Machine Learning Methods for Gene Expression Data

Day 3

Models

Regularizatio

Linear Classifiers

GLMs

I D A

Naive Bayes

References

Machine Learning Methods for Gene Expression Data

Day 3

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Outline

Machine Learning Methods for Gene Expression Data

Day 3

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Regularizatio

Linear

GLMs

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1 Linear Models

2 Regularization

3 Linear Classifiers

4 GLMs

5 LDA

6 Naive Bayes

Linear Models for Univariate Feature Selection

Machine Learning Methods for Gene Expression Data

Day 3

Linear Models

Regularization

Linear Classifiers

GLMs

Naive Bayes

References

Most common univariate filter is *t*-test (or *F*-test if more than 2 groups) originating from model

$$\mathbb{P}(X_g = x \mid Y = y) = \frac{1}{\sqrt{2\pi\sigma_g^2}} \exp\left[\frac{(x - \mu_{yg})^2}{2\sigma_g^2}\right]$$

Now considering conditional probabilities of X_g given Y ...

- ... and considering each X_g separately!
 - "Univariate" analysis; not necessarily realistic, but tractable.

Linear Models

Machine Learning Methods for Gene Expression Data

Day 3

Linear Models

Regularizat

Linear Classifier

GLMs

LDA

Naive Baye

References

Alternately, this model may be described by

$$X_g = \mu_{0g} + (\mu_{1g} - \mu_{0g})Y + \sigma_g \epsilon_g$$

with $\epsilon_g \sim \mathcal{N}(0,1)$, implying

$$\mathbb{E}(X_g \mid Y = y) = \mu_{yg}$$

$$\mathbb{V}(X_g \mid Y = y) = \sigma_g^2$$

Can then rank features based on $|t_g|$ where:

$$t_{g} = \frac{\hat{\mu}_{0g} - \hat{\mu}_{1g}}{\hat{\sigma}_{g}\sqrt{\frac{1}{n_{0}} + \frac{1}{n_{1}}}}$$

Linear Models

Machine Learning Methods for Gene Expression Data

Day 3

Linear Models

Regularizati

Linear Classifiers

GLMs

LDA

ivalve baye

References

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While X_g is linear in Y here,

 \blacktriangleright it is linearity in μ_{yg} that makes "linear model" in statistics.

Linear Models

Machine Learning Methods for Gene Expression Data

Models

Linear

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Naive Bay

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While X_{σ} is linear in Y here,

- \blacktriangleright it is linearity in μ_{yg} that makes "linear model" in statistics.
- With a few modifications linear models can model $Y \mid \mathbf{X}$ too.

Student's t-Test and Pearson's Correlation

Machine Learning Methods for Gene Expression Data

Day 3

Linear Models

Regularizati

Linear Classifiers

GLMs

...

References

Mentioned correlation coefficients when discussing clustering (though there correlations were between samples).

 t_g related to Pearson correlation between X_g and outcome Y:

$$t_{g} = \sqrt{n-2} \frac{\hat{\rho}(\mathbf{x}_{g}, \mathbf{y})}{\sqrt{1 - \hat{\rho}^{2}(\mathbf{x}_{g}, \mathbf{y})}}$$

or

$$\hat{\rho}(\mathbf{x}_g, \mathbf{y}) = \frac{\delta \mathbf{x}_g \cdot \delta \mathbf{y}}{\|\delta \mathbf{x}_g\| \|\delta \mathbf{y}\|}$$
$$= \frac{t_g}{\sqrt{n - 2 + t_g^2}}$$

where
$$\delta \mathbf{x} = \mathbf{x} - \sum_{i=1}^{n} x_i$$
.

Student's *t*-Test and Pearson's Correlation

Machine Learning Methods for Gene Expression Data

Day 3

Linear Models

Regularizati

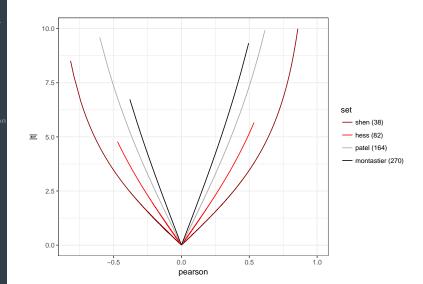
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Naive Baye

Reference



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Machine Learning Methods for Gene Expression Data

Day 3

Linear Models

Regularizatio

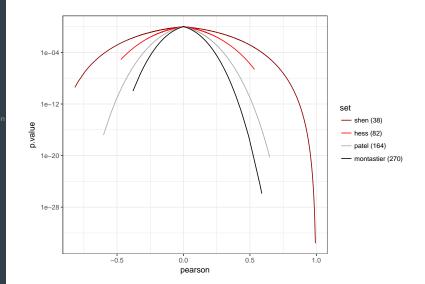
Linear

GLMs

LDA

Naive Baye

References



Multivariate Filtration by Linear Algebra

Machine Learning Methods for Gene Expression Data

Day 3

Linear Models

Linear

GLMs

Naive Baye

References

Let's say we know we like gene g.

 \blacktriangleright How can we find another gene h that complements g well?

Consider the **projection operator** $proj_u$:

$$\mathsf{proj}_{\mathbf{u}}(\mathbf{v}) = \frac{\mathbf{u} \cdot \mathbf{v}}{\|\mathbf{u}\|^2} \ \mathbf{u}$$

and the associated operator for "projecting **u** out"

$$\mathsf{projout}_{\mathbf{u}}(\mathbf{v}) = \mathbf{v} - \frac{\mathbf{u} \cdot \mathbf{v}}{\|\mathbf{u}\|^2} \mathbf{u}$$

The gene h which maximizes R^2 when paired with g in the linear model $Y = \beta_0 + \beta_g X_g + \beta_h X_h$ can be found as

$$h = \mathop{\arg\max}_{h'} \hat{\rho} \left(\mathop{\mathsf{projout}}_{\delta \mathbf{x}_g} (\delta \mathbf{x}_{h'}), \; \mathop{\mathsf{projout}}_{\delta \mathbf{x}_g} (\delta \mathbf{y}) \right)$$

Multivariate Filtration by Linear Algebra

Machine Learning Methods for Gene Expression Data

Day 3

Linear Models

Linear

CLMa

LDA

Naive Bayes

References

In matrix notation,

$$\mathsf{projout}_{\mathbf{u}}(\underline{X}) = \underline{P}^{(\mathbf{u})}\underline{X}$$
 $\mathsf{projout}_{\mathbf{u}}(\mathbf{y}) = \underline{P}^{(\mathbf{u})}\mathbf{y}$

where

$$p_{ij}^{(\mathbf{u})} = \delta_{ij} - \frac{u_i u_j}{\|\mathbf{u}\|^2}$$

Now we're back to sorting genes by correlation except that

- \blacktriangleright we've replaced X and Y
- with $\underline{\mathbf{P}}^{(\delta \mathbf{x}_g)} \delta \underline{\mathbf{X}}$ and $\underline{\mathbf{P}}^{(\delta \mathbf{x}_g)} \delta \mathbf{y}$.

Implemented in MaclearnUtilities.(R|py) as gramSchmidtSelect.

* $\delta \mathbf{x}$ and $\delta \mathbf{y}$ are again defined by:

$$\delta x_{ig} = x_{ig} - \sum_{i=1}^{n} x_{ig}$$
$$\delta y_{i} = y_{i} - \sum_{i=1}^{n} y_{i}$$

Regularization

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Day 3

Linear Models

Regularization

Classifier

GLMs

LDA

Naive Baye

References

Linear models can be made robust in high-dimensional settings using **regularization**.

Unregularized linear regression uses maximum likelihood to fit $\beta_{\mathbf{g}}$

▶ fit by ordinary least-squares (OLS) estimator:

$$\hat{eta}_{\mathsf{OLS}} = \operatorname*{\mathsf{arg\,min}}_{eta} \sum_{i} \left(y_i - eta \cdot \mathbf{x}_i \right)^2$$

Bayesian derivation of OLS uses uniform prior on β .

Regularization

Machine Learning Methods for Gene Expression Data

Day 3

Linear Models

Regularization Linear

GI Me

LDA

Naive Baye

References

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ight)^2$$

Bayesian derivation of OLS uses uniform prior on β .

If Gaussian prior (L2 regression) imposed on β , maximum a posteriori (MAP) estimator is (Park & Casella (2008)):

$$\hat{\boldsymbol{\beta}}_{L2} = \underset{\boldsymbol{\beta}}{\operatorname{arg\,min}} \left\{ \sum_{i} (y_i - \boldsymbol{\beta} \cdot \mathbf{x}_i)^2 + \phi_2 \sum_{g} \beta_g^2 \right\}$$

where regularization parameter ϕ_2 determined by width of prior.

Lasso regression

Machine Learning Methods for Gene Expression Data

Day 3

Linear Models

Regularization

Linear Classifiers

GLMs

LDA

Naive Baye

References

Alternatively, Laplace prior for β yields MAP estimator (Park & Casella (2008)):

$$\hat{\boldsymbol{\beta}}_{L1} = \underset{\boldsymbol{\beta}}{\operatorname{arg\,min}} \left\{ \sum_{i} (y_i - \boldsymbol{\beta} \cdot \mathbf{x}_i)^2 + \phi_1 \sum_{g} |\beta_g| \right\}$$

where now ϕ_1 is determined by width of the Laplace prior.

Lasso regression

Machine Learning Methods for Gene Expression Data

Day 3

Linear Models

Regularization

Linear Classifier

GI Ms

LDA

ivalve Bayes

References

Alternatively, Laplace prior for β yields MAP estimator (Park & Casella (2008)):

$$\hat{\boldsymbol{\beta}}_{\mathsf{L}1} = \operatorname*{arg\,min}_{\boldsymbol{\beta}} \left\{ \sum_{i} (y_i - \boldsymbol{\beta} \cdot \mathbf{x}_i)^2 + \phi_1 \sum_{g} |\beta_g| \right\}$$

where now ϕ_1 is determined by width of the Laplace prior.

As ϕ_1 is increased,

- progressively more β_g set to zero,
- ▶ de-selecting the corresponding features (Tibshirani (1996))
 - ▶ L1, or LASSO, regression has embedded feature selection.

Lasso regression

Machine Learning Methods for Gene Expression Data

Day 3

Linear Models

Regularization

Linear Classifier

GLIV

Naive Baye

References

Alternatively, Laplace prior for β yields MAP estimator (Park & Casella (2008)):

$$\hat{\boldsymbol{\beta}}_{\mathsf{L}1} = \operatorname*{arg\,min}_{\boldsymbol{\beta}} \left\{ \sum_{i} \left(y_{i} - \boldsymbol{\beta} \cdot \mathbf{x}_{i} \right)^{2} + \phi_{1} \sum_{\mathcal{g}} \left| \beta_{\mathcal{g}} \right| \right\}$$

where now ϕ_1 is determined by width of the Laplace prior.

As ϕ_1 is increased,

- progressively more β_g set to zero,
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 - ▶ L1, or LASSO, regression has embedded feature selection.

Both L1/LASSO and L2/ridge logistic regression are implemented in the R package glmnet

▶ function glmnet using argument family="binomial".

Day 3

Linear Models

Linear

Linear Classifiers

GLMs

LDA

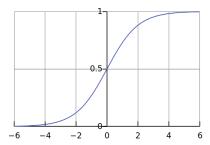
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References

In the context of classification, "linear model" often means

$$\mathbb{P}(Y = 1 \mid \mathbf{X} = \mathbf{x}) = \text{expit}(\beta_0 + \boldsymbol{\beta} \cdot \mathbf{x})$$

where expit: $\mathbb{R} \to (0,1)$ defined by $\text{expit}(u) = \frac{\exp(u)}{1 + \exp(u)}$ is the logistic, or inverse-logit, function.



Linear Classifiers

Machine Learning Methods for Gene Expression Data

Day 3

Linear Models

Linear

Linear Classifiers

GLIVI

Naive Baye

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Two main classes of such linear classification models:

- 1. linear discriminant analysis (LDA) (generative): adds assumption $\mathbb{P}(\mathbf{X}=\mathbf{x}|Y=y)\sim\mathcal{N}(\boldsymbol{\mu}_y,\boldsymbol{\Sigma})$, fit by maximizing joint likelihood $\mathbb{P}(\mathbf{X},y)$.
- 2. **logistic regression** (discriminative): makes no explicit distributional assumptions about X, instead fit by maximizing conditional $\mathbb{P}(Y \mid X)$ over training set.

Logistic Regression

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Day 3

Linear Models

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Classifier:

GLMs

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References

Logistic regression is widely used because:

- 1. it's familiar,
- 2. it has often worked well in the past,
- 3. $\beta_0 + \beta \cdot \mathbf{x} = \operatorname{logit}(\mathbb{P}(y \mid \mathbf{x})) = \operatorname{log}\left(\frac{\mathbb{P}}{1-\mathbb{P}}\right)$ may be interpreted as a "log-odds;" such quantities are often used in many areas of statistics,
- 4. there are very efficient techniques for fitting logistic regression models, and
- logistic regression is a type of generalized linear model (GLM); GLMs have many useful statistical properties that make them easy to work with and to extend (e.g., by adding regularization).

Note that (4) above may be seen as largely a consequence of (5).

Machine Learning Methods for Gene Expression Data

Day 3

Linear Models

Regularizatio

Classifier

GLMs

LDA

Maive Baye

References

Linear discriminant analysis assumes that the class densities $P(X = x \mid Y = y)$ are Gaussian with

- common covariance matrix Σ
- but distinct, class-dependent, means μ_y :

$$X \mid_{Y=y} \sim \mathcal{N}(\mu_y, \Sigma)$$

Day 3

Linear Models

Linear

Classifiers

LDA

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References

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$$X \mid_{Y=y} \sim \mathcal{N}(\mu_y, \Sigma)$$

or

$$\mathbb{P}(\mathbf{X} = \mathbf{x} \mid Y = y) = \frac{1}{(2\pi)^{p/2} |\Sigma|^{1/2}} \exp\left[-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_y)^\mathsf{T} \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu}_y)\right]$$

Day 3

Linear Models

Linear

CLASSIII

LDA

Naive Bayes

References

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Then, applying Bayes' theorem

$$\mathbb{P}(Y = y \mid \mathbf{X} = \mathbf{x}) = \frac{\pi_y \, \mathbb{P}(\mathbf{X} = \mathbf{x} \mid Y = y)}{\sum_{y'} \pi_{y'} \, \mathbb{P}(\mathbf{X} = \mathbf{x} \mid Y = y')}$$

where $\pi_{y} = \mathbb{P}(Y = y)$ is the prior probability of class y.

LDA Compared to Logistic Regression

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Day 3

Linear Models

Regularizatio

Linear Classifier

GLMs

LDA

ivalve Bayes

References

Shrinkage and Diagonal LDA

Machine Learning Methods for Gene Expression Data

Day 3

Linear Models

Regularizatio

Linear

GLMs

LDA

Naive Baye

References

LDA can also be regularized:

Modify maximum likelihood covariance matrix Σ ,

- ▶ by shrinking off-diagonal entries Σ_{gh} towards 0.
- ► Degree of shrinkage is regularization parameter.
- ► Implemented in R package sda

Shrinkage and Diagonal LDA

Machine Learning Methods for Gene Expression Data

Day 3

Linear Models

..

Classifiers

GLMs

LDA

Naive Bayes

References

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In most extreme form, shrinkage LDA sets all $\Sigma_{gh}=0$ $(g \neq h)$.

Thus Σ becomes diagonal matrix: diagonal LDA or **DLDA**.

- ▶ DLDA has been been found to be particularly useful for gene expression data (Dudoit et al. (2002)).
- ▶ Implemented in R packages sparsediscrim.

Day 3

Models

Regularizati

Linear Classifiers

GLMs

LDA

Naive Bayes

References

"Naive Bayes" describes a family of classification methods sharing a common assumption:

$$\mathbb{P}(X = x \mid Y = y) = \prod_{g} \mathbb{P}(X_g = X_g \mid Y = y)$$

which can be substituted into Bayes' formula to yield:

$$\mathbb{P}(Y = y \mid \mathbf{X} = \mathbf{x}) = \frac{\pi_y \prod_g \mathbb{P}(X_g = x_g \mid Y = y)}{\sum_{y'} \pi_{y'} \prod_g \mathbb{P}(X_g = x_g \mid Y = y')}$$

Naive Bayes

Machine Learning Methods for Gene Expression Data

Day 3

Models

Regularizati

Linear Classifiers

GLMs

LDA

Naive Bayes

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DLDA is actually a form of naive Bayes classification in which the additional assumption of linearity is posed.

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Day 3

Models

Regularizatio

Linear

GLMs

Naive Bayes

References

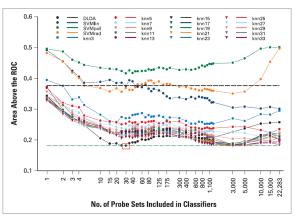


Fig 1. Mean area above the receiver operating characteristic (ROC curves plotted against the number of top genes included in the classifiers. Complete 5-fold cross validation results (means over the 100 iterations) for 20 classifier algorithms including different numbers of probe sets 39 gene sets) are shown. Green and black horizontal dotted lines indicate the mean 14-280 for the nominally best Diagonal Linear Discriminant Analysis (DLDA) classifier with 30 probe sets that was selected for independent validation, polynomial kernels (SSVM). and Knearest neitings

Taken from Hess et al. (2006).

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Day 3

Linear Models

Regulariza

Linear Classifiers

GLMs

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Naive Bayes

References

The conditional independence assumption is basically never true, but:

 frequently not enough data to accurately assess true inter-feature covariance, so that attempts to do so just lead to overfitting, and

Machine Learning Methods for Gene Expression Data

Day 3

Models

Regularizatio

Linear Classifiers

GLMs

LDA

Naive Bayes

References

The conditional independence assumption is basically never true, but:

- 1. frequently not enough data to accurately assess true inter-feature covariance, so that attempts to do so just lead to overfitting, and
- while this assumption tends to lead to overconfident classifiers—probability scores very near 0 or 1 even when wrong—it still often leads to accurate classifiers—most calls aren't wrong.

Machine Learning Methods for Gene Expression Data

Day 3

Models

Regularizatio

Linear Classifiers

GLMs

LDA

Naive Bayes

References

The conditional independence assumption is basically never true, but:

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- while this assumption tends to lead to overconfident classifiers—probability scores very near 0 or 1 even when wrong—it still often leads to accurate classifiers—most calls aren't wrong.
- 3. Naive Bayes methods work well when either:
 - features truly are independent within each class or
 - features are very tightly correlated (may actually be more relevant in gene expression context) (Rish *et al.* (2001)).

Bias-Variance Tradeoff

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Day 3

Linear Models

Regularizati

Linear Classifier

GLMs

Naive Bayes

Reference

From Wikipedia (http://en.wikipedia.org/wiki/Bias-variance_tradeoff):

The bias-variance tradeoff (or dilemma) is the problem of simultaneously minimizing two sources of error that prevent supervised learning algorithms from generalizing beyond their training set:

bias error from erroneous assumptions in the learning algorithm. High bias can cause an algorithm to miss the relevant relations between features and target outputs (underfitting).

variance error from sensitivity to small fluctuations in the training set. High variance can cause **overfitting**: modeling the random noise in the training data, rather than the intended outputs.

References I

Machine Learning Methods for Gene Expression Data

Day 3

Models

Regularizat

Linear Classifiers

GLMs

LDA

Naive Bayes

References

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