

ENERGY DATA SCIENCE

Statistical models for Time series

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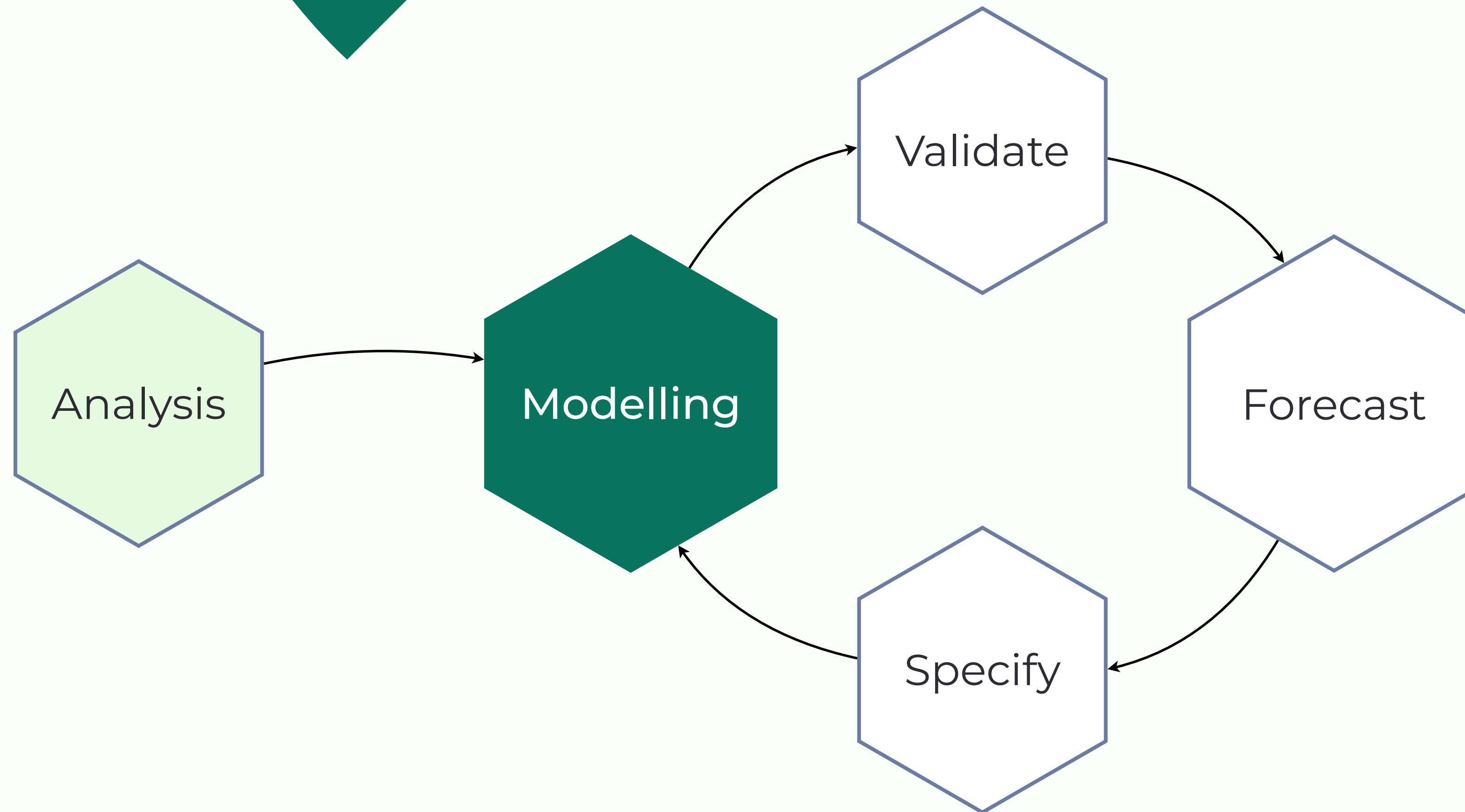
PREVIOUSLY IN COURSE ...

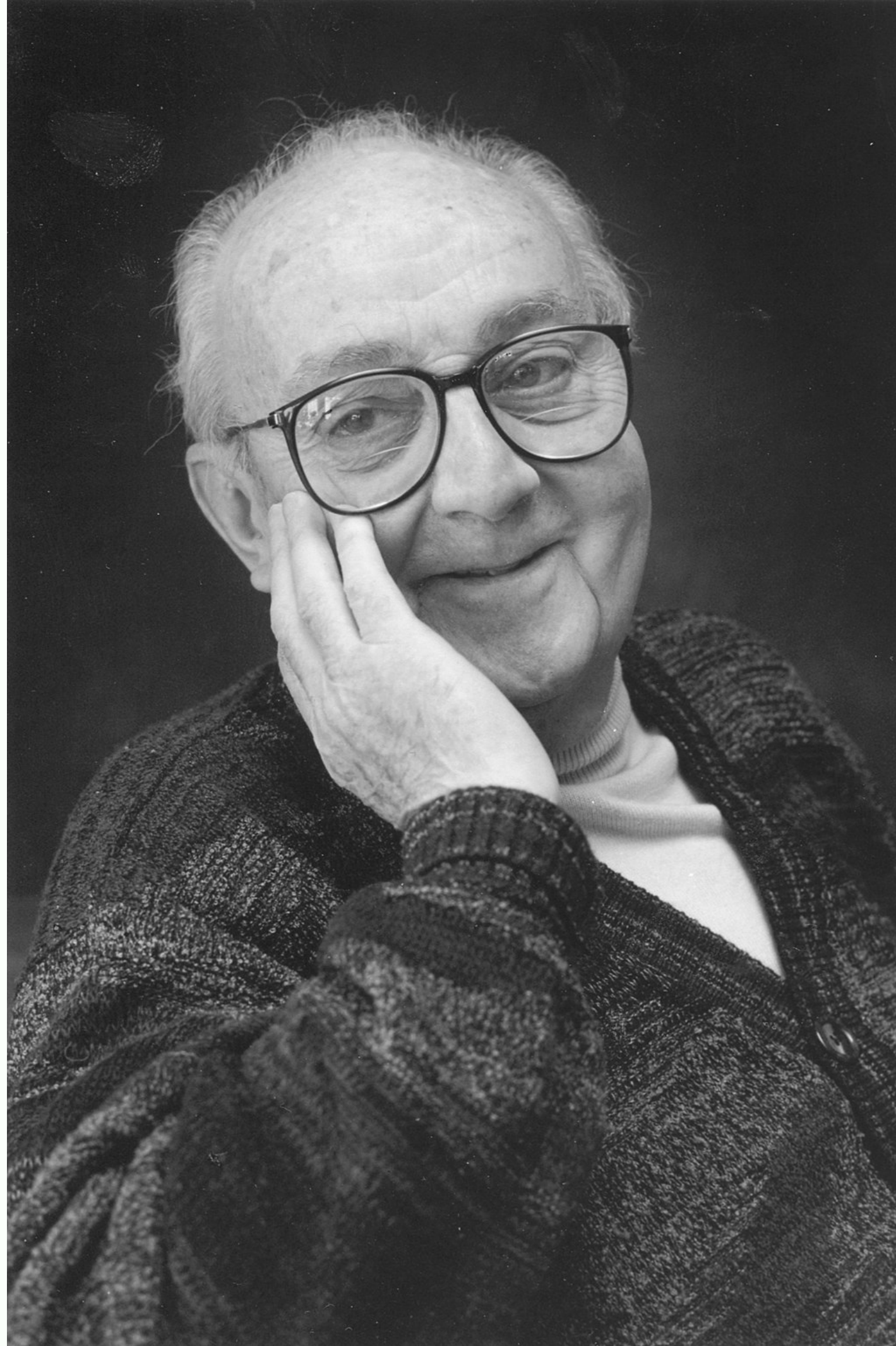
Key takeaways:

- Time series and why they are so special
- Components: trend, seasonality, cycle, remainder
- Classical decomposition

Train ML models, evaluate performance, and make predictions.

06 MODELLING





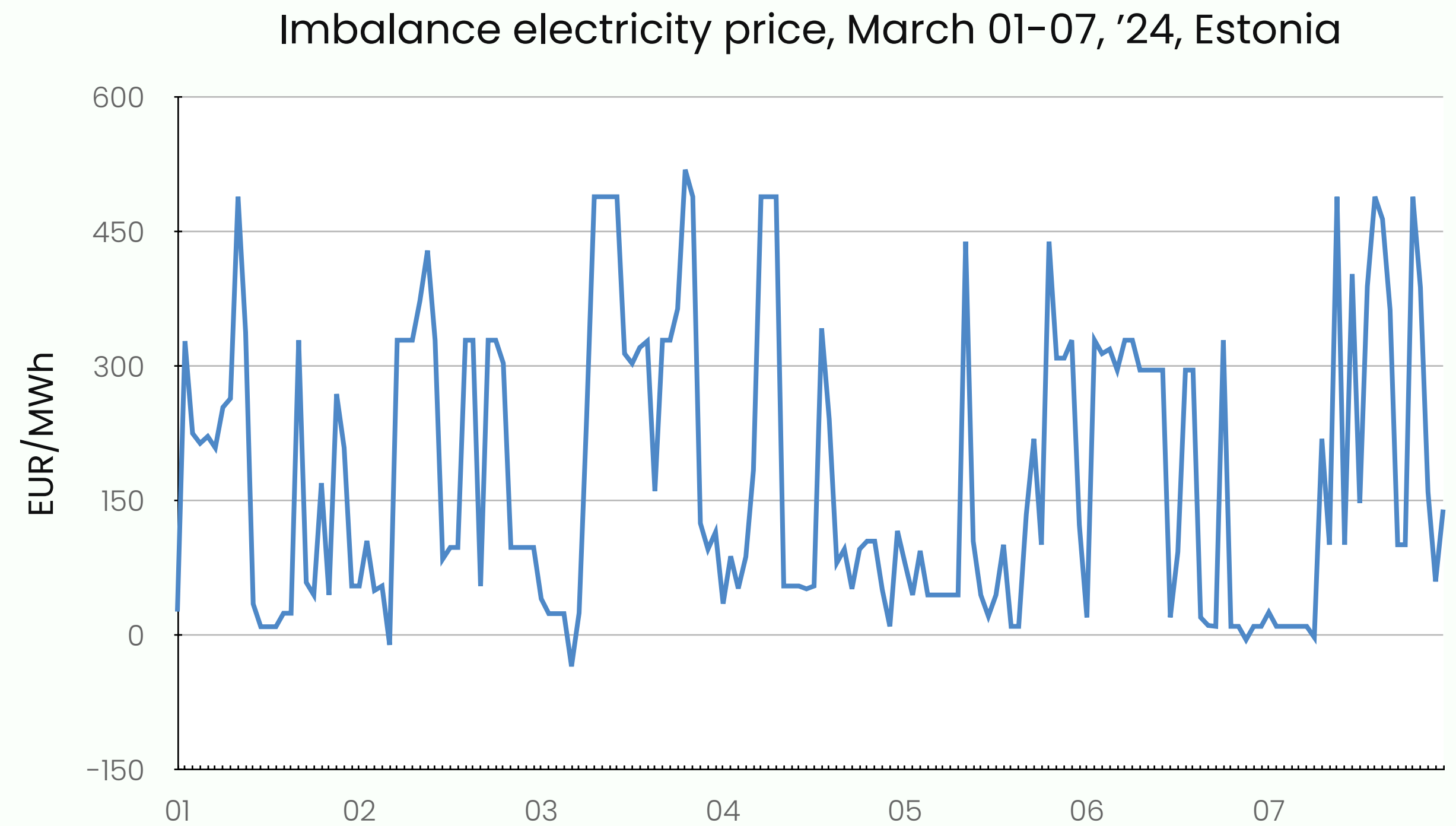
*"All models are wrong, but some
models are useful."*

George E. P. Box (1919–2013)

Stationarity

CONSISTENCY

- ✓ Cannot have independence, but want some consistency.
- ✓ Distribution depends **only** on **difference in time** and **not** on location in time.



CONSISTENCY

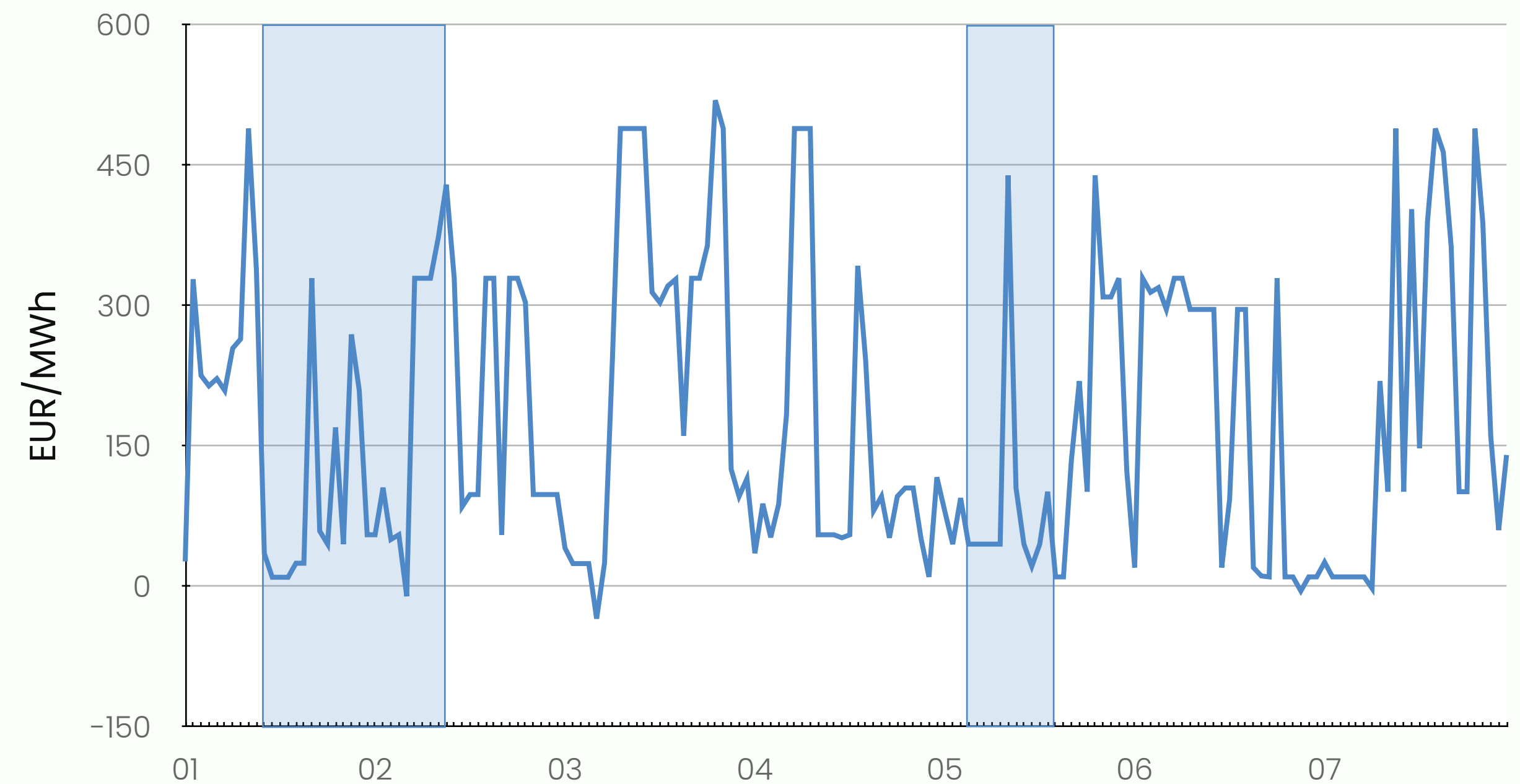
- ✓ Cannot have independence, but want some consistency.
- ✓ Distribution depends **only** on **difference in time** and **not** on location in time.

Different widths



Different distributions

Imbalance electricity price, March 01-07, '24, Estonia



STRONG STATIONARITY

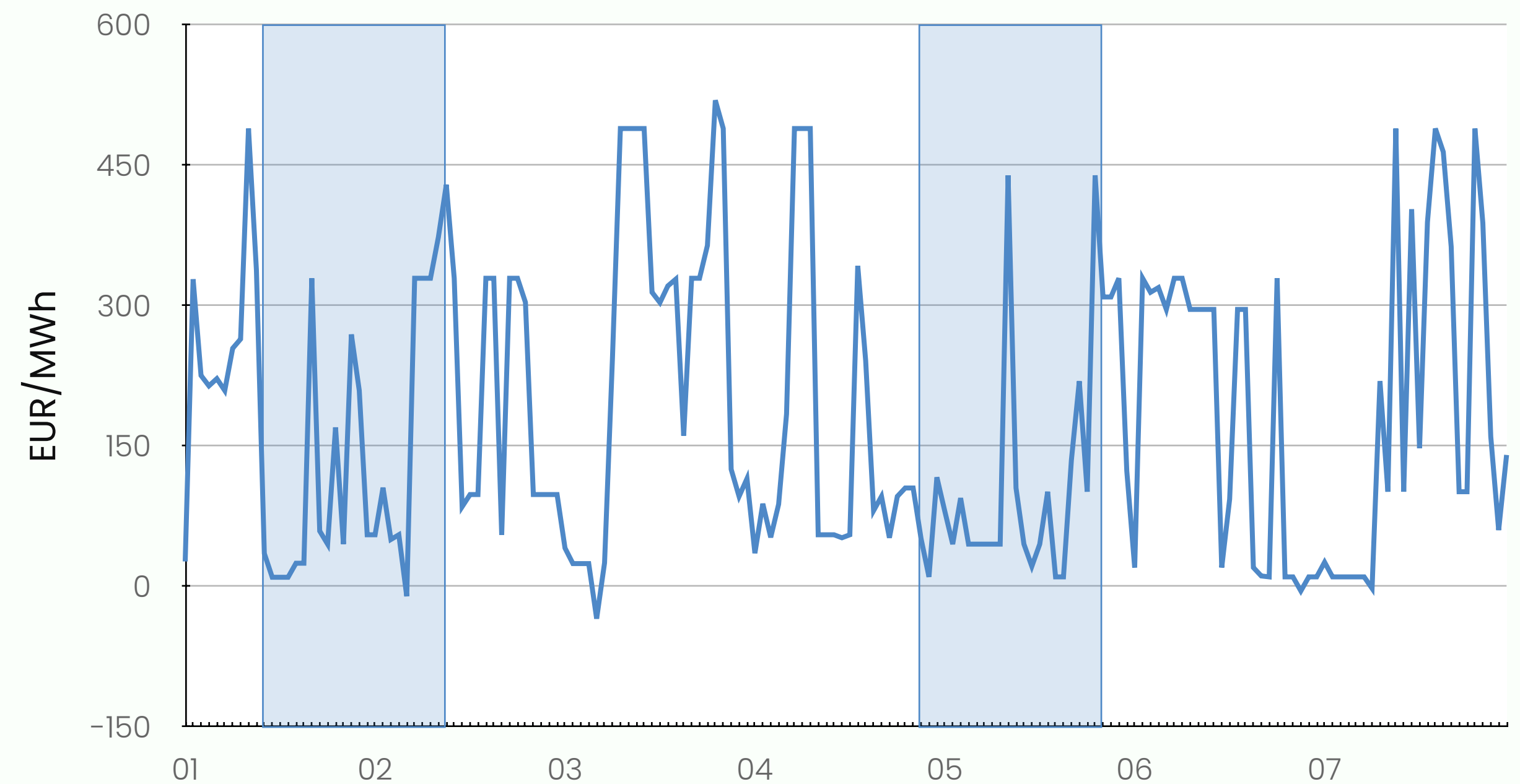
Distribution depends **only** on difference in time and **not** on location in time.

Same widths (any size)



Expect same distributions

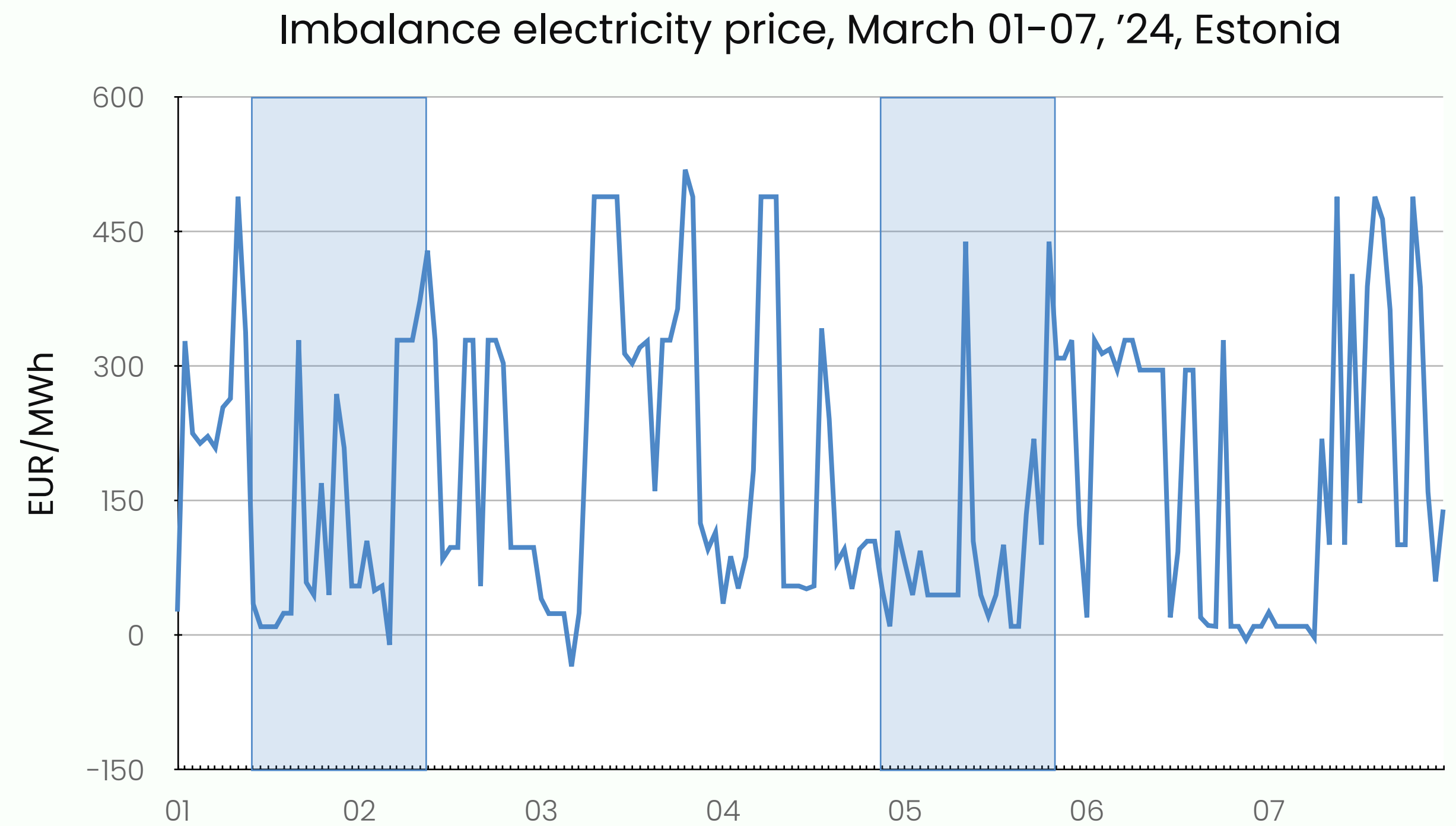
Imbalance electricity price, March 01-07, '24, Estonia



WEAK STATIONARITY

Properties depend only on difference in time,
not on location in time:

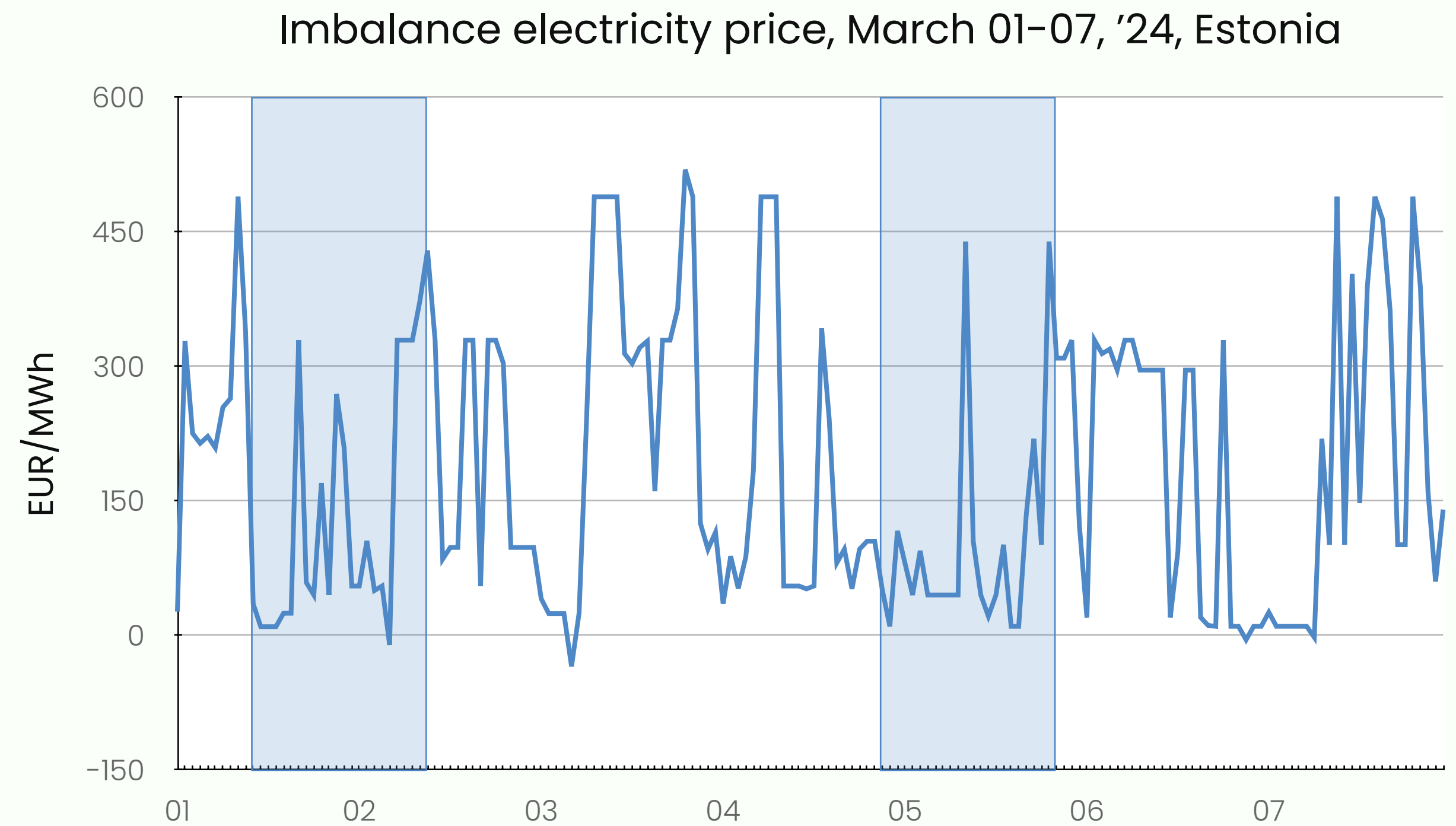
- ✓ mean is constant
- ✓ variation is constant
- ✓ no seasonality



WEAK STATIONARITY

Properties depend only on difference in time,
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- ✓ no seasonality



In most cases we only need **weak** stationarity.

WHY IS IT IMPORTANT?

Statistical properties do not change over the time:

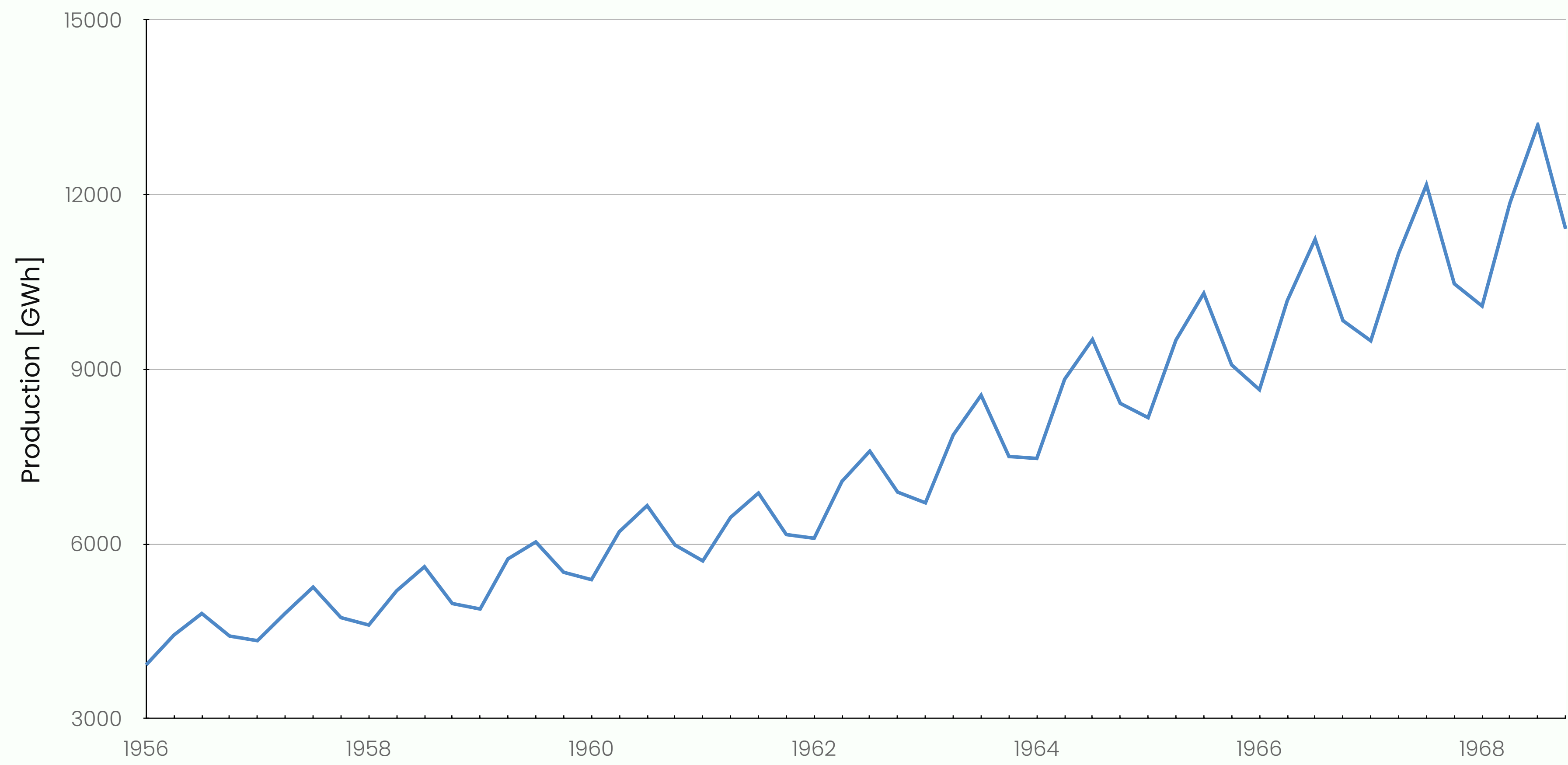
- Easy to analyse
- Easy to model

Easy to forecast.

Many existing statistical methods are applicable only when the data are stationary (e.g., ARMA-family models).

METHOD 1: VISUAL INSPECTION

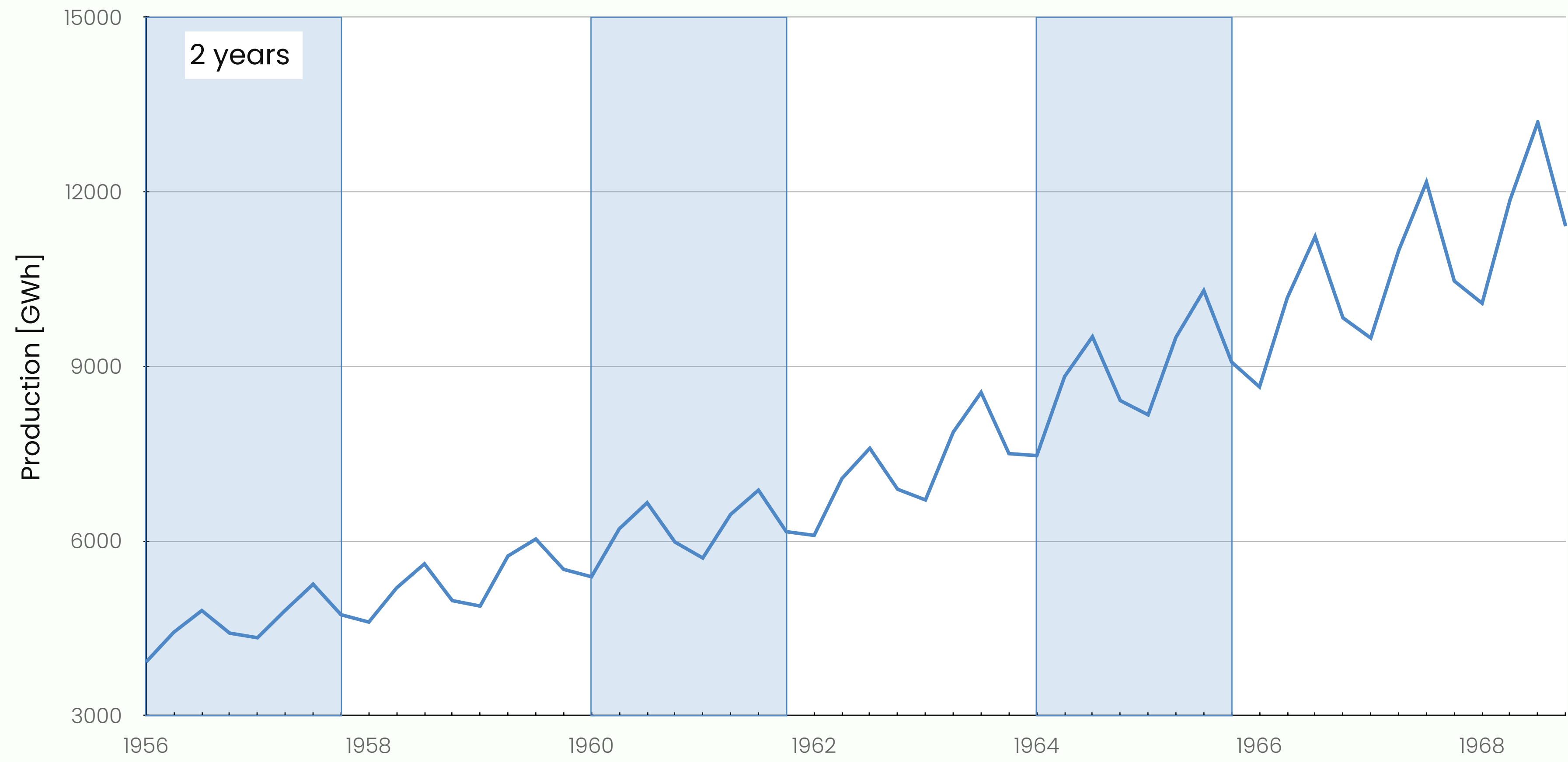
Quarterly electricity production, Australia



Stationary
Non-stationary
Mean (trend)
Variation
Seasonal

METHOD 1: VISUAL INSPECTION

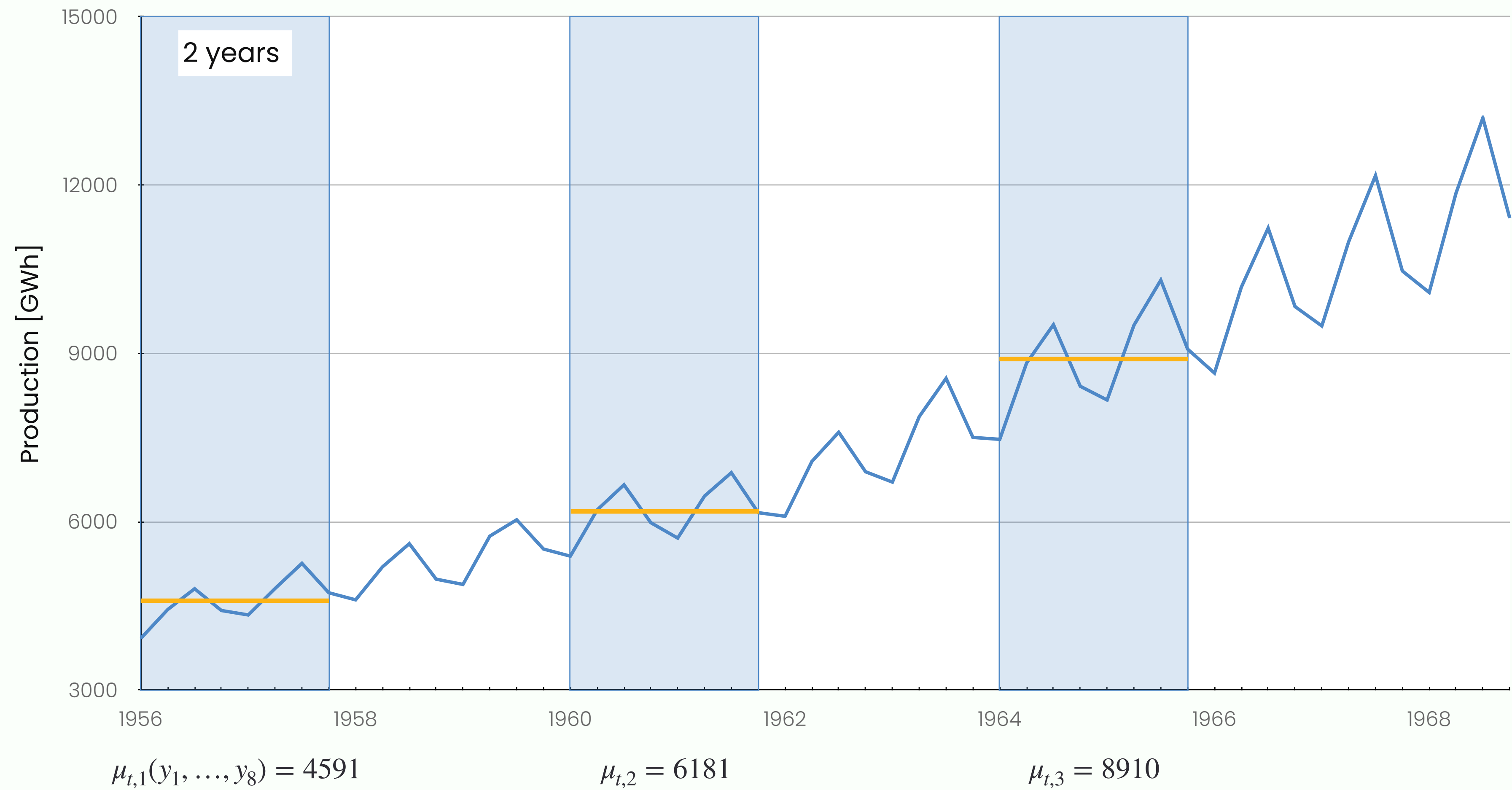
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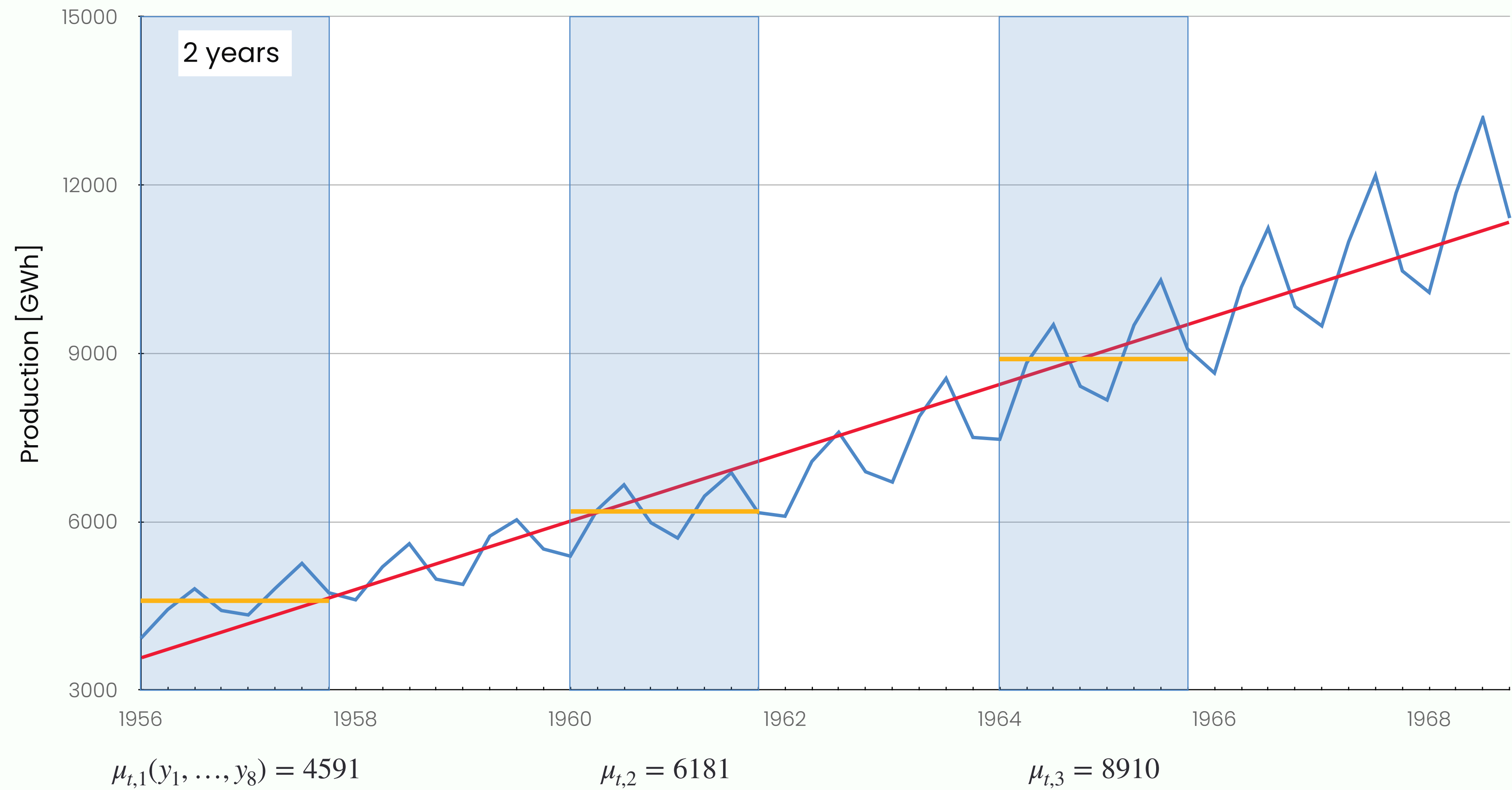
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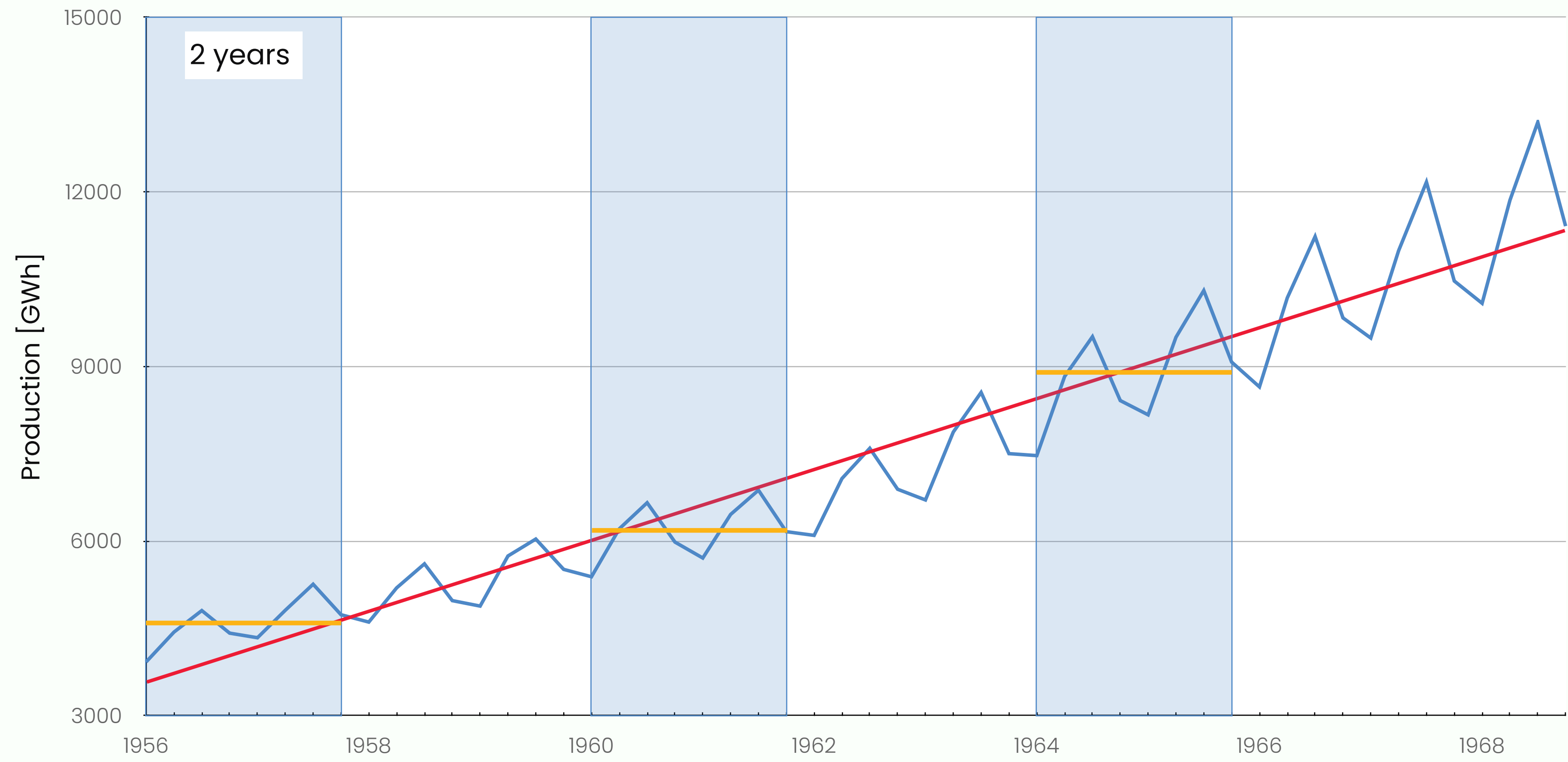
Quarterly electricity production, Australia



- Stationary
- Non-stationary
- Mean (trend) ✗
- Variation
- Seasonal

METHOD 1: VISUAL INSPECTION

Quarterly electricity production, Australia



$$\mu_{t,1}(y_1, \dots, y_8) = 4591$$

$$\sigma_{t,1}^2 = 401$$

$$\mu_{t,2} = 6181$$

$$\sigma_{t,2}^2 = 490$$

$$\mu_{t,3} = 8910$$

$$\sigma_{t,3}^2 = 892$$

Stationary

Non-stationary

Mean (trend)

×

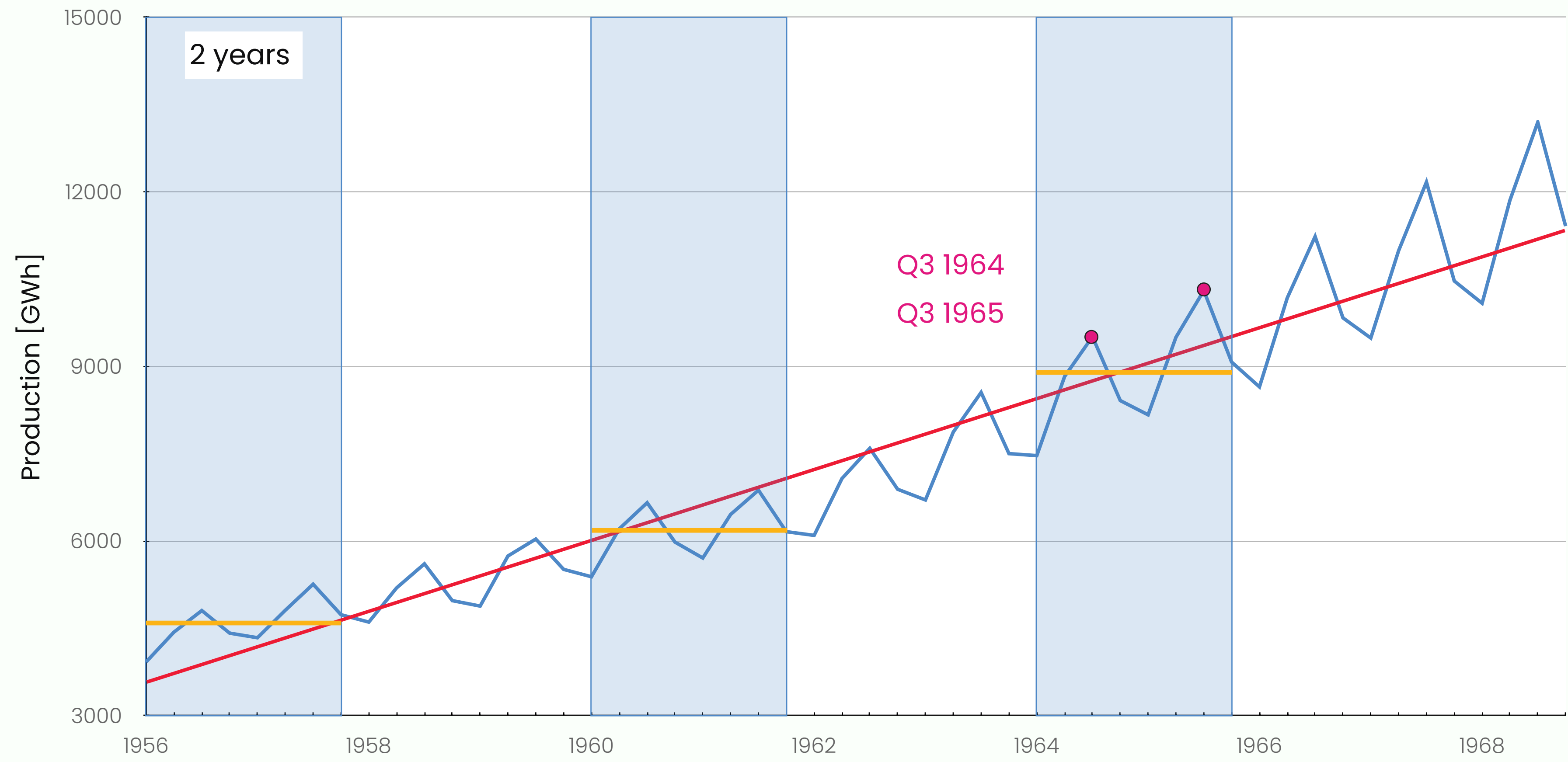
Variation

×

Seasonal

METHOD 1: VISUAL INSPECTION

Quarterly electricity production, Australia



- Stationary
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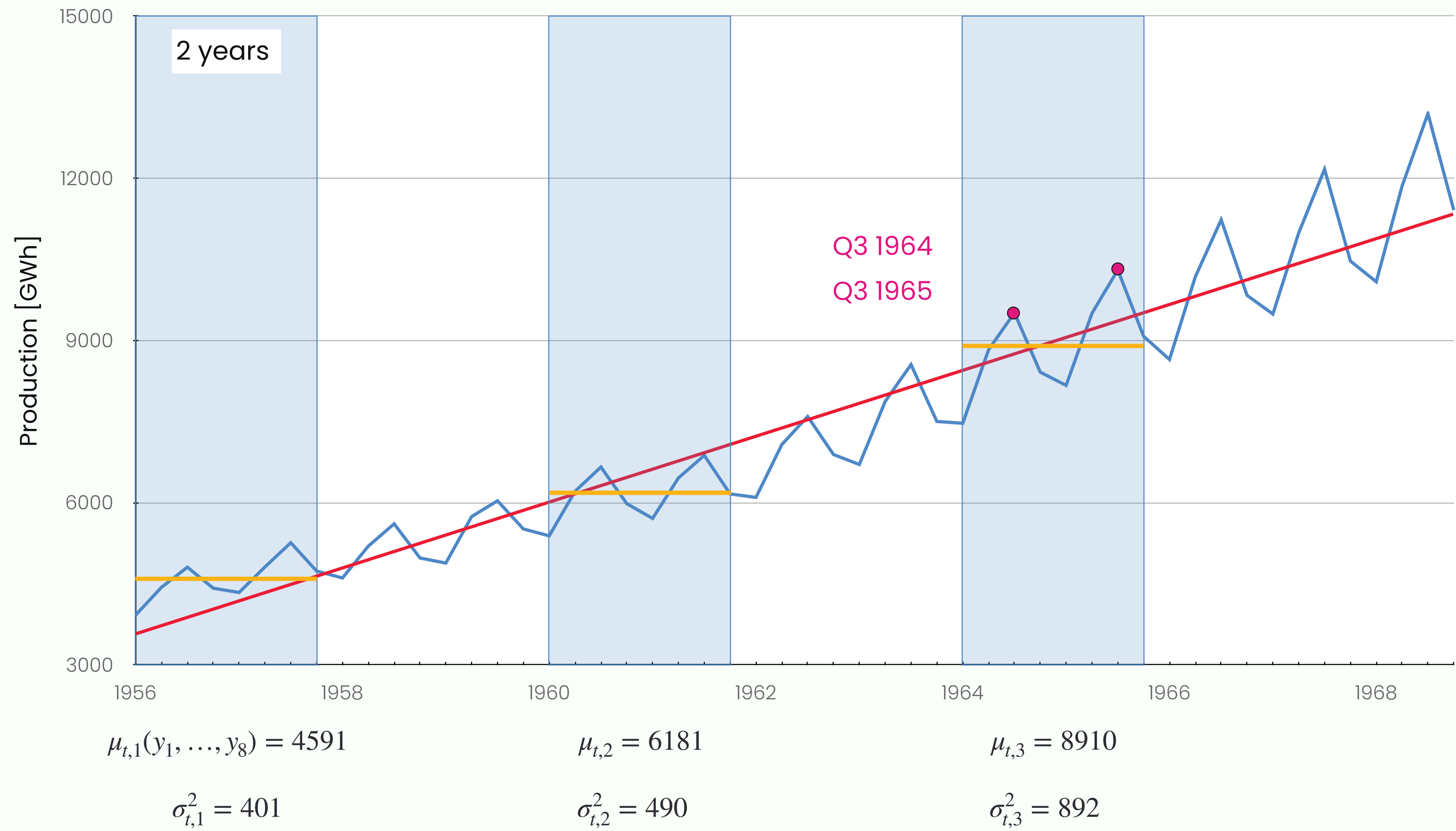
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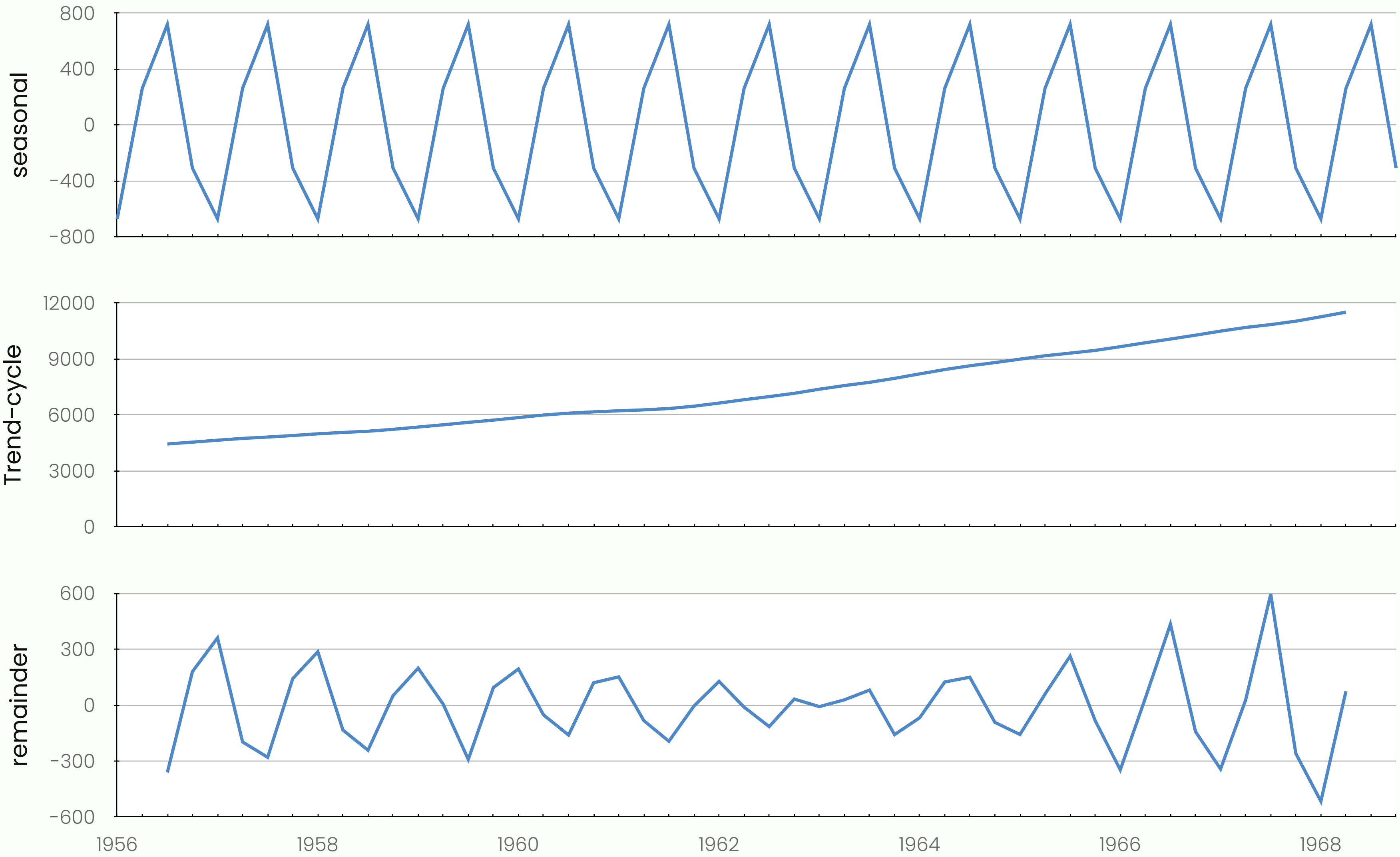


- Stationary
- Non-stationary ✓
- Mean (trend) ✗
- Variation ✗
- Seasonal ✗

METHOD 2: PARAMETRIC TESTS

- Dickey-Fuller test
- Augmented Dickey-Fuller (ADF) test
- Phillips-Perrone test
- Kwiatkowski-Phillips-Schmidt-Shin (KPSS) test

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Original data

ADF Statistic: 2.233387

p-value: 0.998909

Critical Values:

1%: -3.578

5%: -2.925

10%: -2.601

Remainder (decomposed)

ADF Statistic: -11.701448

p-value: 0.0000

Critical Values:

1%: -3.585

5%: -2.928

10%: -2.602

ACF & PACF

AUTOCORRELATION FUNCTION (ACF)

ACF measures the self-similarity of the signal over different delay times (*lagged values*).

ACF: INTUITION

Forecast average monthly electricity price.

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Forecast average monthly electricity price.

y_t December

y_{t-1} November

y_{t-2} October

⋮

ACF: INTUITION

Forecast average monthly electricity price.

| | | | | | | | | | |
|-----------|----------|--|--|--|-----------|--|-----------|--|-------|
| y_t | December | | | | y_{t-2} | | y_{t-1} | | y_t |
| y_{t-1} | November | | | | | | | | |
| y_{t-2} | October | | | | | | | | |
| | ⋮ | | | | | | | | |

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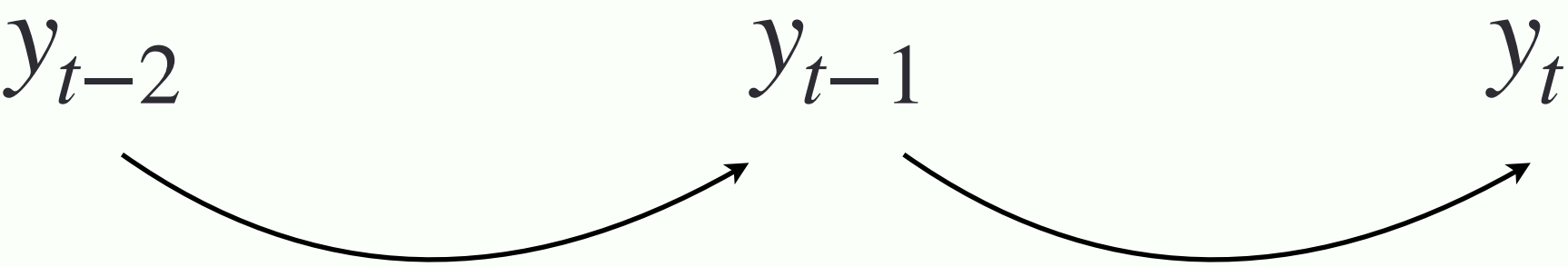
| | | | | | | | | | |
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| y_{t-2} | October | | | | | | | | |
| | ⋮ | | | | | | | | |

$\text{corr}(y_{t-2}, y_t)$

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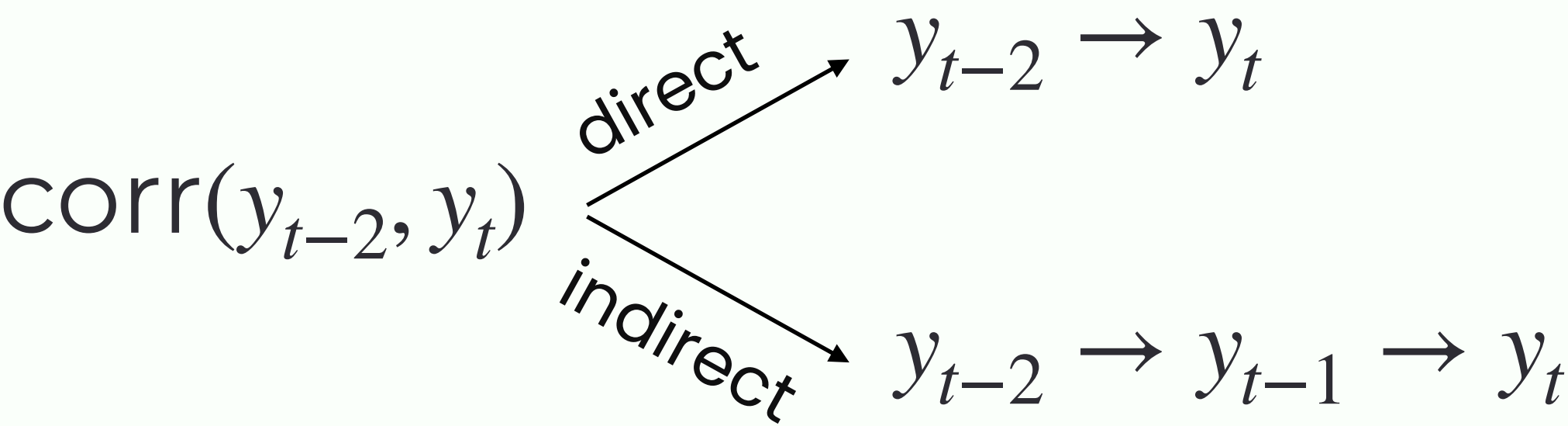
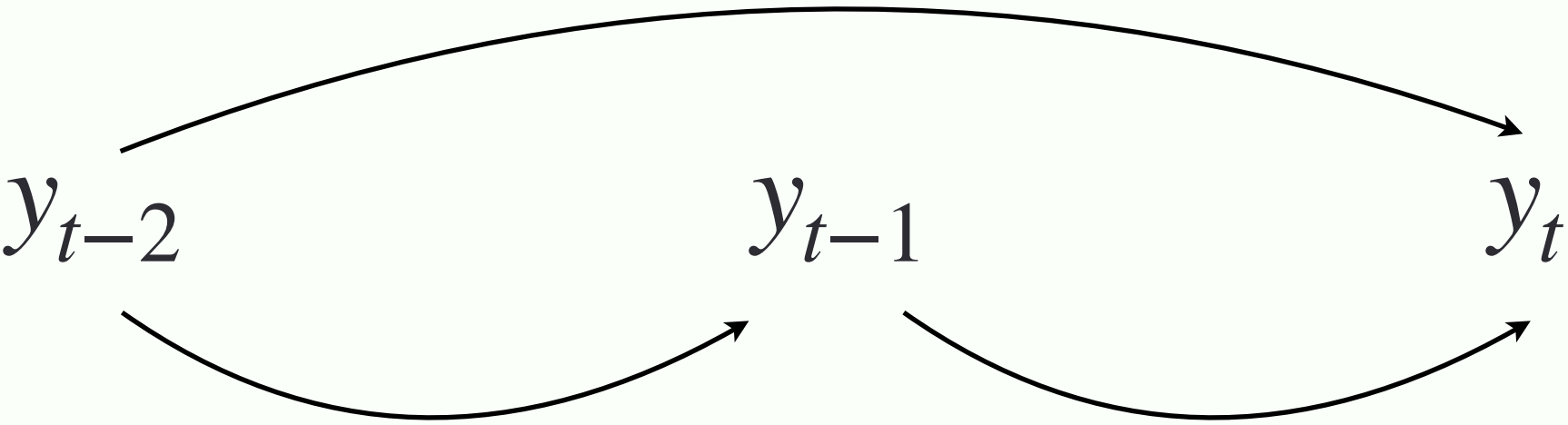


$\text{corr}(y_{t-2}, y_t)$ *indirect* $\rightarrow y_{t-2} \rightarrow y_{t-1} \rightarrow y_t$

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Forecast average monthly electricity price.

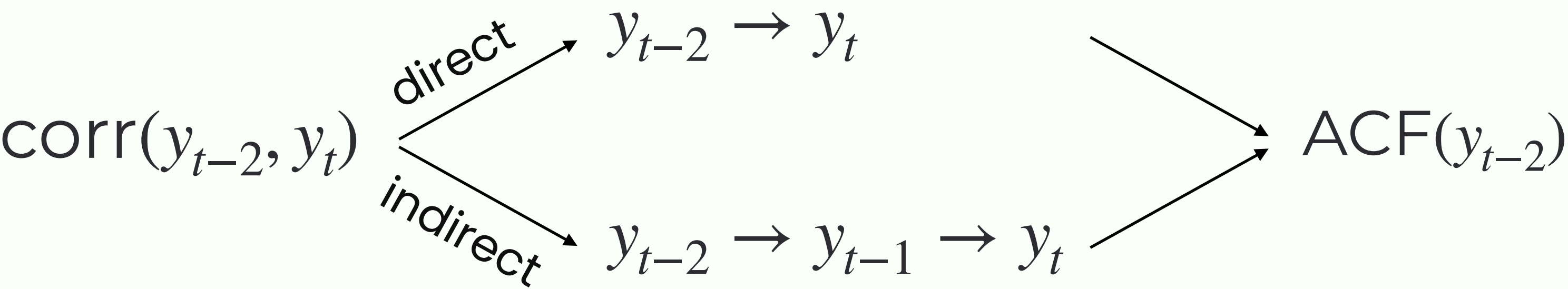
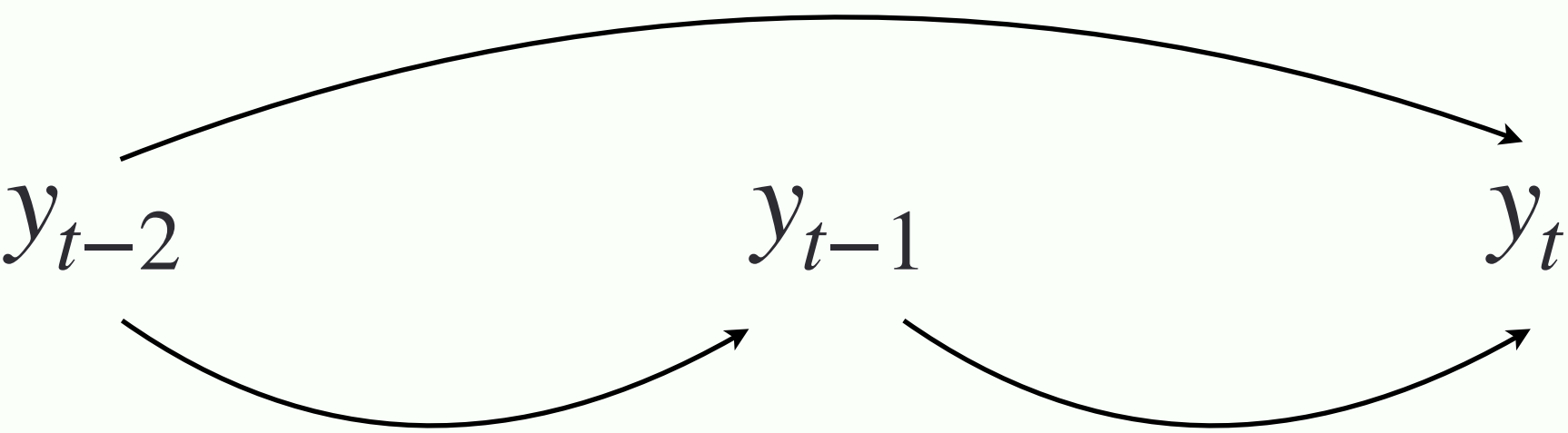
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⋮



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Forecast average monthly electricity price.

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⋮



MATHEMATICALLY SPEAKING

Let r_i denote the autocorrelation coefficient.

The value of r_i can be calculated as:

$$r_i = \frac{\sum_{t=i+1}^T (y_t - \bar{y})(y_{t-i} - \bar{y})}{\sum_{t=1}^T (y_t - \bar{y})^2},$$

where T is the length of the time series, and \bar{y} is the average of a time series signal.

AUTOCORRELATION PLOT

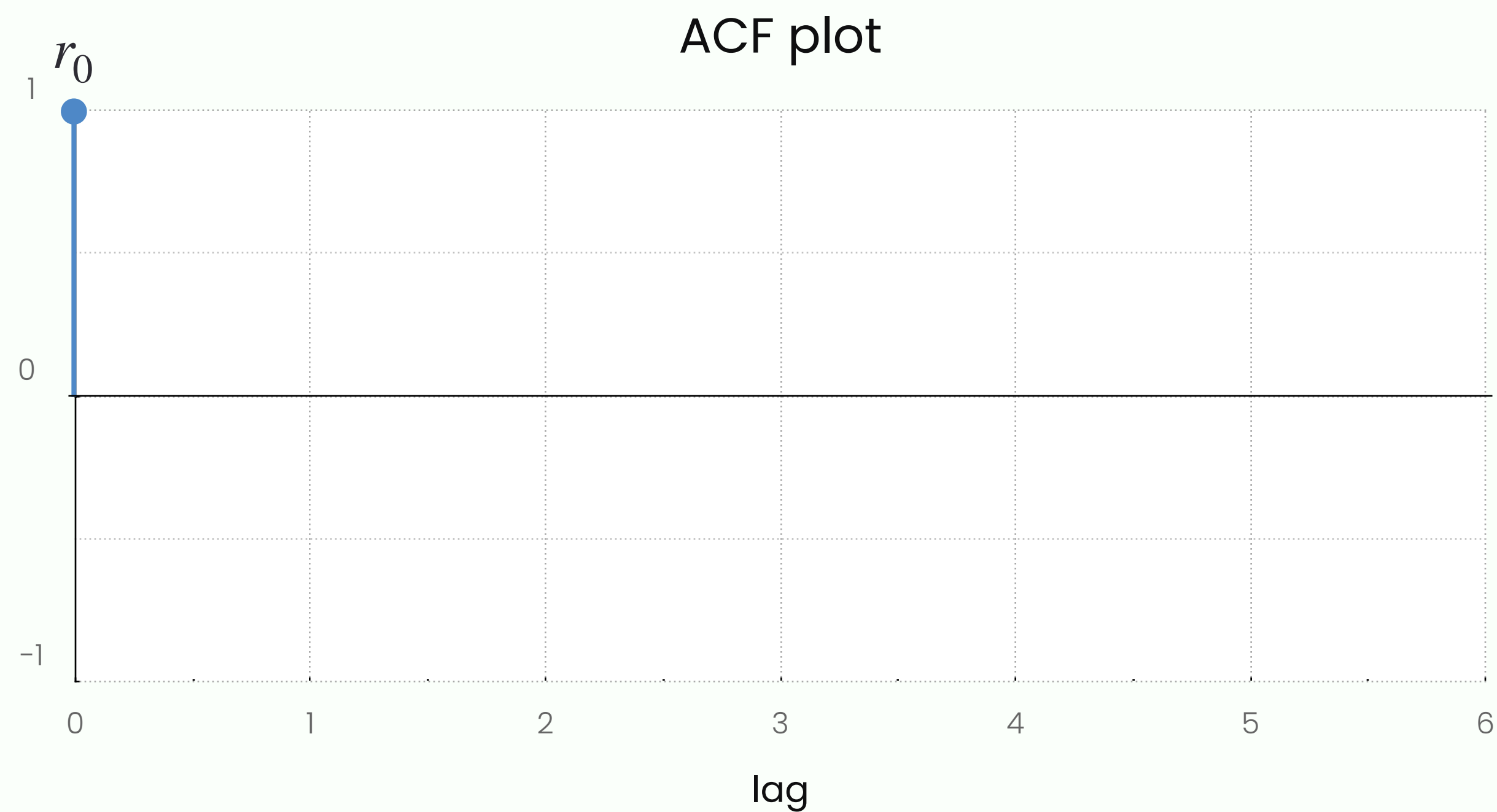
Autocorrelation coefficient r_i measures the relationship between y_t and y_{t-i} .

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For example,

- r_0 measures the relationship between y_t and itself;

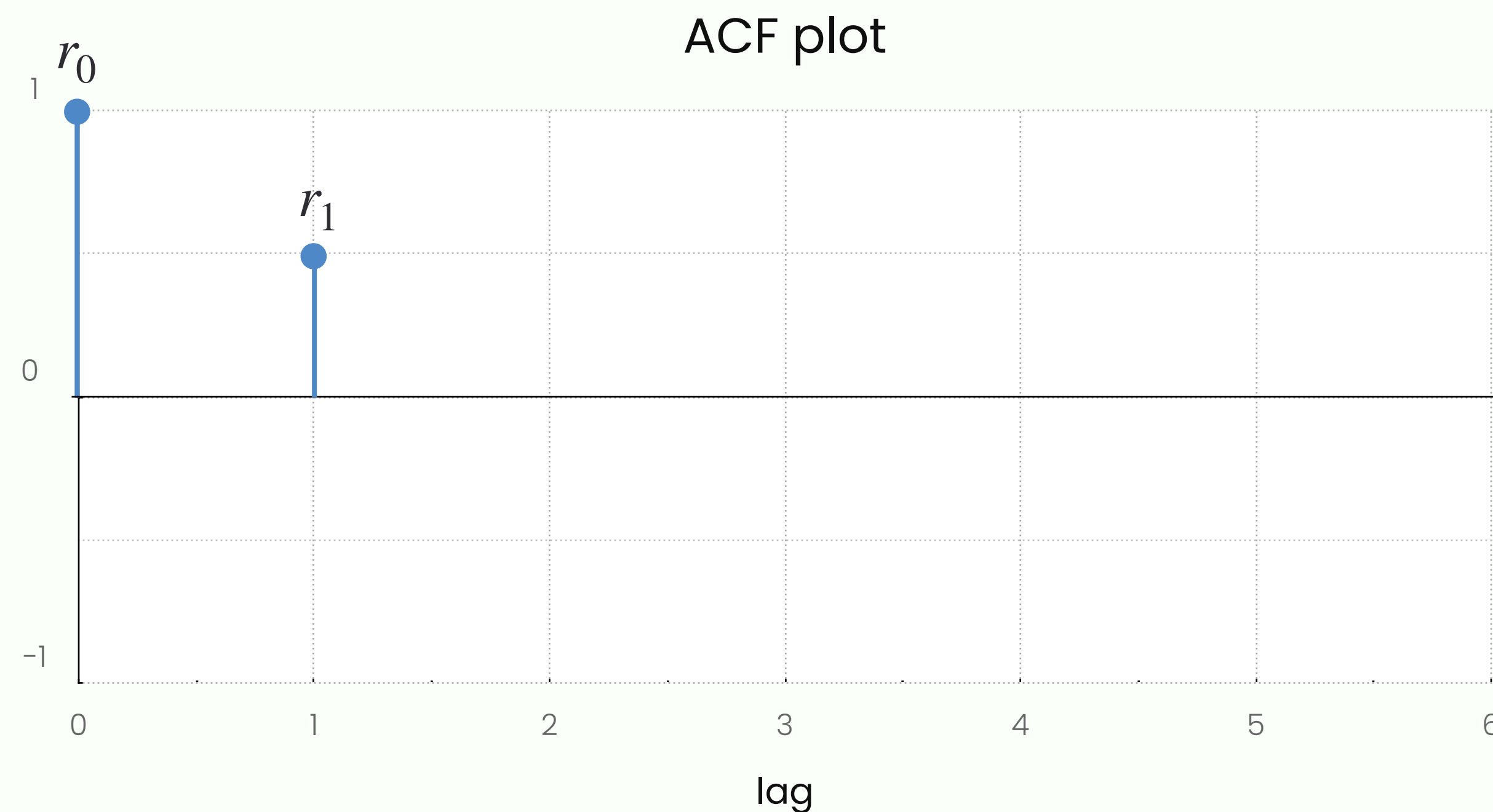


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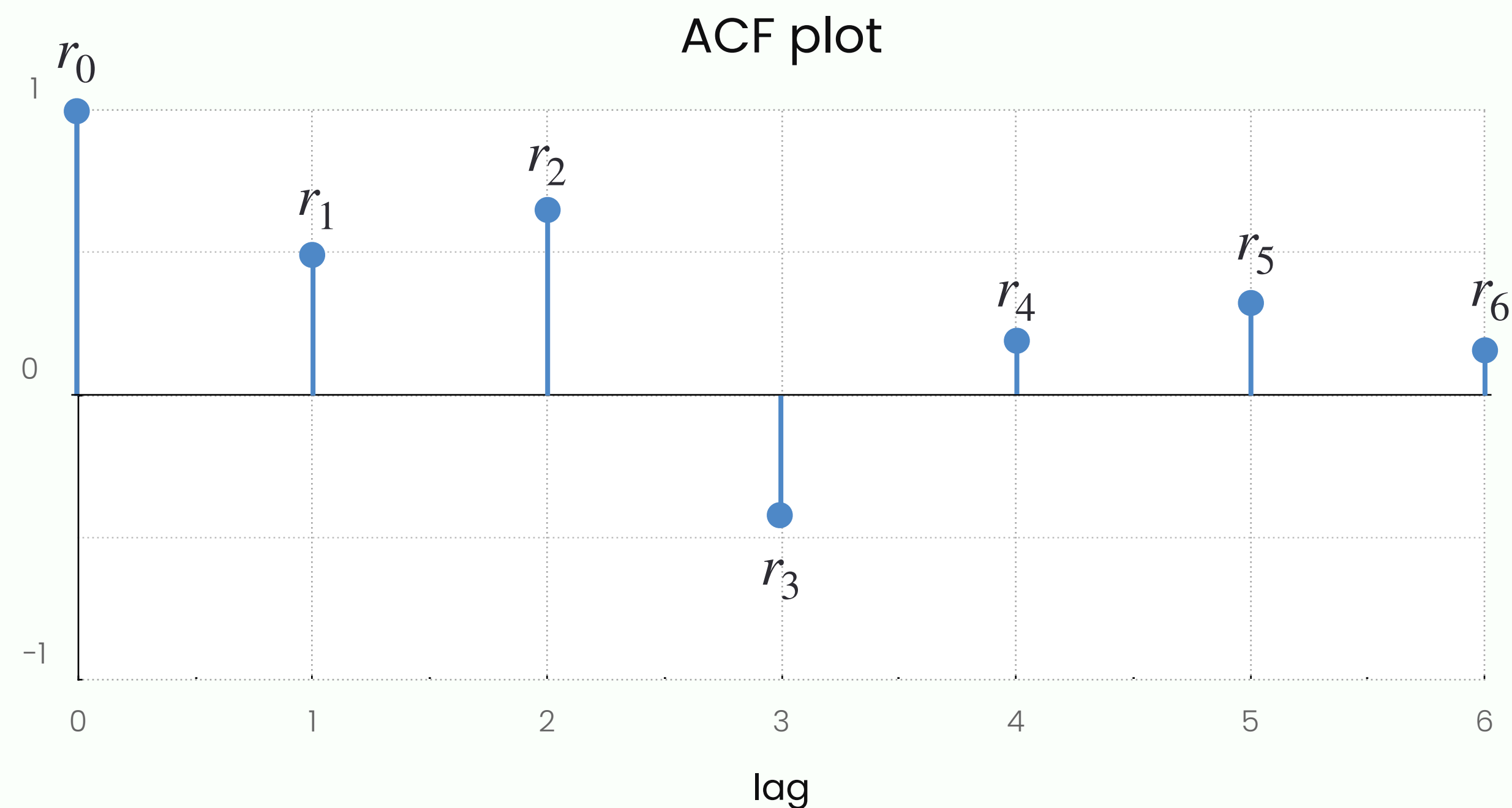


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Autocorrelation coefficient r_i measures the relationship between y_t and y_{t-i} .

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- r_1 between y_t and y_{t-1} ;
- r_2 between y_t and y_{t-2} ;
- ...

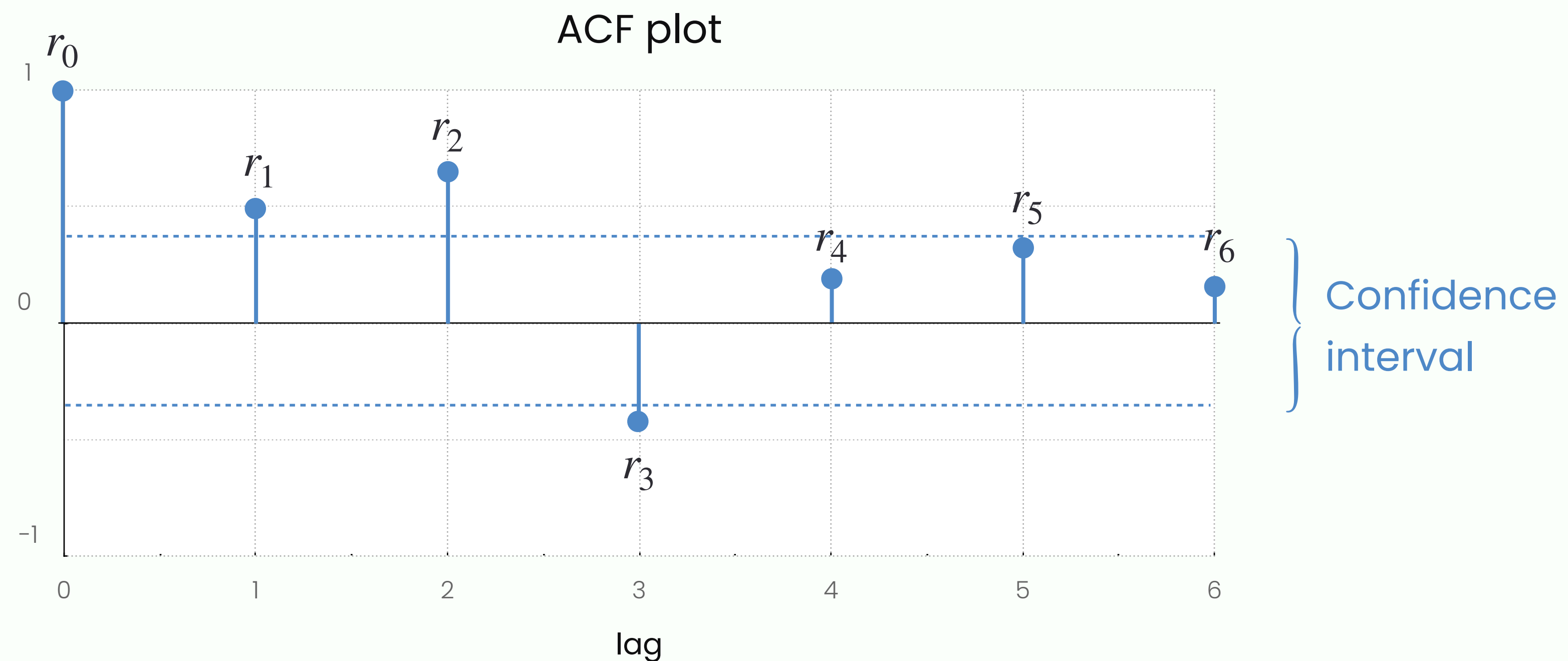


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- ...



INTERPRETATION

Positive ACF at lag i indicates a positive correlation between the current observation (y_t) and the observation at lag i (y_{t-i}).

Negative ACF at lag i indicates a negative correlation between the current observation (y_t) and the observation at lag i (y_{t-i}).

Decay in ACF as lag increases often signifies the presence of a trend or seasonality in the time series.

Significant ACF values at certain lags may suggest potential patterns or relationships in the time series.

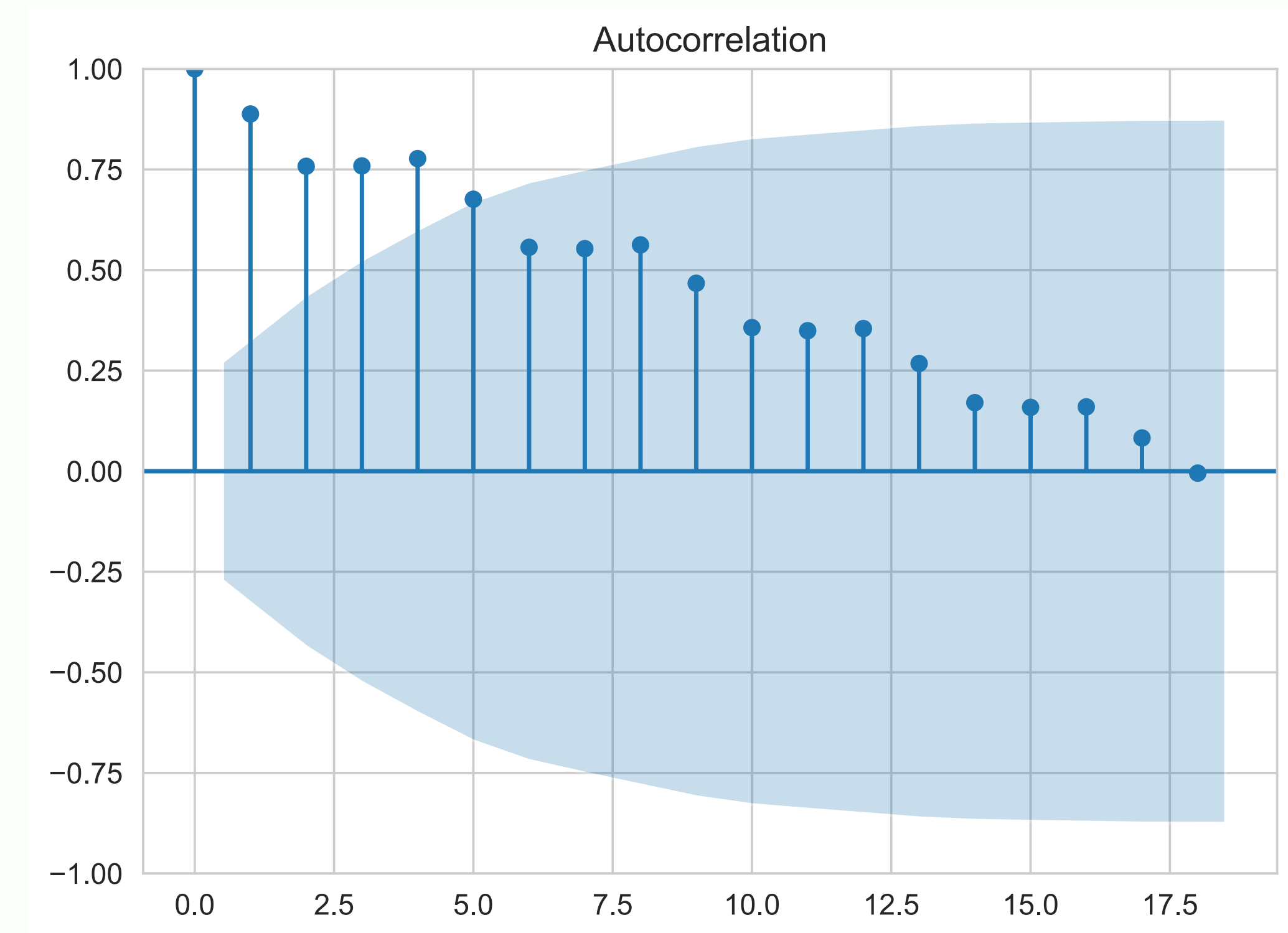
ACF: EXAMPLE

Recall original Australian quarterly electricity production data.

| lag (1Q) | 0 | 1 | 2 | 3 | 4 | 5 |
|---------------|---|--------|--------|--------|-------|-------|
| AC coef (r_i) | 1 | 0.8879 | 0.7579 | 0.7588 | 0.777 | 0.676 |

When data have a trend, the autocorrelations for small lags tend to be large and positive because observations nearby in time are also nearby in value.

That's why ACF for the non-stationary data does not make sense.



ACF: DATA TRANSFORMATION

Seasonal difference (m is the number of seasons):

$$y'_t = y_t - y_{t-m} = (1 - B^m)y_t$$

Difference:

$$y''_t = y'_t - y'_{t-1} = (1 - B)y'_t$$

ACF: DATA TRANSFORMATION

Seasonal difference (m is the number of seasons):

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Difference:

$$y''_t = y'_t - y'_{t-1} = (1 - B)y_t$$

The final transformation (for $m = 4$) is then given as:

$$\begin{aligned}(1 - B)(1 - B^4) y_t &= (1 - B - B^4 + B^5) y_t \\ &= y_t - y_{t-1} - y_{t-4} + y_{t-5}\end{aligned}$$

ACF: DATA TRANSFORMATION (2)

When both seasonal and first differences are applied:

- Order does not matter—result will be the same.
- Try seasonal first as there might no need for further first difference.

ACF: DATA TRANSFORMATION (2)

When both seasonal and first differences are applied:

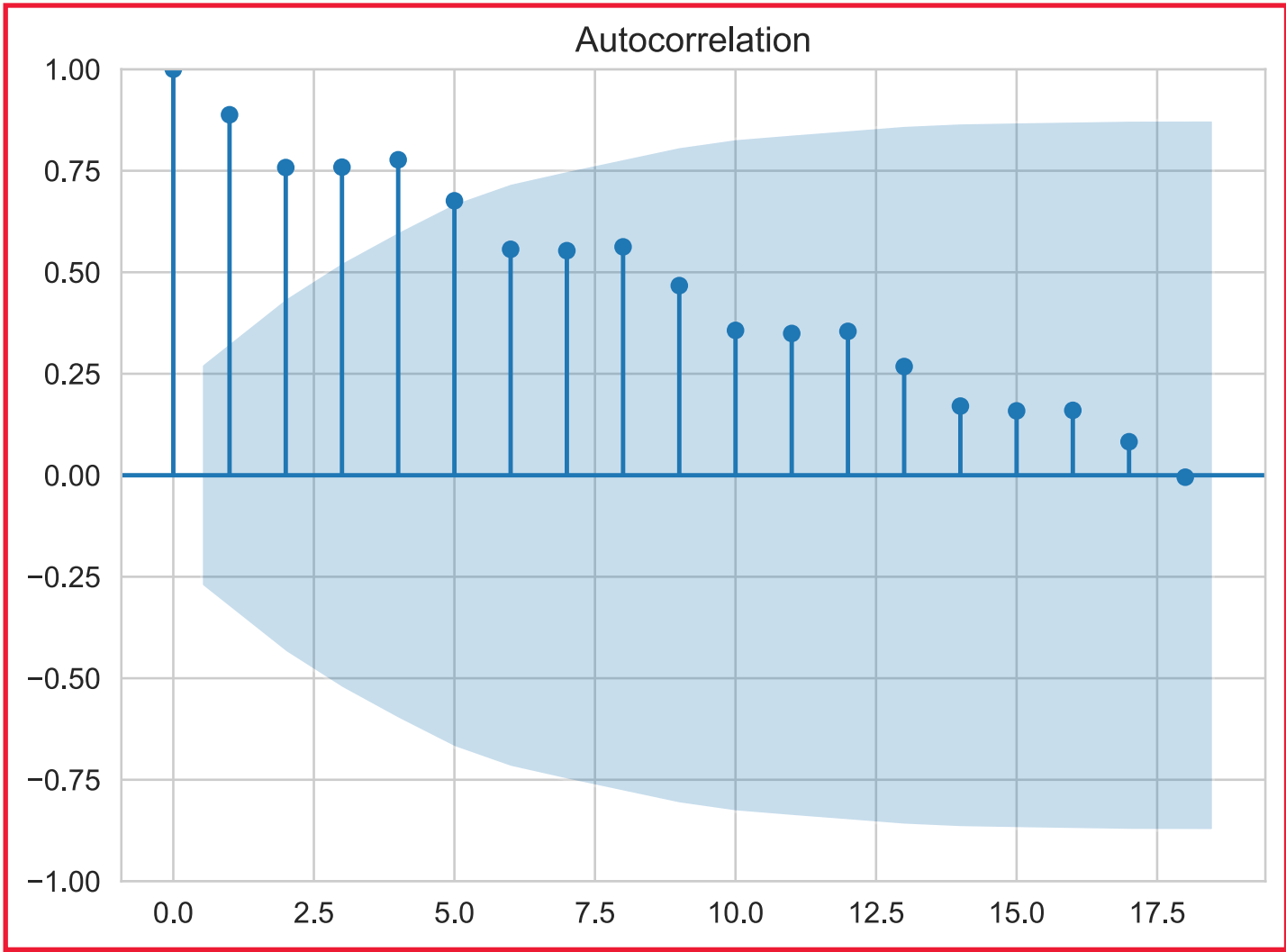
- Order does not matter—result will be the same.
- Try seasonal first as there might no need for further first difference.

NB! Differences must be interpretable.

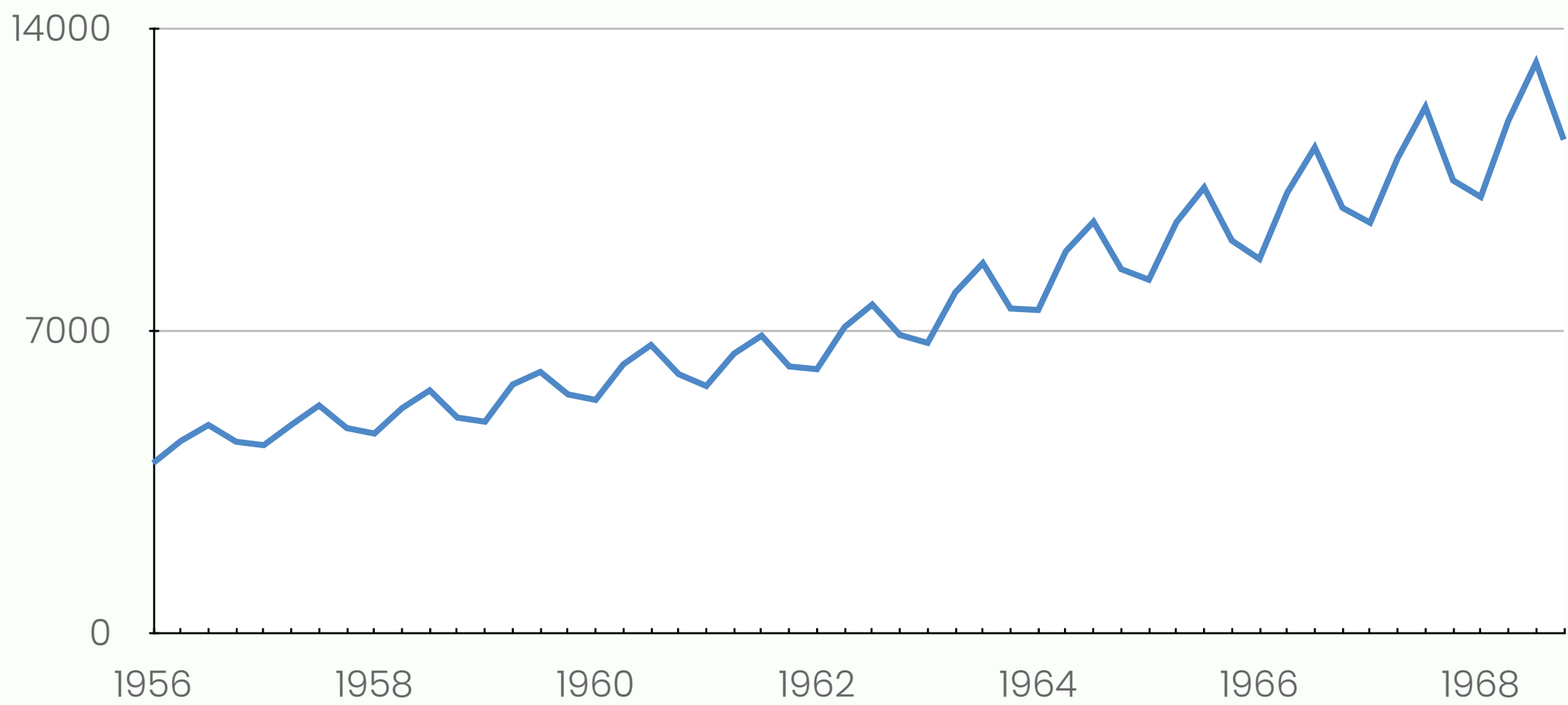
- ✓ First differences are the change between close observations.
- ✓ Seasonal differences are the change between corresponding observations.
- ✓ But, for example, $(1 - B^3)$ difference (lag3) for quarterly data is not interpretable.

ACF: EXAMPLE

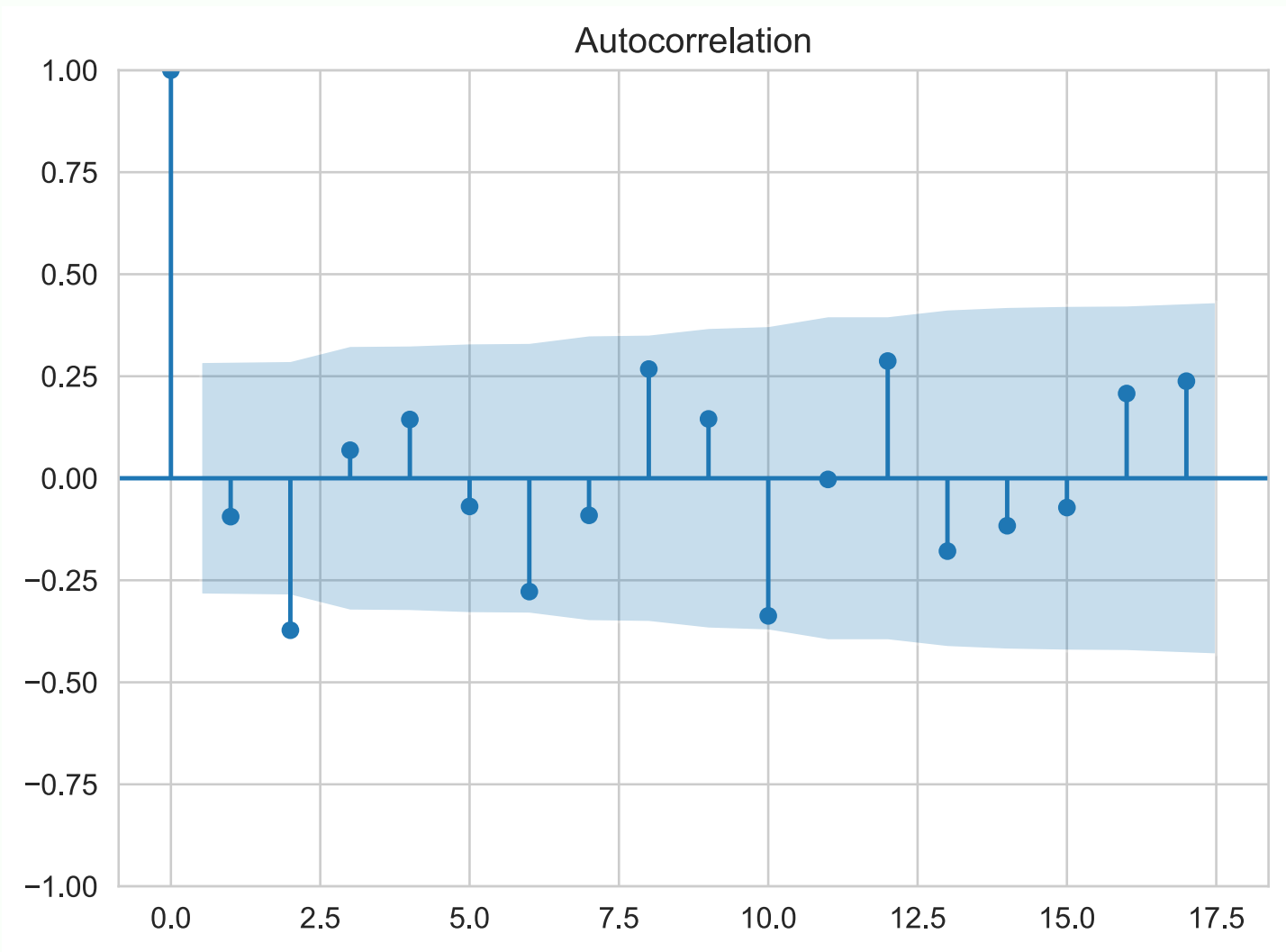
ORIGINAL



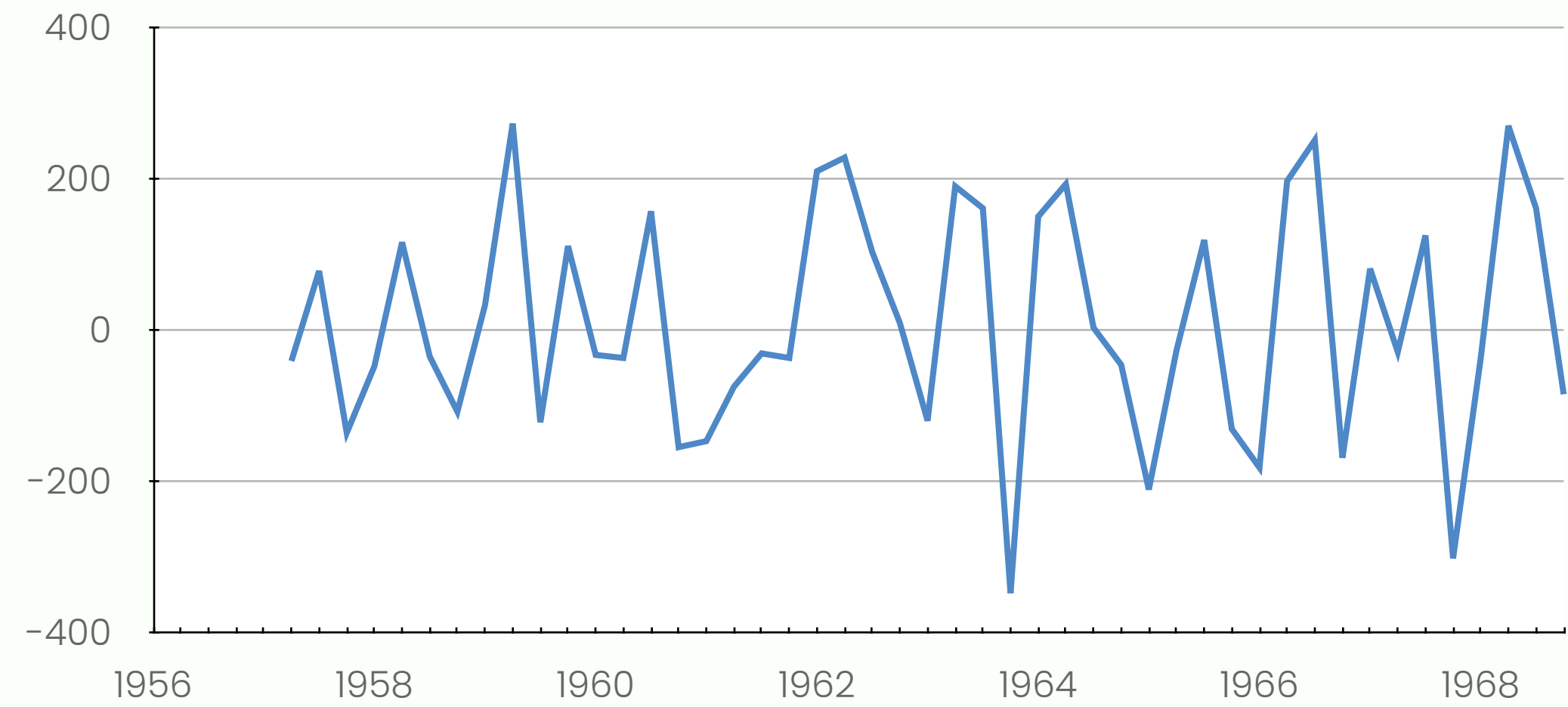
Production [GWh]



TRANSFORMED
SEAS + 1st DIFF



[GWh]



PARTIAL ACF (PACF)

Partial ACF measures the correlation between a variable and its lagged values, excluding the linear influence of the intermediate lags.

PACF: INTUITION

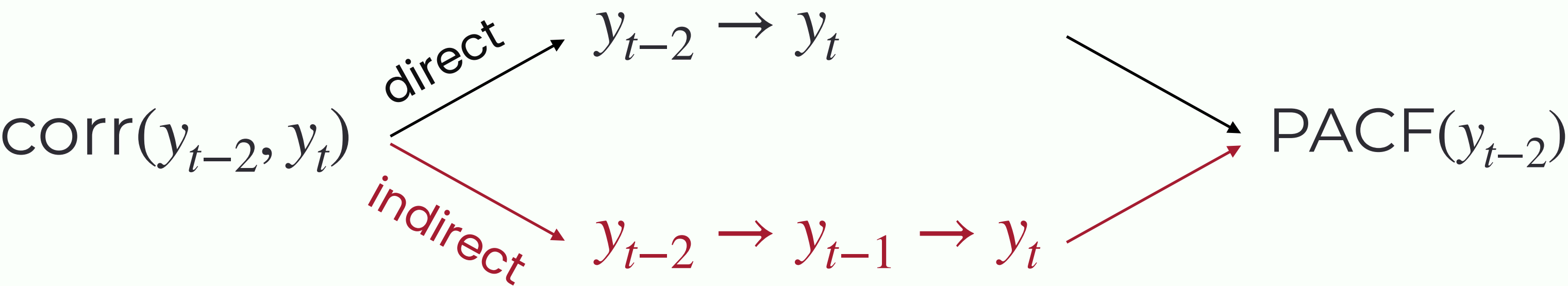
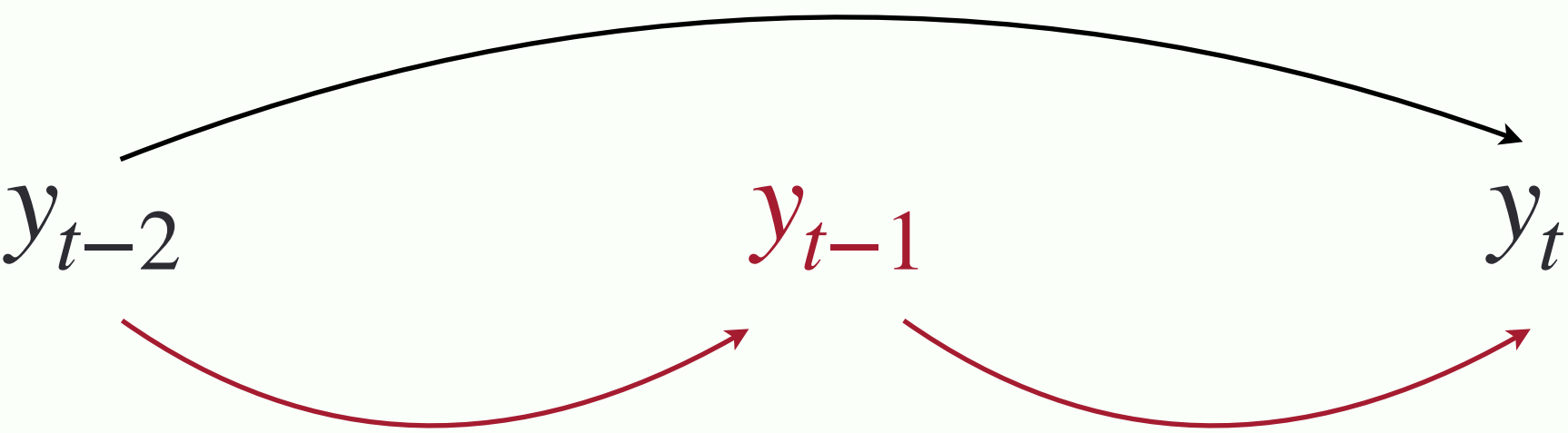
Forecast average monthly electricity price.



PACF: INTUITION

Forecast average monthly electricity price.

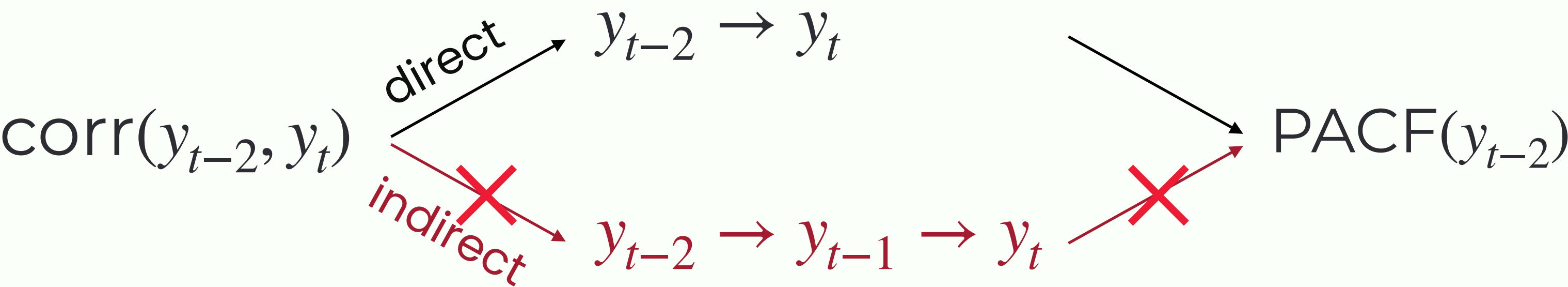
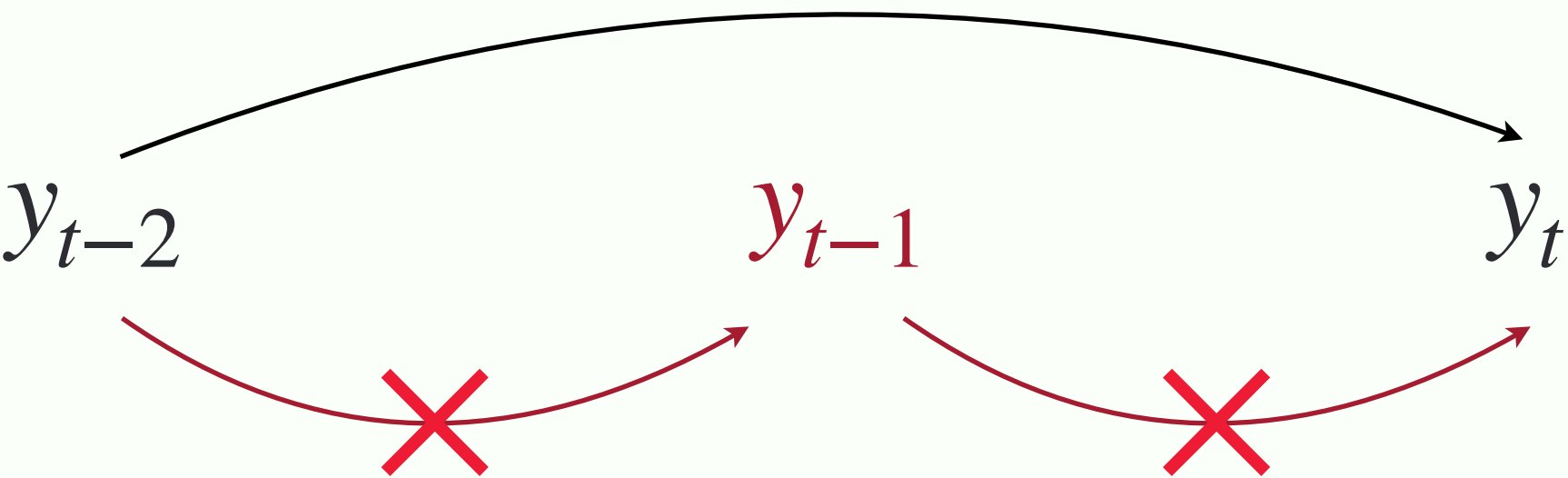
- y_t December
- y_{t-1} November
- y_{t-2} October
- ⋮



PACF: INTUITION

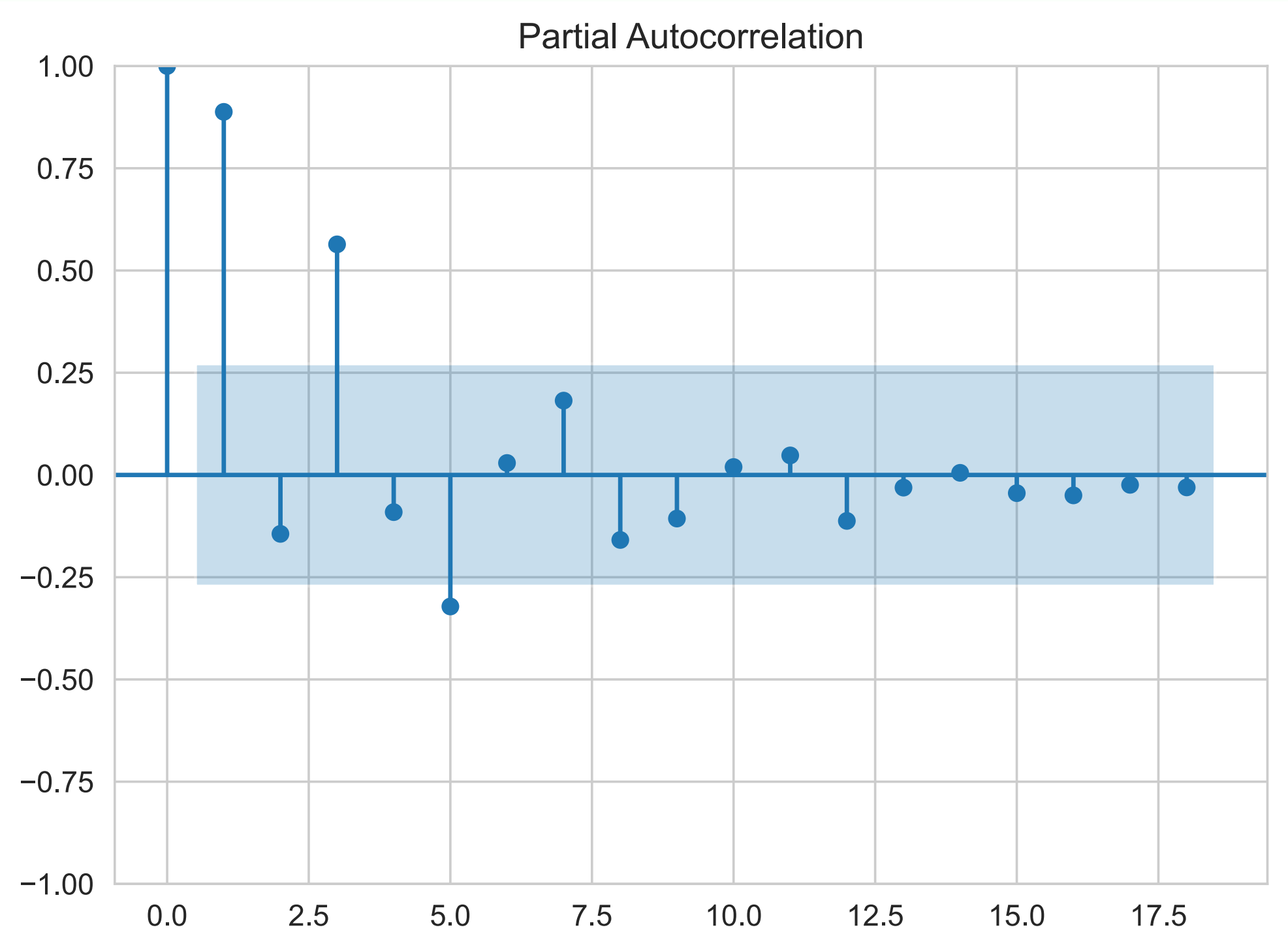
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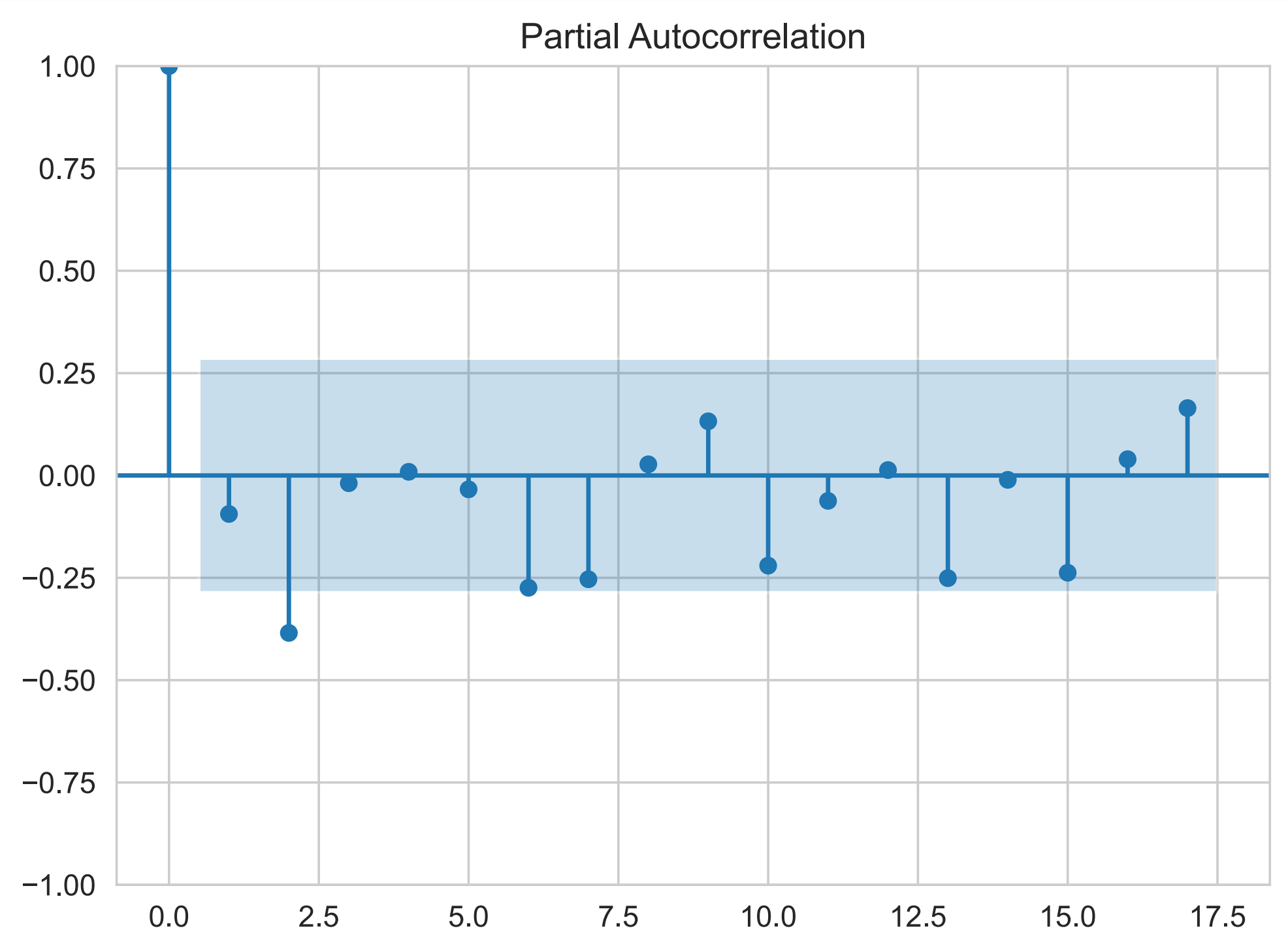
EXAMPLE

ORIGINAL



TRANSFORMED

SEAS + 1st DIFF

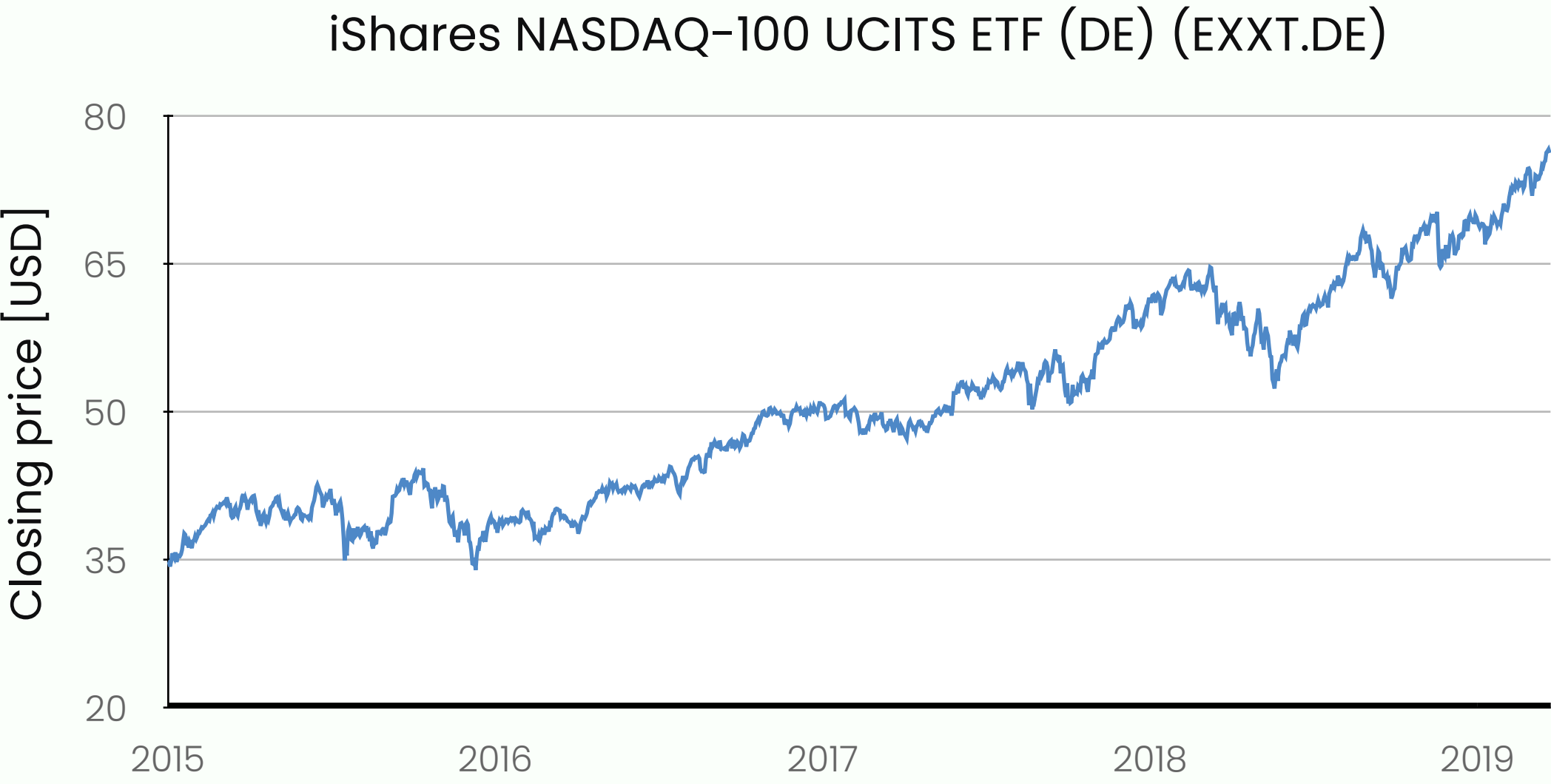


ACF VS PACF

Autocorrelation (ACF) and Partial Autocorrelation (PACF) are both measures used in time series analysis to understand the relationships between observations at different time points.

| ACF | PACF |
|---|--|
| Shows order of a moving average (MA) process | Shows order of an autoregressive (AR) process |
| Represents the overall correlation structure of the time series | Highlights the direct relationships between observations at specific lags |
| Measures the linear relationship between an observation and its previous observations at different lags | Measures the direct linear relationship between an observation and its previous observations at a specific lag, excluding the contributions from intermediate lags |

ACF & PACF AT STOCK MARKET

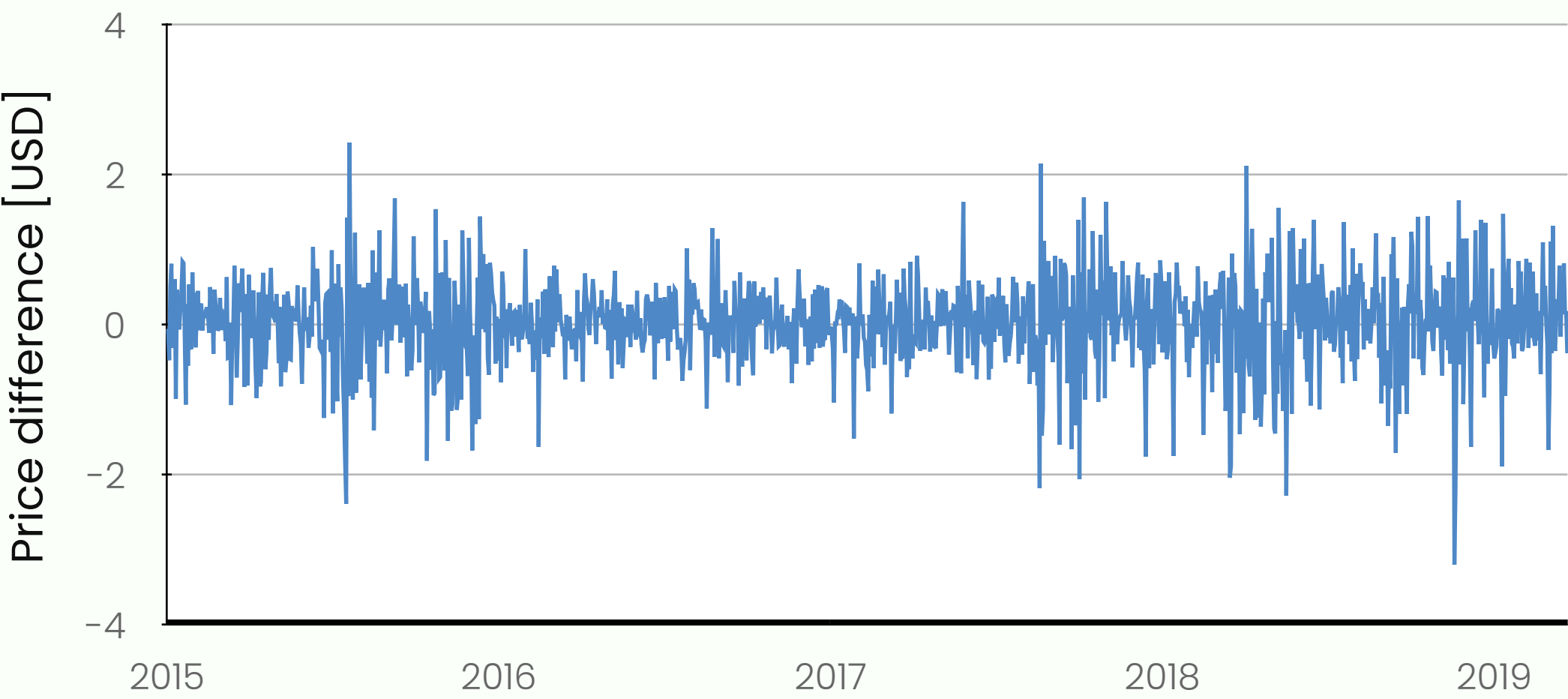


ACF & PACF AT STOCK MARKET

iShares NASDAQ-100 UCITS ETF (DE) (EXXT.DE)

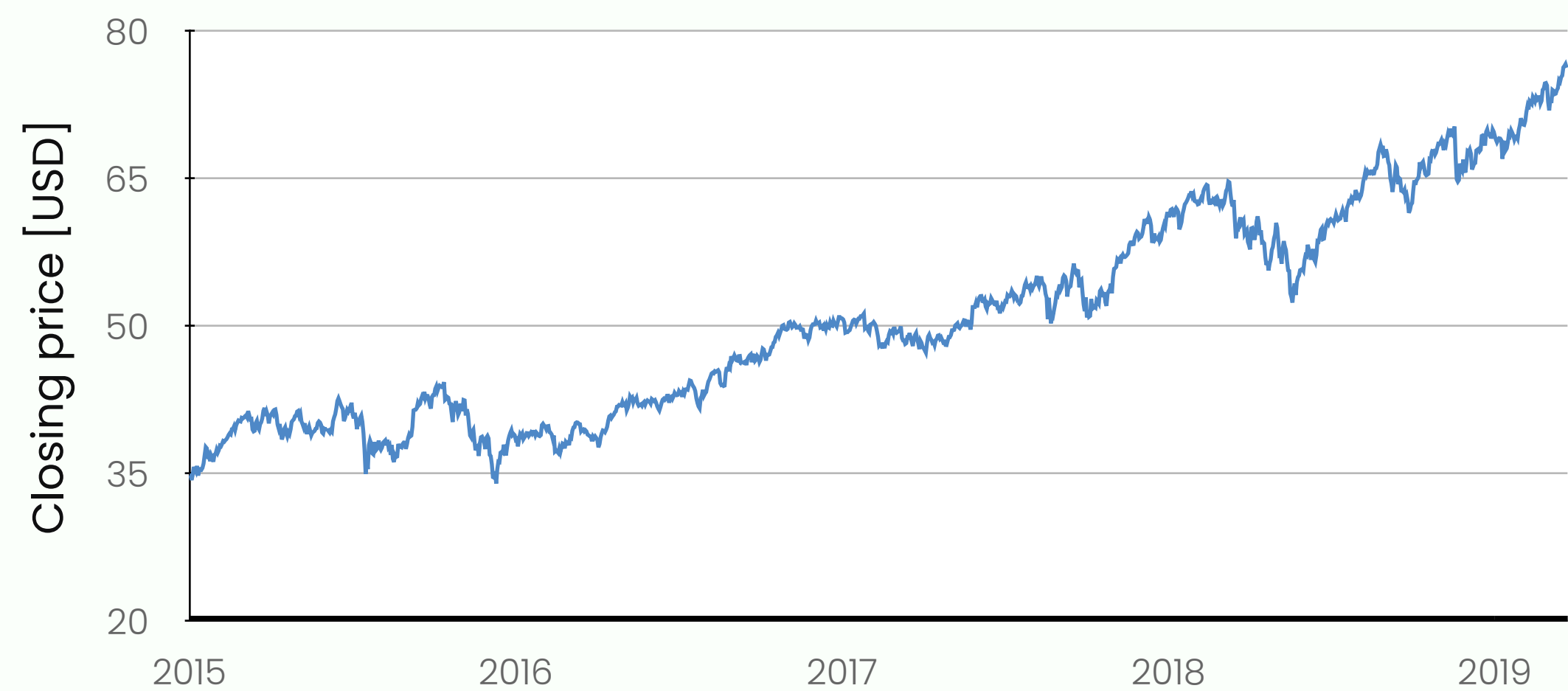


$(1 - B)y_t$

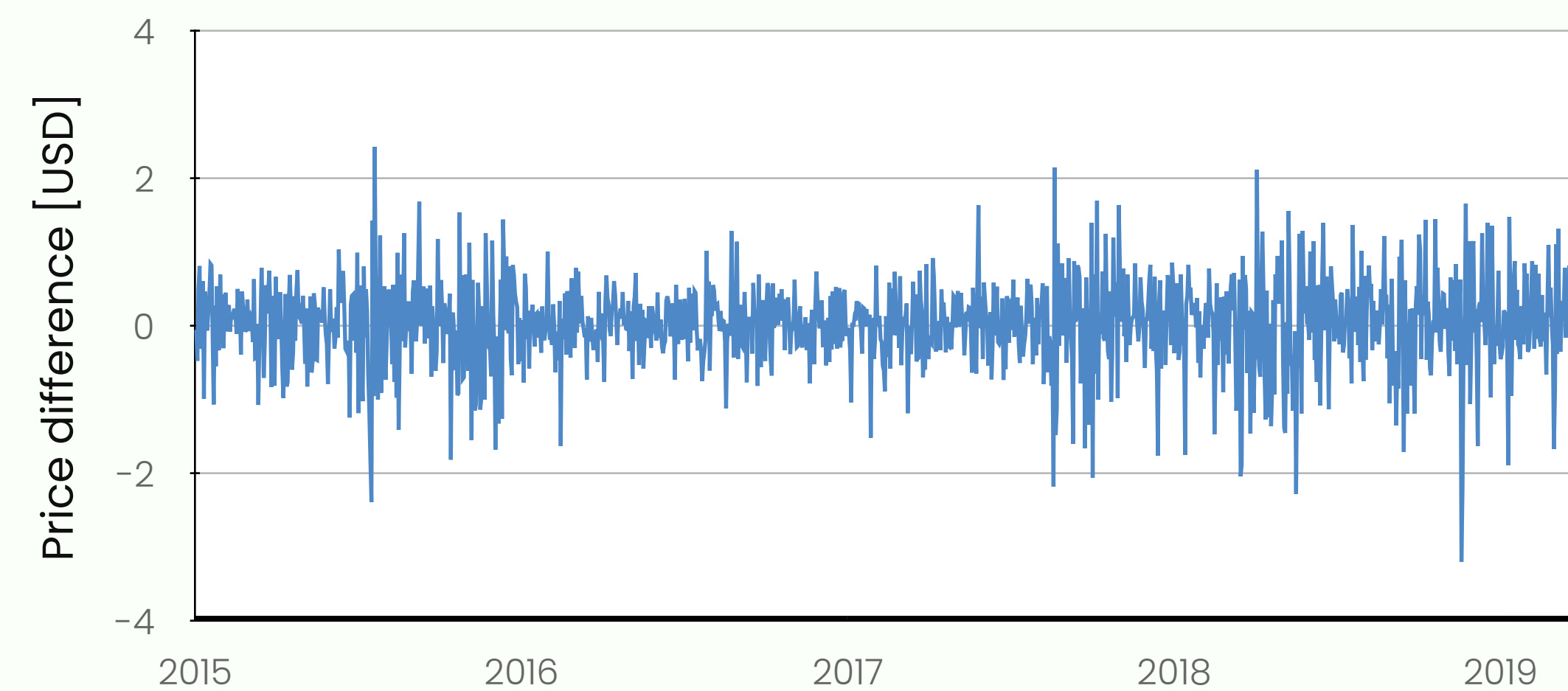


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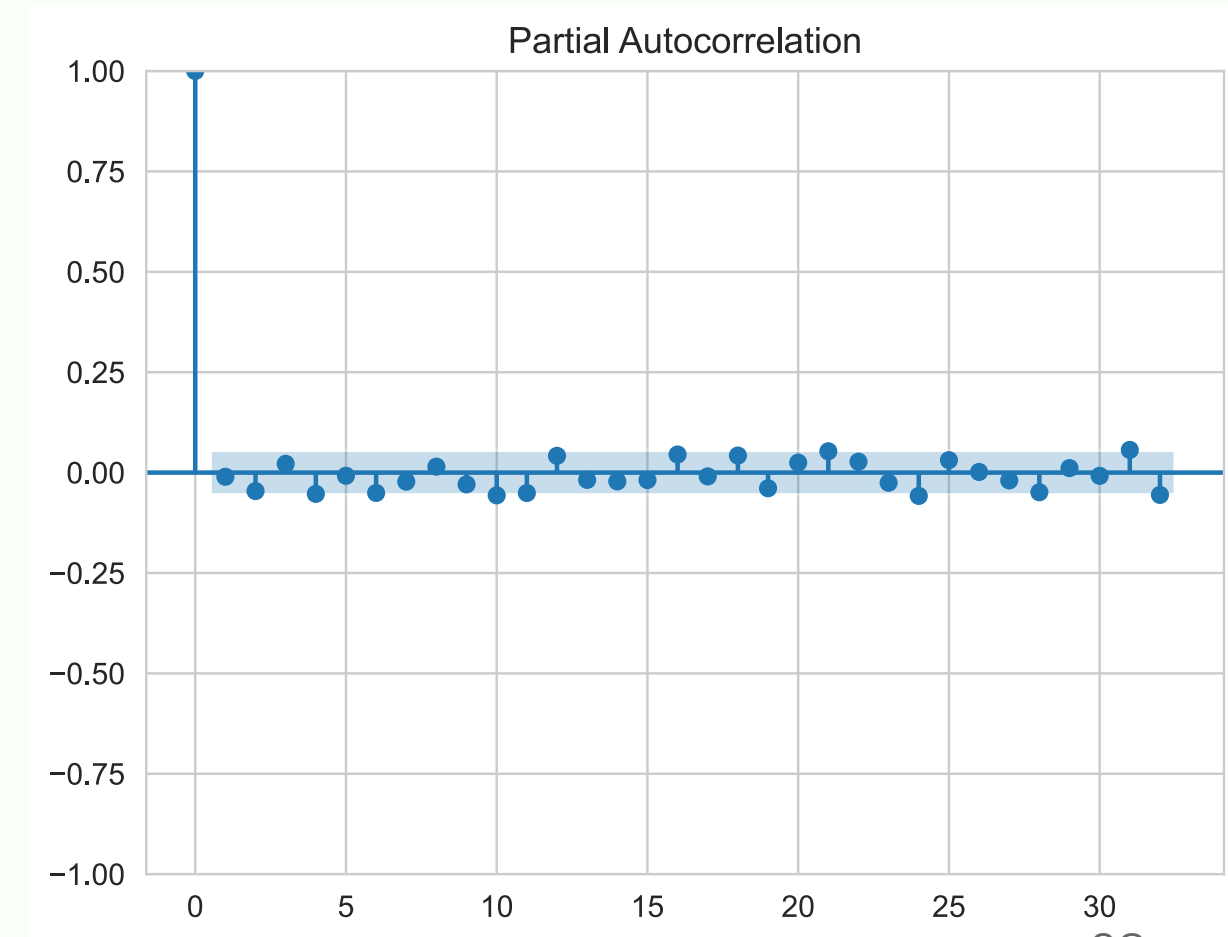
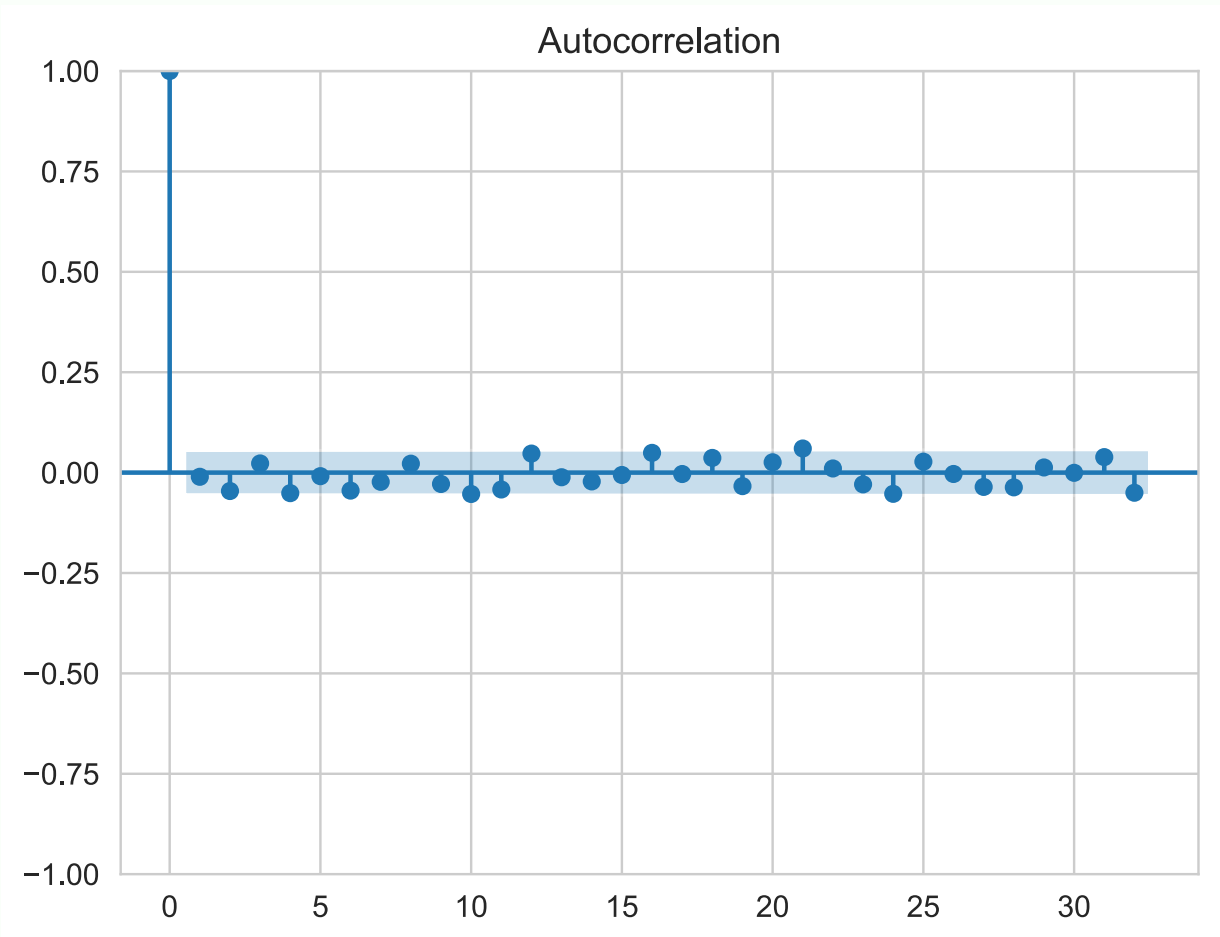
iShares NASDAQ-100 UCITS ETF (DE) (EXXT.DE)



$(1 - B)y_t$



No significant lags in both ACF and PACF



MODELLING METHODS

TRADITIONAL (STATISTICAL)

- Simple methods
- Regression methods
- ARMA family methods
- Exponential smoothing
- Kalman filter

MACHINE LEARNING

- Artificial neural networks
- Support vector machines
- Fuzzy logic
- Decision tree
- Genetic algorithm
- Knowledge-based expert systems
- LSTM

Autoregressive Models

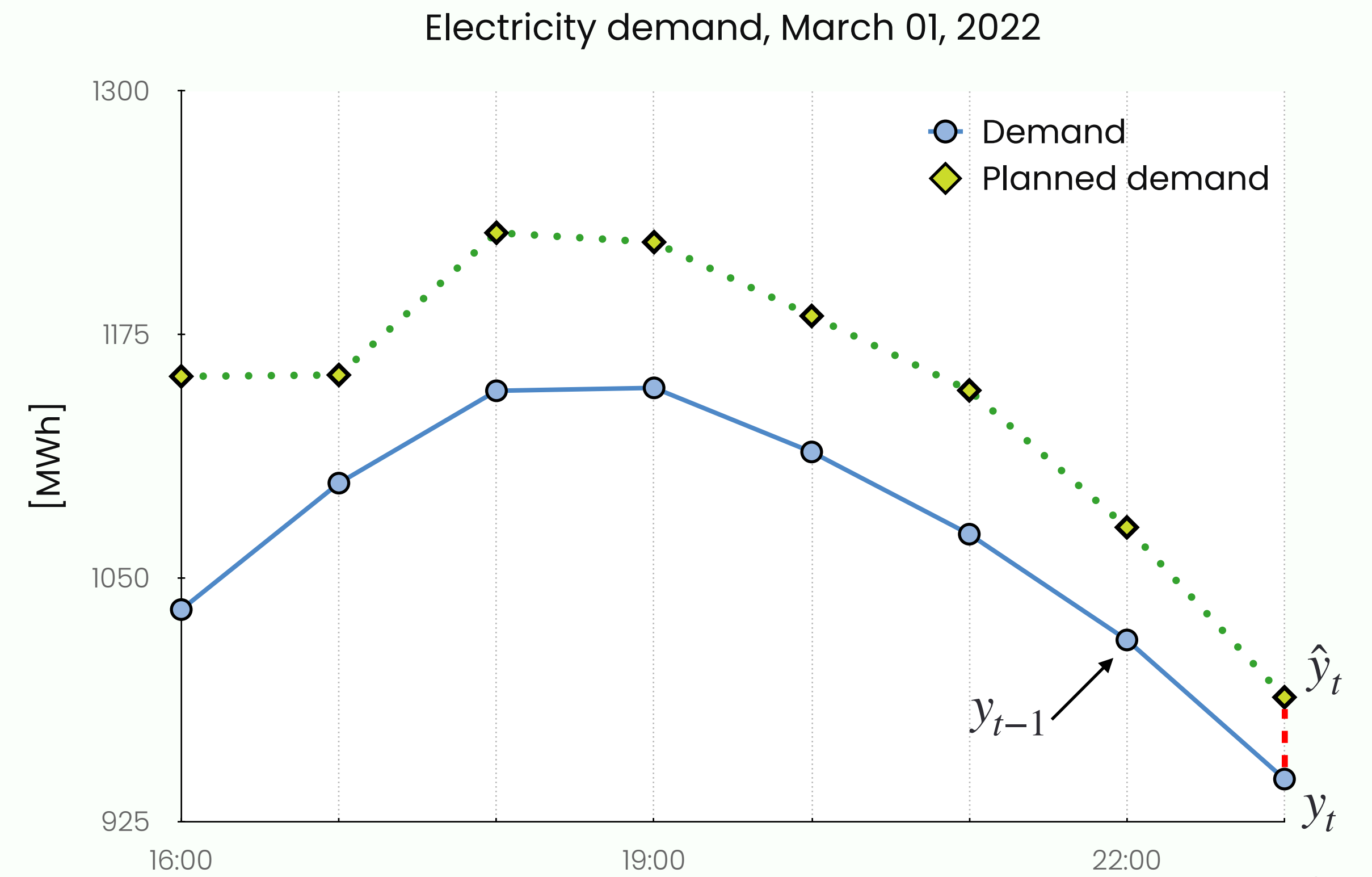
AUTOREGRESSION

The term *autoregression* indicates that it is a regression of the variable against itself (*past* values of the *same* variable).

AR(1) MODEL

For example, autoregressive model of order 1 (or AR(1)) can be written as

$$y_t = c + \phi y_{t-1} + \epsilon_t.$$



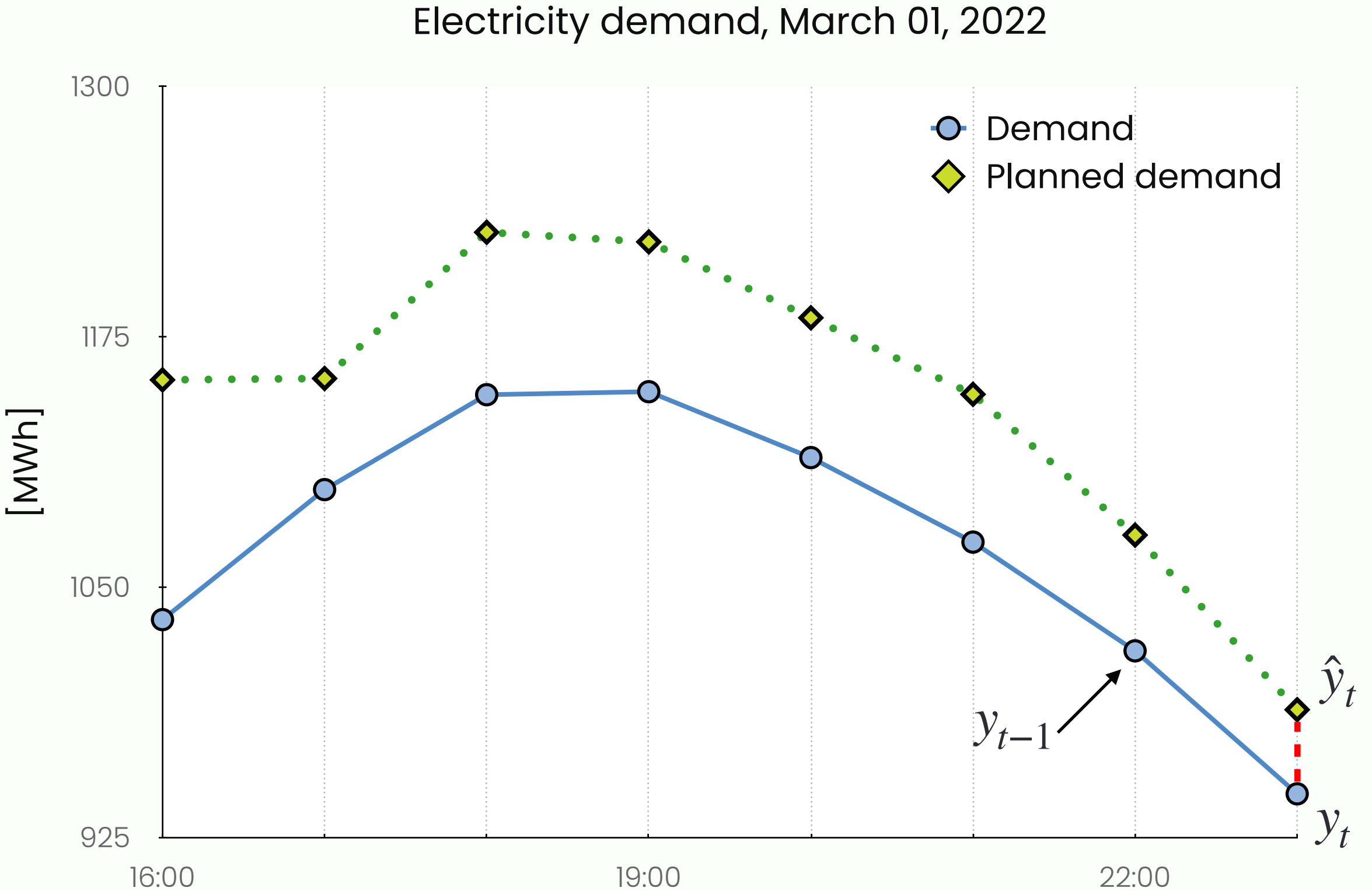
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target

↑



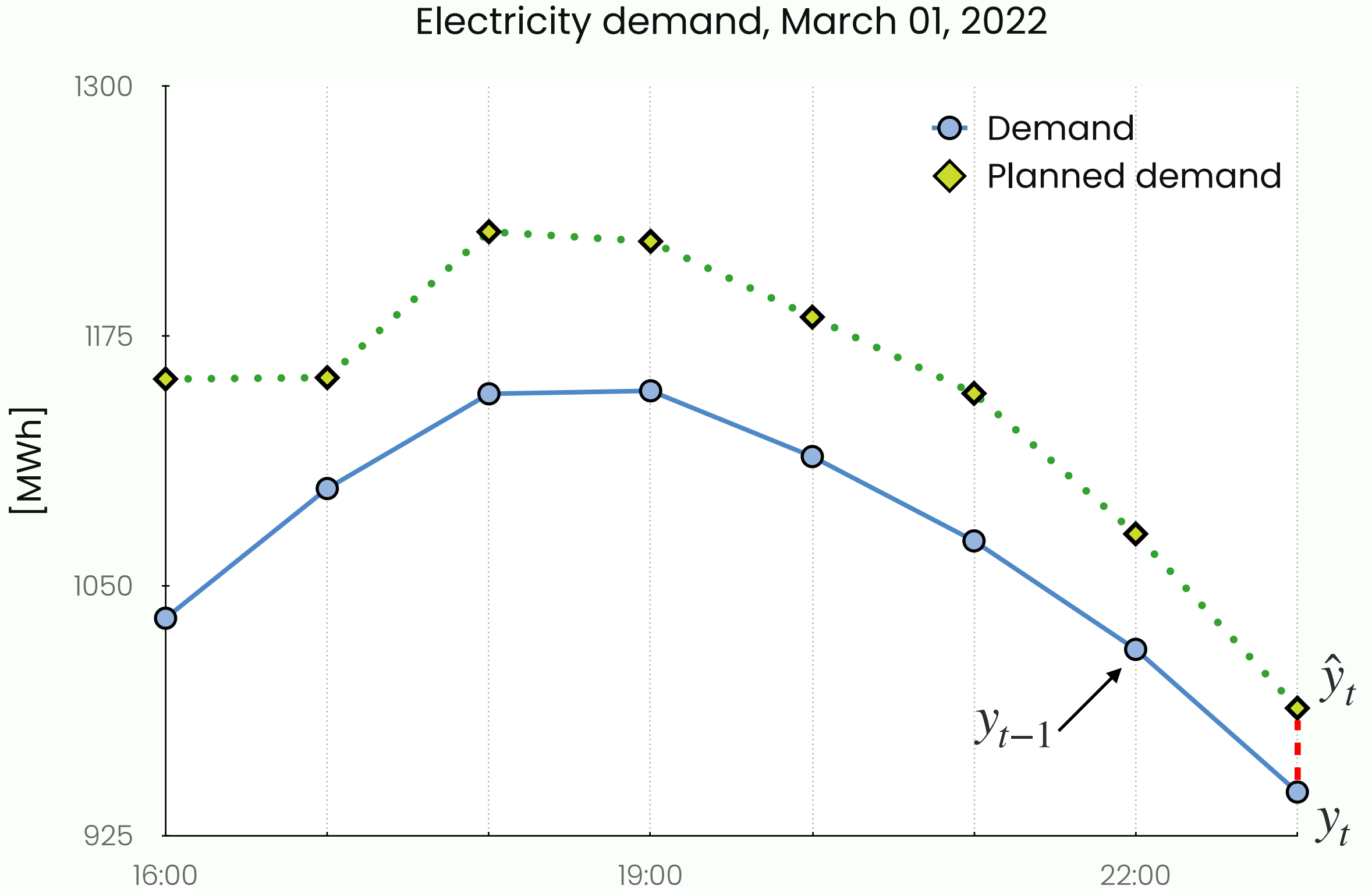
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intercept coefficient

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target



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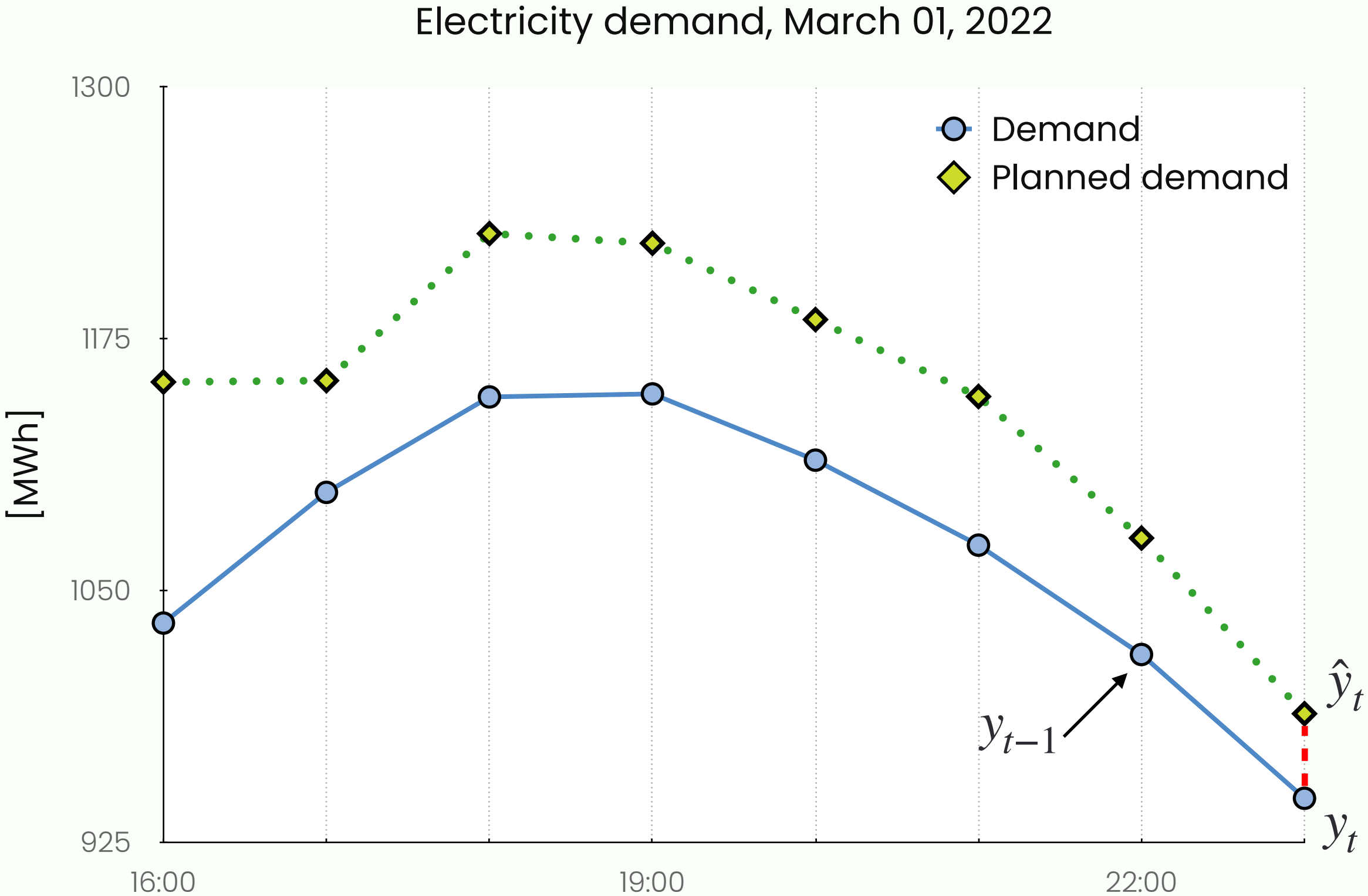
intercept coefficient error

$$y_t = c + \phi y_{t-1} + \epsilon_t$$

target

Practical observation (there are no perfect models):

$$\hat{y}_t = c + \phi y_{t-1} + \phi_2 y_{t-2}$$



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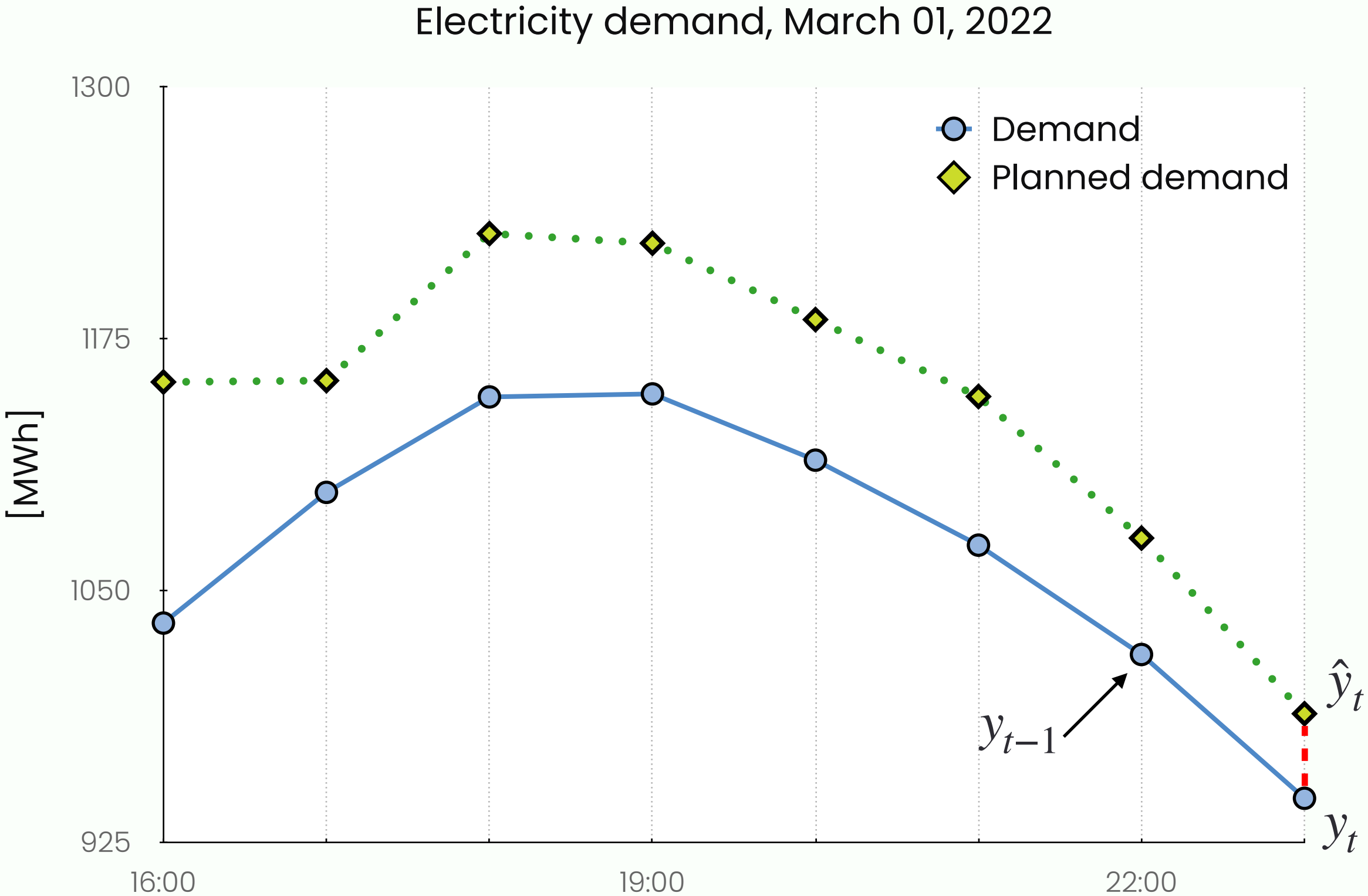
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target lagged target

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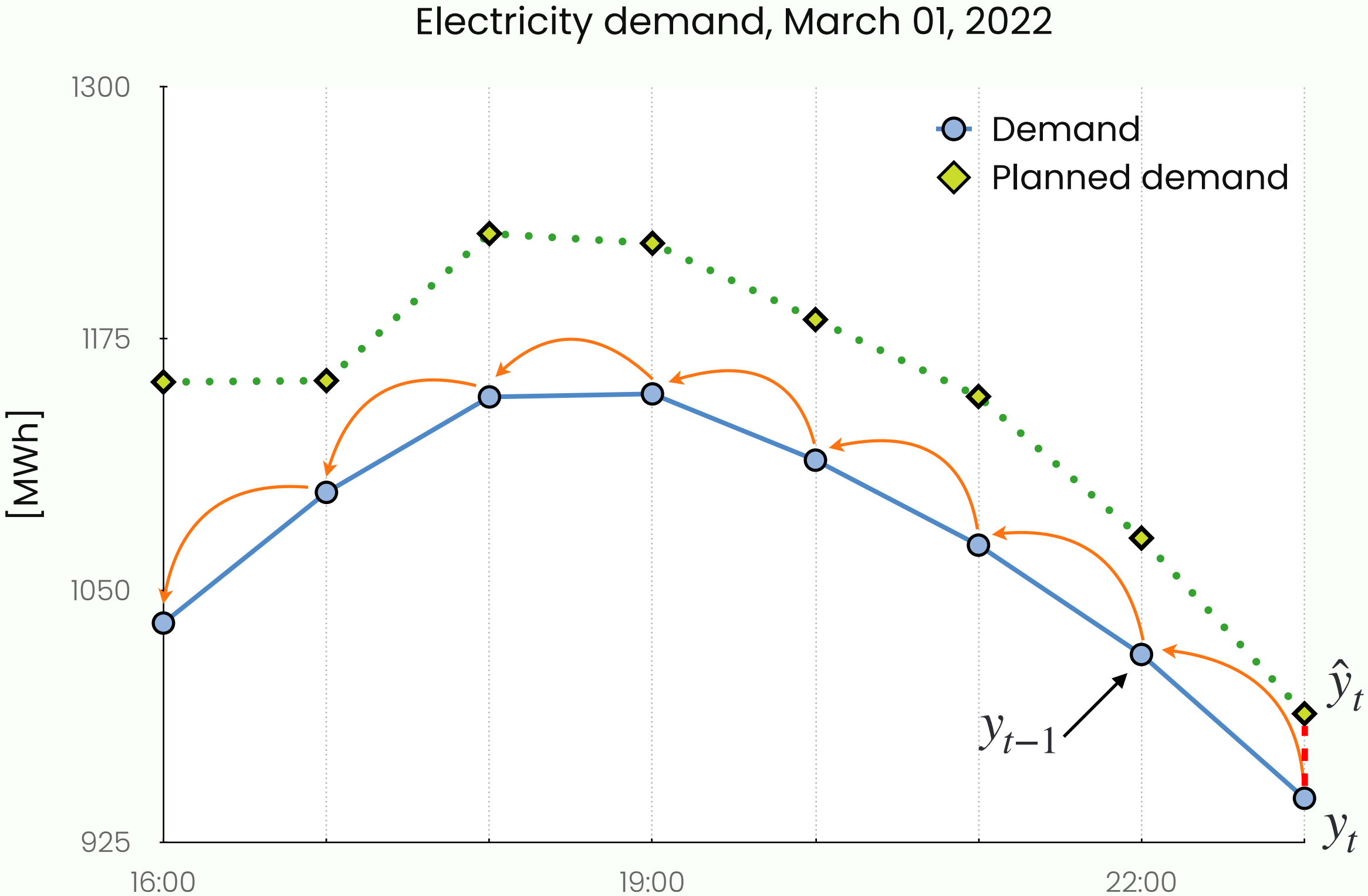
intercept coefficient error

$$y_t = c + \phi y_{t-1} + \epsilon_t$$

target lagged target

Practical observation (there are no perfect models):

$$\hat{y}_t = c + \phi y_{t-1} + \phi_2 y_{t-2}$$



LONG MEMORY MODELS

Recursion goes back until the beginning of the series.

$$y_t = c + \phi y_{t-1} + \epsilon_t$$

$$y_{t-1} = c + \phi y_{t-2} + \epsilon_{t-1}$$

$$y_{t-2} = c + \phi y_{t-3} + \epsilon_{t-2}$$

LONG MEMORY MODELS

Recursion goes back until the beginning of the series.

$$\begin{array}{l} \boxed{y_t = c + \phi y_{t-1} + \epsilon_t} \\ y_{t-1} = c + \phi y_{t-2} + \epsilon_{t-1} \\ y_{t-2} = c + \phi y_{t-3} + \epsilon_{t-2} \end{array} \quad \begin{array}{l} \nearrow \\ \nearrow \end{array} \quad y_t = c + \phi(c + \phi y_{t-2} + \epsilon_{t-1}) + \epsilon_t$$

LONG MEMORY MODELS

Recursion goes back until the beginning of the series.

$$\begin{array}{lcl} \boxed{y_t = c + \phi y_{t-1} + \epsilon_t} & \nearrow & y_t = c + \phi(c + \phi y_{t-2} + \epsilon_{t-1}) + \epsilon_t \\ y_{t-1} = c + \phi y_{t-2} + \epsilon_{t-1} & \nearrow & \downarrow \text{simplify} \\ y_{t-2} = c + \phi y_{t-3} + \epsilon_{t-2} & & y_t = c^* + \phi^2 y_{t-2} + \phi \epsilon_{t-1} + \epsilon_t \end{array}$$

LONG MEMORY MODELS

Recursion goes back until the beginning of the series.

The diagram illustrates the recursive substitution process for the equation $y_t = c + \phi y_{t-1} + \epsilon_t$. It shows how the term y_{t-1} is replaced by its own equation, and then y_{t-2} is replaced, leading to a simplified form with a constant c^* .

$$\begin{array}{lcl} \boxed{y_t = c + \phi y_{t-1} + \epsilon_t} & \xrightarrow{\quad} & y_t = c + \phi(c + \phi y_{t-2} + \epsilon_{t-1}) + \epsilon_t \\ y_{t-1} = c + \phi y_{t-2} + \epsilon_{t-1} & \xrightarrow{\quad} & \downarrow \text{simplify} \\ y_{t-2} = c + \phi y_{t-3} + \epsilon_{t-2} & \xrightarrow{\quad} & y_t = c^* + \phi^2 y_{t-2} + \phi \epsilon_{t-1} + \epsilon_t \\ & & \downarrow \\ & & y_t = c^* + \phi^3 y_{t-3} + \phi^2 \epsilon_{t-2} + \phi \epsilon_{t-1} + \epsilon_t \end{array}$$

LONG MEMORY MODELS

Recursion goes back until the beginning of the series.

$$\boxed{y_t = c + \phi y_{t-1} + \epsilon_t}$$
$$y_{t-1} = c + \phi y_{t-2} + \epsilon_{t-1}$$
$$y_{t-2} = c + \phi y_{t-3} + \epsilon_{t-2}$$
$$y_t = c + \phi(c + \phi y_{t-2} + \epsilon_{t-1}) + \epsilon_t$$

simplify

$$y_t = c^* + \phi^2 y_{t-2} + \phi \epsilon_{t-1} + \epsilon_t$$
$$y_t = c^* + \phi^3 y_{t-3} + \phi^2 \epsilon_{t-2} + \phi \epsilon_{t-1} + \epsilon_t$$

⌚

$$y_t = \frac{c}{1 - \phi} + \phi^t y_1 + \phi^{t-1} \epsilon_2 + \phi^{t-2} \epsilon_3 + \cdots + \epsilon_t$$

LONG MEMORY MODELS

Recursion goes back until the beginning of the series.

$$\begin{array}{lcl}
 \boxed{y_t = c + \phi y_{t-1} + \epsilon_t} & \nearrow & y_t = c + \phi(c + \phi y_{t-2} + \epsilon_{t-1}) + \epsilon_t \\
 y_{t-1} = c + \phi y_{t-2} + \epsilon_{t-1} & \nearrow & \downarrow \text{simplify} \\
 y_{t-2} = c + \phi y_{t-3} + \epsilon_{t-2} & \longrightarrow & y_t = c^* + \phi^2 y_{t-2} + \phi \epsilon_{t-1} + \epsilon_t \\
 & & \downarrow \\
 & & y_t = c^* + \phi^3 y_{t-3} + \phi^2 \epsilon_{t-2} + \phi \epsilon_{t-1} + \epsilon_t \\
 & & \downarrow \text{⌚} \\
 & & y_t = \frac{c}{1 - \phi} + \phi^t y_1 + \phi^{t-1} \epsilon_2 + \phi^{t-2} \epsilon_3 + \dots + \epsilon_t
 \end{array}$$

✓ Effects from past have little effect on the present if $|\phi| < 1$.

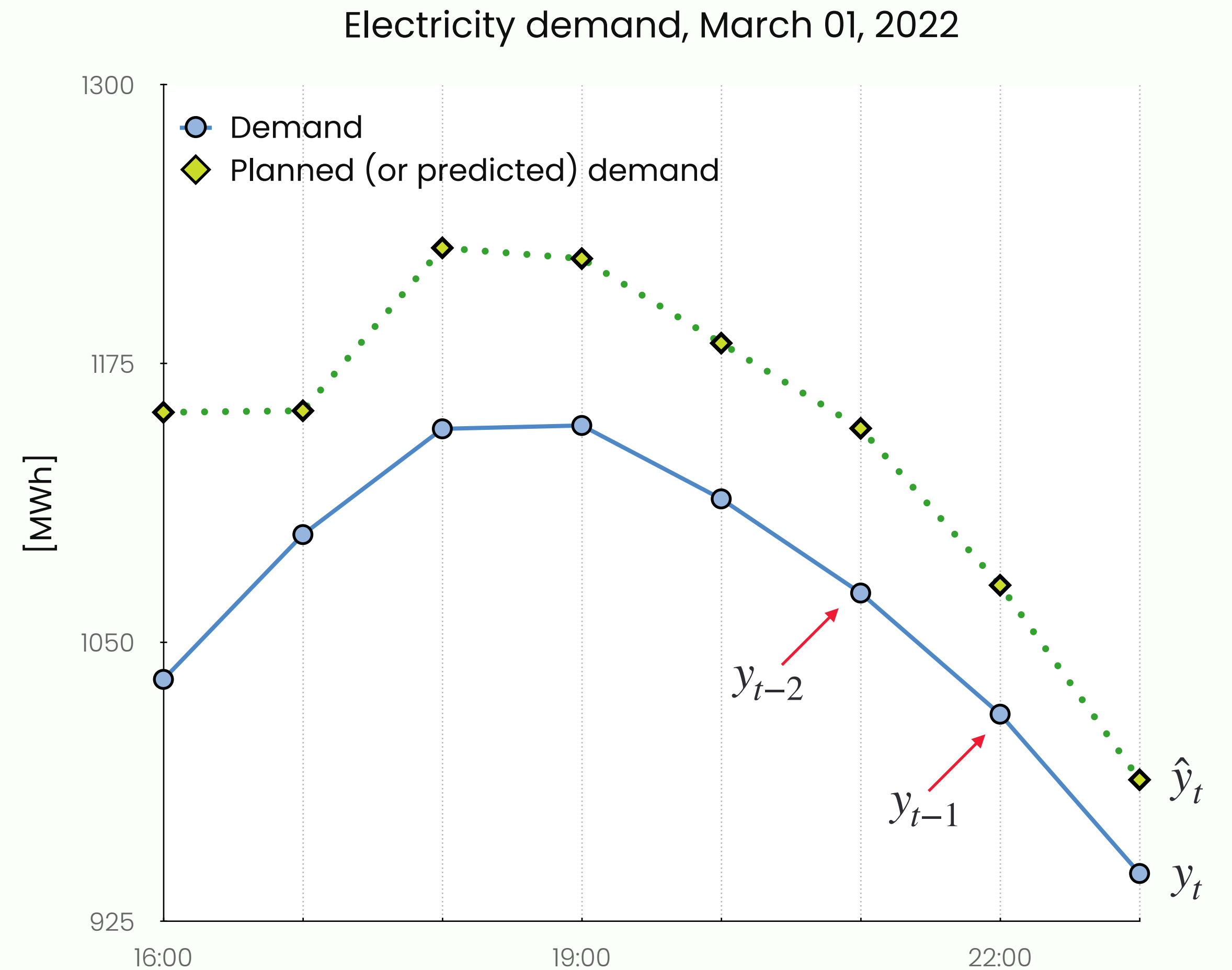
✓ Stationarity – the dependence of previous observations declines over time.

AR(2) MODEL

A linear function of 2 past values or AR(2) can be written as

$$y_t = c + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \epsilon_t$$

lagged targets



AUTOREGRESSIVE (AR) MODEL

An autoregressive model of order p can be written as

$$y_t = c + \sum_{i=1}^p \phi_i y_{t-i} + \epsilon_t = c + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \dots + \phi_p y_{t-p} + \epsilon_t$$

where $\epsilon_t \sim wn(0, \sigma_w^2)$ is white noise, ϕ_1, \dots, ϕ_p are parameters, c is intercept.

We refer to this as an $AR(p)$ model, an autoregressive model of order p .

MA(2) MODEL

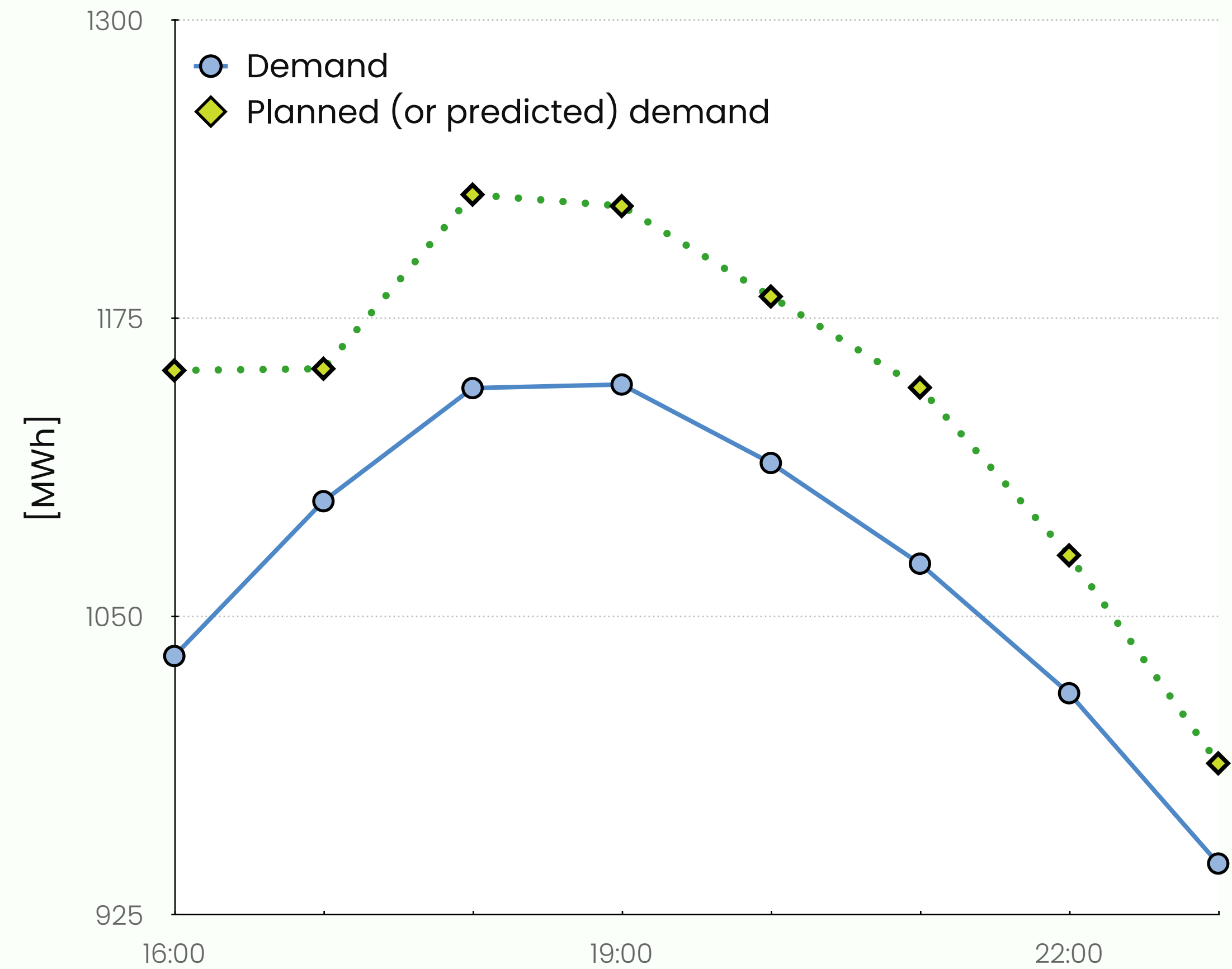
For example, moving average model MA(2) can be written as

$$y_t = \mu + \epsilon_t + \theta_1 \epsilon_{t-1} + \theta_2 \epsilon_{t-2}.$$

MA(2) MODEL

For example, moving average model MA(2) can be written as

$$y_t = \mu + \epsilon_t + \theta_1 \epsilon_{t-1} + \theta_2 \epsilon_{t-2}.$$



MA(2) MODEL

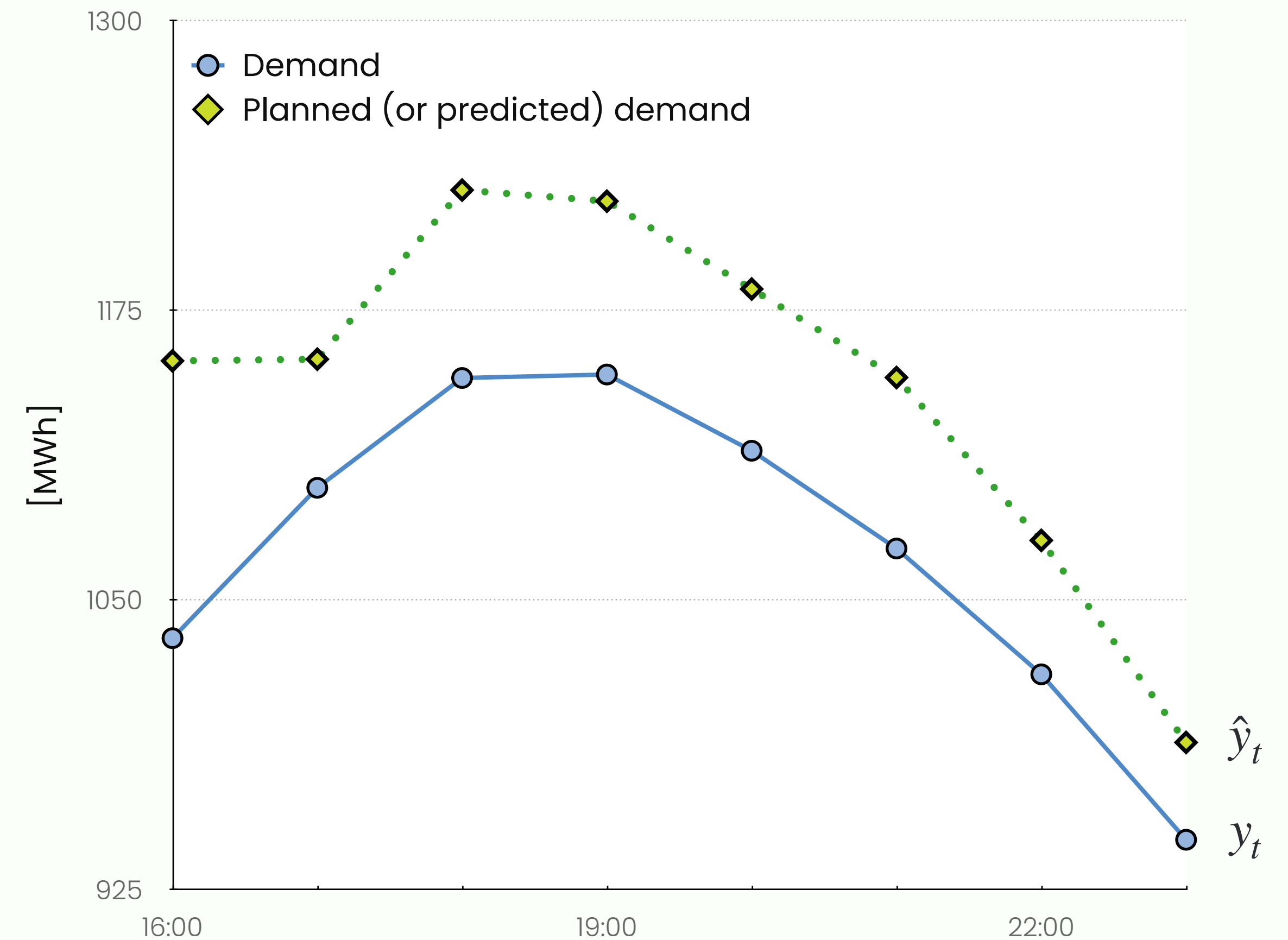
For example, moving average model MA(2) can be written as

$$y_t = \mu + \epsilon_t + \theta_1 \epsilon_{t-1} + \theta_2 \epsilon_{t-2}.$$

MA helps to correct the error made in the previous time step.

Practical observation (we do not have access to the current error ϵ_t):

$$\hat{y}_t = \mu + \theta_1 \epsilon_{t-1} + \theta_2 \epsilon_{t-2}.$$



MA(2) MODEL

For example, moving average model MA(2) can be written as

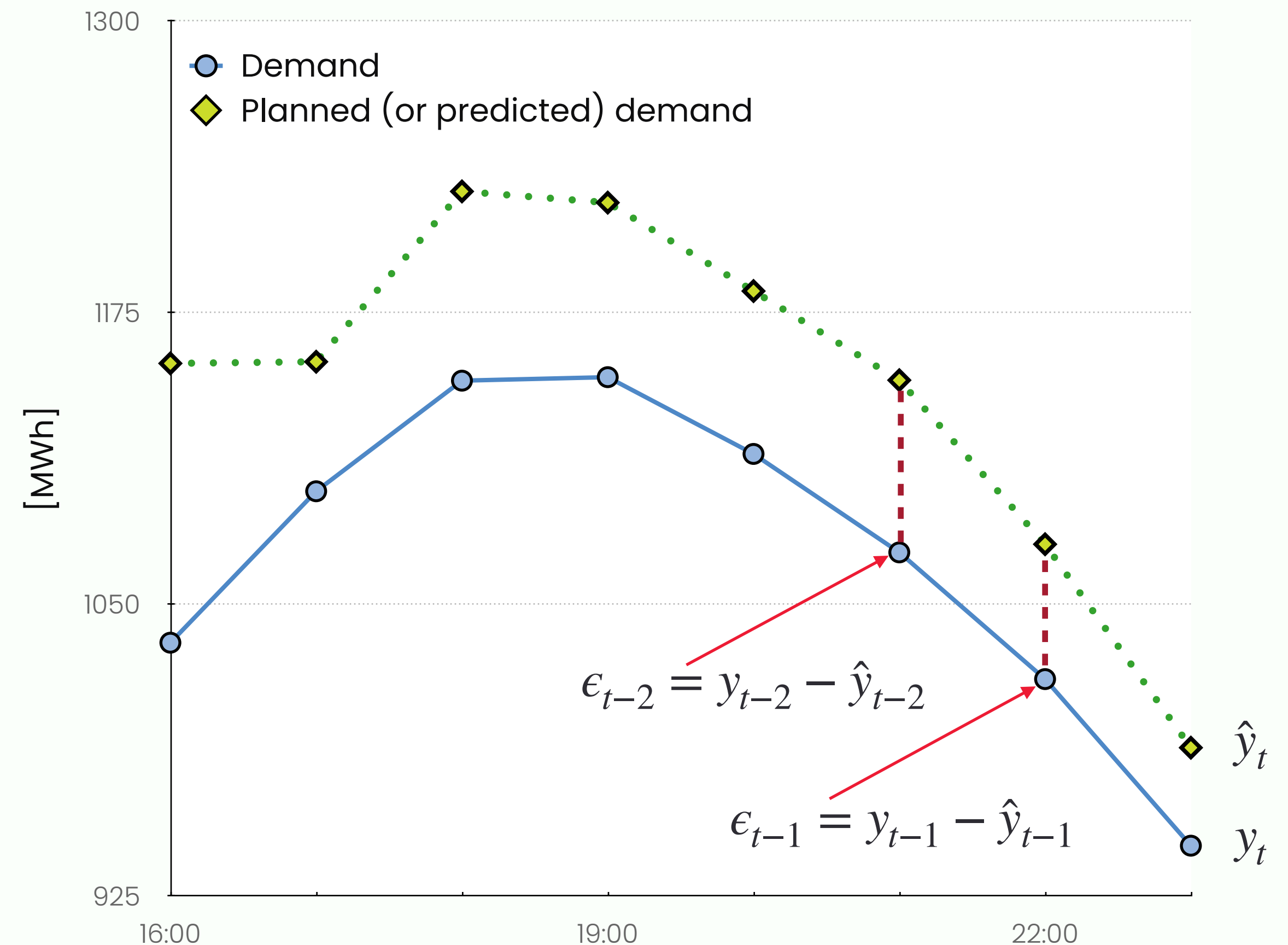
$$y_t = \mu + \epsilon_t + \theta_1 \epsilon_{t-1} + \theta_2 \epsilon_{t-2}$$

lagged errors

MA helps to correct the error made in the previous time step.

Practical observation (we do not have access to the current error ϵ_t):

$$\hat{y}_t = \mu + \theta_1 \epsilon_{t-1} + \theta_2 \epsilon_{t-2}$$



MOVING AVERAGE (MA) MODEL

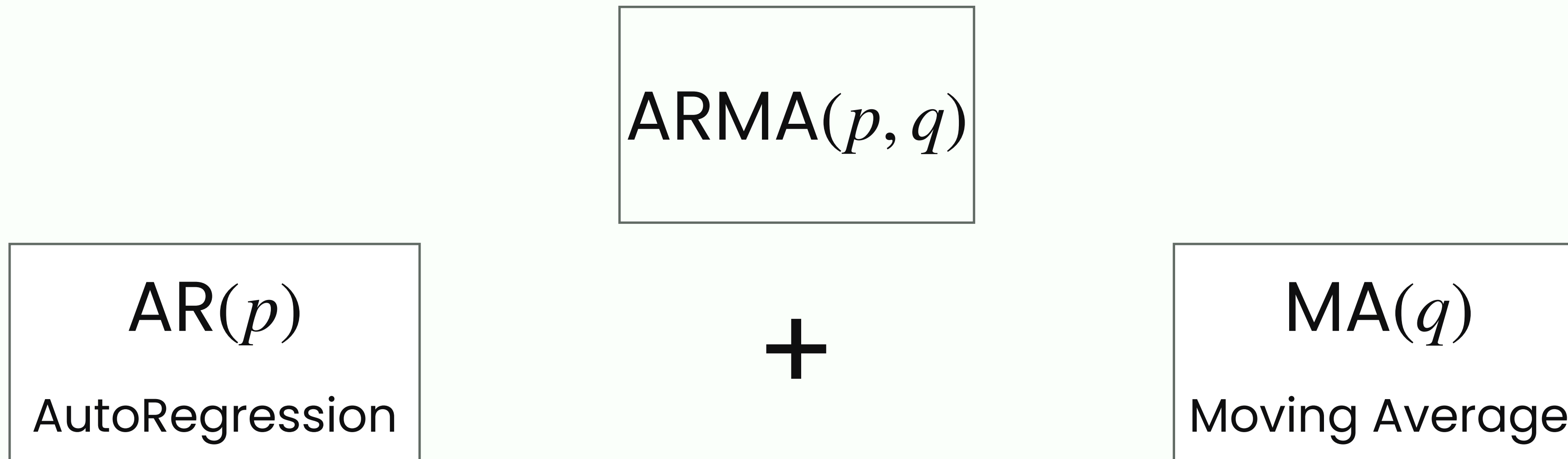
Rather than using past values of the target variable in a regression, a moving average model uses past errors and can be written as

$$y_t = \mu + \epsilon_t + \sum_{j=1}^q \theta_j \epsilon_{t-j} = \mu + \epsilon_t + \theta_1 \epsilon_{t-1} + \theta_2 \epsilon_{t-2} + \cdots + \theta_q \epsilon_{t-q},$$

where $\epsilon_t \sim wn(0, \sigma_w^2)$ is white noise, $\theta_1, \dots, \theta_q$ are parameters, μ is the mean of the series.

We refer to this as an MA(q) model, a moving average model of order q .

AUTOREGRESSIVE MOVING AVERAGE (ARMA)



AR models a series based solely on the past values in the series (lags).

MA models a series based solely on the past errors in the series (error lags).

The family of methods is reach:

SARIMA, VARIMA, FARIMA, ARCH, VAR, etc.

ARMA: FULL MODEL

The full model can be written as

$$\begin{aligned} y_t &= c + \epsilon_t + \sum_{i=1}^p \phi_i y_{t-i} + \sum_{j=1}^q \theta_j \epsilon_{t-j} \\ &= c + \epsilon_t + \phi_1 y_{t-1} + \cdots + \phi_p y_{t-p} + \theta_1 \epsilon_{t-1} + \cdots + \theta_q \epsilon_{t-q}. \end{aligned}$$

We call this an **ARMA**(p, q) model, where

- p is the number of lag observations,
- q is the order of the moving average part.

ARMA: EXAMPLE

For example, ARMA(2,1) can be written as

$$\hat{y}_t = \phi_1 y_{t-1} + \phi_2 y_{t-2} + \theta_1 \epsilon_{t-1}$$

or using backward shift (lag) notation as

$$(1 - \phi_1 B - \phi_2 B^2) \hat{y}_t = (1 + \theta_1 B) \epsilon_t.$$

ESTIMATING ARMA ORDERS

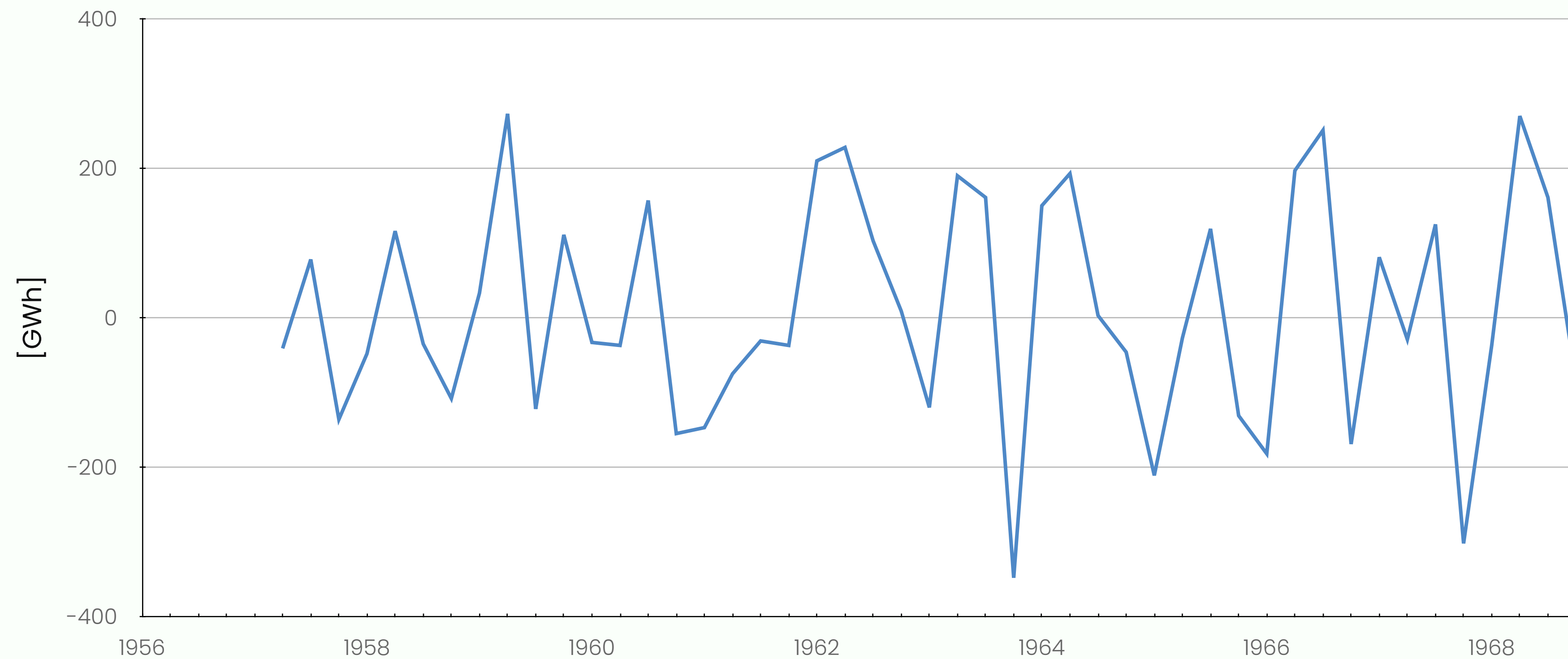
Note that the selection for p and q is not unique, and one can use:

- ACF and PACF plots
- information criteria (IC) (Akaike's IC or Bayesian IC)
- performance metrics (MSE or RMSE)

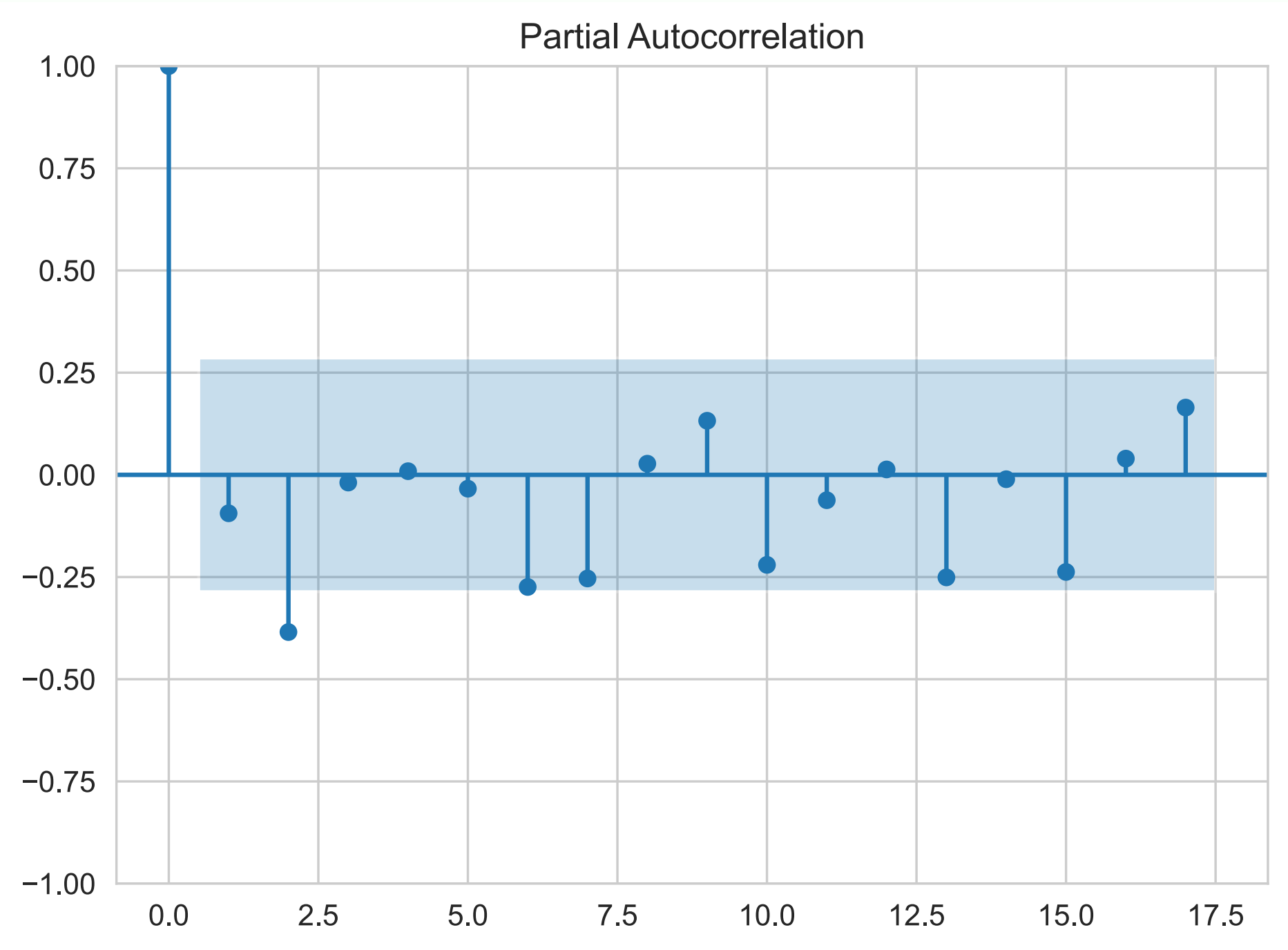
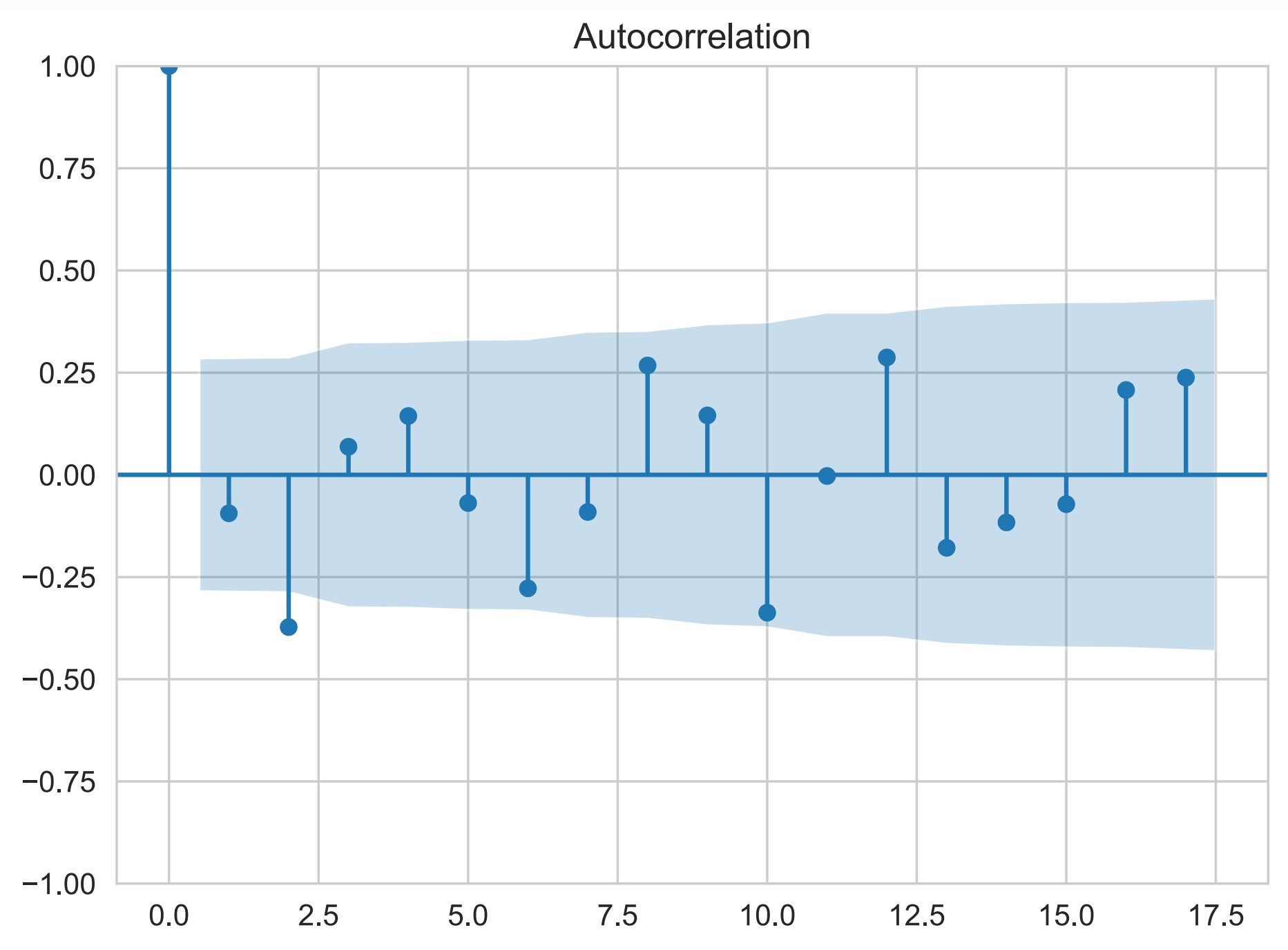
| AR(p)/MA(q) | 0 | 1 | 2 | 3 | ... |
|-------------|---|---|---|---|-----|
| 0 | | | | | |
| 1 | | | | | |
| 2 | | | | | |
| 3 | | | | | |
| ... | | | | | |

ILLUSTRATIVE EXAMPLE

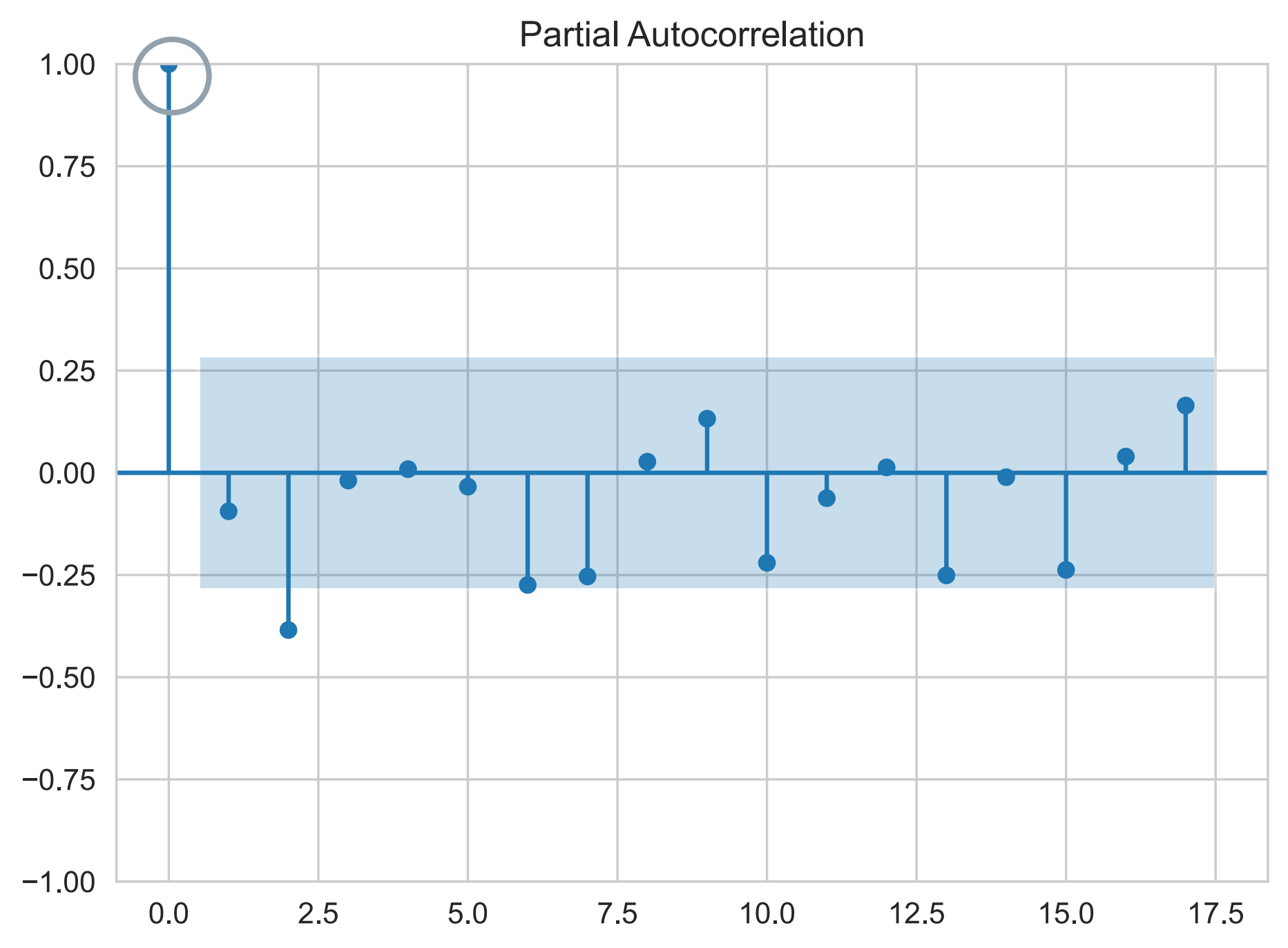
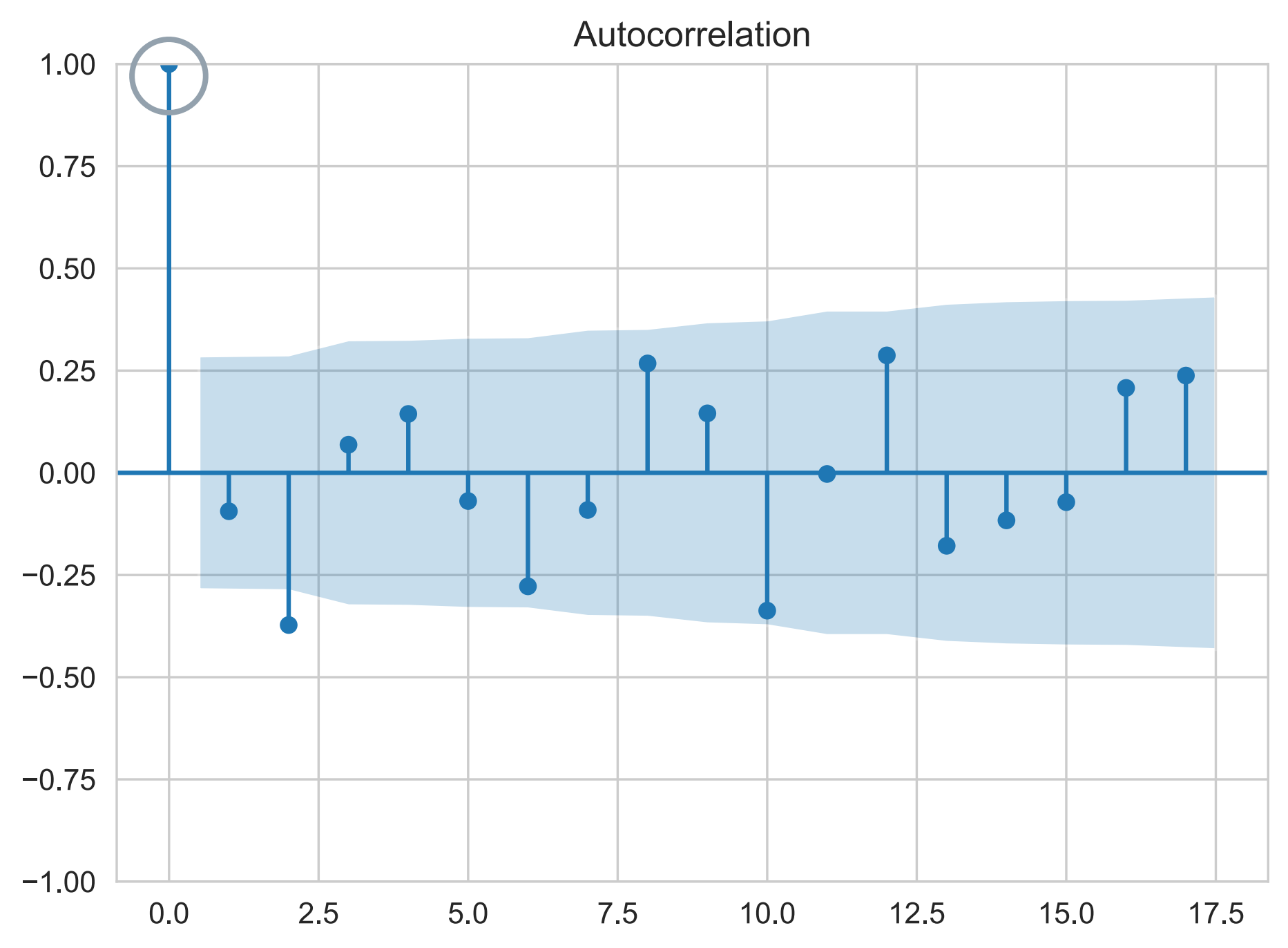
Consider Australian electricity production data after seasonal and first difference transformation $(1 - B)(1 - B^4) y_t$.



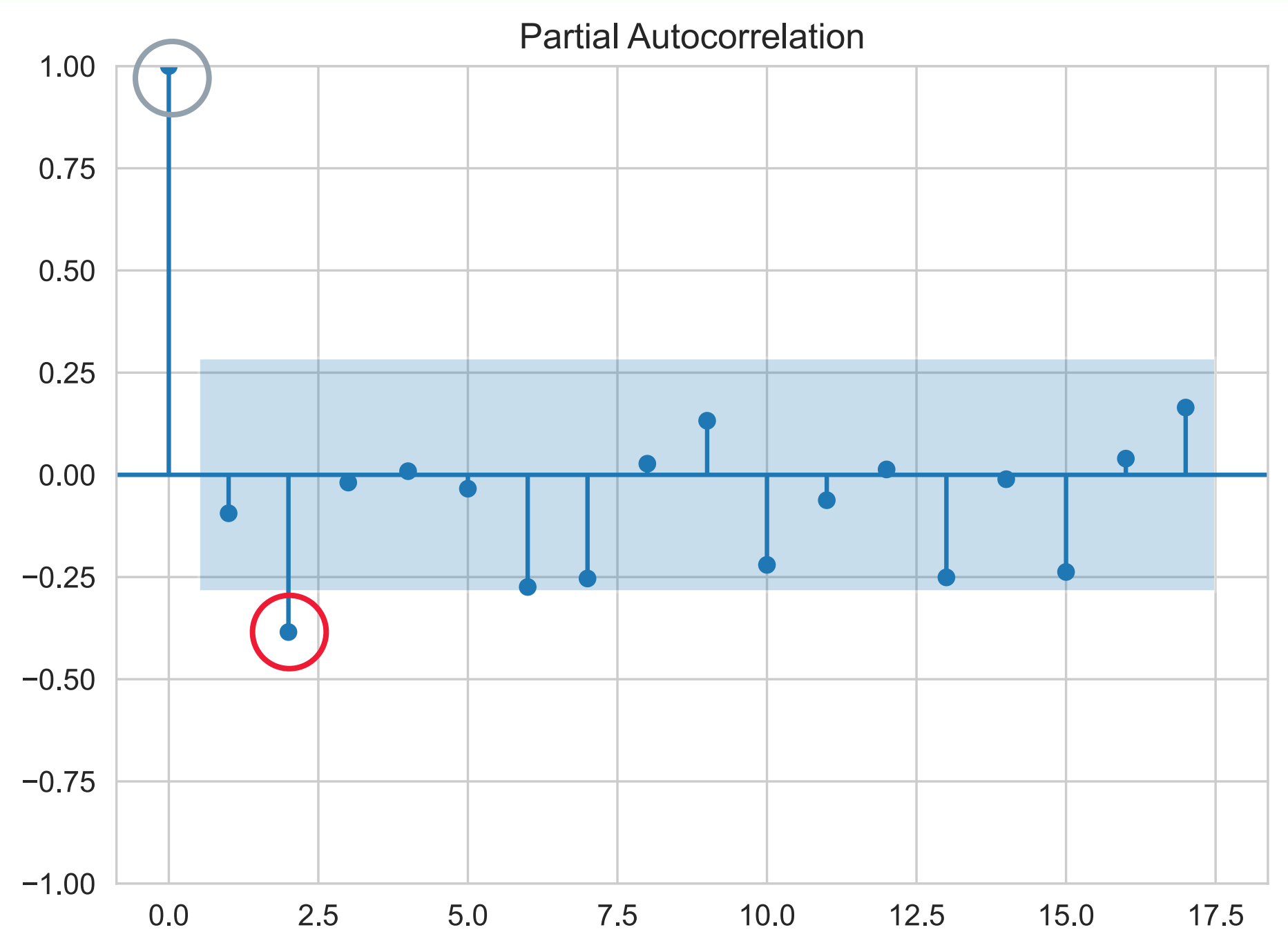
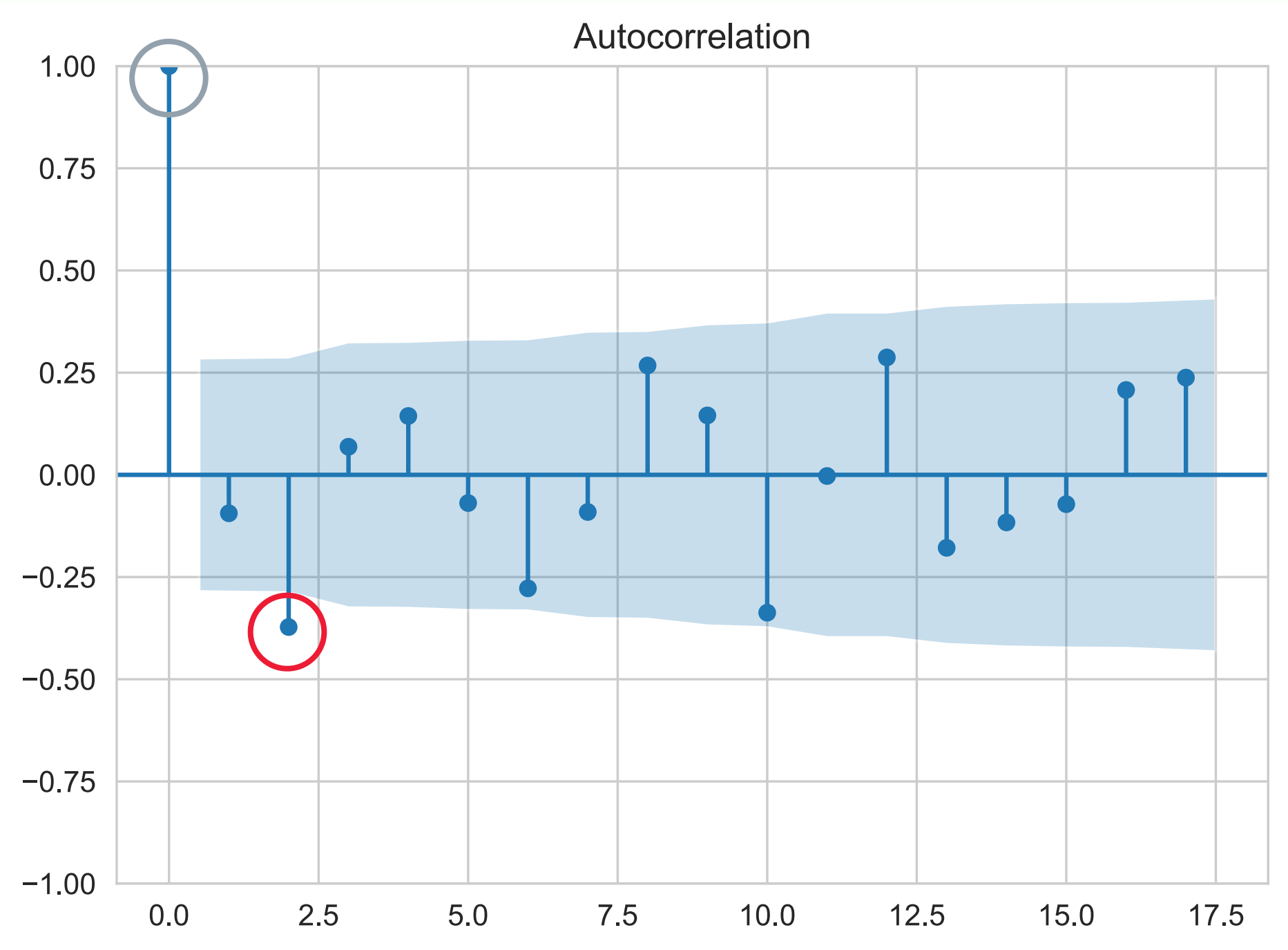
ACF & PACF



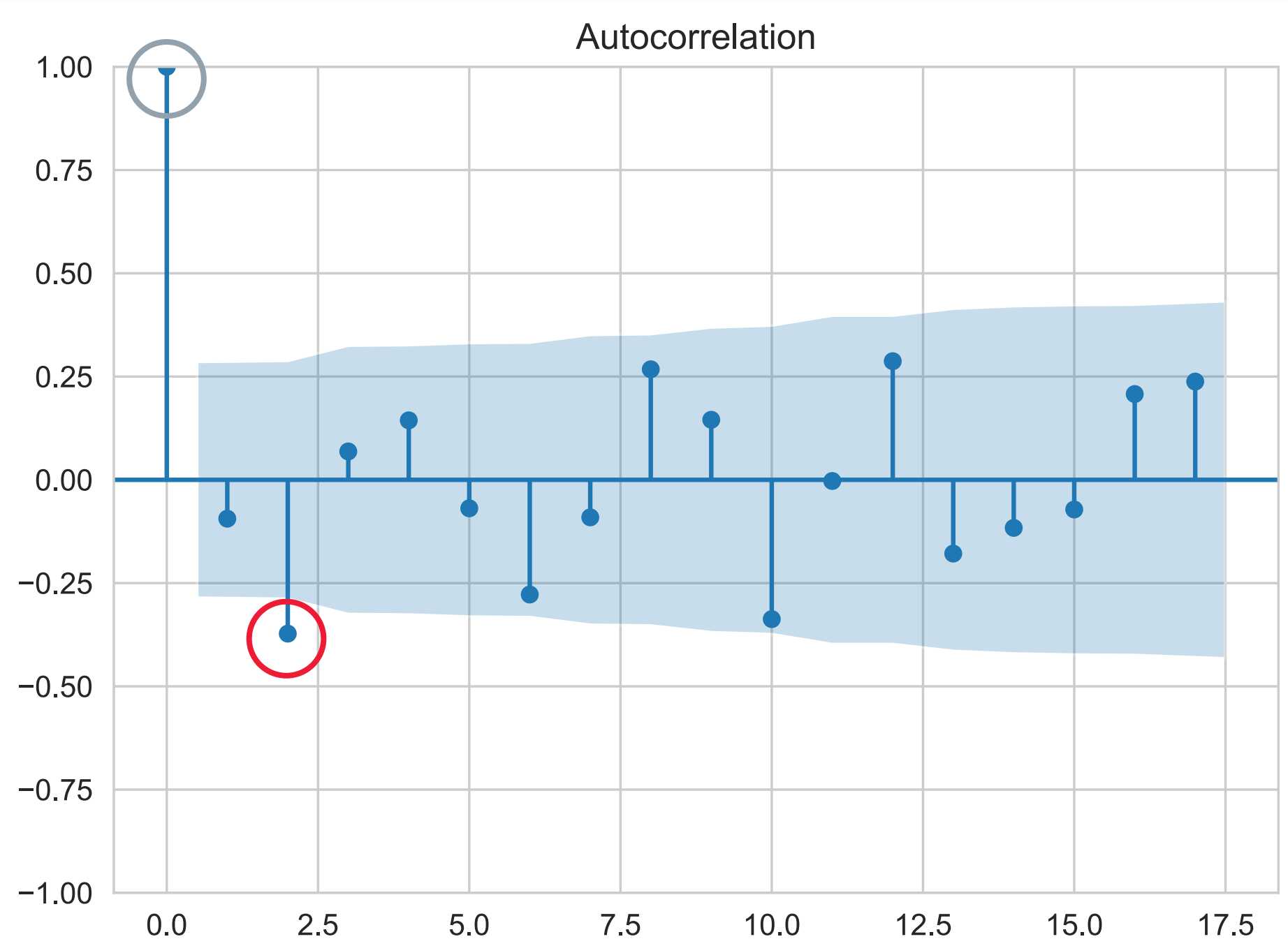
ACF & PACF



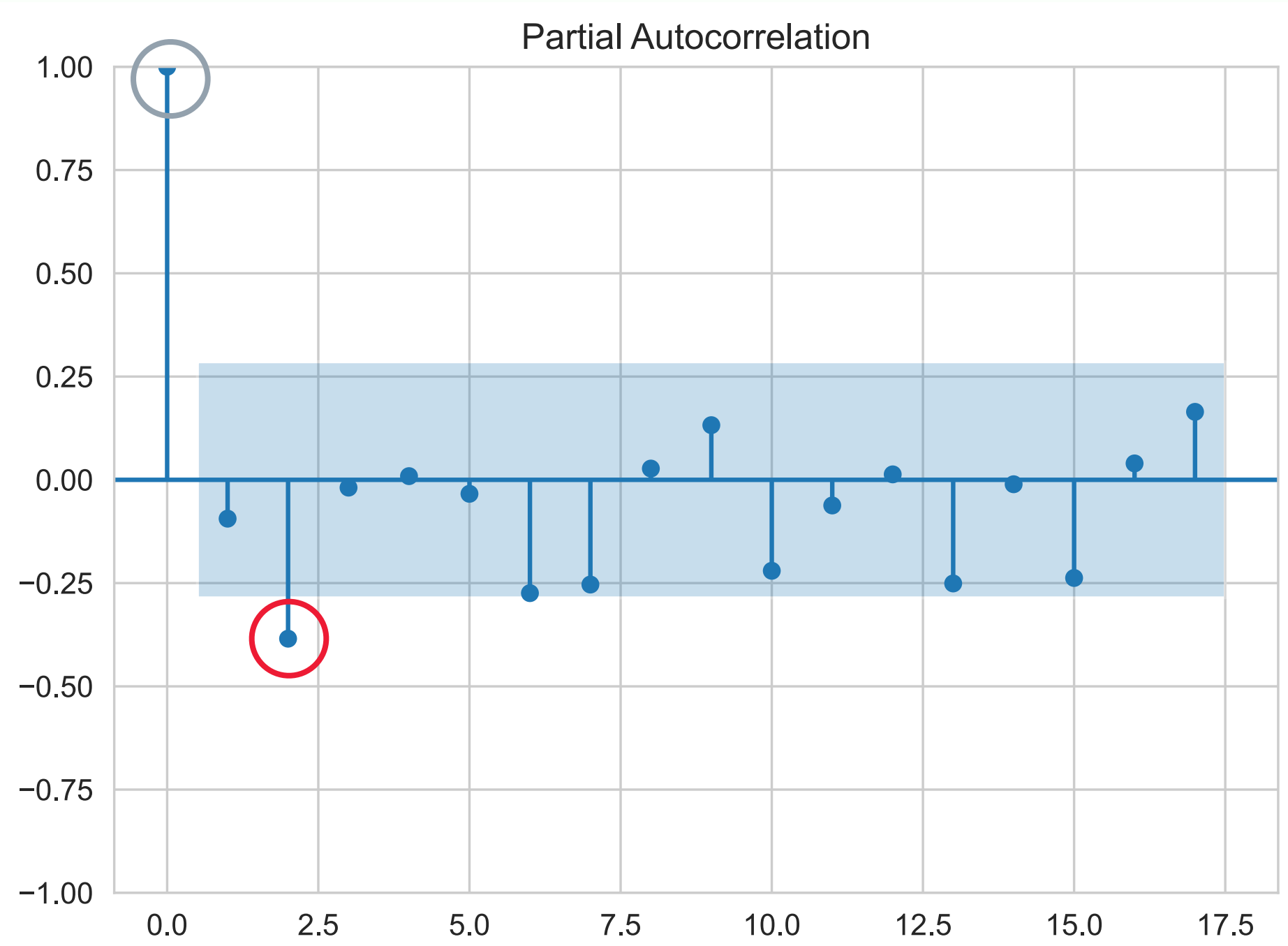
ACF & PACF



ACF & PACF



Based on ACF q option is:
MA(2)



Based on PACF p option is:
AR(2)

ESTIMATION

| SARIMAX Results | | | | | | |
|-------------------------|------------------|-------------------|----------|-------|----------|----------|
| ===== | | | | | | |
| Dep. Variable: | electricity | No. Observations: | 44 | | | |
| Model: | ARIMA(2, 0, 2) | Log Likelihood | -277.270 | | | |
| Date: | Tue, 22 Jul 2025 | AIC | 566.539 | | | |
| Time: | 15:12:58 | BIC | 577.244 | | | |
| Sample: | 06-01-1957 | HQIC | 570.509 | | | |
| | - 03-01-1968 | | | | | |
| Covariance Type: | opg | | | | | |
| ===== | | | | | | |
| | coef | std err | z | P> z | [0.025 | 0.975] |
| ----- | | | | | | |
| const | 4.6923 | 18.744 | 0.250 | 0.802 | -32.044 | 41.429 |
| ar.L1 | 0.0975 | 0.217 | 0.449 | 0.653 | -0.328 | 0.523 |
| ar.L2 | -0.9488 | 0.144 | -6.567 | 0.000 | -1.232 | -0.666 |
| ma.L1 | -0.2016 | 0.333 | -0.606 | 0.544 | -0.853 | 0.450 |
| ma.L2 | 0.8692 | 0.273 | 3.189 | 0.001 | 0.335 | 1.404 |
| sigma2 | 1.791e+04 | 6557.285 | 2.732 | 0.006 | 5060.488 | 3.08e+04 |
| ===== | | | | | | |
| Ljung-Box (L1) (Q): | 0.23 | Jarque-Bera (JB): | 1.23 | | | |
| Prob(Q): | 0.63 | Prob(JB): | 0.54 | | | |
| Heteroskedasticity (H): | 1.39 | Skew: | -0.04 | | | |
| Prob(H) (two-sided): | 0.53 | Kurtosis: | 2.18 | | | |
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ESTIMATION

SARIMAX Results

Dep. Variable:

electricity

No. Observations:

44

Model:

ARIMA(2, 0, 2)

Log Likelihood

-277.270

Date:

Tue, 22 Jul 2025

AIC

566.539

Time:

15:12:58

BIC

577.244

Sample:

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ESTIMATION

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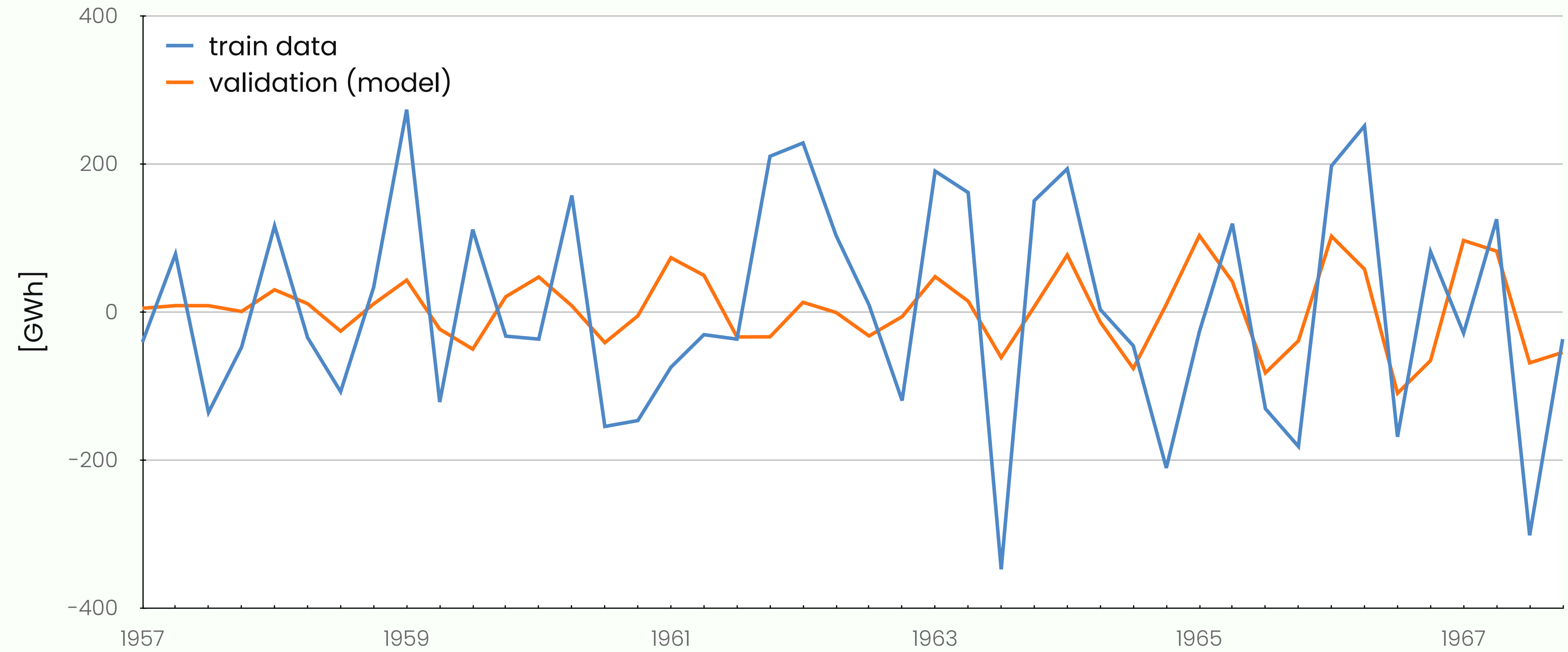
ESTIMATION

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| Prob(H) (two-sided): | 0.53 | Kurtosis: | 2.18 | | | |
| ===== | | | | | | |

The final ARMA(2,2) model (with B^1 lags and intercept excluded) reads as:

$$\hat{y}_t = -0.9488y_{t-2} + 0.8692\epsilon_{t-2}.$$

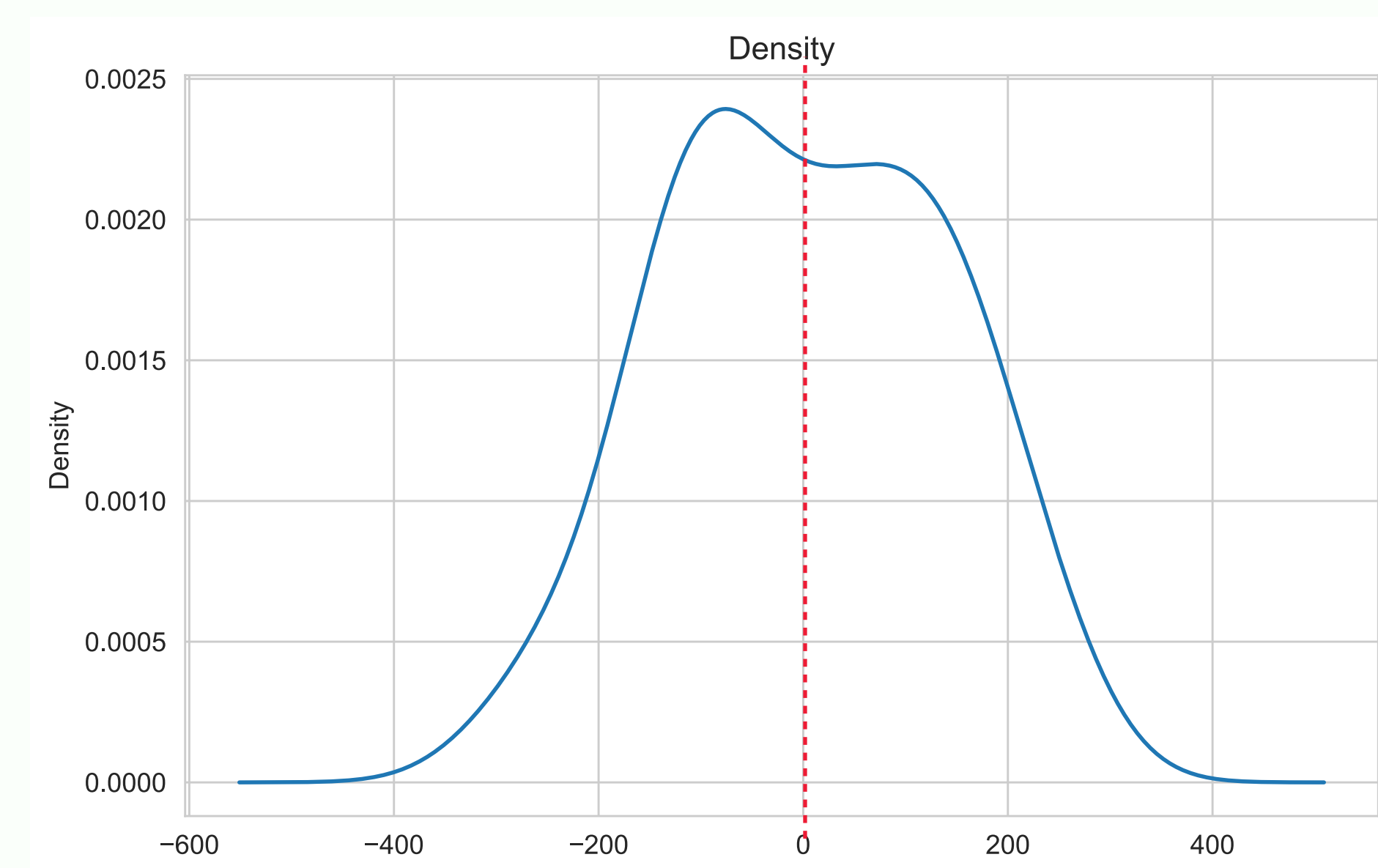
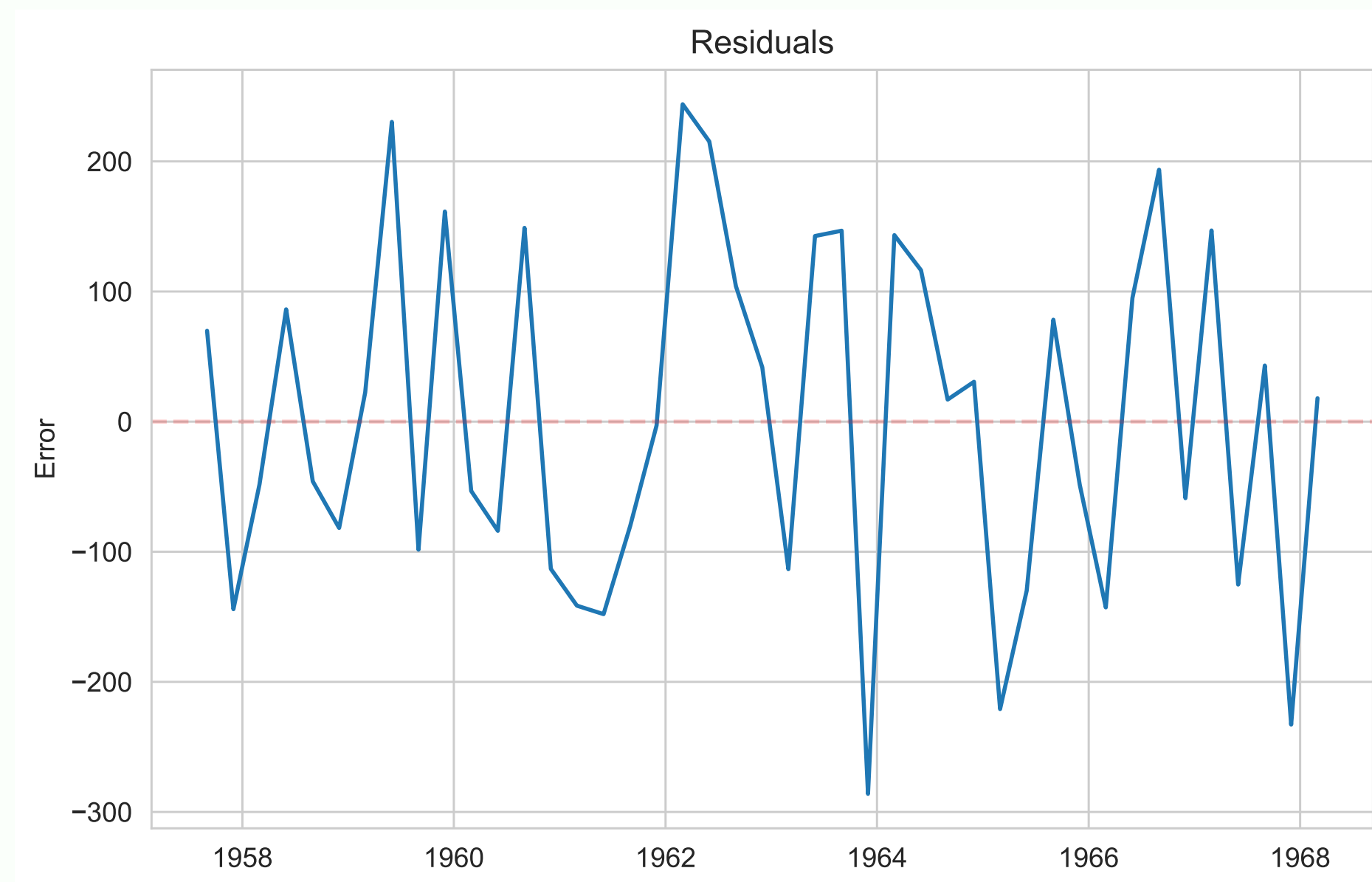
VALIDATION: TRAIN DATA



RMSE = 131.19

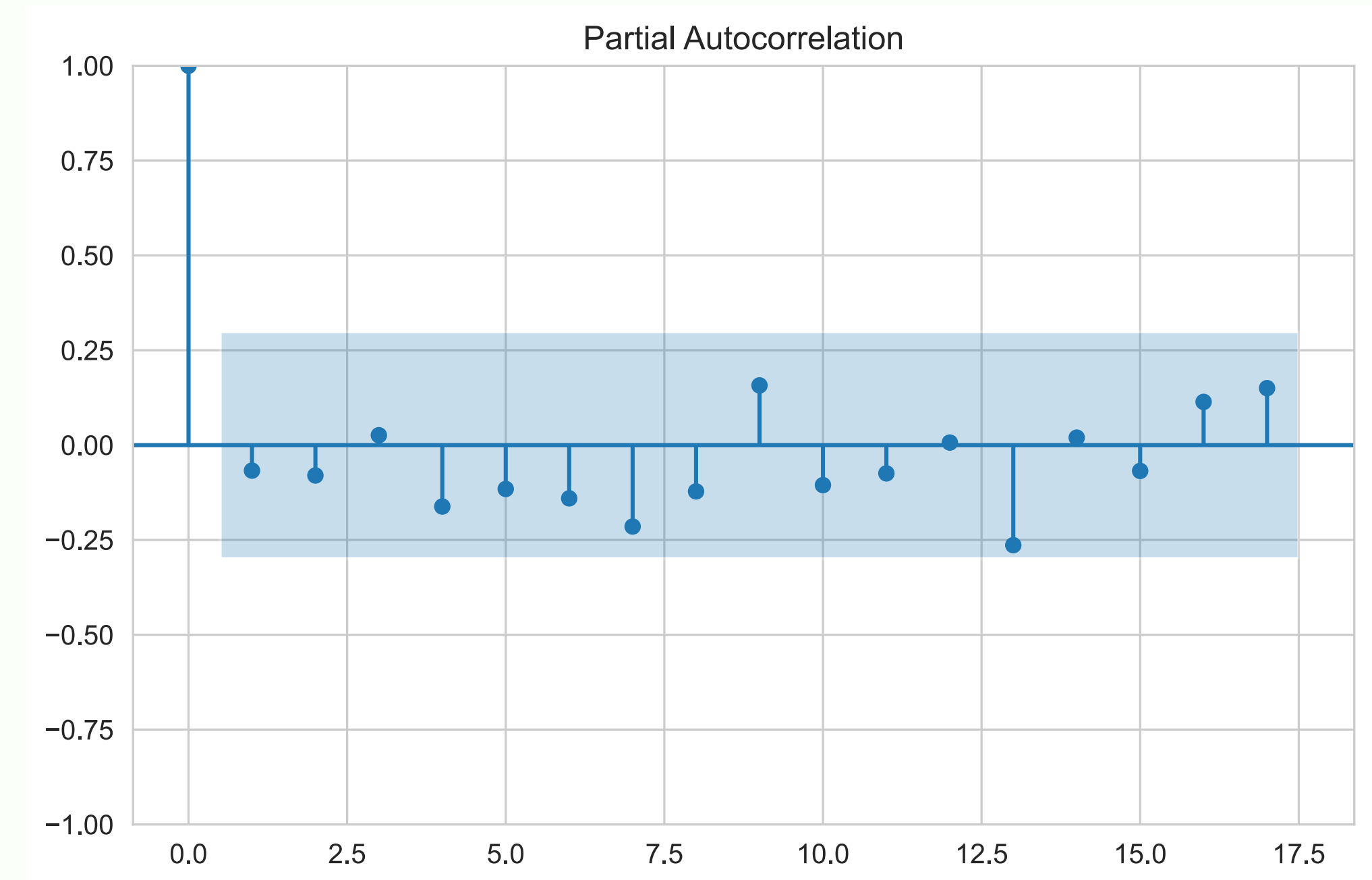
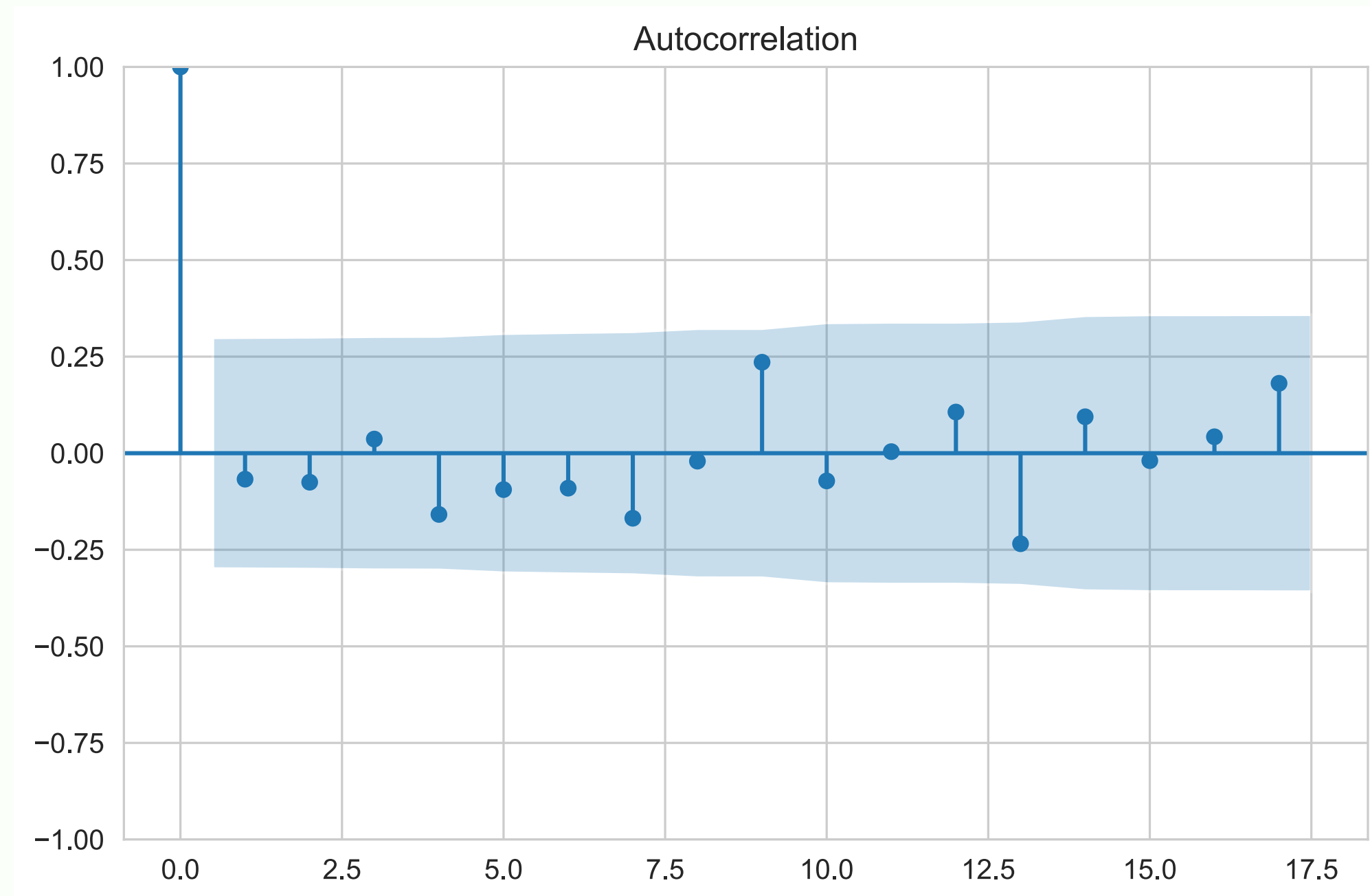
VALIDATION: RESIDUALS

The residuals look random in general, and their density looks kind of normally distributed with a mean close to 0.



VALIDATION: RESIDUALS (2)

ACF and PACF of the residuals



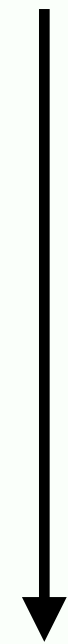
These show that the residuals are close to white noise, and the model can be further utilised.

ARMA EXTENDED

$$\text{ARMA}(p, q)$$

ARMA EXTENDED

$$\text{ARMA}(p, q)$$

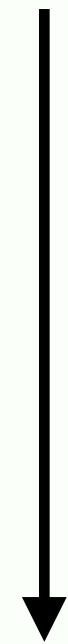


d is the degree of first differencing

$$\text{ARIMA}(p, d, q)$$

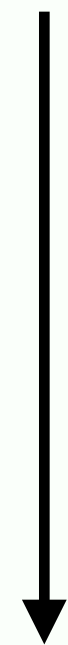
ARMA EXTENDED

$$\text{ARMA}(p, q)$$



d is the degree of first differencing

$$\text{ARIMA}(p, d, q)$$



Seasonal: $\text{AR}(P)$, first differencing (D), and $\text{MA}(Q)$ parts
 m is the number of observations per year

$$\text{SARIMA}(p, d, q)(P, D, Q)_m$$

OUTRO

Evaluating Time Series models

TIME SERIES VS CROSS-SECTIONAL

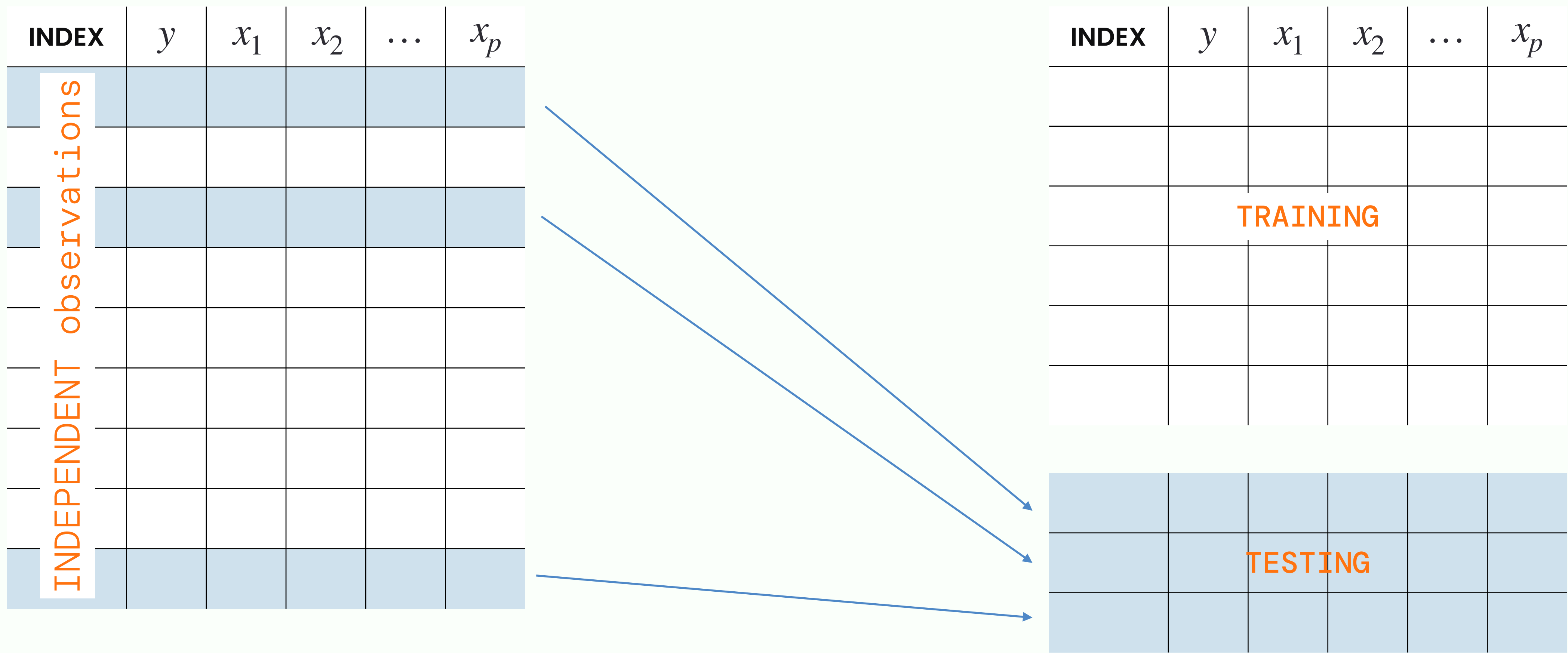
Time Series

Set of ordered data values
consistently collected over an
interval of time.

Cross-sectional

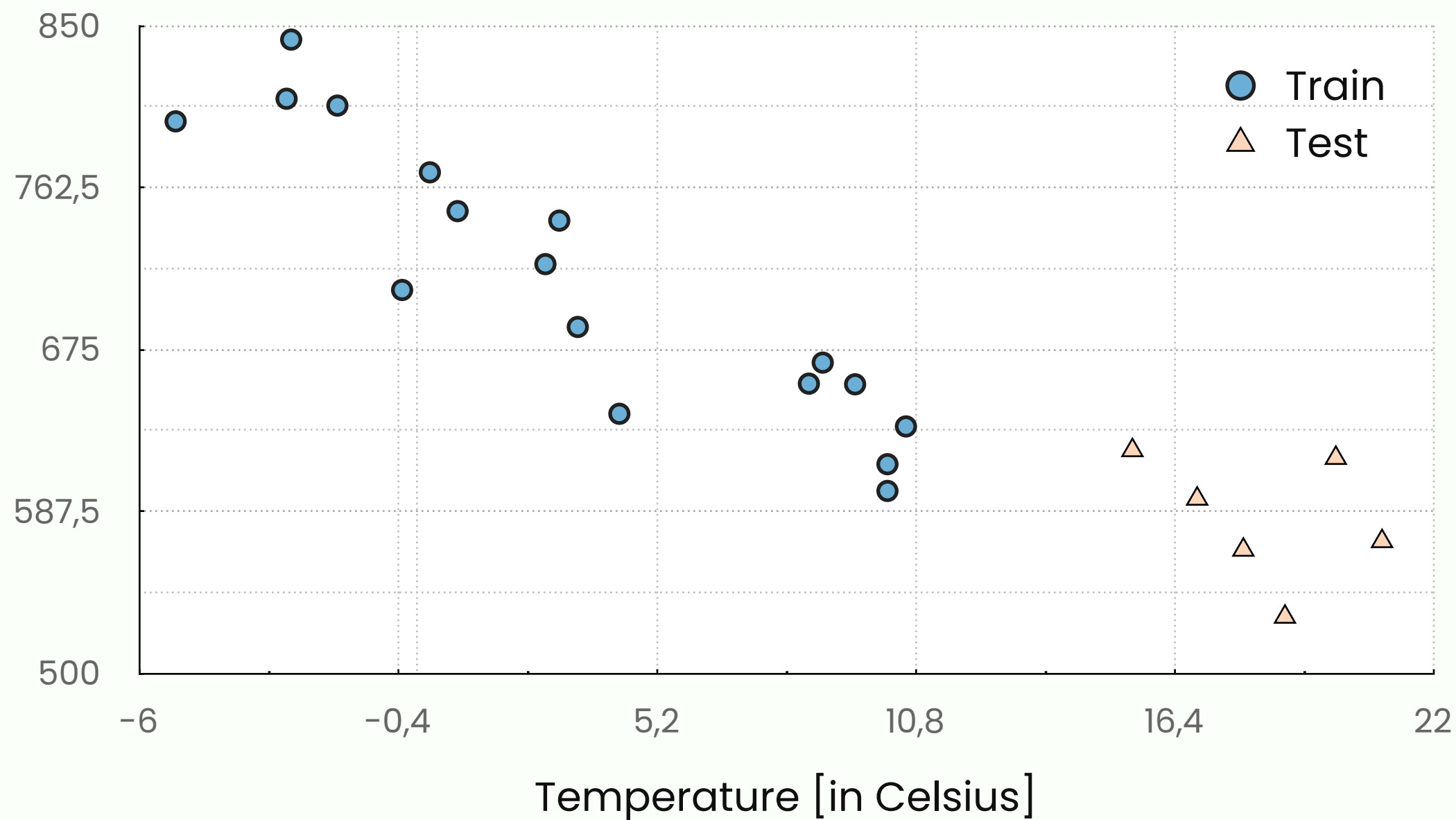
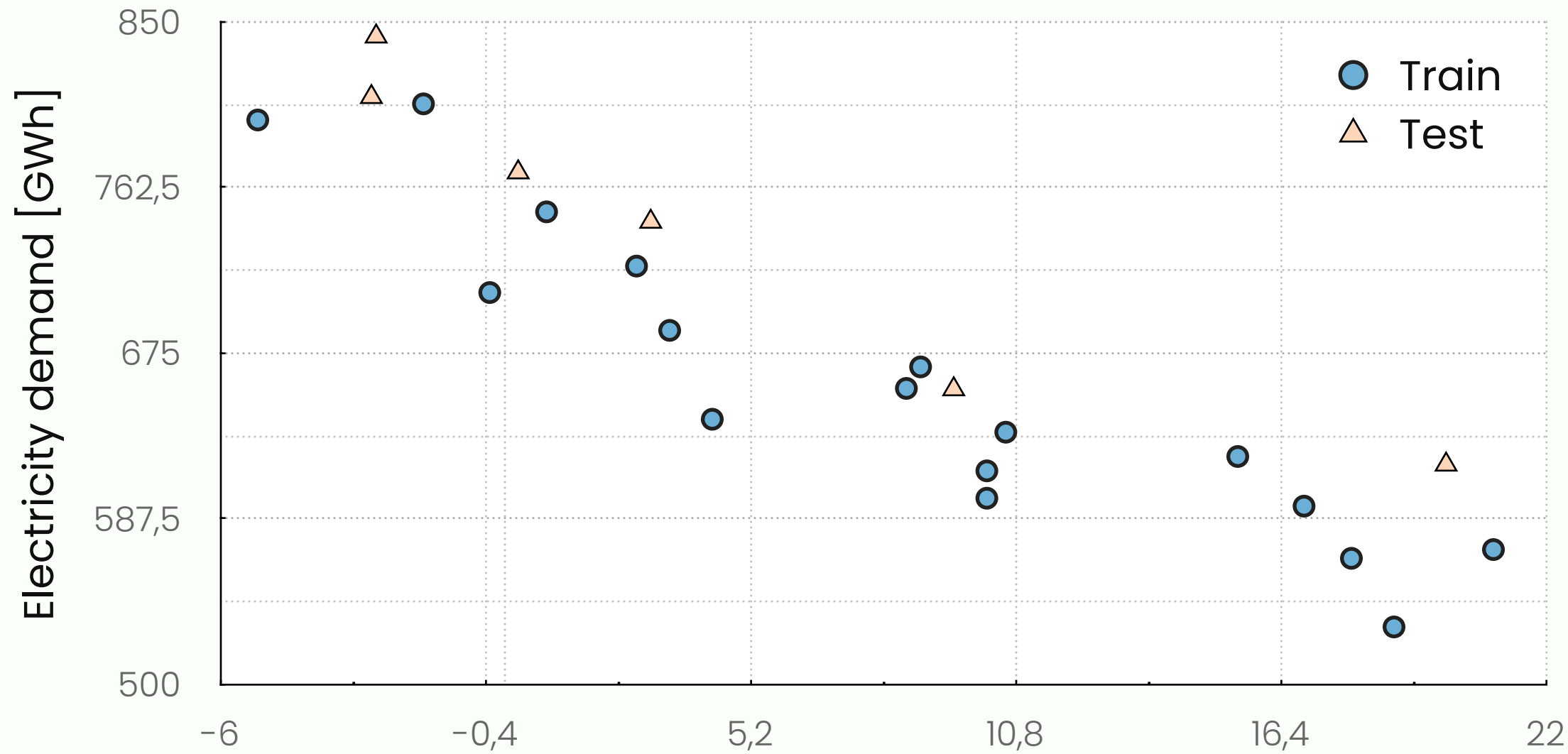
Set of data values collected at a
fixed point in time, or where time
is of *no significance*.

CS: RANDOM SPLIT FOR TRAINING/TESTING



(UN)LUCKY SPLIT?

| y | x_1 |
|------|-------|
| -2.7 | 842 |
| ... | ... |
| 0.3 | 771 |
| ... | ... |
| 3.1 | 745 |
| ... | ... |
| 9.5 | 656 |
| ... | ... |
| -2.8 | 810 |



| y | x_1 |
|------|-------|
| ... | ... |
| 18.8 | 530 |
| 20.9 | 571 |
| 15.5 | 620 |
| ... | ... |
| 16.9 | 594 |
| 17.9 | 567 |
| 19.9 | 616 |
| ... | ... |

K-FOLD CROSS-VALIDATION

[illegible]

| INDEX | y | x_1 | x_2 | \dots | x_p |
|--------------|-----------------|-------|-------|---------|-------|
| | | | | | |
| | | | | | |
| | TRAINING | | | | |
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| | | | | | |
| | TESTING | | | | |
| | | | | | |

K-FOLD CROSS-VALIDATION

[illegible]

| INDEX | y | x_1 | x_2 | \dots | x_p |
|--------------|-----------------|-------|-------|---------|-------|
| | | | | | |
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| | TRAINING | | | | |
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| | TESTING | | | | |
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[illegible]

K-FOLD CROSS-VALIDATION

[illegible][illegible][illegible][illegible]

• • •

k -fold cross-validation

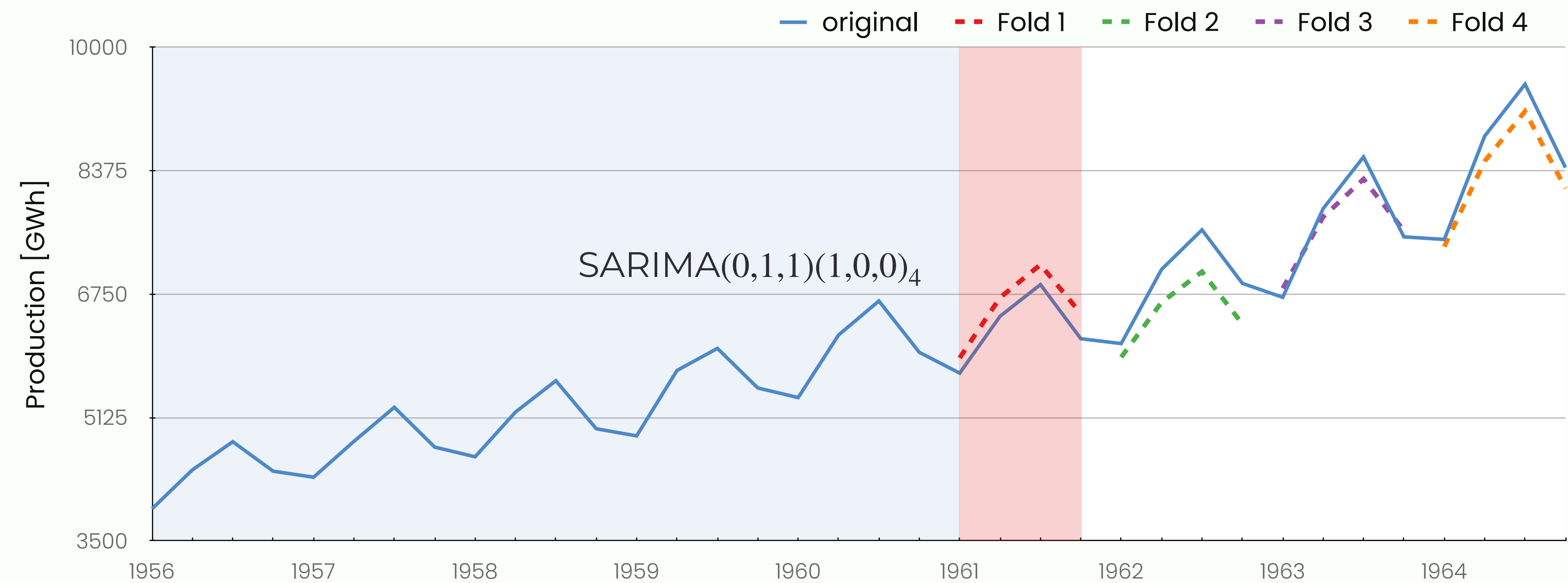
~~TS: RANDOM~~ SPLIT FOR TRAINING/TESTING

[illegible]

| INDEX | y | x_1 | x_2 | \dots | x_p |
|-------|-----|-------|-------|---------|-------|
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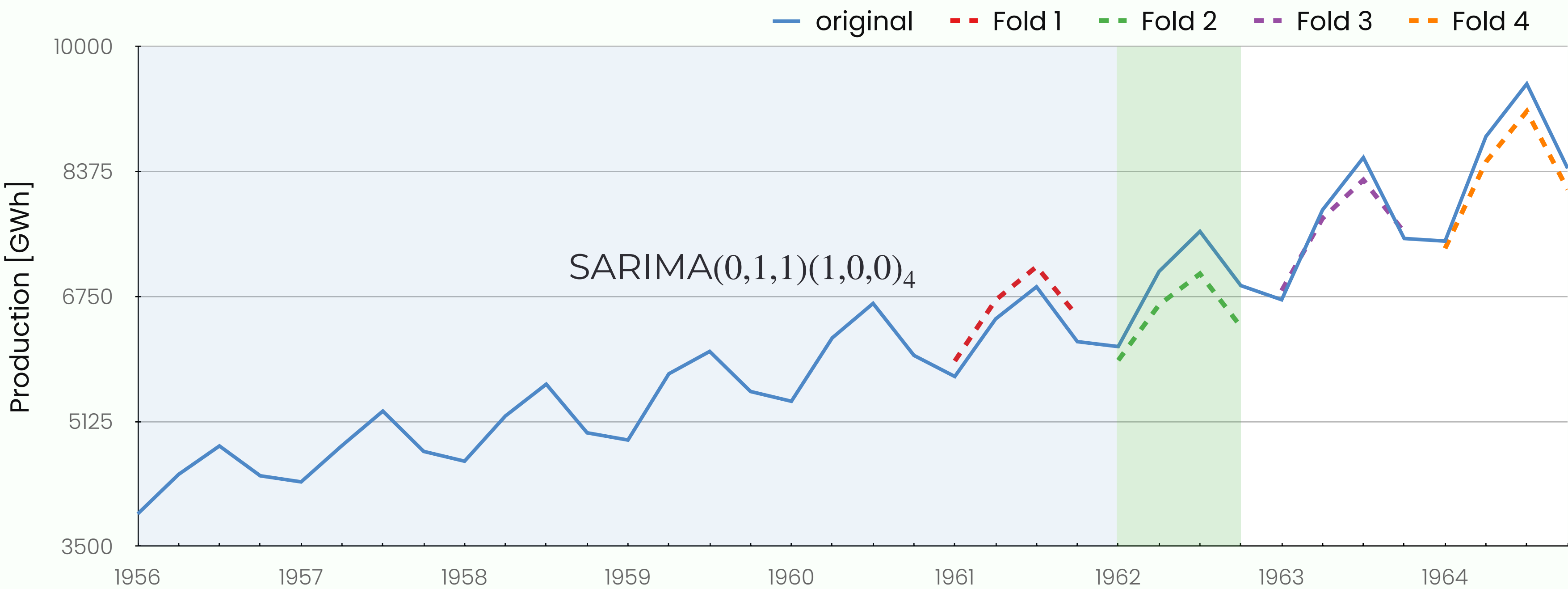
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WALK-FORWARD VALIDATION



| | Q1'56-Q4'60 | Q1-Q4'61 | Q1-Q4'62 | Q1-Q4'63 | Q1-Q4'64 | SARIMA (0,1,1)(1,0,0) ₄ | SARIMA (0,1,2)(1,0,0) ₄ |
|----------|-------------|----------|----------|----------|----------|---------------------------------------|---------------------------------------|
| Fold 1 | train | test | | | | 4.13 | 2.62 |
| Fold 2 | train | train | test | | | 6.02 | 3.69 |
| Fold 3 | train | train | train | test | | 1.91 | 3.14 |
| Fold 4 | train | train | train | train | test | 3 | 4.66 |
| avg MAPE | | | | | | 3.76 | 3.53 |

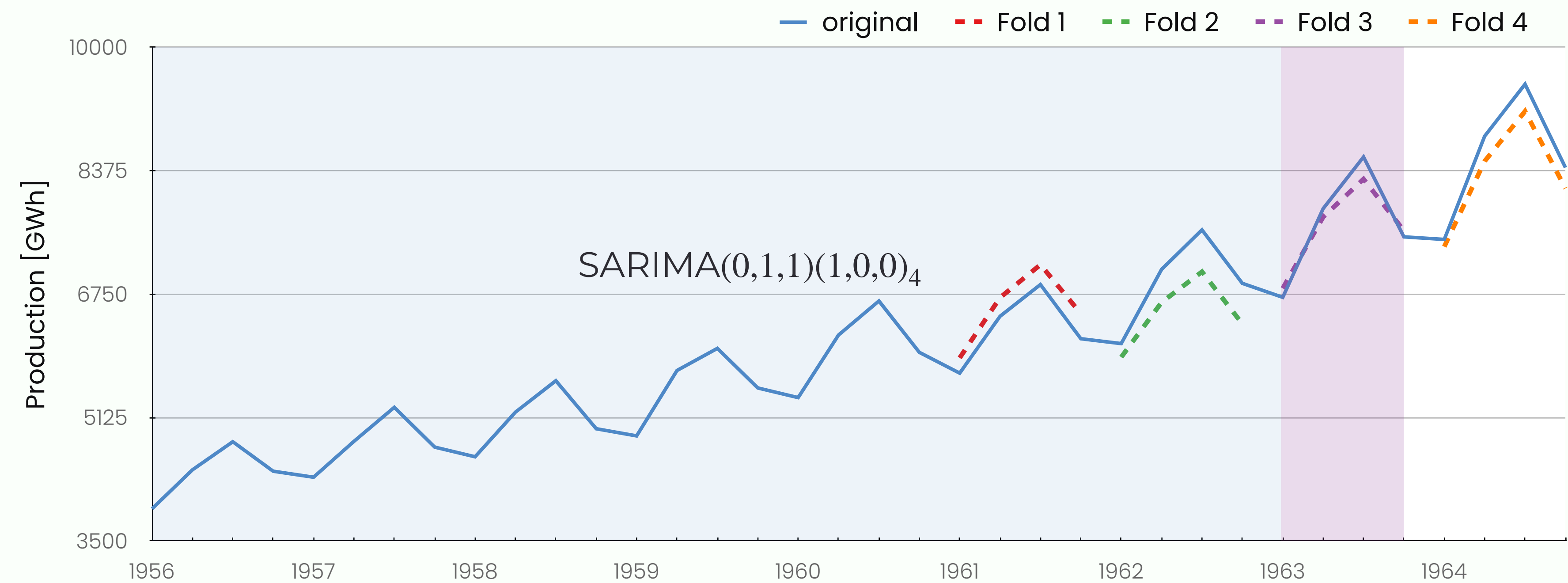
WALK-FORWARD VALIDATION



| | Q1'56-Q4'60 | Q1-Q4'61 | Q1-Q4'62 | Q1-Q4'63 | Q1-Q4'64 | SARIMA (0,1,1)(1,0,0) ₄ | SARIMA (0,1,2)(1,0,0) ₄ |
|----------|-------------|----------|----------|----------|----------|---------------------------------------|---------------------------------------|
| Fold 1 | train | test | | | | 4.13 | 2.62 |
| Fold 2 | train | train | test | | | 6.02 | 3.69 |
| Fold 3 | train | train | train | test | | 1.91 | 3.14 |
| Fold 4 | train | train | train | train | test | 3 | 4.66 |
| avg MAPE | | | | | | 3.76 | 3.53 |

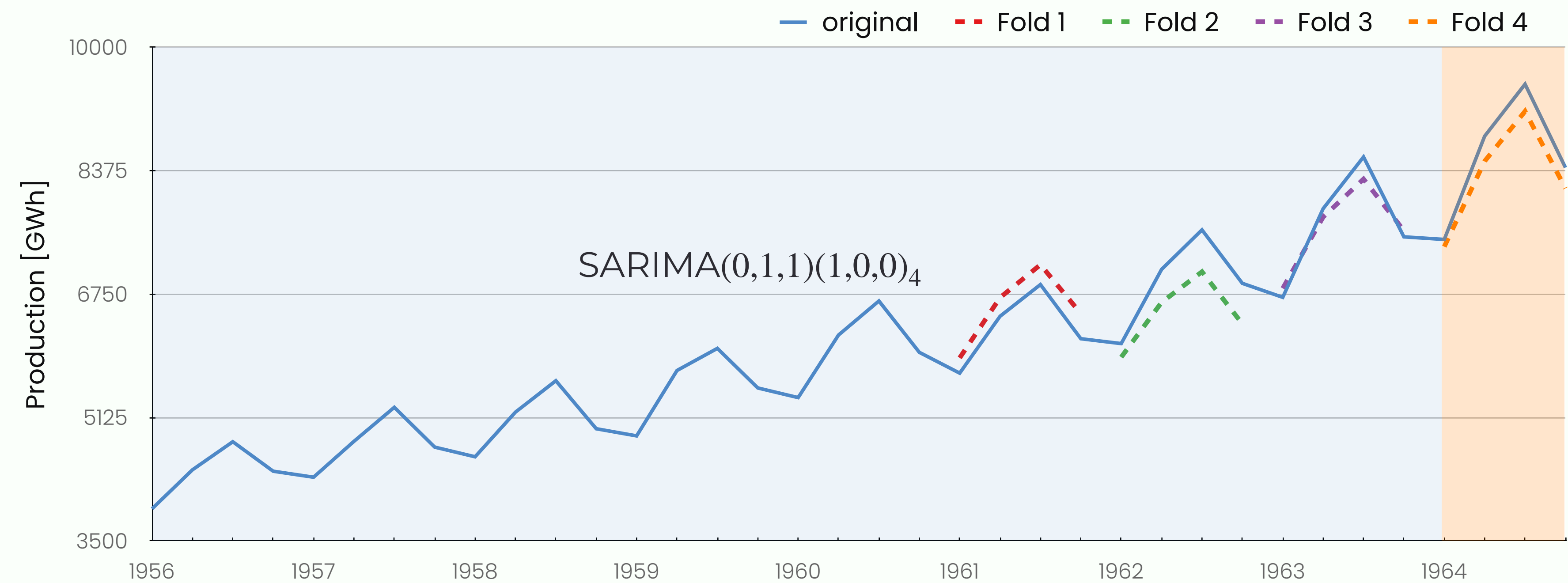
Alternative method is **sliding window validation** which uses fixed-size window for training.

WALK-FORWARD VALIDATION



| | Q1'56-Q4'60 | Q1-Q4'61 | Q1-Q4'62 | Q1-Q4'63 | Q1-Q4'64 | SARIMA (0,1,1)(1,0,0) ₄ | SARIMA (0,1,2)(1,0,0) ₄ |
|----------|-------------|----------|----------|----------|----------|---------------------------------------|---------------------------------------|
| Fold 1 | train | test | | | | 4.13 | 2.62 |
| Fold 2 | train | train | test | | | 6.02 | 3.69 |
| Fold 3 | train | train | train | test | | 1.91 | 3.14 |
| Fold 4 | train | train | train | train | test | 3 | 4.66 |
| avg MAPE | | | | | | 3.76 | 3.53 |

WALK-FORWARD VALIDATION



| | Q1'56-Q4'60 | Q1-Q4'61 | Q1-Q4'62 | Q1-Q4'63 | Q1-Q4'64 | SARIMA (0,1,1)(1,0,0) ₄ | SARIMA (0,1,2)(1,0,0) ₄ |
|----------|-------------|----------|----------|----------|----------|---------------------------------------|---------------------------------------|
| Fold 1 | train | test | | | | 4.13 | 2.62 |
| Fold 2 | train | train | test | | | 6.02 | 3.69 |
| Fold 3 | train | train | train | test | | 1.91 | 3.14 |
| Fold 4 | train | train | train | train | test | 3 | 4.66 |
| avg MAPE | | | | | | 3.76 | 3.53 |

HOME ACTIVITIES & BRAIN EXERCISE

- * Try other orders for ARMA model and compare results.
- * Try different lengths of a dataset.
- * Try other transformations (incl. variation stabilisation) and compare results.
- * Given a time series y_t of the length T . Compare the results calculated by

$$r_i = \text{ACF}(y_{t-i})$$

and

$$\text{Pearson's coefficient } r_i = \text{corr}(y_t, y_{t-i})$$

for an arbitrary $i \in \{1, \dots, T-1\}$.

Thank you!

Questions?