the factors f1,f2,f3. Formally, an estimable function in our context is a matrix operator whose row space is contained in the row space of D. That is, an estimable function may be written as a matrix. Like the contrasts argument to 1m. However, the lfe package uses an R-function instead. That is, felm is called first, it uses the Frisch-Waugh-Lovell theorem to project out the $D\alpha$ term from (1) (see [4, Remark 3.2]):

```
est <- felm(y ~ x1 + x2 + G(f1) + G(f2) + G(f3))
```

This yields the parameters for x1 and x2, i.e. β . To find $\hat{\alpha}$, the parameters for the levels of f1,f2,f3, getfe solves a certain linear system (see [4, eq. (14)]):

$$(2) D\gamma = \rho$$

where the vector ρ can be computed when we have $\hat{\beta}$. This does not identify γ uniquely, we have to apply an estimable function to γ . The estimable function F is characterized by the property that $F\gamma_1 = F\gamma_2$ whenever γ_1 and γ_2 are solutions to equation (2). Rather than coding F as a matrix, **lfe** codes it as a function. It is of course possible to let the function apply a matrix, so this is not a material distinction. So, let's look at an example of how an estimable function may be made:

```
library(lfe)
x1 <- rnorm(100)
f1 <- sample(7, 100, replace = TRUE)
f2 <- sample(8, 100, replace = TRUE)/8
f3 <- sample(10, 100, replace = TRUE)/10
e1 <- sin(f1) + 0.02 * f2^2 + 0.17 * f3^3 + rnorm(100)
y <- 2.5 * x1 + (e1 - mean(e1))</pre>
```

```
summary(est \leftarrow felm(y \sim x1 + G(f1) + G(f2) + G(f3)))
##
## Call:
##
     felm(formula = y ~ x1 + G(f1) + G(f2) + G(f3))
## Residuals:
     Min
               1Q Median
                              3Q
##
## -1.8869 -0.7252 -0.0788 0.7558 2.3050
## Coefficients:
##
     Estimate Std. Error t value Pr(>|t|)
## x1 2.354 0.112
                           21 <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.08 on 76 degrees of freedom
## Multiple R-squared: 0.901 Adjusted R-squared: 0.869
## F-statistic: 29.9 on 23 and 76 DF, p-value: <2e-16
## *** Standard errors may be too high due to more than 2 groups and exactDOF=FALSE
```

In this case, with 3 factors we can not be certain that it is sufficient with a single reference in two of the factors, but we try it as an exercise. (**lfe** does not include an intercept, it is subsumed in one of the factors, so it should tentatively be sufficient with a reference for the two others).

The input to our estimable function is a solution γ of equation (2). The argument addnames is a logical, set to TRUE when the function should add names to the resulting vector. The coefficients is ordered the same way as the levels in the factors. We should pick a single reference in factors f2,f3, subtract these, and add the sum to the first factor:

```
ef <- function(gamma, addnames) {</pre>
   ref2 <- gamma[[8]]
    ref3 <- gamma[[16]]
    gamma[1:7] <- gamma[1:7] + ref2 + ref3
    gamma[8:15] <- gamma[8:15] - ref2
    gamma[16:25] <- gamma[16:25] - ref3
    if (addnames) {
        names(gamma) <- c(paste("f1", 1:7, sep = "."), paste("f2", 1:8, sep = "."),</pre>
            paste("f3", 1:10, sep = "."))
    gamma
is.estimable(ef, fe = est$fe)
## [1] TRUE
getfe(est, ef = ef)
           effect
## f1.1 0.855296
## f1.2 0.323044
```

```
## f1.3 -0.146409
## f1.4
        -1.304527
## f1.5
        -1.210151
## f1.6
        -0.852878
## f1.7
         -0.646233
## f2.1
          0.000000
## f2.2
          0.002498
## f2.3
         -0.602877
## f2.4
          1.133586
## f2.5
          0.346222
## f2.6
         -0.043524
## f2.7
          0.425861
## f2.8
          0.445270
## f3.1
          0.000000
## f3.2
          0.068918
## f3.3
          0.587690
## f3.4
          0.295037
## f3.5
        -0.052250
## f3.6
          0.618679
## f3.7
         -0.212498
## f3.8 -0.017318
## f3.9 -0.571390
## f3.10 0.782764
```

We may compare this to the default estimable function, which picks a reference in each connected component as defined by the two first factors.

```
getfe(est)
##
             effect obs comp fe
                                    idx
## f1.1
             0.5323
                      16
                            1 f1
                                      1
                                      2
## f1.2
             0.0000
                      17
                            1 f1
## f1.3
            -0.4695
                                      3
                      15
                            1 f1
## f1.4
            -1.6276
                      12
                            1 f1
                                      4
## f1.5
                      12
            -1.5332
                            1 f1
                                      5
## f1.6
            -1.1759
                      15
                            1 f1
                                      6
## f1.7
            -0.9693
                      13
                            1 f1
                                      7
## f2.0.125
             0.6181
                      10
                            1 f2 0.125
## f2.0.25
             0.6206
                      16
                            1 f2 0.25
## f2.0.375
             0.0152
                      15
                            1 f2 0.375
## f2.0.5
             1.7517
                      13
                            1 f2
                                    0.5
## f2.0.625
             0.9643
                      12
                             1 f2 0.625
## f2.0.75
             0.5746
                      14
                            1 f2 0.75
## f2.0.875
             1.0439
                      10
                            1 f2 0.875
## f2.1
             1.0634
                             1 f2
                      10
                                      1
## f3.0.1
            -0.2950
                       5
                             2 f3
                                    0.1
## f3.0.2
            -0.2261
                       9
                             2 f3
                                    0.2
## f3.0.3
             0.2927
                      10
                             2 f3
                                    0.3
## f3.0.4
                             2 f3
             0.0000
                      13
                                    0.4
## f3.0.5
            -0.3473 11
                            2 f3
                                    0.5
```

```
## f3.0.6
             0.3236
                       8
                            2 f3
                                   0.6
## f3.0.7
            -0.5075
                       8
                            2 f3
                                    0.7
## f3.0.8
            -0.3124
                     13
                            2 f3
                                    0.8
## f3.0.9
            -0.8664 12
                            2 f3
                                    0.9
## f3.1
             0.4877 11
                            2 f3
                                      1
```

We see that the default has some more information. It uses the level names, and some more information, added like this:

```
efactory(est)
## function (v, addnames)
## {
       esum <- sum(v[extrarefs])</pre>
##
       df <- v[refsubs]</pre>
##
       sub <- ifelse(is.na(df), 0, df)</pre>
##
##
       df <- v[refsuba]</pre>
       add <- ifelse(is.na(df), 0, df + esum)
##
##
       v \leftarrow v - sub + add
       if (addnames) {
##
##
            names(v) <- nm
##
            attr(v, "extra") <- list(obs = obs, comp = comp, fe = fef,
##
                idx = idx)
##
       }
##
       V
## }
## <bytecode: 0x4cde800>
## <environment: 0x5883a98>
```

I.e. when asked to provide level names, it is also possible to add additional information as a list (or data.frame) as an attribute 'extra'. The vectors extrarefs,refsubs,refsuba etc. are precomputed by efactory for speed efficiency.

Here is the above example, but we create an intercept instead, and don't report the zero-coefficients, so that it closely resembles the output from 1m

```
f1 <- factor(f1)
f2 <- factor(f2)
f3 <- factor(f3)
ef <- function(gamma, addnames) {
    ref1 <- gamma[[1]]
    ref2 <- gamma[[8]]
    ref3 <- gamma[[16]]
    # put the intercept in the first coordinate
    gamma[[1]] <- ref1 + ref2 + ref3
    gamma[2:7] <- gamma[2:7] - ref1
    gamma[8:14] <- gamma[9:15] - ref2
    gamma[15:23] <- gamma[17:25] - ref3
    length(gamma) <- 23
    if (addnames) {
        names(gamma) <- c("(Intercept)", paste("f1", levels(f1)[2:7], sep = ""),</pre>
```

```
paste("f2", levels(f2)[2:8], sep = ""), paste("f3", levels(f3)[2:10],
               sep = "")
    gamma
getfe(est, ef = ef, bN = 1000, se = TRUE)
                effect
## (Intercept) 0.855296 0.6581
## f12
              -0.532252 0.3754
## f13
              -1.001705 0.3997
## f14
              -2.159823 0.4704
              -2.065447 0.4174
## f15
              -1.708174 0.3850
## f16
## f17
              -1.501529 0.4339
## f20.25
             0.002498 0.4268
## f20.375
              -0.602877 0.4465
## f20.5
              1.133586 0.4565
## f20.625
              0.346222 0.4795
## f20.75
              -0.043524 0.4329
## f20.875
               0.425861 0.5289
## f21
               0.445270 0.4801
## f30.2
             0.068918 0.6143
## f30.3
             0.587690 0.6301
## f30.4
              0.295037 0.6098
## f30.5
              -0.052250 0.5948
## f30.6
              0.618679 0.6528
## f30.7
              -0.212498 0.6463
              -0.017318 0.5840
## f30.8
## f30.9
              -0.571390 0.5806
## f31
              0.782764 0.6231
# compare with lm
summary(lm(y ~ x1 + f1 + f2 + f3))
##
## Call:
## lm(formula = y ~ x1 + f1 + f2 + f3)
## Residuals:
     Min
             1Q Median
                              3Q
## -1.8869 -0.7252 -0.0788 0.7558 2.3050
##
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 0.8553 0.7140 1.20 0.23471
## x1
               2.3545
                          0.1120 21.03 < 2e-16 ***
## f12
               -0.5323
                          0.4013 -1.33 0.18876
               -1.0017
                          0.4246 -2.36 0.02089 *
## f13
## f14
           -2.1598 0.5065 -4.26 5.7e-05 ***
```

```
-4.51 2.3e-05 ***
## f15
                -2.0654
                            0.4582
## f16
                -1.7082
                            0.4175
                                     -4.09
                                            0.00011 ***
## f17
                -1.5015
                            0.4526
                                     -3.32 0.00139 **
## f20.25
                 0.0025
                            0.4697
                                      0.01
                                            0.99577
## f20.375
                -0.6029
                            0.4837
                                     -1.25
                                            0.21646
## f20.5
                 1.1336
                            0.5020
                                      2.26 0.02682 *
## f20.625
                 0.3462
                            0.5039
                                      0.69
                                            0.49413
## f20.75
                -0.0435
                            0.4809
                                     -0.09
                                            0.92813
## f20.875
                 0.4259
                            0.5581
                                      0.76
                                            0.44782
                            0.5357
## f21
                 0.4453
                                      0.83 0.40848
## f30.2
                 0.0689
                            0.6414
                                      0.11
                                            0.91472
## f30.3
                 0.5877
                            0.6331
                                      0.93
                                            0.35622
## f30.4
                 0.2950
                            0.6367
                                      0.46
                                            0.64443
## f30.5
                -0.0522
                                     -0.08 0.93352
                            0.6243
## f30.6
                 0.6187
                            0.6661
                                      0.93
                                            0.35591
## f30.7
                -0.2125
                            0.6435
                                     -0.33
                                            0.74213
## f30.8
                -0.0173
                            0.6010
                                     -0.03 0.97709
## f30.9
                -0.5714
                            0.6172
                                     -0.93 0.35747
## f31
                 0.7828
                            0.6395
                                      1.22 0.22476
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 1.08 on 76 degrees of freedom
## Multiple R-squared: 0.901, Adjusted R-squared: 0.87
## F-statistic: 29.9 on 23 and 76 DF, p-value: <2e-16
```

5. Non-estimability

We consider another example. To ensure spurious relations there are almost as many factor levels as there are observations, and it will be hard to find enough estimable function to interpret all the coefficients. The coefficient for $\mathtt{x1}$ is still estimated, but with a large standard error. Note that this is an illustration of non-obvious non-estimability which may occur in much larger datasets, the author does not endorse using this kind of model for the kind of data you find below.

```
set.seed(128)
x1 <- rnorm(25)
f1 <- sample(9, length(x1), replace = TRUE)
f2 <- sample(8, length(x1), replace = TRUE)
f3 <- sample(8, length(x1), replace = TRUE)
e1 <- sin(f1) + 0.02 * f2^2 + 0.17 * f3^3 + rnorm(length(x1))
y <- 2.5 * x1 + (e1 - mean(e1))
summary(est <- felm(y ~ x1 + G(f1) + G(f2) + G(f3)))
##
## Call:
## felm(formula = y ~ x1 + G(f1) + G(f2) + G(f3))
##
## Residuals:</pre>
```

```
##
     Min 1Q Median
                           3Q
                                 Max
## -0.870 -0.368 0.000 0.148
##
## Coefficients:
##
     Estimate Std. Error t value Pr(>|t|)
         1.41
                    0.47
                               3
                                    0.03 *
## x1
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 1.3 on 5 degrees of freedom
## Multiple R-squared:
                       1 Adjusted R-squared: 0.999
## F-statistic: 933 on 19 and 5 DF, p-value: 1.34e-07
## *** Standard errors may be too high due to more than 2 groups and exactDOF=FALSE
```

The default estimable function fails, and the coefficients from getfe are not useable. getfe yields a warning in this case.

```
ef <- efactory(est)
is.estimable(ef, est$fe)
## Warning: non-estimable function, largest error 0.6 in coordinate 4
("f1.5")
## [1] FALSE</pre>
```

Indeed, the rank-deficiency is larger than expected. There are more spurious relations between the factors than what can be accounted for by looking at components in the two first factors. In this low-dimensional example we may find the matrix D of equation (2), and its (column) rank deficiency is larger than 2.

```
f1 <- factor(f1)
f2 <- factor(f2)
f3 <- factor(f3)
D <- t(do.call("rBind", lapply(list(f1, f2, f3), as, Class = "sparseMatrix")))
dim(D)
## [1] 25 21
as.integer(rankMatrix(D))
## [1] 18</pre>
```

Alternatively we can use an internal function in lfe for finding the rank deficiency directly.

```
lfe:::rankDefic(list(f1, f2, f3))
## [1] 3
```

This rank-deficiency also has an impact on the standard errors computed by felm. If the rank-deficiency is small relative to the degrees of freedom the standard errors are scaled slightly upwards if we ignore the rank deficiency, but if it is large, the impact on the standard errors can be substantial. The above mentioned rank-computation procedure can be activated by specifying exactDOF=TRUE in the call to felm, but it may be time-consuming if the factors have many levels. Computing the rank does not in itself help us find estimable functions for getfe.

```
summary(est \leftarrow felm(y \sim x1 + G(f1) + G(f2) + G(f3), exactDOF = TRUE))
##
## Call:
##
     felm(formula = y ~ x1 + G(f1) + G(f2) + G(f3), exactDOF = TRUE)
## Residuals:
             10 Median
    Min
                           3Q
## -0.870 -0.368 0.000 0.148 1.884
##
## Coefficients:
##
     Estimate Std. Error t value Pr(>|t|)
                   0.429
                            3.28
## x1
       1.408
                                  0.017 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.19 on 6 degrees of freedom
## Multiple R-squared: 1 Adjusted R-squared: 0.999
## F-statistic: 1.18e+03 on 18 and 6 DF, p-value: 3.7e-09
```

We can get an idea what happens if we keep the dummies for f3. In this case, with 2 factors, lfe will partition the dataset into connected components and account for all the multicollinearities among the factors f1 and f2 just as above, but this is not sufficient. The interpretation of the resulting coefficients is not straightforward.

```
summary(est \leftarrow felm(y ~ x1 + G(f1) + G(f2) + f3, exactDOF = TRUE))
##
## Call:
     felm(formula = y ~ x1 + G(f1) + G(f2) + f3, exactDOF = TRUE)
##
##
## Residuals:
##
        Min
                   1Q
                         Median
                                      3Q
## -8.70e-01 -3.68e-01 -3.87e-08 1.48e-01 1.88e+00
##
## Coefficients:
      Estimate Std. Error t value Pr(>|t|)
## x1
         1.408
                   0.429
                          3.28
                                    0.017 *
## f32
         0.697
                    1.058
                            0.66
                                    0.535
## f33
        4.231
                   1.544 2.74
                                    0.034 *
## f34
                       NA
                             NA
           NA
                                       NA
## f35 19.489
                  1.160 16.81 2.8e-06 ***
## f36
       35.684
                   2.264
                           15.76 4.1e-06 ***
## f37 57.772
                  1.185
                           48.75 5.0e-09 ***
## f38 86.498
                   1.226
                           70.53 5.5e-10 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.19 on 6 degrees of freedom
```

```
## Multiple R-squared: 1 Adjusted R-squared: 0.999
## F-statistic: 1.18e+03 on 18 and 6 DF, p-value: 3.7e-09
getfe(est)
         effect obs comp fe idx
## f1.1 -34.5246
                 4
                      1 f1
## f1.2 -33.3982
                 5
                      1 f1
                            2
## f1.3 -34.6698
                6
                    1 f1
                            3
## f1.5 -27.3871
                     1 f1
                1
                            5
                     1 f1
## f1.6 -35.7423
                5
                            6
## f1.7 -36.4468
                     1 f1
                            7
                 1
## f1.9 -32.2795
               3
                    1 f1
                            9
## f2.2 -0.4037
               3
                     1 f2
                            2
## f2.3
        1.8848
                      1 f2
                            3
                3
## f2.5 -2.6807
                 2
                     1 f2
                            5
## f2.6 0.0000
                            6
               8 1 f2
## f2.7
         0.8925
                4
                      1 f2
                            7
## f2.8 0.9553 5
                      1 f2
                            8
```

In this particular example, we may use a different order of the factors, and we see that by partitioning the dataset on the factors f1,f3 instead of f1,f2, there are 2 connected components (the factor f2 gets its own comp-code, but this is not a graph theoretic component number, it merely indicates that there is a separate reference among these).

```
summary(est <- felm(y ~ x1 + G(f1) + G(f3) + G(f2), exactDOF = TRUE))
##
## Call:
##
     felm(formula = y ~ x1 + G(f1) + G(f3) + G(f2), exactDOF = TRUE)
##
## Residuals:
##
                   1Q
                         Median
                                      30
        Min
## -8.70e-01 -3.68e-01 -4.24e-11 1.48e-01 1.88e+00
## Coefficients:
## Estimate Std. Error t value Pr(>|t|)
## x1
       1.408
                   0.429
                           3.28
                                   0.017 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 1.19 on 6 degrees of freedom
## Multiple R-squared: 1 Adjusted R-squared: 0.999
## F-statistic: 1.18e+03 on 18 and 6 DF, p-value: 3.7e-09
is.estimable(efactory(est), est$fe)
## [1] TRUE
getfe(est)
         effect obs comp fe idx
## f1.1
         0.1452 4 1 f1
                             1
## f1.2 1.2716 5 1 f1
```

```
0.0000
                          1 f1
                                  3
## f1.3
                     6
## f1.5
           0.0000
                     1
                           2 f1
                                  5
## f1.6
         -1.0725
                     5
                          1 f1
                                  6
                          1 f1
## f1.7
         -1.7770
                                  7
                     1
## f1.9
           2.3903
                     3
                          1 f1
                                  9
## f3.1 -34.6698
                     4
                          1 f3
                                  1
## f3.2 -33.9730
                          1 f3
                                  2
## f3.3 -30.4384
                          1 f3
                                  3
                     1
## f3.4 -27.3871
                          2 f3
                                  4
                     1
## f3.5 -15.1809
                                  5
                     4
                          1 f3
          1.0138
## f3.6
                     2
                          1 f3
                                  6
## f3.7
         23.1023
                                  7
                     3
                          1 f3
## f3.8
         51.8278
                     6
                          1 f3
                                  8
                                  2
## f2.2
         -0.4037
                     3
                          3 f2
## f2.3
           1.8848
                     3
                          3 f2
                                  3
                     2
                                  5
## f2.5
          -2.6807
                          3 f2
## f2.6
           0.0000
                     8
                          3 f2
                                  6
## f2.7
           0.8925
                     4
                           3 f2
                                  7
## f2.8
           0.9553
                     5
                          3 f2
                                  8
```

Below is the same estimation in lm. We see that the coefficient for x1 is identical to the one from felm, but there is no obvious relation between e.g. the coefficients for f1; the difference f15-f16 is not the same for lm and felm. Since these are in different components, they are not comparable. But of course, if we compare in the same component, e.g. f16-f17 or take a combination which actually occurs in the dataset, it is unique (estimable):

```
data.frame(f1, f2, f3)[1, ]
## f1 f2 f3
## 1 1 6 7
```

I.e. if we add the coefficients f1.1 + f2.6 + f3.7 and include the intercept for lm, we will get the same number for both lm and felm.

```
summary(est <- lm(y ~ x1 + f1 + f2 + f3))
##
## Call:
## lm(formula = y ~ x1 + f1 + f2 + f3)
##
## Residuals:
##
      Min
              1Q Median
                             3Q
                                   Max
## -0.870 -0.368 0.000
                         0.148
                                 1.884
##
## Coefficients: (1 not defined because of singularities)
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept)
               -34.928
                              0.981
                                    -35.59
                                             3.3e-08 ***
## x1
                  1.408
                              0.429
                                       3.28
                                              0.0168 *
## f12
                              1.354
                                       0.83
                                              0.4373
                  1.126
                 -0.145
                                      -0.12
## f13
                              1.162
                                              0.9047
## f15
                  7.138
                             1.822
                                     3.92
                                              0.0078 **
```

```
-1.218
## f16
                              1.412
                                       -0.86
                                               0.4215
## f17
                  -1.922
                              2.240
                                       -0.86
                                               0.4237
## f19
                  2.245
                              1.925
                                       1.17
                                               0.2877
## f23
                   2.289
                                       1.47
                              1.562
                                               0.1932
## f25
                  -2.277
                              1.455
                                       -1.56
                                               0.1687
## f26
                   0.404
                              1.187
                                       0.34
                                               0.7454
## f27
                   1.296
                              1.528
                                       0.85
                                               0.4289
## f28
                   1.359
                              1.477
                                       0.92
                                               0.3930
## f32
                   0.697
                              1.058
                                       0.66
                                               0.5346
## f33
                   4.231
                              1.544
                                        2.74
                                               0.0337
## f34
                      NA
                                 NA
                                          NA
                                                   NA
## f35
                  19.489
                              1.160
                                       16.81
                                              2.8e-06 ***
## f36
                  35.684
                              2.264
                                       15.76
                                              4.1e-06 ***
                                       48.75
                                              5.0e-09 ***
## f37
                  57.772
                              1.185
## f38
                  86.498
                              1.226
                                       70.53
                                              5.5e-10 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 1.19 on 6 degrees of freedom
## Multiple R-squared:
                            1, Adjusted R-squared:
## F-statistic: 1.18e+03 on 18 and 6 DF, p-value: 3.7e-09
```

6. Weeks-Williams partitions

There is a partial solution to the non-estimability problem in [8]. Their idea is to partition the dataset into components in which all differences between factor levels are estimable. The components are connected components of a subgraph of an e-dimensional grid graph where e is the number of factors. That is, a graph is constructed with the observations as vertices, two observations are adjacent (in a graph theoretic sense) if they differ in at most one of the factors. The dataset is then partitioned into (graph theoretic) connected components. It's a finer partitioning than the above, and consequently introduces more reference levels than is necessary for identification. I.e. it does not find all estimable functions, but in some cases (e.g. in [7]) the largest component will be sufficiently large for proper analysis. It is of course always a question whether such an endogeneous selection of observations will yield a dataset which results in unbiased coefficients. This partitioning can be done by the compfactor function with argument WW=TRUE:

```
fe <- list(f1, f2, f3)
wwcomp <- compfactor(fe, WW = TRUE)</pre>
```

It has more levels than the rank deficiency

```
lfe:::rankDefic(fe)
## [1] 3
nlevels(wwcomp)
## [1] 10
```

and each of its components are contained in a component of the previously considered components, no matter which two factors we consider. For the case of two factors, the concepts coincide.

```
nlevels(interaction(compfactor(fe), wwcomp))
## [1] 10
# pick the largest component:
wwdata <- data.frame(y, x1, f1, f2, f3)[wwcomp == 1, ]
print(wwdata)
##
                   x1 f1 f2 f3
          У
## 2
      52.91 0.482612 2
                         6
      54.99 0.004553 2
                          6
## 11 -13.28
             0.322709
                       2
                         8
                             5
## 15 54.62 0.684795
                       2
                         7
                             8
## 16 -35.09 -0.883670 3
## 17 56.13 0.722127
                       9 7
## 19 -14.60 -0.563255
                       2
                             5
## 20 -18.94 -1.644455
                       6
                         6 5
## 21 -34.37 0.359966
                      3 2 2
## 25 52.13 0.589060 3 6 8
```

Though, in this particular example, there are more parameters than there are observations, so an estimation would not be feasible.

efactory cannot easily be modified to produce an estimable function corresponding to WW components. The reason is that efactory, and the logic in getfe, work on partitions of factor levels, not on partitions of the dataset, these are the same for the two-factor case.

WW partitions have the property that if you pick any two of the factors and partition a WW-component into the previously mentioned non-WW partitions, there will be only one component, hence you may use any of the estimable functions from efactory on each partition. That is, a way to use WW partitions with lfe is to do the whole analysis on the largest WW-component. felm may still be used on the whole dataset, and it may yield different results than what you get by analysing the largest WW-component.

Here is a larger example:

```
set.seed(135)
x <- rnorm(10000)
f1 <- sample(1000, length(x), replace = TRUE)
f2 <- (f1 + sample(18, length(x), replace = TRUE))%%500
f3 <- (f2 + sample(9, length(x), replace = TRUE))%%500
y <- x + 1e-04 * f1 + sin(f2^2) + cos(f3)^3 + 0.5 * rnorm(length(x))
dataset <- data.frame(y, x, f1, f2, f3)
summary(est <- felm(y ~ x + G(f1) + G(f2) + G(f3), data = dataset, exactDOF = TRUE))
##
## Call:
## felm(formula = y ~ x + G(f1) + G(f2) + G(f3), data = dataset, exactDOF = TRUE)
##
## Residuals:</pre>
```

```
## Min 1Q Median 3Q Max
## -1.929087 -0.305385 0.000243 0.305398 1.921153
##
## Coefficients:
## Estimate Std. Error t value Pr(>|t|)
## x 0.99991 0.00565 177 <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.505 on 8001 degrees of freedom
## Multiple R-squared: 0.903 Adjusted R-squared: 0.879
## F-statistic: 37.5 on 1998 and 8001 DF, p-value: <2e-16</pre>
```

We count the number of connected components in f1,f2, and see that this is sufficient to ensure estimability

```
nlevels(est$cfactor)
## [1] 1
is.estimable(efactory(est), est$fe)
## [1] TRUE
nrow(alpha <- getfe(est))
## [1] 2000</pre>
```

It has rank deficiency one less than the number of factors :

```
lfe:::rankDefic(est$fe)
## [1] 2
```

Then we analyse the largest WW-component

```
wwcomp <- compfactor(est$fe, WW = TRUE)</pre>
nlevels(wwcomp)
## [1] 958
wwset <- dataset[wwcomp == 1, ]</pre>
nrow(wwset)
## [1] 2394
summary(wwest <- felm(y \sim x + G(f1) + G(f2) + G(f3), data = wwset, exactDOF = TRUE))
## Call:
     felm(formula = y ~ x + G(f1) + G(f2) + G(f3), data = wwset, exactDOF = TRUE)
##
##
## Residuals:
        Min
                    1Q
                          Median
## -1.882864 -0.294976 0.000194 0.294391 2.037462
## Coefficients:
## Estimate Std. Error t value Pr(>|t|)
## x 1.003
                0.012 83.5 <2e-16 ***
```

```
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.514 on 1788 degrees of freedom
## Multiple R-squared: 0.911 Adjusted R-squared: 0.881
## F-statistic: 30.3 on 605 and 1788 DF, p-value: <2e-16
```

We see that we get the same coefficient for \mathbf{x} in this case. This is not surprising, there is no obvious reason to believe that our selection of observations is skewed in this randomly created dataset.

This one has the same rank deficiency:

```
lfe:::rankDefic(wwest$fe)
## [1] 2
```

but a smaller number of identifiable coefficients.

```
nrow(wwalpha <- getfe(wwest))
## [1] 607</pre>
```

We may compare effects which are common to the two methods:

```
head(alpha)
##
         effect obs comp fe idx
                      1 f1
## f1.1 -0.5542
                6
## f1.2 -0.2756 13
                       1 f1
## f1.3 -0.5763
                       1 f1
                              3
                6
## f1.4 -0.5413
                8
                       1 f1
                              4
## f1.5 -0.5697 13
                              5
                       1 f1
## f1.6 -0.4250
                       1 f1
head(wwalpha)
          effect obs comp fe idx
## f1.1 -1.27887
                 3
                        1 f1
                               1
## f1.2 -0.34865
                 7
                        1 f1
                               2
## f1.3 -0.95121
                 4
                        1 f1
                               3
## f1.4 -1.04719
                  1
                        1 f1
                               4
## f1.6 -0.89773
                  4
                        1 f1
                               6
## f1.7 0.04965
                        1 f1
```

but there is no obvious relation between e.g. $\tt f1.1-f1.2$, they are very different in the two estimations. The coefficients are from different datasets, and the standard errors are large (≈ 0.7) with this few observations for each factor level. The number of identified coefficients for each factor varies (these figures contain the two references):

```
table(wwalpha[, "fe"])
##
## f1 f2 f3
## 310 148 149
```

References

- J. M. Abowd, R. H. Creecy, and F. Kramarz, Computing Person and Firm Effects Using Linked Longitudinal Employer-Employee Data, Tech. Report TP-2002-06, U.S. Census Bureau, 2002.
- [2] J. M. Abowd, F. Kramarz, and D. N. Margolis, High Wage Workers and High Wage Firms, Econometrica 67 (1999), no. 2, 251–333.
- [3] J. A. Eccleston and A. Hedayat, On the Theory of Connected Designs: Characterization and Optimality, Ann. Statist. 2 (1974), 1238–1255.
- [4] S. Gaure, OLS with Multiple High Dimensional Category Variables, Computational Statistics and Data Analysis 66 (2013), 8–18.
- [5] J. D. Godolphin and E. J. Godolphin, On the connectivity of row-column designs, Util. Math. 60 (2001), 51–65.
- [6] L.L. Kupper, J.M. Janis, I.A. Salama, C.N. Yoshizawa, and B.G. Greenberg, Age-Period-Cohort Analysis: An Illustration of the Problems in Assessing Interaction in One Observation Per Cell Data, Commun. Statist.-Theor. Meth. 12 (1983), no. 23, 2779–2807.
- [7] S.M. Torres, P. Portugal, J.T. Addison, and P. Guimarães, The Sources of Wage Variation: A Three-Way High-Dimensional Fixed Effects Regression Model., IZA Discussion Paper 7276, Institute for the Study of Labor (IZA), March 2013.
- [8] D.L. Weeks and D.R. Williams, A Note on the Determination of Connectedness in an N-Way Cross Classification, Technometrics 6 (1964), no. 3, 319–324.

RAGNAR FRISCH CENTRE FOR ECONOMIC RESEARCH, OSLO, NORWAY