# L1 and L2 Penalized Regression Models

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# 1 Introduction

This short note explains the use of the *penalized* package. The package is designed for penalized estimation in generalized linear models. At this moment, the only supported models are linear regression, logistic regression and the Cox proportional hazards model, but others are likely to be included in the future. As to penalties, the package allows an L1 absolute value ("lasso") penalty (Tibshirani, 1996, 1997), an L2 quadratic ("ridge") penalty (Hoerl and Kennard, 1970; Le Cessie and van Houwelingen, 1992; Verweij and Van Houwelingen, 1994), or a combination of the two (the "naive elastic net" of Zou and Hastie, 2005). The package also includes facilities for likelihood cross-validation and for optimization of the tuning parameter.

L1 and L2 penalized estimation methods shrink the estimates of the regression coefficients towards zero relative to the maximum likelihood estimates. The purpose of this shrinkage is to prevent overfit arising due either collinearity of the covariates or high-dimensionality. Although both shrinkage methods, the effects of L1 and L2 penalization are quite different in practice. Applying an L2 penalty tends to result in all small but non-zero regression coefficients, whereas applying an L1 penalty tends to result in many regression coefficients shrunk exactly to zero and a few other regression coefficients with comparatively little shrinkage. Combining L1 and L2 penalties tends to give a result in between, with fewer regression coefficients set to zero than in a pure L1 setting, and more shrinkage to the other coefficients. The amount of shrinkage is determined by tuning parameters  $\lambda_1$  and  $\lambda_2$ , for L1 and L2 respectively, where a value of zero always means no shrinkage (= maximum likelihood estimation) and a value of infinity means infinite shrinkage (= setting all regression coefficients to zero). For more details about the methods, please refer to the above-mentioned papers.

It is important to note that shrinkage methods are generally not invariant to the relative scaling of the covariates. Before fitting a model, it is prudent to consider if the covariates already have a natural scaling relative to each other or whether they should be standardized.

The main algorithm for L1 penalized estimation that used in this package will be documented in a forthcoming paper. It has been combined with ideas from Eilers et al. (2001) and Van Houwelingen et al. (2005) for efficient L2 penalized estimation.

## 2 Penalized likelihood estimation

The basic function of the package is the penalized function, which performs penalized estimation for fixed values of  $\lambda_1$  and  $\lambda_2$ . Its syntax has been loosely modeled on that of the functions glm (package stats) and coxph (package survival), but it is slightly more flexible. Two main input types are allowed: one using formula objects, one using matrices.

## 2.1 the nki70 data

As example data we use the 70 gene signature of Van 't Veer et al. (2002) in the gene expression data set of Van de Vijver et al. (2002).

```
> library(penalized)
> data(nki70)
```

This loads a *data frame* with 295 breast cancer patients and 77 covariates. The first two indicate the survival time and event status (time is in months), the next five are clinical covariates (diameter of the tumor, lymph node status, estrogen receptor status, grade of the tumor and age of the patients), and the other 70 are gene expression measurements of the 70 molecular markers.

#### 2.2 the penalized function

To fit a model to predict survival (Surv(time, event)) with the two markers "DIAPH3" and "NUSAP1" at  $\lambda_1 = 0$  and  $\lambda_2 = 1$ , we can say (all are equivalent)

The covariates may be specified in the second function argument (penalized) as a formula object with an open left hand side, as in the first line. Alternatively, they may be specified as a matrix, as in the second line. If, as here, they are supplied as a data frame, they are coerced to a matrix.

For consistency with glm and coxph the third option is also allowed, in which the covariates are included in the first function argument.

Use attach to avoid specifying the data argument every time.

> attach(nki70)

#### 2.3 choice of lambda

It is difficult to say in advance which value of lambda1 or lambda2 to use. The penalized package offers ways of finding optimal values using cross-validation. This is explained in Section ??

Note that for small values of lambda1 or lambda2 the algorithm be very slow, may fail to converge or may run into numerical problems, especially in high-dimensional data. When this happens, increase the value of lambda1 or lambda2.

#### 2.4 standardization

If the covariates are not on the same scale in a natural way, it is advisable to standardize them. The function argument *standardize* (default: FALSE) standardizes the covariates to unit second central moment before applying penalization. This standardization makes sure that each covariates is affected more or less equally by the penalization.

The fitted regression coefficients that the function returns have been scaled back and correspond to the original scale of the covariates.

#### 2.5 penfit objects

The penalized function returns a *penfit* object, from which useful information can be extracted.

#### > coefficients(fit)

DIAPH3 NUSAP1 -0.003347245 1.610876235

#### > residuals(fit)

```
 \begin{smallmatrix} 1 \end{smallmatrix} \rbrack -0.12993363 \quad 0.71048110 \quad -0.35170595 \quad -0.20835122 \quad -0.42640209 \quad -0.36211082 
  \begin{bmatrix} 7 \end{bmatrix} \quad 0.74649177 \quad -0.61721035 \quad 0.73673593 \quad -0.44704599 \quad -0.22460950 \quad 0.68845952 
     0.93996917 -0.67997913 0.94948604 0.16863313 -0.59195437 -0.26745471
 [13]
     0.91449205 -1.23676097 -0.40968870 0.91061049 -0.55575761 -0.32862313
 [31]
      0.60671164 \ -0.40183288 \ \ 0.80661766 \ -0.31719394 \ \ \ 0.94870766 \ \ \ 0.20208607
 [37] -0.48991812 -0.50785174 -0.49283996 0.78584340 0.55295465 -0.46627185
 [43] -0.20266445 -0.16518805 -0.39959140 -0.32196958 -0.53846156 -0.48486697
      0.77151870 \ -0.53925553 \ \ 0.86070571 \ \ 0.95689622 \ \ 0.75463968 \ -0.29199390
 [49]
 [55] -0.13588499 -0.24660422 -0.57657142 -0.40936722 -0.28811049 0.58716502
 [61] 0.88842941 0.77270364 0.79608318 -0.11263132 0.87925059 -0.28498404
 [67] -0.26210767 -0.37312428 0.97771185 -0.44077885 -0.51232607 0.00000000
 [73] 0.97445478 -0.24820082 -0.16619852 -0.22048742 -0.20562994 -0.50146186
 [79] 0.69349646 -0.10674793 0.99328316 -0.20115664 -0.46919776 -0.40857248
 [85] -0.29487005 -0.24853058 -0.25164899 0.81512763 -0.47900003 -0.27256394
 [91] 0.95656111 -0.63983679 0.67564267 -0.55777432 0.73463695 0.75003384
[109] -0.42835134 -0.39221302 0.57616189 0.89271966 -0.26544486 -0.23038969
[115] -0.21655779 -0.19855736 -0.19964839 -0.61347376 -0.33562178 -0.13084269
[121] -0.21060922 -0.15174868 -0.14392871 -0.28763617 0.70907738 0.69706807
[127] -0.17691026 0.81423467 -0.79709884 0.49514854 -1.10379660 -0.63325914
[139] -0.54941392 0.53020161 -0.42178369 -0.20314618 -0.11758312 -0.33595144
```

#### > fitted.values(fit)

```
[1] 0.4023261 1.0605204 0.8671254 0.6451380 1.3203100 1.1783128 0.7849620 [8] 1.3615191 1.2242175 0.5909803 0.7647654 1.3883721 1.0340419 0.8989103 [15] 1.7786194 1.0990400 0.7825445 0.8281459 1.1737099 1.5011198 0.8214706 [22] 0.7294193 1.2669400 1.0719473 1.3030135 1.6349579 0.6322409 2.0758688 [29] 0.8576578 0.5071387 0.9163257 0.6201177 1.3351461 0.6023242 1.4370682 [36] 1.3727073 0.7560528 0.9643671 0.8478672 1.3240965 0.6898906 0.8854104 [43] 0.4184755 0.3410916 0.8251038 0.6648249 1.1118524 1.0011866 0.7434808 [50] 1.1134918 1.3343857 0.5883087 1.7971514 0.6029291 0.9952950 0.5439902 [57] 1.0948599 1.3938408 0.9809778 0.9106826 0.9271613 0.7739138 0.5612994 [64] 0.2776907 1.0749732 0.8824236 0.8529014 1.2141507 1.0512123 1.3648261 [71] 1.5863647 0.7354024 1.8171231 0.8763932 0.5658838 0.6827165 0.7001425 [78] 1.7074108 1.3106208 0.3634625 0.9588171 0.7102809 0.8909665 1.3911348 [85] 1.0039932 0.5482396 0.55551186 1.0327885 1.0566377 0.6012553 0.4499945
```

```
[92] 1.4114313 0.7996978 1.2304078 1.0900148 1.2708685 0.3427188 0.8566860 [99] 0.9425167 0.4720855 2.0220257 1.4321562 0.8803293 1.9796288 0.3323755 [106] 1.0126127 0.7093202 0.3891949 1.3263456 1.2762657 2.2573646 1.3253340 [113] 0.8637606 0.7133779 0.6705488 0.7273243 1.0150449 2.0887963 1.1427474 [120] 0.4455019 0.7170963 0.5166840 0.4900581 0.9793628 1.0272432 0.6255146 [127] 0.5477843 1.2122066 1.0537388 1.3141934 1.7034040 1.0894401 1.0894112 [134] 1.1700266 1.1336016 1.4561498 0.7064038 1.4981711 1.8706811 0.8921070 [141] 0.8709279 0.6610400 0.3640840 1.1438698
```

> basehaz(fit)

A "breslow" object with 1 survival curve and 50 time points.

See help(penfit) for more information.

## 2.6 unpenalized covariates

In some situations not all covariates are subject to a penalty. Any additional covariates that should be included in the model without being penalized can be specified separately. using the third function argument (unpenalized). For example

```
> fit <- penalized(Surv(time, event), nki70[, 8:77], ~ER, lambda2 = 1)</pre>
```

This adds estrogen receptor status as an unpenalized covariate.

In rare cases each covariate may have to penalized in a different way, or some covariates have to be given an L2 penalty and others an L1 penalty. In those cases, the arguments lambda1 and lambda2 may be supplied as vectors of the same length as the number of covariates in the function argument penalized.

#### 2.7 factors

If some of the factors included in the *formula* object *penalized* are *factor*, these are automatically made into dummy variables, as in glm and coxph, but in a special way that is more appropriate for penalized regression.

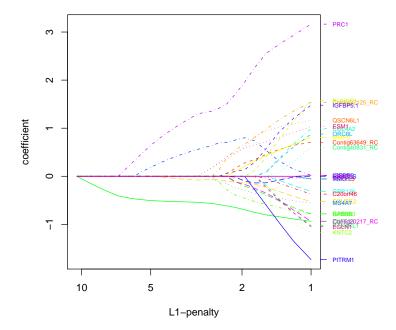
Unordered factors are turned into as many dummy variables as the factor has levels. This ensures a symmetric treatment of all levels and guarantees that the fit does not depend on the ordering of the levels. See help(contr.none) for details.

Ordered factors are turned into dummy variables that code for the difference between successive levels (one dummy less than the number of levels). L2 penalization on such factors therefore leads to small successive differences; L1 penalization leads to ranges of successive levels with identical effects.

To override the automatic choice of contrasts, use C (package stats).

#### 2.8 fitting in steps

In some cases it may be interesting to visualize the effect of changing the tuning parameter lambda1 or lambda2 on the values of the fitted regression coefficients. This can be done using the function argument steps in combination with the plotpath function. At this moment, this functionality is only available for visualizing the effect of lambda1.



When using the *steps* argument, the function starts fitting the model at the maximal value of  $\lambda_1$ , that is the smallest value that shrinks all regression coefficients to zero. From that value it continues fitting the model for *steps* successively decreasing values of  $\lambda_1$  until the specified value of lambda1 is reached.

If the argument *steps* is supplied to penalized, the function returns a *list* of *penfit* objects. These can be accessed individually or their coefficients can be plotted using plotpath.

# 3 Cross-validation and optimization

Cross-validation can be used to assess the predictive quality of the penalized prediction model or to compare the predictive ability of different values of the tuning parameter.

The *penalized* package uses likelihood cross-validation for all models. Likelihood cross-validation has some advantages over other optimization criteria: it tends to be a continuous function of the tuning parameter; it can be defined in a general way for almost any model, and it does not require calculation the effective dimension of a model, which is problematic in L1 penalized models. For

the Cox proportional hazards model, the package uses cross-validated log partial likelihood (Verweij and Van Houwelingen, 1993), which is a natural extension of the cross-validated log likelihood to the Cox model.

Five functions are available for calculating the cross-validated log likelihood and for optimizing the cross-validated log likelihood with respect to the tuning parameters. They have largely the same arguments. See help(cvl) for an overview.

#### 3.1 cross-validation

The function cvl calculates the cross-validated log likelihood for fixed values of  $\lambda_1$  and  $\lambda_2$ .

It accepts the same arguments as penalized (except steps: see profL1 below) as well as the fold argument. This will usually be a single number k to indicate k-fold cross-validation. In that case, the allocation of the subjects to the folds is random. Alternatively, fold can give the precise allocation of the subjects into the folds by giving a vector of the length of the number of subjects with values form 1 to k, each indicating the fold of the corresponding subject. The default is to do leave-one-out cross-validation.

The function cvl returns a names *list* with four elements:

cvl the cross-validated log likelihood.

fold the fold allocation used; this may serve as input to a next call to cvl to ensure comparability.

predictions the prediction made on each left-out subject. The format depends on the model used. In logistic regression this is just a vector of probabilities. In the Cox model this is a collection of predicted survival curves (a breslow object). In the linear model this is a collection of predicted means and predicted standard deviations (the latter are the maximum likelihood estimates of  $\sigma^2$ ).

```
fullfit the fit on the full data (a penfit object)
```

#### 3.2 breslow objects

The *breslow* class is defined in the *penalized* package to store estimated survival curves. They are used for the predictions in cross-validation and for the baseline hazard in the **penalized** function. See help(breslow) for details.

#### > fit\$predictions

A "breslow" object with 144 survival curves and 51 time points.

#### > time(fit\$predictions)

```
0.0000000 0.3531828
                           0.6488706 0.9363276
                                                  0.9609856
 [1]
                                                              1.2101300
 [7]
     1.3880903
                 1.5003422
                            1.6098563
                                       1.6125941
                                                   1.7166324
                                                              1.7330595
                            1.9739904
[13]
     1.9466119
                 1.9657769
                                        2.2231348
                                                   2.2970568
                                                              2.3353867
Г197
     2.3408624
                 2.6146475
                            2.6803559
                                       2.6967830
                                                   2.8117728
                                                              2.8528405
[25]
     3.1211499
                 3.2197125
                            3.4195756
                                       3.4387406
                                                   3.6550308
                                                              3.9151266
[31]
     4.2190281
                 4.4462697
                            4.6214921
                                        4.6625599
                                                   4.9719370
                                                              5.1170431
[37]
     6.5653662
                 6.9952088
                            8.1286790
                                       8.3039014
                                                   8.5284052
                                                              8.5612594
[43]
     8.9253936
                8.9883641
                           9.9986311 11.2114990 11.7399042 12.4654346
[49] 14.0123203 17.4209446 17.6591376
```

# > as.matrix(fit\$predictions)[1:2, ]

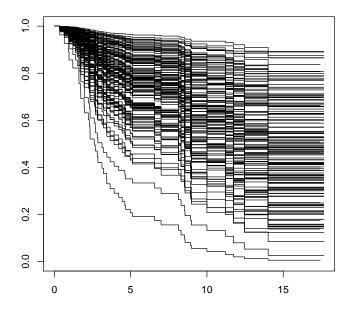
```
0 0.353182752 0.648870637 0.9363276 0.960985626 1.210130048 1.388090349
125 1
         0.998685
                    0.9973606 0.9959858
                                           0.9946032
                                                        0.9931895
                                                                     0.9931895
127 1
         0.989432
                    0.9788884 0.9681762
                                           0.9574688
                                                        0.9466753
                                                                     0.9356891
    1.500342231 1.609856263 1.612594114 1.716632444 1.733059548 1.94661191
125
      0.9931895
                  0.9917613
                               0.9902746
                                           0.9887726
                                                        0.9872397
                                                                   0.9856904
      0.9246843
                  0.9246843
                               0.9134758
                                           0.9022695
                                                        0.8909130
                                                                   0.8794765
    1.965776865 1.973990418 2.223134839 2.29705681 2.335386721 2.340862423
125
      0.9856904
                  0.9841091
                               0.9824831
                                          0.9808350
                                                       0.9791760
                                                                    0.9774997
127
      0.8681047
                  0.8565477
                               0.8446641
                                          0.8328476
                                                       0.8210618
                                                                    0.8092300
    2.614647502 2.680355921 2.696783025 2.811772758 2.85284052 3.121149897
125
      0.9757876
                  0.9740095
                               0.9740095
                                           0.9721435
                                                       0.9702287
                                                                    0.9682503
127
      0.7971578
                  0.7847284
                               0.7722717
                                           0.7594478
                                                      0.7462694
                                                                    0.7329316
    3.219712526 3.419575633 3.438740589 3.655030801 3.915126626 4.219028063
                                           0.9599871
125
      0.9662446
                  0.9641873
                               0.9621165
                                                        0.9578311
                                                                     0.9556501
      0.7194082
                  0.7194082
                               0.7059928
                                           0.6925437
                                                        0.6791403
127
                                                                     0.6657654
    4.446269678 4.621492129 4.66255989 4.971937029 5.117043121 6.565366188
125
      0.9534276
                  0.9511821
                              0.9488971
                                          0.9465588
                                                       0.9441326
                                                                    0.9409462
      0.6657654
                  0.6524527
                              0.6390528
                                          0.6390528
                                                       0.6253918
                                                                    0.6078132
    6.995208761 8.128678987 8.303901437 8.528405202 8.561259411 8.925393566
      0.9372360
                  0.9320941
                               0.9268622
                                           0.9215675
125
                                                        0.9161356
                                                                     0.9103283
      0.5879861
                  0.5644805
                               0.5415987
                                           0.5192460
                                                        0.4971154
127
                                                                     0.4741562
    8.988364134 9.998631075 11.21149897 11.73990418 12.46543463 14.01232033
125
      0.9042337
                  0.8959155
                               0.8854297
                                           0.8725759
                                                        0.8578365
                                                                     0.8366626
                               0.3850881
                                           0.3390516
                                                        0.2849061
                                                                     0.2102364
127
      0.4509372
                  0.4202807
    17.42094456 17.65913758
125
      0.8366626
                          NA
```

0.2102364

0.2102364

127

<sup>&</sup>gt; plot(fit\$predictions)



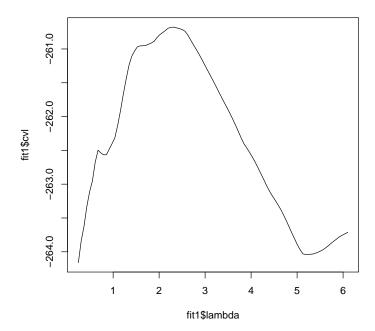
#### 3.3 profiling the cross-validated log likelihood

The functions profL1 and profL2 can be used to examine the effect of the parameters  $\lambda_1$  and  $\lambda_2$  on the cross-validated log likelihood. The profL1 can be used to vary  $\lambda_1$  while keeping  $\lambda_2$  fixed, vice versa for profL2.

The minimum and maximum values between which the cross-validated log likelihood is to be profiled can be given as minlambda1 and maxlambda1 or minlambda2 and maxlambda2, respectively. The default value of minlambda1 and minlambda2 is at zero and the default value of maxlambda1 is at the maximal value of  $\lambda_1$ , that is the smallest value that shrinks all regression coefficients to zero, but maxlambda2 has no default.

The number of steps between the minimal and maximal values can be given in the *steps* argument (default 100). These steps are equally spaced if the argument log is FALSE or equally spaced on the log scale if the argument log is TRUE. Note that the default value of log differs between profL1 (FALSE) and profL2 (FALSE). If log is TRUE, minlambda1 or minlambda2 must be given by the user as the default value is not usable.

By default, the profiling is stopped prematurely when the cross-validated log likelihood drops below the cross-validated log likelihood of the null model with all penalized regression coefficients equal to zero. This is done because it avoids lengthy calculations at small values of  $\lambda$  when the models are most likely not interesting. The automatic stopping can be controlled using the option minsteps (default steps/5). The algorithm only considers stopping prematurely after it has done at least minsteps steps. Setting minsteps=steps cancels the



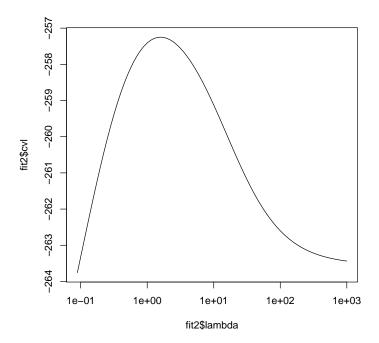
automatic stopping.

The functions profL1 and profL2 return a named list with the same elements as returned by cvl, but each of cvl, predictions, fullfit is now a *vector* or a *list* (as appropriate) as multiple cross-validated likelihoods were calculated. An additional vector lambda is returned which lists the values of  $\lambda_1$  or  $\lambda_2$  at which the cross-validated likelihood was calculated.

The allocation of the subjects into cross-validation folds is done only once, so that all cross-validated likelihoods are done with the same allocation. This makes the cross-validated log likelihoods more comparable. As in cvl the allocation is returned in fold.

## 3.4 optimizing the cross-validated likelihood

Often we are not interested in the whole profile of the cross-validated likelihood, but only in the optimum. The functions optL1 and optL2 can be used to find the optimal value of  $\lambda_1$  or  $\lambda_2$ .



The algorithm used for the optimization is the Brent algorithm for minimization without derivatives (Brent, 1973, see also help(optimize)). When using this algorithm, it is important to realize that this algorithm is guaranteed to work only for unimodal functions and that it may converge to a local maximum. This is especially relevant for L1 optimization, as the cross-validated likelihood as a function of  $\lambda_1$  very often has several local maxima. It is recommended to only use optL1 in combination with profL1 to prevent convergence to the wrong optimum. The cross-validated likelihood as a function of  $\lambda_2$ , on the other hand, is far better behaved and practically never has local maxima. The function optL2 can safely be used without combining it with profL2.

The functions optL1 and optL2 take the same arguments as cv1, and some additional ones.

The arguments minlambda1 and maxlambda1, and minlambda2 and maxlambda2 can be used to specify the range between which the cross-validated log likelihood is to be optimized. Both arguments can be left out in both functions, but supplying them can improve convergence speed. In optL1, the parameter range can be use to ensure that the function converges to the right maximum. In optL2 the user also can supply only one of minlambda2 and maxlambda2 to give the algorithm advance information of the order of magnitude of  $\lambda_2$ . In this case, the algorithm may search an find the optimum below minlambda2 or above maxlambda2.

The functions optL1 and optL2 return a named list just as cvl, with an additional element lambda which returns the optimum found. The returned

- cvl, predictions, fullfit all relate to the optimal  $\lambda$  found.
- > opt1 <- optL1(Surv(time, event), nki70[, 50:70], fold = fit1\$fold)
- > opt1\$lambda
- [1] 2.319363
- > opt1\$cv1
- [1] -260.6773
- > opt2 <- optL2(Surv(time, event), nki70[, 50:70], fold = fit2\$fold)

## References

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