Restricted permutations; using the permute Package

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Abstract

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1. Introduction

In classical frequentist statistics, the significance of a relationship or model is determined by reference to a null distribution for the test statistic. This distribution is derived mathematically and the probability of achieving a test statistic as large or larger if the null hypothesis were true is looked-up from this null distribution. In deriving this probability, some assumptions about the data or the errors are made. If these assumptions are violated, then the validity of the derived p-value may be questioned.

An alternative to deriving the null distribution from theory is to generate a null distribution for the test statistic by randomly shuffling the data in some manner, refitting the model and deriving values for the test statistic for the permuted data. The level of significance of the test can be computed as the proportion of values of the test statistic from the null distribution that are equal to or larger than the observed value.

In many data sets, simply shuffling the data at random is inappropriate; under the null hypothesis, that data are not freely exchangeable. If there is temporal or spatial correlation, or the samples are clustered in some way, such as multiple samples collected from each of a number of fields. The **permute** was designed to provide facilities for generating these restricted permutations for use in randomisation tests.

2. Simple randomisation

As an illustration of both randomisation and the use of the **permute** package we consider a small data set of mandible length measurements on specimens of the golden jackal (*Canis aureus*) from the British Museum of Natural History, London, UK. These data were collected as part of a study comparing prehistoric and modern canids (Higham *et al.* 1980), and were analysed by Manly (2007). There are ten measurements of mandible length on both male and female specimens. The data are available in the jackal data frame supplied with **permute**.

```
R> require(permute)
R> data(jackal)
R> jackal
Length Sex
1 120 Male
```

```
2
      107
             Male
3
             Male
      110
4
      116
             Male
5
      114
             Male
6
             Male
      111
7
      113
             Male
8
      117
             Male
9
             Male
      114
      112
             Male
10
      110 Female
11
12
      111 Female
13
      107 Female
14
      108 Female
      110 Female
15
16
      105 Female
      107 Female
17
18
      106 Female
19
      111 Female
20
      111 Female
```

The interest is whether there is a difference in the mean mandible length between male and female golden jackals. The null hypothesis is that there is zero difference in mandible length between the two sexes or that females have larger mandible. The alternative hypothesis is that males have larger mandibles. The usual statistical test of this hypothesis is a one-sided t test, which can be applied using t.test()

The observed t is 3.484 with 18 df. The probability of observing a value this large or larger if the null hypothesis were true is 0.0013. Several assumptions have been made in deriving this p-value, namely

- 1. random sampling of individuals from the populations of interest,
- 2. equal population standard deviations for males and females, and
- 3. that the mandible lengths are normally distributed within the sexes.

Assumption 1 is unlikely to be valid for museum specimens such as these, that have been collected in some unknown manner. Assumption 2 may be valid, Fisher's F-test and a Fligner-Killeen test both suggest that the standard deviations of the two populations do not differ significantly

This assumption may be relaxed using var.equal = FALSE (the default) in our call to t.test(), to employ Welch's modification for unequal variances. Assumption 3 may be valid, but with such a small sample we are able to reliably test this.

A randomisation test of the same hypothesis can be performed by randomly allocating ten of the mandible lengths to the male group and the remaining lengths to the female group. This randomisation is justified under the null hypothesis because the observed difference in mean mandible length between the two sexes is just a typical value for the difference in a sample if there were no difference in the population. An appropriate test statistic needs to be selected. We could use the t statistic as derived in the t-test. Alternatively, we could base our randomisation test on the difference of means D_i (male - female).

The main function in **permute** for providing random permutations is **shuffle()**. We can write our own randomisation test for the **jackal** data by first creating a function to compute the difference of means for two groups

```
R> meanDif <- function(x, grp) {
+     mean(x[grp == "Male"]) - mean(x[grp == "Female"])
+ }</pre>
```

which can be used in a simple for() loop to generate the null distribution for the difference of means. First, we allocate some storage to hold the null difference of means; here we use 4999 random permutations so allocate a vector of length 5000. Then we iterate, randomly generating an ordering of the Sex vector and computing the difference means for that permutation.

The observed difference of means was added to the null distribution, because under the null hypothesis the observed allocation of mandible lengths to male and female jackals is just one of the possible random allocations.

The null distribution of D_i can be visualised using a histogram, as shown in Figure 1. The observed difference of means (4.8) is indicated by the red tick mark.

```
R> hist(Djackal, main = "",
+ xlab = expression("Mean difference (Male - Female) in mm"))
R> rug(Djackal[5000], col = "red", lwd = 2)
```

The number of values in the randomisation distribution equal to or larger than the observed difference is

```
R> (Dbig <- sum(Djackal >= Djackal[5000]))
```

[1] 12

giving a permutational p-value of

```
R> Dbig/length(Djackal)
```

[1] 0.0024

which is comparable with that determined from the frequentist t-test, and indicate strong evidence against the null hypothesis of no difference.

In total there ${}^{20}C_{10} = 184,756$ possible allocations of the 20 observations to two groups of ten

```
R> choose(20, 10)
```

[1] 184756

so we have only evaluated a small proportion of these in the randomisation test.

The main workhorse function we used above was shuffle(). In this example, we could have used the base R function sample() to generate the randomised indices perm that were used to permute the Sex factor. Where shuffle() comes into it's own is for generating permutation indices from restricted permutation designs.

3. The shuffle() and shuffleSet() functions

In the previous section I introduced the shuffle() function to generate permutation indices for use in a randomisation test. Now we will take a closer look at shuffle() and explore the various restricted permutation designs from which it can generate permutation indices.

shuffle() has two arguments: i) n, the number of observations in the data set to be permuted, and ii) control, a list that defines the permutation design describing how the samples should be permuted.

```
R> args(shuffle)
function (n, control = permControl())
NULL
```

A series of convenience functions are provided that allow the user to set-up even quite complex permutation designs with little effort. The user only needs to specify the aspects of the design they require and the convenience functions ensure all configuration choices are set and passed on to shuffle(). The main convenience function is permControl(), which return a list specifying all the options available for controlling the sorts of permutations returned by shuffle()

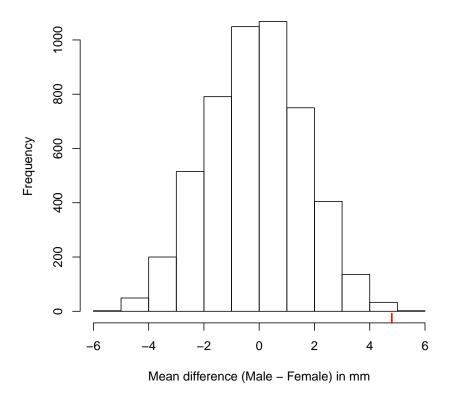


Figure 1: Distribution of the difference of mean mandible length in random allocations, ten to each sex.

R> str(permControl())

List of 10

\$ strata : NULL
\$ nperm : num 199
\$ complete : logi FALSE
\$ within :List of 5
..\$ type : chr "free"
..\$ constant: logi FALSE
..\$ mirror : logi FALSE
..\$ ncol : NULL

..\$ nrow : NULL
\$ blocks :List of 4
..\$ type : chr "none"
..\$ mirror: logi FALSE

..\$ ncol : NULL
..\$ nrow : NULL

\$ maxperm : num 9999 \$ minperm : num 99 \$ all.perms : NULL

\$ observed : logi FALSE

```
$ name.strata: chr "NULL"
- attr(*, "class")= chr "permControl"
```

The defaults describe a random permutation design where all objects are freely exchangeable. Using these defaults, shuffle(10) amounts to sample(1:10, 10, replace = FALSE):

```
R> set.seed(2)
R> (r1 <- shuffle(10))

[1] 2 7 5 10 6 8 1 3 4 9

R> set.seed(2)
R> (r2 <- sample(1:10, 10, replace = FALSE))

[1] 2 7 5 10 6 8 1 3 4 9

R> all.equal(r1, r2)

[1] TRUE
```

3.1. Generating restricted permutations

Several types of permutation are available in **permute**:

- Free permutation of objects
- Time series or line transect designs, where the temporal or spatial ordering is preserved.
- Spatial grid designs, where the spatial ordering is preserved in both coordinate directions
- Permutation of blocks or groups of samples.

The first three of these can be nested within the levels of a factor or to the levels of that factor, or to both. Such flexibility allows the analysis of split-plot designs using permutation tests.

permControl() is used to set up the design from which shuffle() will draw a permutation. permControl() has two main arguments that specify how samples are permuted within blocks of samples or at the block level itself. These are within and blocks. Two convenience functions, Within() and Blocks() can be used to set the various options for permutation.

For example, to permute the observations 1:10 assuming a time series design for the entire set of observations, the following control object would be used

```
R> set.seed(4)
R> x <- 1:10
R> CTRL <- permControl(within = Within(type = "series"))
R> perm <- shuffle(10, control = CTRL)
R> perm

[1] 7 8 9 10 1 2 3 4 5 6

R> x[perm] ## equivalent

[1] 7 8 9 10 1 2 3 4 5 6
```

It is assumed that the observations are in temporal or transect order. We only specified the type of permutation within blocks, the remaining options were set to their defaults via Within().

A more complex design, with three blocks, and a 3 by 3 spatial grid arrangement within each block can be created as follows

Visualising the permutation as the 3 matrices may help illustrate how the data have been shuffled

```
R> ## Original
R> lapply(split(1:27, block), matrix, ncol = 3)
$'1'
     [,1] [,2] [,3]
[1,]
             4
        1
[2,]
        2
              5
                   8
[3,]
        3
              6
                   9
$'2'
     [,1] [,2] [,3]
[1,]
       10
            13
                  16
[2,]
       11
             14
                  17
[3,]
            15
       12
                  18
$'3'
     [,1] [,2] [,3]
[1,]
       19
            22
                  25
[2,]
            23
                  26
       20
[3,]
       21
            24
                  27
R> ## Shuffled
R> lapply(split(perm, block), matrix, ncol = 3)
$'1'
     [,1] [,2] [,3]
[1,]
        6
             9
                   3
              7
[2,]
        4
                   1
[3,]
        5
             8
                   2
$'2'
     [,1] [,2] [,3]
[1,]
       14
            17
                  11
[2,]
       15
            18
                  12
[3,]
            16
       13
                  10
```

```
$'3'
[,1] [,2] [,3]
[1,] 22 25 19
[2,] 23 26 20
[3,] 24 27 21
```

In the first grid, the lower-left corner of the grid was set to row 2 and column 2 of the original, to row 1 and column 2 in the second grid, and to row 3 column 2 in the third grid.

To have the same permutation within each level of block, use the constant argument of the Within() function, setting it to TRUE

```
R> set.seed(4)
R> CTRL <- permControl(strata = block,</pre>
                         within = Within(type = "grid", ncol = 3, nrow = 3,
                                           constant = TRUE))
R> perm2 <- shuffle(length(block), control = CTRL)</pre>
R> lapply(split(perm2, block), matrix, ncol = 3)
$'1'
      [,1] [,2] [,3]
[1,]
        6
              9
                    3
              7
[2,]
         4
                    1
[3,]
         5
              8
                    2
$'2'
      [,1] [,2] [,3]
[1,]
        15
             18
                   12
[2,]
        13
             16
                   10
[3,]
             17
        14
                   11
$'3'
      [,1] [,2] [,3]
[1,]
        24
             27
                   21
[2,]
        22
             25
                   19
[3,]
       23
             26
                   20
```

3.2. Generating sets of permutations with shuffleSet()

There are several reasons why one might wish to generate a set of n permutations instead of repeatedly generating permutations one at a time. Interpreting the permutation design happens each time shuffle() is called. This is an unnecessary computational burden, especially if you want to perform tests with large numbers of permutations. Furthermore, having the set of permutations available allows for expedited use with other functions, they can be iterated over using for loops or the apply family of functions, and the set of permutations can be exported for use outside of R.

The shuffleSet() function allows the generation of sets of permutations from any of the designs available in **permute**. shuffleSet() takes an additional argument to that of shuffle(), nset, which is the number of permutations required for the set. Internally, shuffle() and shuffleSet() are very similar, with the major difference being that shuffleSet() arranges repeated calls to the workhorse permutation-generating functions with only the overhead associated with interpreting the permutation design once. shuffleSet() returns a matrix where the rows represent different permutations in the set.

As an illustration, consider again the simple time series example from earlier. Here I generate a set of 5 permutations from the design, with the results returned as a matrix

```
R> set.seed(4)
R> CTRL <- permControl(within = Within(type = "series"))</pre>
R> pset <- shuffleSet(10, nset = 5, control = CTRL)</pre>
R> pset
     [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [,9] [,10]
[1,]
                                    2
                              1
[2,]
                                   7
                              6
                                         8
                                                   10
                                                          1
[3,]
        4
             5
                   6
                        7
                              8
                                   9
                                        10
                                                          3
                                              1
                        7
[4,]
             5
                   6
                              8
                                        10
                                              1
                                                          3
[5,]
       10
```

Computational details

R> toLatex(sessionInfo())

- R version 2.13.1 Patched (2011-09-13 r57007), x86_64-unknown-linux-gnu
- Locale: LC_CTYPE=en_GB.utf8, LC_NUMERIC=C, LC_TIME=en_GB.utf8, LC_COLLATE=C, LC_MONETARY=C, LC_MESSAGES=en_GB.utf8, LC_PAPER=en_GB.utf8, LC_NAME=C, LC_ADDRESS=C, LC_TELEPHONE=C, LC_MEASUREMENT=en_GB.utf8, LC_IDENTIFICATION=C
- Base packages: base, datasets, grDevices, graphics, methods, stats, utils
- Other packages: permute 0.6-2
- Loaded via a namespace (and not attached): tools 2.13.1

References

Higham C, Kijngam A, Manly B (1980). "An analysis of prehistoric canid remains from Thailand." *Journal of Archaeological Science*, **7**, 149–165.

Manly B (2007). Randomization, bootstrap and Monte Carlo methods in biology. 3rd edition. Chapman & Hall/CRC, Boca Raton.