

# Assessing County-level Vulnerability of Opioid Overdose Mortality in Georgia

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## Abstract

**Background** Finding local indicators of opioid overdose vulnerability can be useful in the prevention and mitigation of opioid-related harms. We assessed county-level vulnerability for opioid overdose mortality in Georgia.

**Methods** We collected a total of 89 plausible indicators of opioid overdose mortality. We used a stepwise dimension reduction process consisting of an empirical review, principal component analysis, and factor analysis before fitting a negative binomial regression model, assessing multicollinearity, and selecting final predictors. Lastly, we created a vulnerability index that used the coefficients from the final model to rank all 159 counties in Georgia.

**Results** Four indicators were significantly associated with opioid overdose mortality: percent of population non-Hispanic White, rate of emergency room (ER) and inpatient visits related to all drugs per 100,000 persons, rate of ER and inpatient visits related to all opioids per 100,000 persons, and the rate of specialty care providers per 100,000 persons.

**Conclusions** Our assessment can be utilized on a state and local level to reduce vulnerability to fatal opioid overdoses by expanding prevention and mitigation resources in identified at-risk counties and using the associated indicators to track changes in the opioid epidemic as it evolves over time.

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## Introduction

In the last 15 years, the United States has witnessed a substantial increase in opioid use and overdose mortality stemming largely from opioid over-prescribing and the increased availability of synthetic opioids like fentanyl (Guy et al., 2017; Vadivelu et al., 2018; Beletsky & Davis, 2017). Approximately 47,000 people in the US died from opioid overdoses in 2018 (NIH, 2020). It is estimated that by 2025, the annual number of deaths attributable to opioids will be approximately 82,000, and a total of 700,000 people in the US will have died from opioid overdoses between 2016 and 2025 (Chen et al., 2019). The opioid epidemic has also led to an increase in injection drug use (IDU). IDU is related to a number of bloodborne infectious diseases, including acute hepatitis C virus (HCV) and human immunodeficiency virus (HIV), where rural areas, in particular, have carried a higher burden of disease (Mack & Ballesteros, 2017). One notable example of an IDU-related disease outbreak happened in Scott County, Indiana, where between 2014 and 2015, 181 individuals in a county of 24,000 were diagnosed with HIV, most as result of injection from the opioid oxycodone (Van Handel et al., 2016; Peters et al., 2016). Lack of access to sterile equipment, substance abuse treatment providers, and HIV services exacerbated this outbreak severely (Peters et al., 2016). Additionally, placing this outbreak in its ecologic context illustrates that the socioeconomic conditions in which the outbreak occurred can be found in numerous other counties in the US (Van Handel et al., 2016). Assessing the

vulnerability of these areas can be a crucial component in the prevention and mitigation of opioid overdose mortality and IDU-related disease.

Aggregate indices of vulnerability and resiliency have become a topic of interest to state departments of health. A vulnerability assessment can help ascertain which geographic regions, such as census tracts, ZIP code tabulation areas (ZCTA), or counties, might be vulnerable or resilient to potentially detrimental health outcomes, such as a natural disaster or a disease outbreak. The Centers for Disease Control and Prevention (CDC) provides its own Social Vulnerability Index (SVI) every two years based on 15 variables obtained from the US Census. These variables are sorted into four themes: socioeconomic status, household composition, race/ethnicity/language, and housing/transportation (CDC, 2020). In response to the Scott County outbreak, CDC created a nationwide vulnerability assessment for HIV and acute HCV related to IDU (Van Handel et al., 2016). Recent literature has also demonstrated the need for tailored vulnerability assessments not only to specific health outcomes but also specific geographies. State-specific assessments of drug-related vulnerability have been conducted in states such as Tennessee, Utah, South Dakota, Louisiana, and Georgia in an effort to include local indicators of drug-related harms and create targeted interventions on a local level (Rickles et al., 2017; Sharareh et al., 2020; Wesner et al., 2020; PRG, 2019; GADPH, 2018). However, methods of variable identification, data collection, data reduction, and vulnerability score ranking vary and are generally underdeveloped.

In this paper, we want to investigate opioid overdose mortality in Georgia counties by utilizing an expanded list of potential indicator variables, which include distinct local data, as compared to the national CDC vulnerability assessment of acute HCV. In Georgia, 63.2 opioid prescriptions were written per 100 persons in 2018, higher than the national rate of 51.4 (NIH, 2020). Although prescription-drug-related deaths in Georgia decreased in 2018, the number of heroin-related deaths continued to rise (NIH, 2020). Synthetic opioid deaths also decreased in 2018, but recent fears of synthetic opioid deaths have been renewed as a result of the COVID-19 pandemic's effects on the opioid drug market and the increasing availability of fentanyl (NIH, 2020; Wakeman, Green, & Rich, 2020). In addition to assessing county-level vulnerability opioid overdose mortality, we also want to investigate how this methodology can be applied for future vulnerability assessments.

## Methods

### Study Design

Based on previous research by CDC and other recent vulnerability assessments of acute HCV and opioid-related harms, we chose an ecologic study design for all 159 counties in Georgia (Van Handel et al.; Rickles et al., 2017).

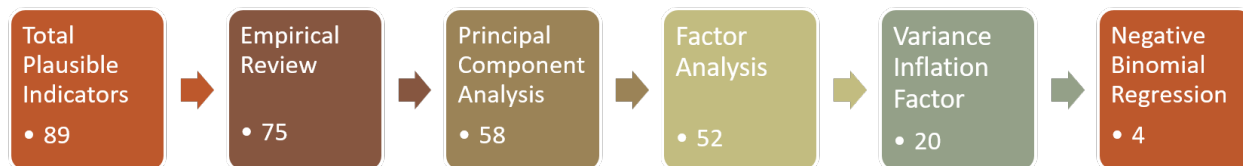


Figure 1: Possible indicators retained at each step of the dimension reduction and regression variable selection

### Primary Outcome

We were interested in assessing the vulnerability of counties in Georgia to opioid overdose mortality. This outcome was measured using counts of fatal opioid overdoses in 2019 from ICD-10 codes (T40.0, T40.1,

T40.2, T40.3, T40.4, and T40.6), which include prescription opioid pain relievers, opioids to treat addiction, heroin, opium, and synthetic opioids. These data are publicly available and provided by Georgia Department of Public Health’s Online Analytical Statistical Information System (OASIS).

## Potential Indicator Variables

Possible county-level indicators of opioid overdose mortality were based largely on previously identified indicators related to IDU listed by CDC, Rickles et al. (2017), and the Policy & Research Group (2019) (Van Handel et al., 2016). Data were collected from 9 sources: the 2019 American Community Survey (ACS), CDC, County Health Rankings, Georgia Department of Public Health (GADPH), Centers for Medicare and Medicaid Services (CMS), AIDSvU, Drug Enforcement Agency (DEA), Substance Abuse and Mental Health Services (SAMHSA), and Association of Religion Data Archives (ARDA), referenced in the first table of the addendum. The study sample consisted of 159 aggregate estimates from the residential populations in each county. A total of 89 variables were collected. All data cleaning, data reduction, and analyses were conducted using R version 4.0.4 (R Core Team, 2021). Estimates for population and income were log transformed. Rates were calculated per 100,000 persons using ACS population estimates, with the exception of teen births, injury-related deaths, church adherence, and opioid dispensing, in which no calculations were needed. National provider information was used to indicate access to provider services. However, county information was not provided in these data, only addresses and ZIP codes. To convert these addresses to counties, we subset the data by state and geocoded provider practice location addresses using three methods. First, if a ZIP code tabulation area (ZCTA) was contained fully within a county, the county would be assigned directly to that zip code using ArcMap. The remaining data were then geocoded sequentially through R packages “tidygeocoder,” which utilizes US Census data, and “ggmap,” which utilizes Google maps, respectively. That is, if an address was not geocoded through “tidygeocoder,” then it would be geocoded through “ggmap.” The geocode coverage was around 99%, and geographic coordinates were converted to counties. Provider addresses that were geocoded as outside Georgia were removed.

Small counts of disease, usually less than 5 cases, are often censored to protect the identity of individuals in counties with smaller populations. However, zeroes are reported when there are no cases in a county. In these incidences, we assumed the number of censored cases to be uniform from 1 to 4 and imputed these values with 2.5, the mean. We did this for counts of chlamydia, gonorrhea, syphilis, and HIV. In the case of Neonatal Abstinence Syndrome, counts less than 10 were censored and zeroes reported, so we imputed these values with 5.5.

## Dimension Reduction

Given the high-dimensional nature of the data with 159 counties and 89 possible indicators, traditional multivariate regression techniques of variable selection might not be adequate in helping us select a parsimonious model with meaningful predictors. It would be useful to utilize other dimension reduction techniques in addressing relatedness and multicollinearity among plausible predictors. Following the example of previous vulnerability assessment literature, we conducted a stepwise dimension reduction process consisting of an empirical review, principal component analysis, and factor analysis (Rickles et al., 2017; Sharareh et al., 2020; PRG, 2019). After an empirical review, we removed 11 variables that were too similar to the outcome (for example, heroin deaths). We also removed county name, Federal Information Processing Standards (FIPS) codes, and opioid overdose mortality, our outcome. We then conducted a principal component analysis, which determined that 69% of the data variance was explained in the first 5 principal components (see Figure 2). Seventeen variables were removed which had no absolute component correlation greater than 0.5 in any of the 5 components (Jolliffe, 2002). We then conducted minimum residual factor analysis on the remaining predictors. Parallel analysis determined that 5 factors, accounting for 80% of the total variance of the remaining data, were sufficient in representing the remaining possible indicators (see Figure 3). Similarly, 6 variables were removed which had no absolute factor loading greater than 0.5. Both PCA and factor analysis were performed using the R package “psych.”

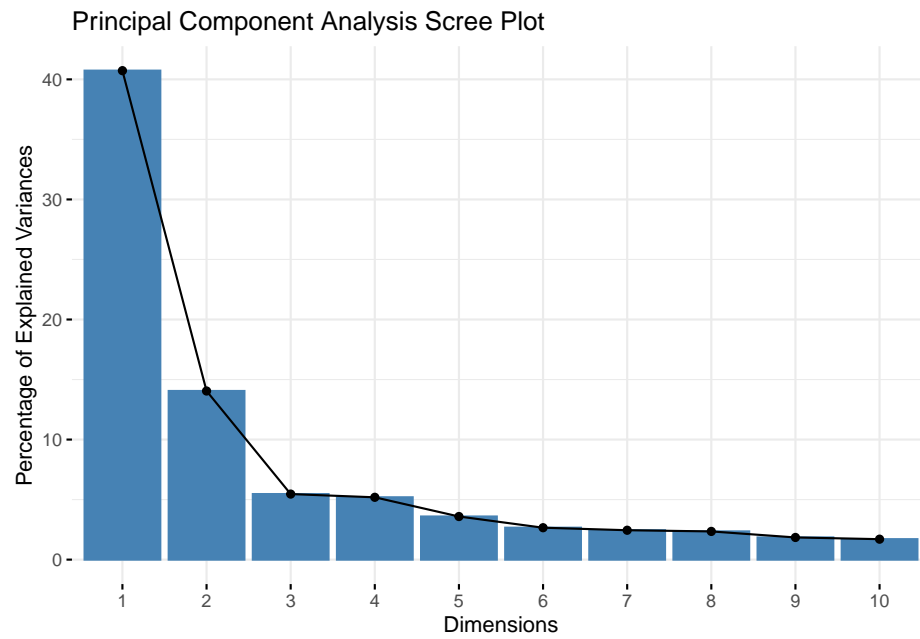


Figure 2: Principal Component Analysis scree plot showing percentage of variance explained in each component

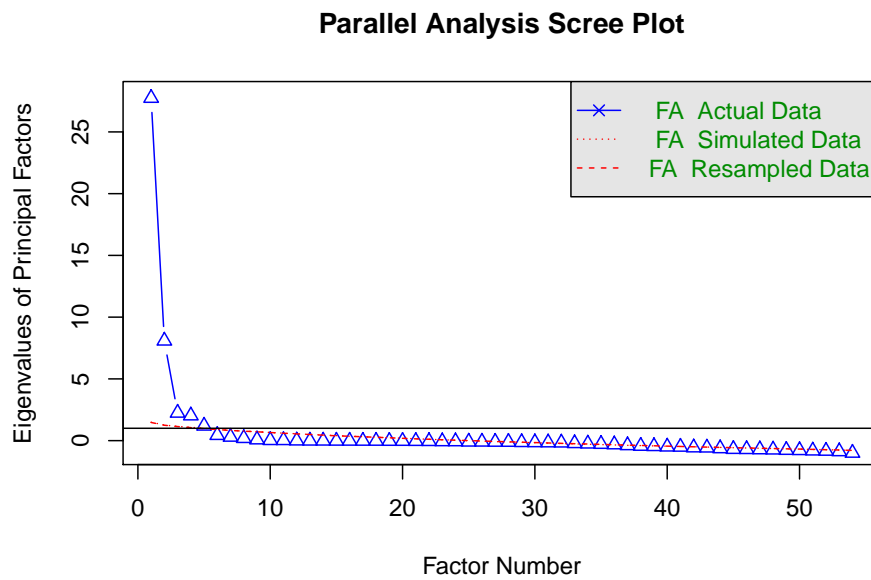


Figure 3: Factor Analysis scree plot for correlation matrix and the eigenvalues for each factor

Table 1: Negative Binomial regression results

	Coefficient	Std. error	P-Value
Intercept	-9.6482226	0.1912	<0.001
Percentage of population non-Hispanic White	0.0082634	0.0025	<0.001
Rate of ER and inpatient visits related to all drugs per 100000	-0.0043837	0.0009	<0.001
Rate of ER and inpatient visits realated to all opioids per 100000	0.0139810	0.0026	<0.001
Rate of specialty care providers per 100000	0.0183558	0.0063	<0.01

### Statistical Modeling and Multicollinearity

Since opioid overdose mortality was indicated with count data, we were interested in fitting a generalized linear model (GLM) with a Poisson distribution, log link, and an offset for log population. However, we conducted a deviance test to examine goodness of fit ( $p = 0.049$ ), which suggested that there was a lack of fit under a Poisson assumption and, therefore, overdispersion in the data. A GLM with a negative binomial distribution, a log link, and an offset for log population was fit instead to address this overdispersion.

We then identified multicollinearity among plausible predictors by calculating the generalized variance-inflation factors (GVIF) for each variable in our model and removing the variable that had the highest GVIF greater than 10. We repeated this process after each variable was removed and recalculated the GVIFs. Thirty-two variables were removed that had GVIFs greater than 10, leaving us with 20 possible predictors. Lastly, stepwise, bidirectional variable selection was performed using Akaike information criteria (AIC).

### Vulnerability Score Ranking

The vulnerability index for opioid overdose mortality was created by multiplying the coefficients from the significant predictors in our final model, with the exception of the intercept, by the respective county values. We then ranked the counties from lowest to highest and transformed the scores so that they took on values from 0 to 1, where 0 indicated lowest vulnerability and 1 indicated highest vulnerability. This was done to increase the interpretability of the scores and make the index more comparable with other known vulnerability indices, such as CDC’s Social Vulnerability Index.

## Results

Our final model yielded 4 significant predictors of opioid overdose mortality: the percentage of population that is non-Hispanic white, rate of ER and inpatient visits related to all drugs per 100,000 persons, rate of ER and inpatient visits related to all opioids per 100,000 persons, and the rate of specialty care providers per 100,000 persons. The regression coefficients and their significance can be found in Table 1.

Although vulnerable counties are found all over the state, we find that they are often clumped together or adjacent to other vulnerable counties (Figure 4). Of the top quintile of vulnerable counties, 17 counties are located in north Georgia (including Catoosa, Dade, and Walker County, counties which comprise the Georgia side of the Chattanooga metropolitan statistical area), 6 in central Georgia (including Richmond County, which contains Augusta, and Harris County, which includes part of the Columbus metropolitan statistical area), and 8 in south Georgia.

Only 2 of the top 31 counties in CDC’s Social Vulnerability Index overlapped with our index, namely, Clay and Candler County (Figure 5). An assessment of county-level vulnerability to opioid overdose was also conducted by GADPH in 2018. While overall regions of vulnerability tended to overlap with this previous assessment, for example in north and southeast Georgia, only 9 of the top 31 counties identified by the GADPH assessment overlapped with our assessment: Bartow, Brantley, Fannin, Pierce, Towns, Haralson, Lincoln, Ware, and Gilmer County.



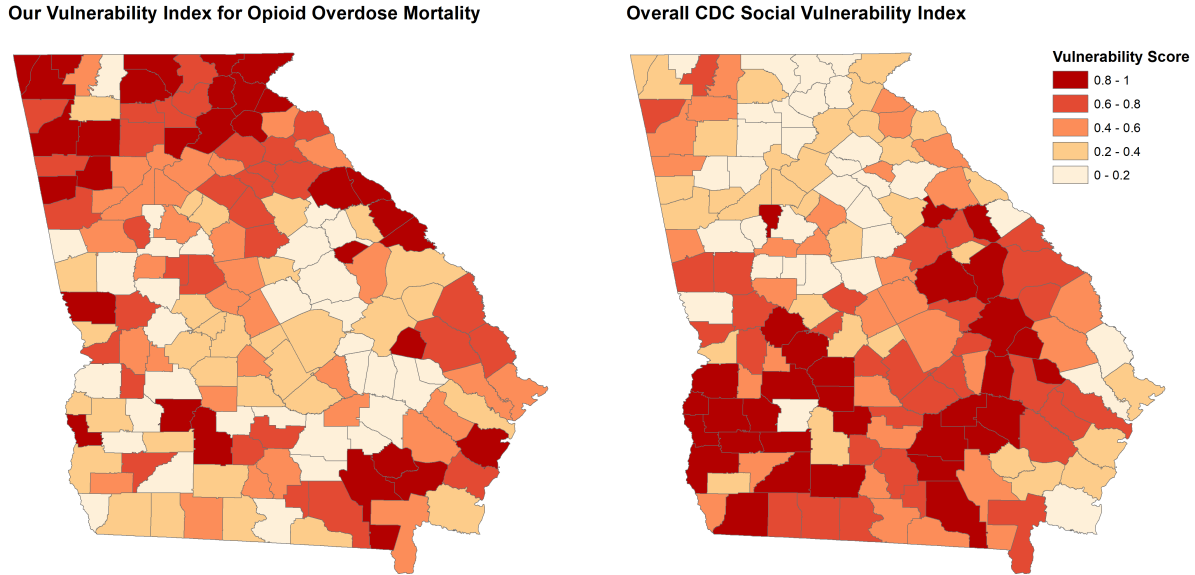


Figure 5: A comparison of our vulnerability index for opioid overdose mortality and CDC’s Overall Social Vulnerability Index in 2018 for Georgia by county

## Discussion

This assessment showcases the importance of adapting and refining current vulnerability index methodology for specific outcomes and specific geographies. Solely using CDC’s Social Vulnerability Index as a reference for opioid overdose mortality, for example, yields a very different map of vulnerability, especially in north Georgia. Like previous literature, we have expanded on the method that Van Handel et al. (2016) developed to identify vulnerable counties in the US to rapid dissemination of HIV and HCV by localizing the geography of interest to the state level, widening the scope of possible indicators, and, additionally, altering the outcome of interest to opioid overdose mortality. As a result, we have identified counties in Georgia that may be more susceptible to fatal opioid overdoses.

The indicators in our final model (Table 1) can be helpful in understanding, predicting, and preventing opioid overdose mortality in Georgia. Opioid-related harms, especially in Georgia, have been more severe in predominantly white communities, which our model reflects (KFF, 2019; Vadivelu et al., 2018). Local rates of ER and inpatient drug-related visits, specifically those attributable to opioids, have potential on an administrative level to predict changes in fatal opioid overdoses. One serious limitation of ER data is that rural counties often lack emergency care providers, forcing individuals to seek emergency care in adjacent counties or to forgo emergency care altogether. This could lead to underrepresenting the vulnerability of rural counties, which are generally more at risk for opioid-related harms (Mack & Ballesteros, 2017). Generally, interpretation of these model coefficients needs to be done with caution. It is counterintuitive to assume that the count of opioid overdose mortality decreases as the rate of ER and inpatient visits related to all drugs in a county increases or as the rate of specialty care providers increases, as our model suggests. As Sharareh et al. (2020) noted, county-level coefficients may instead be proxy indicators for other relationships that were either not available, not measurable, or not included in the model. Additional research into these indicators and their relationship to opioid overdose deaths would be helpful in better characterizing vulnerability to opioid overdose mortality.

This vulnerability information can be utilized on a state and local level to reduce vulnerability to fatal opioid overdoses by expanding prevention and mitigation resources in identified at-risk counties. This includes expanding access to emergency care and substance use care and ensuring adequate access to medication-assisted treatment, such as buprenorphine/naloxone to prevent an overdose and naloxone in the case of an overdose. Additionally, the opioid epidemic has the potential to change rapidly. Although the number of

opioids prescribed in Georgia is decreasing, the number of opioid-related deaths is increasing, and newfound fears have arisen regarding COVID-19's effect on illicit opioid markets, particularly the possible increased accessibility of stronger synthetic opioids (Wakeman, Green, & Rich, 2020).

The results of this vulnerability assessment are only generalizable to Georgia in 2019, but the method used in this paper, in tandem with the methods of similar vulnerability literature, is applicable to future development of aggregate indices of vulnerability and resiliency, especially within the context of administrative data. Employing the methods used here to create routine assessments of county vulnerability to fatal opioid overdoses would also help us ascertain whether vulnerability indicators and vulnerable counties in Georgia remain consistent over time or change as the epidemic evolves. Moreover, all data obtained and utilized in this paper were publicly available. Anyone could use the plausible indicators that were gathered to further investigate fatal opioid overdoses in Georgia. Additionally, future researchers could alter the list of plausible predictors to investigate other vulnerability outcomes on a county level in Georgia, like vulnerability to tornado deaths. In the future, we may also want to consider spatial autocorrelation among counties, especially since it appears that vulnerable counties might cluster together. Alternate approaches could also be explored regarding dimension reduction techniques, such as penalized regression and random forest.

## Limitations

When comparing the results of this paper to the assessment done by GADPH in 2018, some considerations should first be noted. First, vulnerability was assessed at two different time periods in the opioid epidemic. The GADPH assessment in 2018 used data from 2016-2017, whereas most of our data was from 2019. Second, the methods in the GADPH assessment were not explicit but were inspired by Van Handel et al. (2016). This makes a comparison of the methods used between assessments difficult, and our work should be viewed as a separate assessment rather than a reassessment. Third, one of the primary outcomes for the GADPH assessment was nonfatal opioid overdoses rather than fatal opioid overdoses, the primary outcome in this assessment.

It should be noted that the vulnerable counties surrounding Chattanooga, Augusta, and Columbus are similar in that they all lie on the border of other states: Tennessee, South Carolina, and Alabama, respectively. In future research, it might be beneficial to include the surrounding border counties to create a more holistic view of vulnerability around these metropolitan statistical areas and better understand how they affect counties in Georgia. The difficulty in this approach is that administrative data are often provided on a state-by-state basis and may be different across states.

Additionally, opioid overdose mortality was mapped using the residency of the deceased individuals rather than the location of death. This is useful within the context of investigating vulnerability through the lens of socioeconomic conditions; however, this might not accurately characterize opioid use, and using death location might help answer different questions regarding the opioid epidemic in Georgia.

Other limitations include imputation. Although chlamydia, gonorrhea, syphilis, HIV, and Neonatal Abstinence Syndrome were not included in the final model, there are limitations in imputing censored low counts, which may have contributed to their elimination during the data reduction process.

Lastly, we were unable to obtain acute HCV data for Georgia counties. This data would have been helpful in understanding vulnerability to IDU-related infectious diseases in Georgia and helpful in comparing identified indicators to previous vulnerability assessment literature, most of which focused on acute HCV.

## Conclusion

We assessed the vulnerability of opioid overdose mortality in Georgia and identified counties that might be more at risk for fatal overdoses. Vulnerable counties were identified all over the state, however, we found groupings in north and southeast Georgia. We also identified vulnerable counties around Augusta, Chattanooga, and Columbus. Identified indicators combined demographic characteristics, provider access, and medical information, but further research is needed to understand the effects of these indicators. Finally, our work highlights the importance of tailored vulnerability assessments by geography and outcome.



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## Addendum

### Total possible indicators and their descriptive statistics

	Overall (N=159)
<b>Population size</b>	
Mean (SD)	65433.000 (141727.090)
Range	1611.000 - 1036200.000
<b>Per capita income</b>	
Mean (SD)	24484.434 (5873.352)
Range	14005.000 - 47163.000
<b>Gini coefficient</b>	
Mean (SD)	0.462 (0.037)
Range	0.376 - 0.601
<b>Percentage uninsured</b>	
Mean (SD)	14.118 (3.477)
Range	6.200 - 23.600
<b>Percentage with no high school diploma</b>	
Mean (SD)	20.694 (9.529)
Range	3.200 - 60.000
<b>Percentage living in poverty</b>	
Mean (SD)	19.645 (6.725)
Range	5.500 - 40.200
<b>Percentage of population unemployed</b>	
Mean (SD)	6.331 (2.392)
Range	1.900 - 13.400
<b>Teen birth rate per 1000</b>	
Mean (SD)	22.472 (35.042)
Range	0.000 - 228.000
<b>Percentage of population with a disability</b>	
Mean (SD)	15.935 (3.320)
Range	6.900 - 25.200
<b>Percentage of female-headed households</b>	
Mean (SD)	3529.824 (7620.563)
Range	127.000 - 56277.000
<b>Percentage of population never married</b>	
Mean (SD)	31.153 (6.833)
Range	17.100 - 54.100
<b>Percentage of population non-Hispanic White</b>	
Mean (SD)	62.042 (17.317)
Range	10.100 - 94.900
<b>Percentage with vehicle access</b>	
Mean (SD)	0.972 (0.017)
Range	0.921 - 0.997
<b>Population aged 18-29</b>	
Mean (SD)	20432.981 (45556.168)
Range	352.000 - 320820.000
<b>Percentage of population aged 18-29</b>	
Mean (SD)	29.141 (3.493)

	Overall (N=159)
Range	16.441 - 39.676
<b>Per capita income (log)</b>	
Mean (SD)	10.080 (0.222)
Range	9.547 - 10.761
<b>Population Density</b>	
Mean (SD)	207.858 (418.213)
Range	8.278 - 2800.370
<b>Population Density (log)</b>	
Mean (SD)	4.391 (1.247)
Range	2.114 - 7.938
<b>Population Decline 2011-2019</b>	
Population Declined	70 (44.0%)
Population Increased	89 (56.0%)
<b>Housing units</b>	
Mean (SD)	26940.107 (57411.745)
Range	1020.000 - 471836.000
<b>Occupied housing units</b>	
Mean (SD)	23640.239 (51957.541)
Range	593.000 - 410576.000
<b>Vacant housing units</b>	
Mean (SD)	3299.868 (6006.672)
Range	310.000 - 61260.000
<b>Mobile homes</b>	
Mean (SD)	2420.604 (1414.333)
Range	413.000 - 7794.000
<b>Percentage of mobile homes</b>	
Mean (SD)	22.074 (12.474)
Range	0.500 - 58.000
<b>Home without phone service</b>	
Mean (SD)	458.811 (902.622)
Range	0.000 - 8458.000
<b>Percentage of homes without phone service</b>	
Mean (SD)	2.358 (1.751)
Range	0.000 - 15.800
<b>Crowded housing units</b>	
Mean (SD)	1.923 (1.186)
Range	0.175 - 6.655
<b>Premature deaths</b>	
N-Miss	5
Mean (SD)	853.383 (1368.836)
Range	41.000 - 10496.000
<b>Years of potential life lost</b>	
N-Miss	5
Mean (SD)	9392.429 (2000.318)
Range	3873.000 - 17028.000
<b>Adults reporting poor/fair health</b>	
Mean (SD)	20.075 (3.776)
Range	13.000 - 33.000
<b>Poor physical health days</b>	
Mean (SD)	3.847 (0.440)
Range	2.800 - 5.100
<b>Poor mental health days</b>	

	Overall (N=159)
Mean (SD)	4.065 (0.296)
Range	3.200 - 5.000
<b>Injury-related deaths</b>	
N-Miss	1
Mean (SD)	211.892 (392.086)
Range	12.000 - 3255.000
<b>Rate of injury-related deaths</b>	
N-Miss	1
Mean (SD)	77.975 (17.419)
Range	34.000 - 137.000
<b>Percentage of adults who smoke</b>	
Mean (SD)	18.296 (2.277)
Range	13.000 - 24.000
<b>Total church adherents</b>	
Mean (SD)	30970.918 (70340.643)
Range	617.000 - 605740.000
<b>Rate of church adherence per 1000</b>	
Mean (SD)	483.413 (134.816)
Range	90.799 - 828.580
<b>Buprenorphine Providers</b>	
Mean (SD)	5.434 (14.186)
Range	0.000 - 119.000
<b>Per capita buprenorphine providers</b>	
Mean (SD)	5.898 (7.807)
Range	0.000 - 59.000
<b>High Intensity Drug Trafficking Area</b>	
HIDTA County	14 (8.8%)
Not HIDTA County	145 (91.2%)
<b>Deaths related to all drugs</b>	
Mean (SD)	8.742 (18.322)
Range	0.000 - 142.000
<b>Rate of death related to all drugs</b>	
Mean (SD)	13.363 (9.815)
Range	0.000 - 53.648
<b>Deaths related to all opioids</b>	
Mean (SD)	5.365 (11.130)
Range	0.000 - 74.000
<b>Rate of death related to all opioids</b>	
Mean (SD)	7.973 (8.219)
Range	0.000 - 38.314
<b>Deaths related to heroin</b>	
Mean (SD)	1.862 (4.637)
Range	0.000 - 36.000
<b>Rate of death related to heroin</b>	
Mean (SD)	2.319 (4.112)
Range	0.000 - 33.378
<b>Deaths related to synthetic opioids other than methadone</b>	
Mean (SD)	2.616 (6.103)
Range	0.000 - 44.000
<b>Rate of death related to synthetic opioids other than methadone</b>	
Mean (SD)	3.317 (4.139)
Range	0.000 - 21.166

	Overall (N=159)
<b>Deaths related to natural, semi-synthetic, and synthetic opioids</b>	
Mean (SD)	4.428 (9.031)
Range	0.000 - 61.000
<b>Rate of death related to natural, semi-synthetic, and synthetic opioids</b>	
Mean (SD)	6.618 (7.382)
Range	0.000 - 38.314
<b>Deaths related to methadone</b>	
Mean (SD)	0.365 (0.767)
Range	0.000 - 4.000
<b>Rate of death related to methadone</b>	
Mean (SD)	0.723 (2.031)
Range	0.000 - 13.178
<b>ER and inpatient visits related to all drugs</b>	
Mean (SD)	156.516 (304.139)
Range	1.000 - 2532.000
<b>Rate of ER and inpatient visits related to all drugs</b>	
Mean (SD)	254.585 (81.708)
Range	18.522 - 448.092
<b>ER and inpatient visits realated to all opioids</b>	
Mean (SD)	32.031 (60.162)
Range	0.000 - 439.000
<b>Rate of ER and inpatient visits realated to all opioids</b>	
Mean (SD)	51.056 (26.474)
Range	0.000 - 138.202
<b>ER and inpatient visits related to heroin</b>	
Mean (SD)	32.031 (60.162)
Range	0.000 - 439.000
<b>Rate of ER and inpatient visits related to heroin</b>	
Mean (SD)	51.056 (26.474)
Range	0.000 - 138.202
<b>Chlamydia and Gonorrhea Cases</b>	
Mean (SD)	183.557 (214.856)
Range	2.500 - 936.000
<b>Rate Chlamydia and Gonorrhea</b>	
Mean (SD)	604.895 (365.198)
Range	0.241 - 1774.064
<b>Syphilis (all forms) Cases</b>	
Mean (SD)	19.333 (52.778)
Range	0.000 - 353.000
<b>Rate of Syphilis (all forms)</b>	
Mean (SD)	27.945 (33.784)
Range	0.000 - 327.937
<b>Rate of HIV Incidence per 100000</b>	
Mean (SD)	18.796 (20.547)
Range	0.000 - 155.183
<b>New HIV cases</b>	
Mean (SD)	15.818 (59.346)
Range	0.000 - 605.000
<b>HIV prevalent cases</b>	
Mean (SD)	322.692 (1418.328)
Range	2.500 - 15127.000
<b>Rate of HIV Prevalence per 100000</b>	

	Overall (N=159)
Mean (SD)	290.675 (209.789)
Range	46.555 - 1459.853
<b>HIV cases related to injection drug use</b>	
Mean (SD)	23.296 (89.088)
Range	0.000 - 964.000
<b>Number of PrEP users</b>	
Mean (SD)	30.560 (140.164)
Range	2.500 - 1520.000
<b>Rate of PrEP users per 100000</b>	
Mean (SD)	30.065 (20.520)
Range	7.330 - 155.183
<b>Opioid dispensing rate per 100</b>	
N-Miss	2
Mean (SD)	48.631 (37.298)
Range	0.000 - 178.000
<b>% patient days with overlapping opioid prescriptions</b>	
Mean (SD)	16.227 (3.295)
Range	8.900 - 34.700
<b>% patient days with overlapping opioid and benzodiazepine prescriptions</b>	
Mean (SD)	14.307 (3.799)
Range	7.400 - 30.200
<b>Mental health services</b>	
Mean (SD)	172.107 (544.946)
Range	0.000 - 4222.000
<b>Primary care providers</b>	
Mean (SD)	107.742 (297.543)
Range	0.000 - 2699.000
<b>Specialty Care providers</b>	
Mean (SD)	4.906 (20.203)
Range	0.000 - 210.000
<b>Urgent care providers</b>	
Mean (SD)	1.522 (4.147)
Range	0.000 - 36.000
<b>Substance use services</b>	
Mean (SD)	0.604 (2.056)
Range	0.000 - 17.000
<b>Count of Neonatal Abstinence Syndrome</b>	
Mean (SD)	10.711 (17.486)
Range	0.000 - 143.000
<b>NCHS Urban-Rural status</b>	
Large Metropolitan	31 (19.5%)
Medium Metropolitan	15 (9.4%)
Micropolitan	28 (17.6%)
Noncore	57 (35.8%)
Small Metropolitan	28 (17.6%)
<b>Rate of HIV cases related to injection drug use per 100000</b>	
Mean (SD)	28.454 (24.209)
Range	0.000 - 155.183
<b>Rate of mental health services per 100000</b>	
Mean (SD)	133.189 (129.527)
Range	0.000 - 717.793
<b>Rate of primary care providers per 100000</b>	

	Overall (N=159)
Mean (SD)	122.075 (84.330)
Range	0.000 - 530.587
<b>Rate of specialty Care providers per 100000</b>	
Mean (SD)	2.582 (5.104)
Range	0.000 - 24.771
<b>Rate of urgent care providers per 100000</b>	
Mean (SD)	1.564 (2.871)
Range	0.000 - 22.080
<b>Rate of substance use services per 100000</b>	
Mean (SD)	0.607 (1.646)
Range	0.000 - 10.319
<b>Count of violent crimes</b>	
Mean (SD)	228.774 (703.307)
Range	0.000 - 7179.000
<b>Count of property crimes</b>	
Mean (SD)	1822.176 (5015.320)
Range	0.000 - 45984.000