pensim Package Example (Version 1.2.6)

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1 Introduction

This package acts as a wrapper to the *penalized* R package to add the following functionality to that package:

- repeated tuning on different data folds, with parallelization to take advantage of multiple processors
- two-dimensional tuning of the Elastic Net
- calculation of cross-validated risk scores through a nested cross-validation procedure, with parallelization to take advantage of multiple processors

It also provides a function for simulation of collinear high-dimensional data with survival or binary response.

This package was developed in support of the study by Waldron et al.[3]. This paper contains greater detail on proper application of the methods provided here. Please cite this paper when using the pensim package in your research, as well as the penalized package[2].

2 Example data

pensim provides example data from a microarray experiment investigating survival of cancer patients with lung adenocarcinomas (Beer et al., 2002)[1]. Load this data and do an initial pre-filter of genes with low IQR:

```
> library(pensim)
> data(beer.exprs)
> data(beer.survival)
> ##select just 250 genes to speed computation, just for the sake of example:
> set.seed(1)
> beer.exprs.sample <- beer.exprs[sample(1:nrow(beer.exprs), 250), ]
> #
> gene.quant <- apply(beer.exprs.sample,1,quantile,probs=0.75)
> dat.filt <- beer.exprs.sample[gene.quant>log2(100),]
> gene.iqr <- apply(dat.filt,1,IQR)
> dat.filt <- as.matrix(dat.filt[gene.iqr>0.5,])
> dat.filt <- t(dat.filt)
> dat.filt <- data.frame(dat.filt)
> #
> library(survival)
> surv.obj <- Surv(beer.survival$os,beer.survival$status)</pre>
```

Note that the expression data are in "wide" format, with one column per predictor (gene). It is recommended to put covariate data in a dataframe object, rather than a matrix.

3 Nested cross-validation

Unbiased estimation of prediction accuracy involves two levels of cross-validation: an outer level for estimating prediction accuracy, and an inner level for model tuning. This process is simplified by the opt.nested.crossval function.

It is recommended first to establish the arguments for the penalized regression by testing on the *penalized* package functions optL1 (for LASSO), optL2 (for Ridge) or cvl (for Elastic Net). Here we use LASSO:

Now pass these arguments to opt.nested.splitval() for cross-validated calculation and assessment of risk scores, with the additional arguments:

- outerfold and nprocessors: number of folds for the outer level of cross-validation, and the number of processors to use for the outer level of cross-validation (see ?opt.nested.crossval)
- optFUN and scaling: which optimization function to use (opt1D for LASSO or Ridge, opt2D for Elastic Net) see ?opt.splitval.
- setpen and nsim: setpen defines whether to do LASSO ("L1") or Ridge ("L2"), nsim defines the number of times to repeat tuning (see ?opt1D. opt2D has different required arguments.)

Ideally nsim would be 50, and outerfold and fold would be 10, but the values below speed computation 200x compared to these recommended values. Note that here we are using the standardize=TRUE argument of optL1 rather than the scaling=TRUE argument of opt.splitval. These two approaches to scaling are roughly equivalent, but the scaling approaches are not the same (scaling=TRUE does z-score, standardize=TRUE scales to unit central L2 norm), and results will not be identical. Also, using standardize=TRUE scales variables but provides coefficients for the original scale, whereas using scaling=TRUE scales variables in the training set then applies the same scales to the test set.

3.1 Summarization and plotting

standardize=TRUE,
trace=FALSE)

Cox fit on the continuous risk predictions:

```
> coxfit.continuous <- coxph(surv.obj~preds)
> summary(coxfit.continuous)

Call:
coxph(formula = surv.obj ~ preds)

n= 86, number of events= 24

coef exp(coef) se(coef) z Pr(>|z|)
preds 0 1 0 NA NA
```

```
exp(coef) exp(-coef) lower .95 upper .95
preds 1 1 1 1 1

Concordance= 0.5 (se = 0)
Rsquare= 0 (max possible= 0.889)
Likelihood ratio test= 0 on 1 df, p=1
Wald test = NaN on 1 df, p=NaN
Score (logrank) test = 0 on 1 df, p=1
```

Dichotomize the cross-validated risk predictions at the median, for visualization:

```
> preds.dichot <- preds > median(preds)
```

Plot the ROC curve:

4 Getting coefficients for model fit on all the data

Finally, we can get coefficients for the model fit on all the data, for future use. Note that nsim should ideally be greater than 1, to train the model using multiple foldings for cross-validation. The output of opt1D or opt2D will be a matrix with one row per simulation. The default behavior in opt.nested.crossval() is to take the simulation with highest cross-validated partial log likelihood (CVL), which is the recommended way to select a model from the multiple simulations.

We can also include unpenalized covariates, if desired. Note that when keeping only one variable for a penalized or unpenalized covariate, indexing a dataframe like [1] instead of doing [,1] preserves the variable name. With [,1] the variable name gets converted to "".

```
> if(require(survivalROC)){
+ nobs <- length(preds)
+ cutoff <- 12
+ preds.roc <- survivalROC(Stime=beer.survival$os,status=beer.survival$status,
+ marker=preds,predict.time=cutoff,span = 0.25*nobs^(-0.20))
+ plot(preds.roc$FP, preds.roc$TP, type="l", xlim=c(0,1), ylim=c(0,1),
+ xlab=paste("FP", "\n", "AUC = ",round(preds.roc$AUC,3)),
+ ylab="TP",main="LASSO predictions\n ROC curve at 12 months")
+ abline(0,1)
+ }</pre>
```

LASSO predictions ROC curve at 12 months

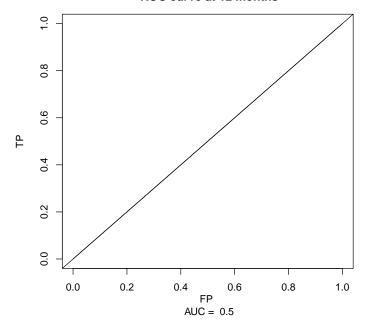


Figure 1: ROC plot of cross-validated continuous risk predictions at 12 months. Note that the predictions are better if you don't randomly select 250 genes to start with! We only did this to ease the load on the CRAN checking servers.

Note the non-zero first coefficient this time, due to it being unpenalized:

```
> beer.coefs[1,1:5]
                           #example output with no unpenalized covariates
                         U21931_at Y00285_s_at
                                                  S78187_at
   11.20907 -115.85524
                            0.00000
                                        0.00000
                                                    0.00000
> beer.coefs.unpen[1,1:5] #example output with first covariate unpenalized
                             U21931_at Y00285_s_at
                      cvl
   5.2235265 -116.8993095
                            -0.1292559
                                          0.0000000
   S78187_at
   0.0000000
```

5 Simulation of collinear high-dimensional data with survival or binary outcome

The *pensim* also provides a convenient means to simulation high-dimensional expression data with (potentially censored) survival outcome or binary outcome which is dependent on specified covariates.

5.1 Binary outcome example

First, generate the data. Here we simulate 20 variables. The first 15 (group "a") are uncorrelated, and have no association with outcome. The final five (group "b") have covariance of 0.8 to each other variable in that group. The response variable is associated with the first variable group "b" (firstonly=TRUE) with a coefficient of 2.

Binary outcomes for $n_s=50$ samples are simulated as a Bernoulli distribution with probability for patient s:

$$p_s = \frac{1}{1 + exp(-\beta X_s)} \tag{1}$$

with $\beta_{s,16} = 0.5$ and all other $\beta_{s,i}$ equal to zero. The code for this simulation is as follows:

Take a look at the simulated data:

> summary(x)

> x\$summary

```
start end cors associations num firstonly a 1 15 0.0 0 15 TRUE b 16 20 0.8 2 5 TRUE
```

A simple logistic model fails at variable selection in this case:

```
> simplemodel <- glm(outcome ~ ., data=x$data,family=binomial)
> summary(simplemodel)
```

Call:

```
glm(formula = outcome ~ ., family = binomial, data = x$data)
```

Deviance Residuals:

Max	3 Q	Median	1Q	Min
3.791e-05	2.100e-08	2.100e-08	-2.100e-08	-4.465e-05

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	75.233	45187.690	0.002	0.999
a.1	94.930	65813.216	0.001	0.999
a.2	7.524	31367.116	0.000	1.000
a.3	81.200	65185.586	0.001	0.999
a.4	42.582	38128.625	0.001	0.999
a.5	-64.709	36980.196	-0.002	0.999
a.6	-57.537	44697.640	-0.001	0.999
a.7	-3.683	37403.857	0.000	1.000
a.8	146.339	66255.304	0.002	0.998
a.9	11.253	45956.861	0.000	1.000
a.10	25.305	35717.718	0.001	0.999
a.11	-85.010	39157.131	-0.002	0.998
a.12	-89.055	41178.391	-0.002	0.998
a.13	-16.231	24105.780	-0.001	0.999
a.14	-53.423	34691.188	-0.002	0.999
a.15	10.269	39541.185	0.000	1.000
b.1	224.568	93229.703	0.002	0.998
b.2	-76.192	56044.269	-0.001	0.999
b.3	-4.161	103161.605	0.000	1.000

```
b.4
                37.759 107257.878
                                      0.000
                                               1.000
                 19.410 74039.554
                                      0.000
                                               1.000
b.5
(Dispersion parameter for binomial family taken to be 1)
    Null deviance: 6.7301e+01 on 49
                                       degrees of freedom
Residual deviance: 1.5078e-08
                                on 29
                                       degrees of freedom
AIC: 42
Number of Fisher Scoring iterations: 25
   But LASSO does a better job, selecting several of the collinear variables in
the "b" group of variables which are associated with outcome:
> lassofit <- opt1D(nsim=3,nprocessors=1,setpen="L1",penalized=x$data[1:20],</pre>
                     response=x$data[,"outcome"],trace=FALSE,fold=10)
> print(lassofit)
           L1
                     cvl (Intercept)
                                                        a.2
                                             a.1
[1,] 1.684304 -29.79442
                           0.6927236 0.40512950 0.1343187
[2,] 5.103898 -30.16723
                           0.5408337 0.01193180 0.0000000
[3,] 4.525225 -29.94452
                           0.5584223 0.07802578 0.0000000
                       a.4
            a.3
                                  a.5 a.6 a.7 a.8 a.9 a.10
[1,] 0.52905955 0.1732812 -0.3337187
                                                           0
[2,] 0.00000000 0.0000000 0.0000000
                                         0
                                             0
                                                 0
                                                           0
                                             0
[3,] 0.01594438 0.0000000 0.0000000
                                                           0
                                                   b.1 b.2
                        a.12 a.13 a.14 a.15
           a.11
[1,] -0.2961634 -0.48388262
                                0
                                     0
                                           0 1.0459555
[2,] 0.0000000 0.00000000
                                0
                                     0
                                           0 0.2896943
                                                          0
[3,] 0.0000000 -0.01284617
                                      0
                                           0 0.3816052
                                                          0
     b.3
               b.4 b.5
[1,]
       0 0.4816783
                      0
[2,]
       0 0.5049148
```

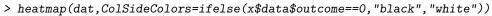
And visualize the data as a heatmap:

0 0.5119789

5.2 Survival outcome example

We simulate these data in the same way, but with response="timetoevent". Here censoring is uniform random between times 2 and 10, generating approximately 34% censoring:

```
> dat <- t(as.matrix(x$data[,-match("outcome",colnames(x$data))]))</pre>
```



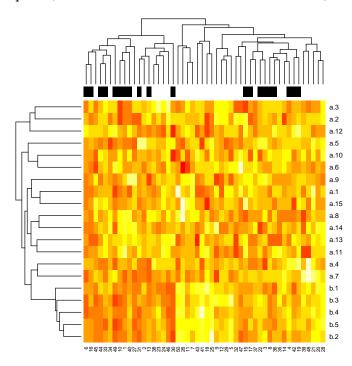


Figure 2: Heatmap of simulated data with binary response.

```
+ firstonly=c(TRUE, TRUE),
+ nsamples=50,
+ censoring=c(2,10),
+ response="timetoevent")

How many events are censored?
> sum(x$data$cens==0)/nrow(x$data)

[1] 0.38

Kaplan-Meier plot of this simulated cohort:
> library(survival)
> surv.obj <- Surv(x$data$time,x$data$cens)
> plot(survfit(surv.obj~1),ylab="Survival probability",xlab="time")
```

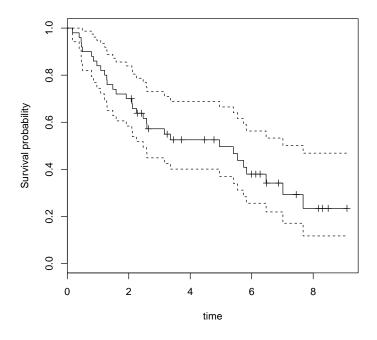


Figure 3: Kaplan-Meier plot of survival of simulated cohort.

References

- [1] David G. Beer, Sharon L.R. Kardia, Chiang-Ching Huang, Thomas J. Giordano, Albert M. Levin, David E. Misek, Lin Lin, Guoan Chen, Tarek G. Gharib, Dafydd G. Thomas, Michelle L. Lizyness, Rork Kuick, Satoru Hayasaka, Jeremy M.G. Taylor, Mark D. Iannettoni, Mark B. Orringer, and Samir Hanash. Gene-expression profiles predict survival of patients with lung adenocarcinoma. Nat Med, 8(8):816–824, 2002.
- [2] Jelle J Goeman. L1 penalized estimation in the cox proportional hazards model. Biometrical Journal. Biometrische Zeitschrift, 52(1):70–84, February 2010. PMID: 19937997.
- [3] Levi Waldron, Melania Pintilie, Ming-Sound Tsao, Frances A Shepherd, Curtis Huttenhower, and Igor Jurisica. Optimized application of penalized regression methods to diverse genomic data. *Bioinformatics*, 27(24):3399–3406, 2011. PMID: 22156367.

> sessionInfo()

R version 2.15.1 (2012-06-22)

Platform: x86_64-unknown-linux-gnu (64-bit)

locale:

[1] LC_CTYPE=en_US.UTF-8 LC_NUMERIC=C
[3] LC_TIME=en_US.UTF-8 LC_COLLATE=C

[5] LC_MONETARY=en_US.UTF-8 LC_MESSAGES=en_US.UTF-8

[7] LC_PAPER=C LC_NAME=C
[9] LC_ADDRESS=C LC_TELEPHONE=C
[11] LC_MEASUREMENT=en_US.UTF-8 LC_IDENTIFICATION=C

attached base packages:

- [1] splines stats graphics grDevices utils
- [6] datasets methods base

other attached packages:

- [1] survivalROC_1.0.3 penalized_0.9-41 pensim_1.2.6
- [4] survival_2.36-14

loaded via a namespace (and not attached):

[1] MASS_7.3-22 snow_0.3-10 tools_2.15.1