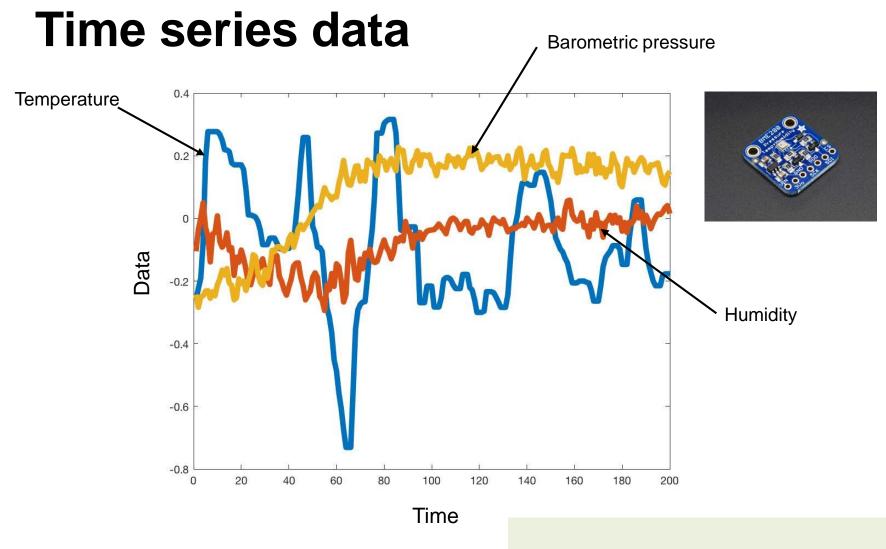
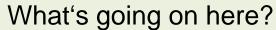
Modeling time series with hidden Markov models

Advanced Machine learning 2017

Nadia Figueroa, Jose Medina and Aude Billard



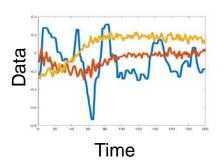




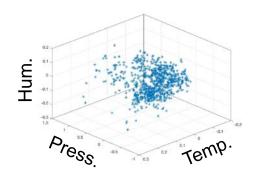


Time series data

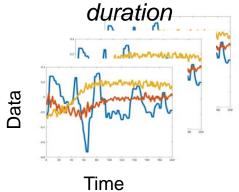
What's the problem setting?



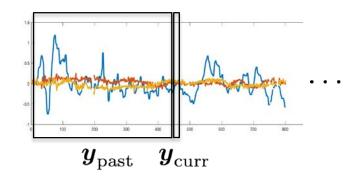
We don't care about time ...



We have several trajectories with identical duration



We have unstructured trajectory(ies)!



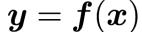
Explicit time dependency

$$egin{aligned} oldsymbol{y} &= oldsymbol{f}(oldsymbol{x},t) \ oldsymbol{y} &= oldsymbol{f}(t) \end{aligned}$$

Consider dependency on the past

$$oldsymbol{y}_{ ext{curr}} = oldsymbol{f}(oldsymbol{y}_{ ext{past}})$$

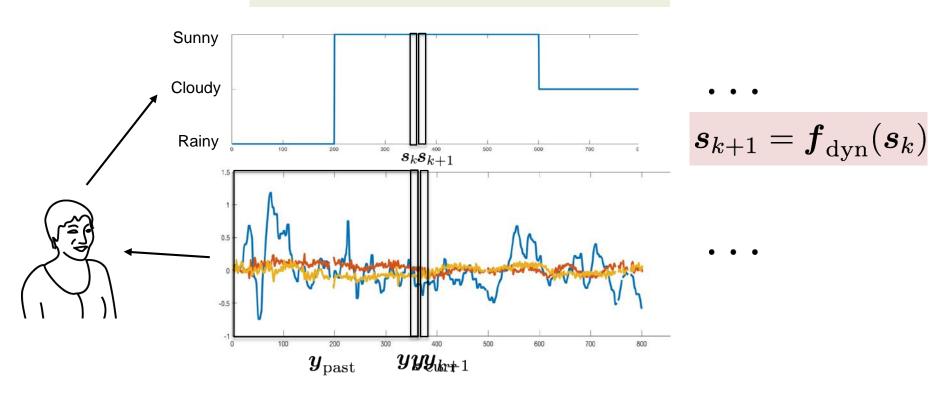
Too complex!





Unstructured time series data

How to simplify this problem?



Consider dependency on the past



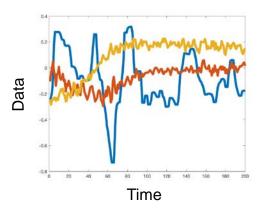
Markov assumption



Outline

First part (10:15 – 11:00):

- Recap on Markov chains
- Hidden Markov Model (HMM)
 - Recognition of time series
 - ML Parameter estimation



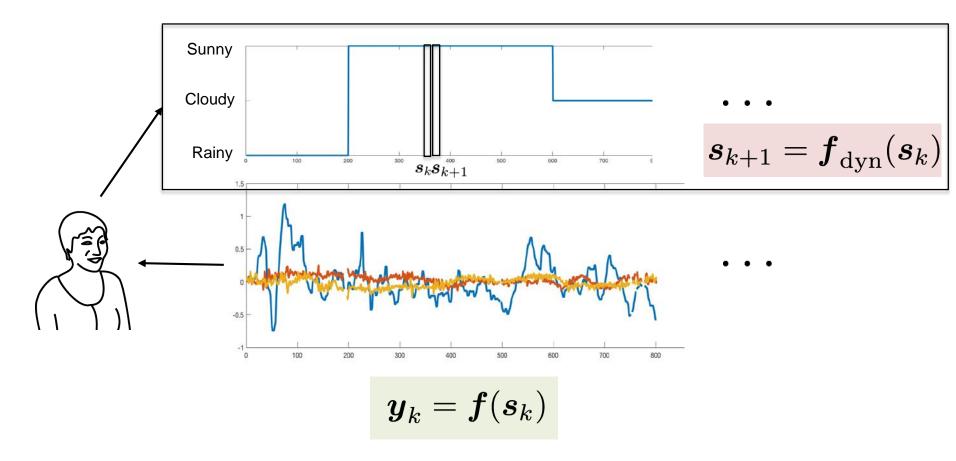
<u>Second part</u> (11:15 – 12:00):

- Time series segmentation
- Bayesian nonparametrics for HMMs

https://github.com/epfl-lasa/ML_toolbox

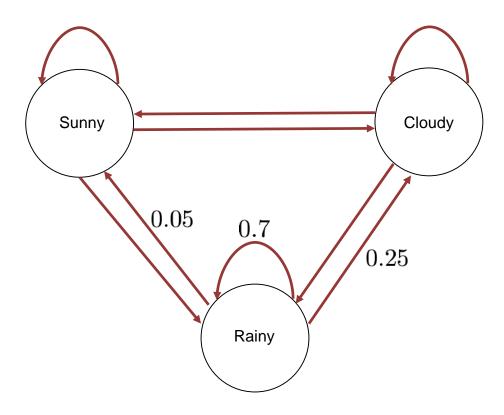


Outline first part





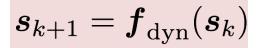
Markov chains



$$P(s_{k+1} = \text{Sunny}|s_k = \text{Rainy}) = 0.05$$

$$P(s_{k+1} = \text{Rainy}|s_k = \text{Rainy}) = 0.7$$

$$P(s_{k+1} = \text{Cloudy}|s_k = \text{Rainy}) = 0.25$$



$$\lambda =$$

Transition

matrix Rainy Sunny

Sunny

 $\begin{bmatrix} 0.9 & 0.08 \end{bmatrix}$

 $0.25 \quad 0.7 \quad 0.05$

0.02

Cloudy

Rainy

0.05

0.25

0.7

Initial probabilities

Cloudy Sunny

Rainy

 $[0.33 \quad 0.33]$

0.33



Likelihood of a Markov chain



$$O = \text{ Sunny } \longrightarrow \text{ Sunny } \longrightarrow \text{ Cloudy }$$

 $\lambda =$

$P(O|\lambda)$?

$$P(\mathbf{s}_0 = \text{Sunny}, \mathbf{s}_1 = \text{Sunny}, \mathbf{s}_2 = \text{Cloudy}) =$$
 $P(\mathbf{s}_0 = \text{Sunny})P(\mathbf{s}_1 = \text{Sunny}|\mathbf{s}_0 = \text{Sunny})$
 $P(\mathbf{s}_2 = \text{Cloudy}|\mathbf{s}_1 = \text{Sunny}) = 0.33 \cdot 0.9 \cdot 0.08$

		<u>Iransition</u>		
	Sunny	matrix	Rainy	
Sunny	0.9	0.08	0.02	
Cloudy	0.25	0.7	0.05	
Rainy	0.05	0.25	0.7	

$P(O|\lambda) = P(\boldsymbol{s}_0) \prod_{k=1}^{T} P(\boldsymbol{s}_k | \boldsymbol{s}_{k-1})$

Initial probabilities

Sunny Cloudy Rainy $\begin{bmatrix} 0.33 & 0.33 & 0.33 \end{bmatrix}$

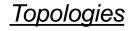


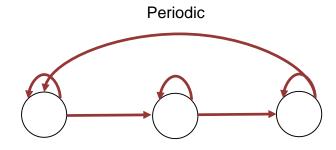
Learning Markov chains

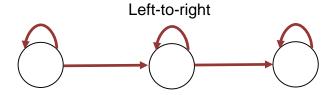


$$O = \operatorname{Sunny} \, o \, \operatorname{Sunny} \, o \, \operatorname{Cloudy}$$

$$\max_{\lambda} \log P(O|\lambda)$$

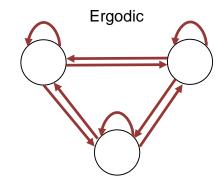






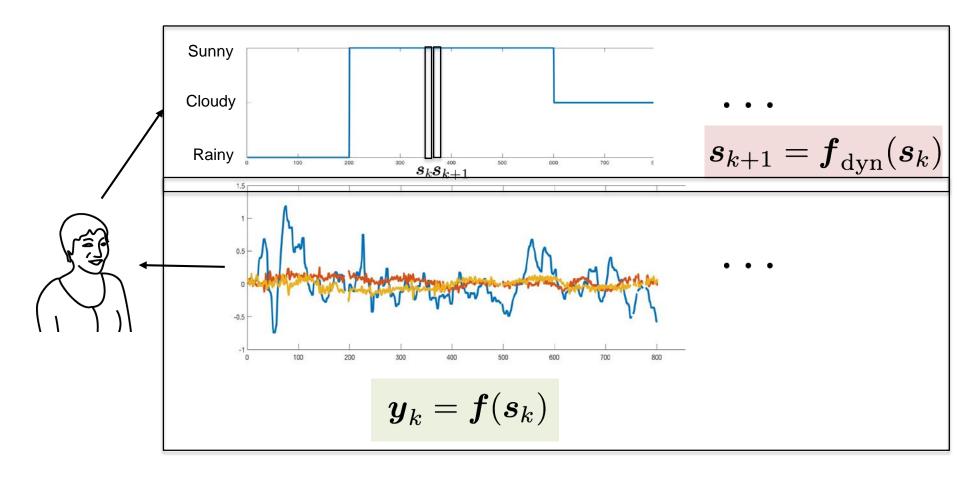
$$\hat{P}(s_{k+1} = \text{Sunny}|s_k = \text{Rainy}) =$$

 $\frac{\text{number of times we transition Rainy} \rightarrow \text{Sunny}}{\text{number of times we observe Sunny}}$



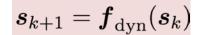


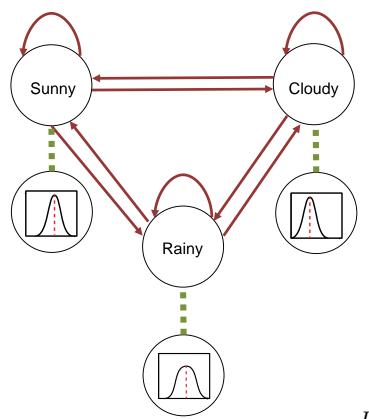
Outline first part





Hidden Markov model





Transition

	Sunny	<u>matrix</u>	Rainy
Sunny	$\boxed{0.9}$	0.08	0.02
Cloudy	0.25	0.7	0.05
	0.05	0.25	0.7

Initial probabilities

Sunny	Cloudy	Rainy
[0.33]	0.33	0.33

$$oldsymbol{y}_k = oldsymbol{f}(oldsymbol{s}_k)$$

$$\begin{split} P(\boldsymbol{y}_k|\boldsymbol{s}_k &= \text{Sunny}) = \mathcal{N}(\boldsymbol{y}_k; \boldsymbol{\mu}_{\text{Sunny}}, \boldsymbol{\Sigma}_{\text{Sunny}}) \\ P(\boldsymbol{y}_k|\boldsymbol{s}_k &= \text{Cloudy}) = \mathcal{N}(\boldsymbol{y}_k; \boldsymbol{\mu}_{\text{Cloudy}}, \boldsymbol{\Sigma}_{\text{Cloudy}}) \\ P(\boldsymbol{y}_k|\boldsymbol{s}_k &= \text{Rainy}) = \mathcal{N}(\boldsymbol{y}_k; \boldsymbol{\mu}_{\text{Rainy}}, \boldsymbol{\Sigma}_{\text{Rainy}}) \end{split}$$



Likelihood of an HMM

$$oldsymbol{s}_{k+1} = oldsymbol{f}_{ ext{dyn}}(oldsymbol{s}_k)$$



$$O = {}^{15}$$

$$oldsymbol{y}_1 \cdots oldsymbol{y}_T$$

Transition

	Sunny	<u>matrix</u>	Rainy
Sunny	$\boxed{0.9}$	0.08	0.02
Cloudy	0.25	0.7	0.05
Rainy	0.05	0.25	0.7

$$P(O|\lambda) = P(oldsymbol{y}_1 \cdots oldsymbol{y}_T | \lambda) =$$

$$\sum_{oldsymbol{s}_1 \dots T} P(oldsymbol{y}_1 \cdots oldsymbol{y}_T, oldsymbol{s}_1 \cdots oldsymbol{s}_T | \lambda)$$

$$\mathbb{D} = \{\text{Sunny, Cloudy, Rainy}\}$$

Initial probabilities

Sunny	Cloudy	Rainy
[0.33]	0.33	0.33

$$\alpha(\boldsymbol{s}_k) = P(\boldsymbol{y}_1 \cdots \boldsymbol{y}_k, \boldsymbol{s}_k)$$

$$P(O|\lambda) = \sum_{\boldsymbol{s}_T} \alpha(\boldsymbol{s}_T)$$

$$oldsymbol{y}_k = oldsymbol{f}(oldsymbol{s}_k)$$

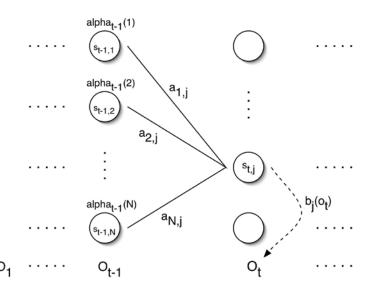
$$\begin{split} P(\boldsymbol{y}_k|\boldsymbol{s}_k &= \text{Sunny}) = \mathcal{N}(\boldsymbol{y}_k; \boldsymbol{\mu}_{\text{Sunny}}, \boldsymbol{\Sigma}_{\text{Sunny}}) \\ P(\boldsymbol{y}_k|\boldsymbol{s}_k &= \text{Cloudy}) = \mathcal{N}(\boldsymbol{y}_k; \boldsymbol{\mu}_{\text{Cloudy}}, \boldsymbol{\Sigma}_{\text{Cloudy}}) \\ P(\boldsymbol{y}_k|\boldsymbol{s}_k &= \text{Rainy}) = \mathcal{N}(\boldsymbol{y}_k; \boldsymbol{\mu}_{\text{Rainy}}, \boldsymbol{\Sigma}_{\text{Rainy}}) \end{split}$$



Likelihood of an HMM

$$oldsymbol{s}_{k+1} = oldsymbol{f}_{ ext{dyn}}(oldsymbol{s}_k)$$

$$\alpha(\boldsymbol{s}_k) = \sum_{\boldsymbol{s}_{k-1}} \alpha(\boldsymbol{s}_{k-1}) P(\boldsymbol{y}_k | \boldsymbol{s}_k) P(\boldsymbol{s}_k | \boldsymbol{s}_{k-1})$$



Transition

	Sunny	<u>matnix</u>	Rainy
Sunny	$\boxed{0.9}$	0.08	0.02
Cloudy	0.25	0.7	0.05
Rainy	0.05	0.25	0.7

Initial probabilities

Sunny Cloudy Rainy
$$\begin{bmatrix} 0.33 & 0.33 & 0.33 \end{bmatrix}$$

$$oldsymbol{y}_k = oldsymbol{f}(oldsymbol{s}_k)$$

$$\alpha(\boldsymbol{s}_k) = P(\boldsymbol{y}_1 \cdots \boldsymbol{y}_k, \boldsymbol{s}_k)$$

$$P(O|\lambda) = \sum_{\boldsymbol{s}_T} \alpha(\boldsymbol{s}_T)$$

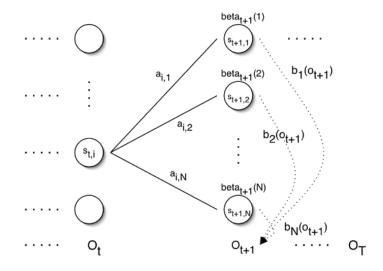
$$egin{aligned} &P(oldsymbol{y}_k|oldsymbol{s}_k = ext{Sunny}) = \mathcal{N}(oldsymbol{y}_k; oldsymbol{\mu}_{ ext{Sunny}}, oldsymbol{\Sigma}_{ ext{Sunny}}) \ &P(oldsymbol{y}_k|oldsymbol{s}_k = ext{Cloudy}) = \mathcal{N}(oldsymbol{y}_k; oldsymbol{\mu}_{ ext{Cloudy}}, oldsymbol{\Sigma}_{ ext{Cloudy}}) \ &P(oldsymbol{y}_k|oldsymbol{s}_k = ext{Rainy}) = \mathcal{N}(oldsymbol{y}_k; oldsymbol{\mu}_{ ext{Rainy}}, oldsymbol{\Sigma}_{ ext{Rainy}}) \end{aligned}$$



Likelihood of an HMM

$$oldsymbol{s}_{k+1} = oldsymbol{f}_{ ext{dyn}}(oldsymbol{s}_k)$$

$$\beta(\boldsymbol{s}_k) = \sum_{\boldsymbol{s}_{k+1}} \beta(\boldsymbol{s}_{k+1}) P(\boldsymbol{y}_{k+1} | \boldsymbol{s}_{k+1}) P(\boldsymbol{s}_{k+1} | \boldsymbol{s}_k)$$



Transition

	Sunny	<u>matnix</u>	Rainy
Sunny	$\boxed{0.9}$	0.08	0.02
Cloudy	0.25	0.7	0.05
Rainy	0.05	0.25	0.7

Initial probabilities

Sunny	Cloudy	Rainy
[0.33]	0.33	0.33

$$oldsymbol{y}_k = oldsymbol{f}(oldsymbol{s}_k)$$

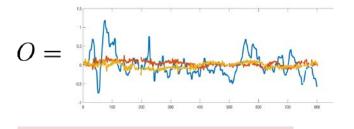
Backward variable

$$\beta(\boldsymbol{s}_k) = P(\boldsymbol{y}_{k+1} \cdots \boldsymbol{y}_T | \boldsymbol{s}_k)$$

$$P(O|\lambda) = \sum_{\boldsymbol{s}_1} \beta(\boldsymbol{s}_1) P(\boldsymbol{s}_1)$$

$$P(oldsymbol{y}_k | oldsymbol{s}_k = ext{Sunny}) = \mathcal{N}(oldsymbol{y}_k; oldsymbol{\mu}_{ ext{Sunny}}, oldsymbol{\Sigma}_{ ext{Sunny}})$$
 $P(oldsymbol{y}_k | oldsymbol{s}_k = ext{Cloudy}) = \mathcal{N}(oldsymbol{y}_k; oldsymbol{\mu}_{ ext{Cloudy}}, oldsymbol{\Sigma}_{ ext{Cloudy}})$
 $P(oldsymbol{y}_k | oldsymbol{s}_k = ext{Rainy}) = \mathcal{N}(oldsymbol{y}_k; oldsymbol{\mu}_{ ext{Rainy}}, oldsymbol{\Sigma}_{ ext{Rainy}})$





$$oldsymbol{y}_1 \cdots oldsymbol{y}_T$$

 $\max_{\lambda} \log P(O|\lambda)$

Baum-Welch algorithm

(Expectation-Maximization for HMMs)

- Iterative solution
- Converges to local minimum

Starting from an initial $\lambda \quad \text{find a} \lambda' \quad \text{such that } P(O|\lambda') \geq P(O|\lambda)$

- E-step: Given an observation sequence and a model, find the probabilities of the states to have produced those observations.
- M-step: Given the output of the E-step, update the model parameters to better fit the observations.



$$O = \begin{bmatrix} 0.5 & 0.5$$

E-step:

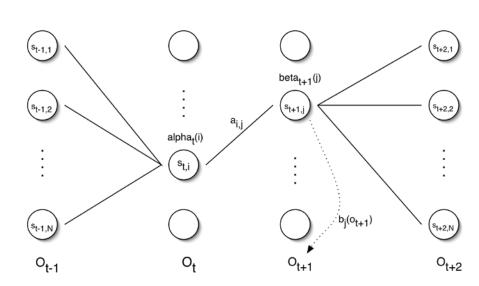
$$\xi_k(i,j) = \frac{\alpha(s_k = i)P(s_k = i, s_{k+1} = j)P(y_k + 1|s_{k+1} = j)\beta(s_{k+1} = j)}{\sum_{s_k} \sum_{s_{k+1}} \alpha(s_k)P(s_k, s_{k+1})P(y_k + 1|s_{k+1})\beta(s_{k+1})}$$

Probability of being in state i at time k and transition to state j

$$\xi_k(i,j) = P(\boldsymbol{s}_k = i, \boldsymbol{s}_k = j | O, \lambda)$$

Probability of being in state i at time k

$$\gamma_k(i) = \sum_{j \in \boldsymbol{s}_{k+1}} \xi_k(i, j)$$





$$oldsymbol{y}_1 \cdots oldsymbol{y}_T$$

M-step:

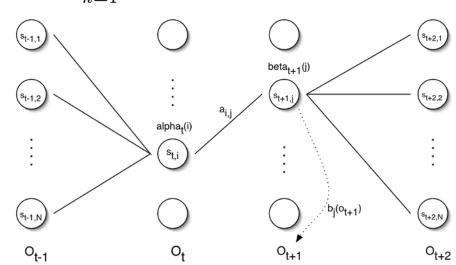
$$\hat{P}(\boldsymbol{s}_{k+1} = \text{Sunny} | \boldsymbol{s}_{k} = \text{Rainy}) = \sum_{\substack{k=1 \ \text{Ntimber of likes we transition Rainy } \rightarrow \text{Sunny} \\ \text{number of times we being k-1}} \sum_{k=1}^{I} \xi_{k}(i,j)$$

Probability of being in state i at time k and transition to state j

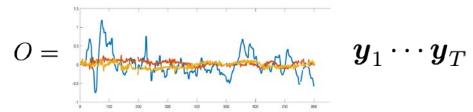
$$\xi_k(i,j) = P(\boldsymbol{s}_k = i, \boldsymbol{s}_k = j | O, \lambda)$$

Probability of being in state i at time k

$$\gamma_k(i) = \sum_{j \in \boldsymbol{s}_{k+1}} \xi_k(i,j)$$







- HMM is a parametric technique (Fixed number of states, fixed topology)
- → Heuristics to determining the optimal number of states

X: dataset; N: number of datapoints; K: number of free parameters

- Aikaike Information Criterion: AIC= $-2 \ln L + 2K$
- Bayesian Information Criterion: $BIC = -2 \ln L + K \ln (N)$

L: maximum likelihood of the model given K parameters

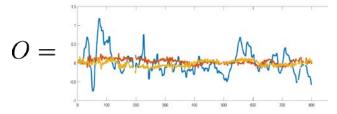
Choosing AIC versus BIC depends on the application:

→ Is the purpose of the analysis to make predictions, or to decide which model best represents reality?

AIC may have better predictive ability than BIC, but BIC finds a computationally more efficient solution.



Applications of HMMs



State estimation:

What is the most probable state/state sequence of the system?

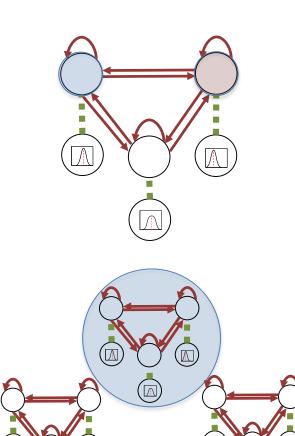
Prediction:

What are the most probable next observations/state of the system?

Model selection:

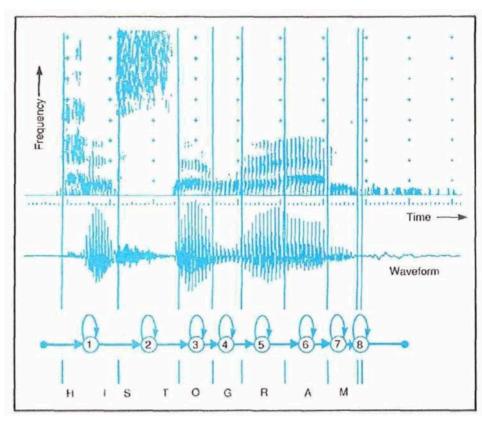
What is the most likely model that represents these observations?





Speech recognition:

- Left-to-right model
- States are phonemes
- Observations in frequency domain

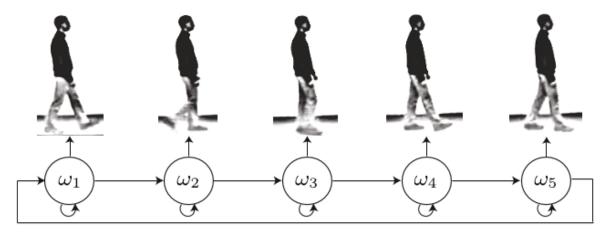


D.B. Paul., Speech Recognition Using Hidden Markov Models, The Lincoln laboratory journal, 1990



Motion prediction:

- Periodic model
- Observations are observed joints
- Simulate/predict walking patterns

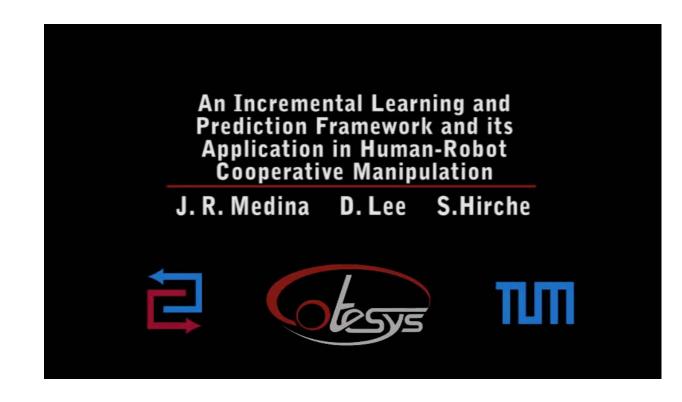


Karg, Michelle, et al. "Human movement analysis: Extension of the f-statistic to time series using hmm." Systems, Man, and Cybernetics (SMC), 2013 IEEE International Conference on. IEEE, 2013.



Motion prediction:

- Left-to-right models
- Autonomous segmentation
- Recognition + prediction





Motion prediction:

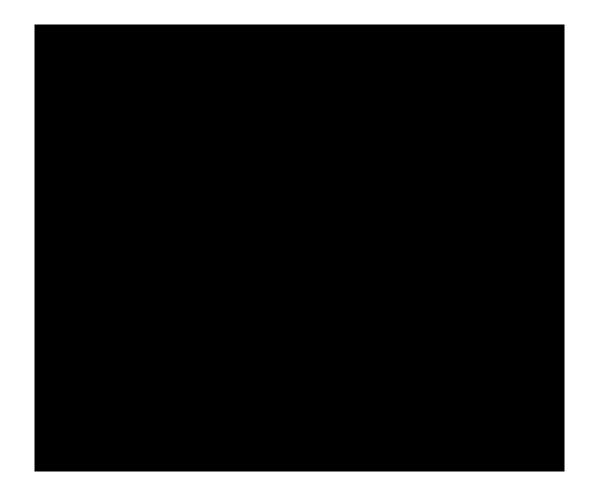
- Left-to-right models
- Autonomous segmentation
- Recognition + prediction





Motion prediction:

- Left-to-right model
- Each state is a dynamical system



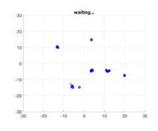


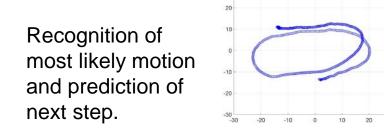
Motion *recognition*:

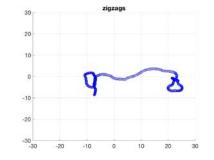
most likely motion and prediction of next step.

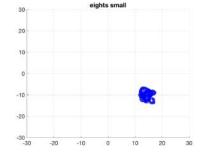
Toy training set

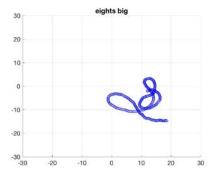
- 1 player
- 7 actions
- 1 Hidden Markov model per action

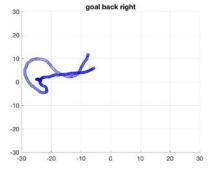


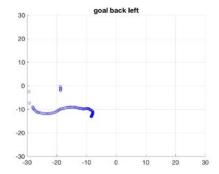












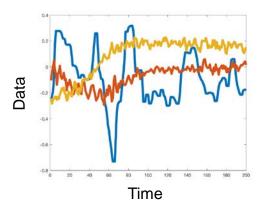


MATLAB demo

Outline

First part (10:15 – 11:00):

- Recap on Markov chains
- Hidden Markov Model (HMM)
 - Recognition of time series
 - ML Parameter estimation



<u>Second part</u> (11:15 – 12:00):

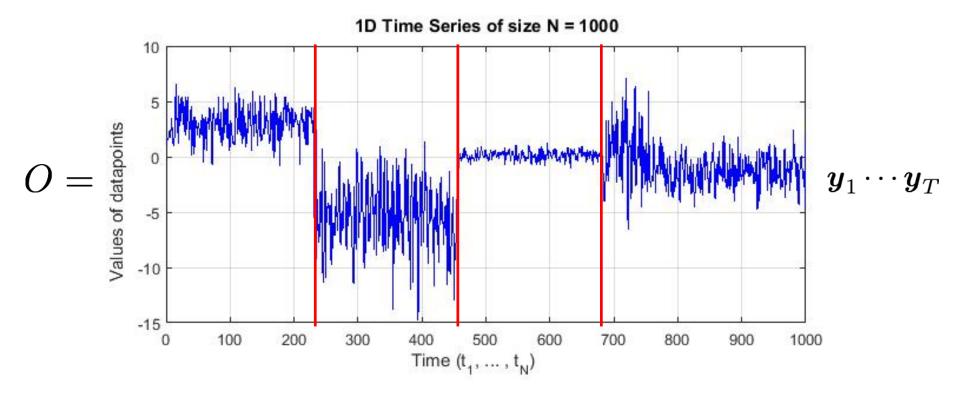
- Time series segmentation
- Bayesian non-parametrics for HMMs

https://github.com/epfl-lasa/ML_toolbox



Time series Segmentation

Times-series = Sequence of discrete segments

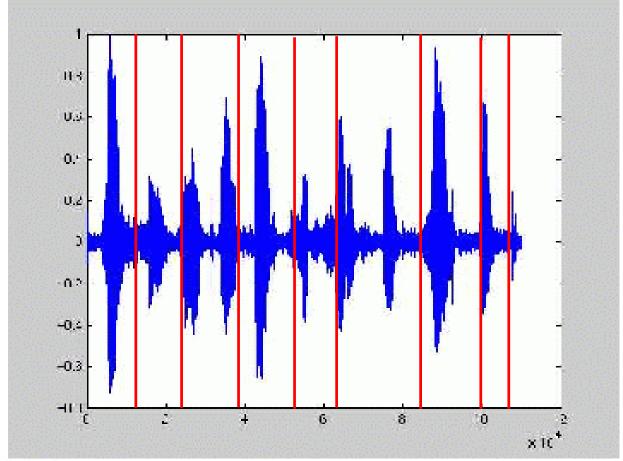


Why is this an important problem?



Segmentation of Speech Signals

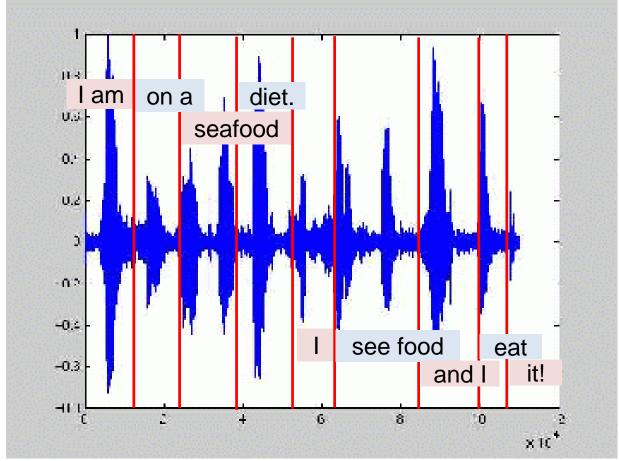
Segmenting a continuous speech signal into sets of distinct words.





Segmentation of Speech Signals

Segmenting a continuous speech signal into sets of distinct words.







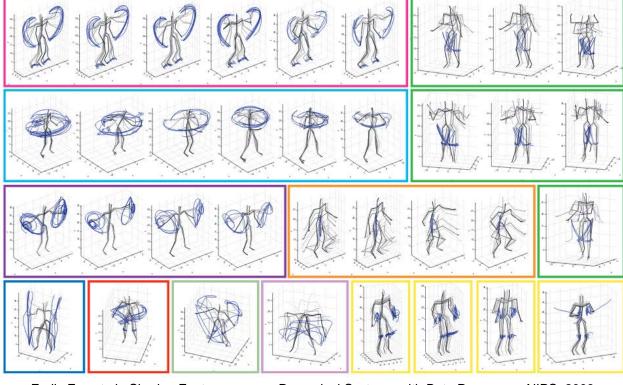
Segmentation of Human Motion Data

Segmention of Continuous Motion Capture data from exercise routines into motion categories

Jumping Jacks

Arm

Circles



Knee

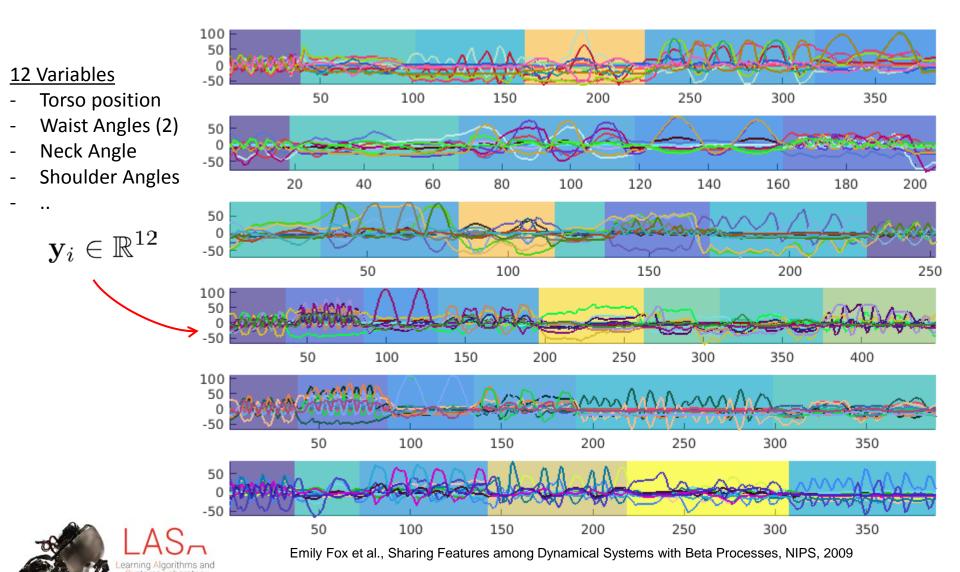
Raises

Squats

Emily Fox et al., Sharing Features among Dynamical Systems with Beta Processes, NIPS, 2009



Segmentation in Human Motion Data



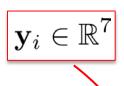
ÉCOLE POLYTECHNIQUE FÉDÉRALE DE LAUSANNE

Segmentation in Robotics

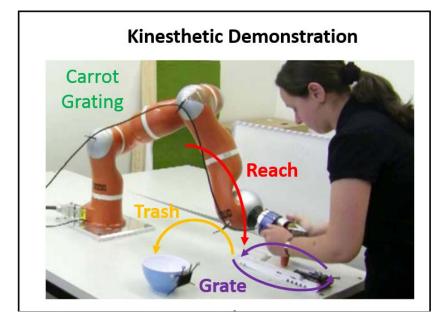
Learning Complex Sequential Tasks from Demonstration

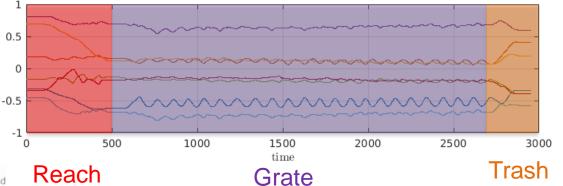
7 Variables

- Position x, y, z
- Orientation q_i,q_j,q_k,q_w









 $\boldsymbol{y}_1 \cdots \boldsymbol{y}_T$

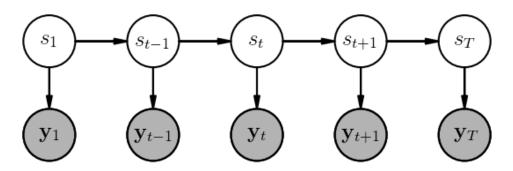
Assumptions:

 The time-series has been generated by a system that transitions between a set of hidden states:

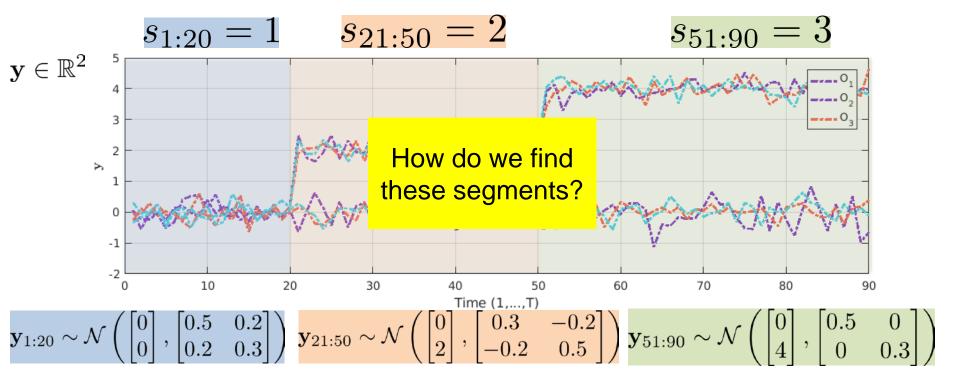
$$s \in \{1, \dots, K\}$$

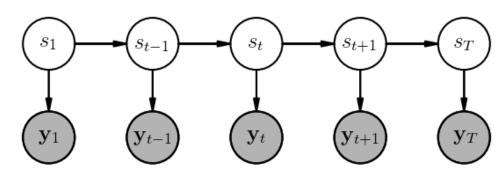
At each time step, a sample is drawn from an emission model associated to the current hidden state:

$$\mathbf{y}_t | s_t = k \quad \sim \mathcal{N}(\theta_{s_t}) \mathbf{z}$$
with $\theta_{s_t} = \{\mu_k, \Sigma_k\}$











Steps for Segmentation with HMM:

1. Learn the HMM parameters through Maximum Likelihood Estimate (MLE):

$$\lambda = \{\pi, A, \Theta\}$$
Initial State Transition Emission Model Probabilities Matrix Parameters

HMM Likelihood

$$P(O|\lambda) = P(\mathbf{y}_1 \cdots \mathbf{y}_T | \lambda) =$$

$$\sum_{\mathbf{s}_1 \dots T \in \mathbb{D}} P(\mathbf{y}_1 \cdots \mathbf{y}_T, \mathbf{s}_1 \cdots \mathbf{s}_T | \lambda)$$

$\max_{\lambda} \ \log P(O|\lambda)$

Baum-Welch algorithm

(Expectation-Maximization for HMMs)

- Iterative solution
- Converges to local minimum

Hyper-parameter:

Number of states possible K



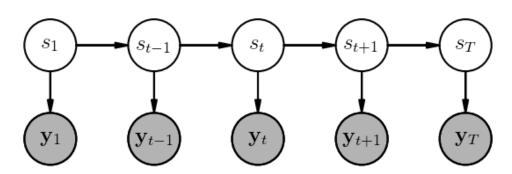
Steps for Segmentation with HMM:

2. Find the most probable sequence of states generating the observations through the **Viterbi algorithm**:

$$S^* = \arg\max_{S} \quad p(Y, S|\Theta)$$

HMM Joint Probability Distribution

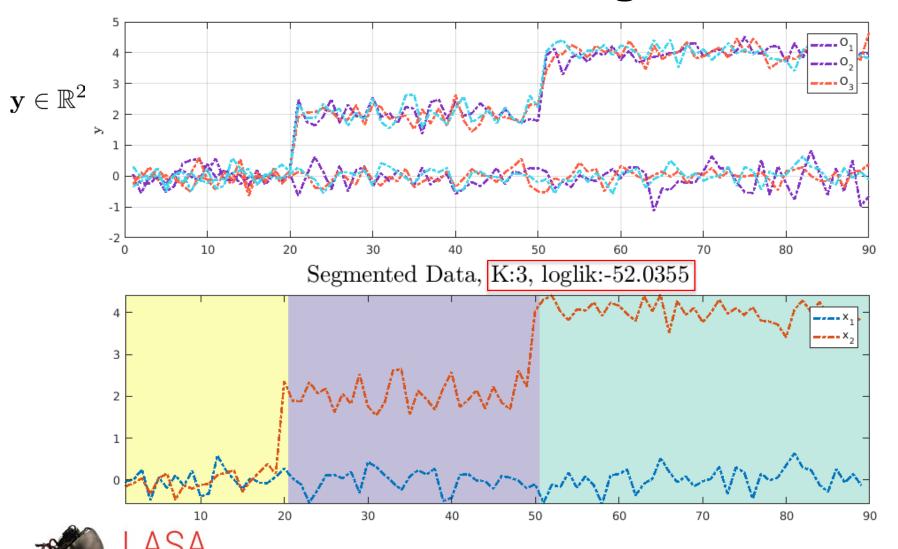
$$p(Y, S|\Theta) = p(s_1|\pi) \left[\prod_{t=2}^{T} p(s_t|s_{t-1}, A) \right] \prod_{t=1}^{T} p(\mathbf{y}_t|s_t, \Theta)$$







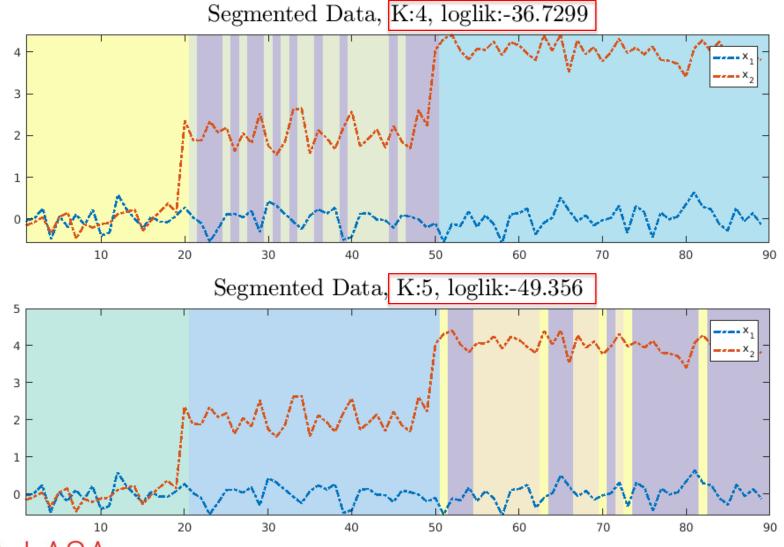
HMM for Time series Segmentation







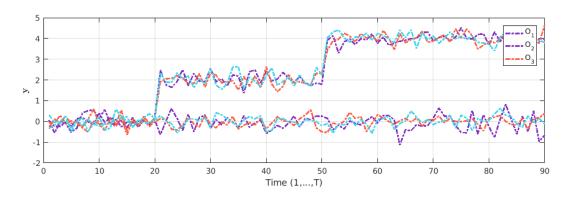
HMM for Time series Segmentation

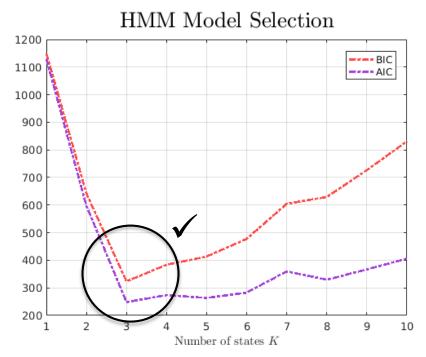






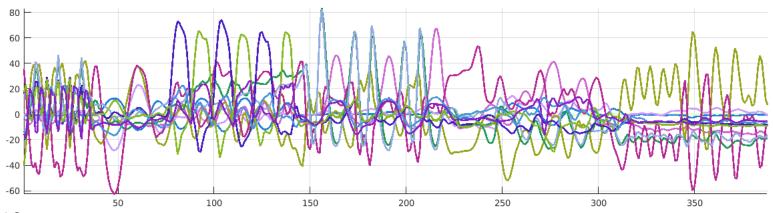
Model Selection for HMMs



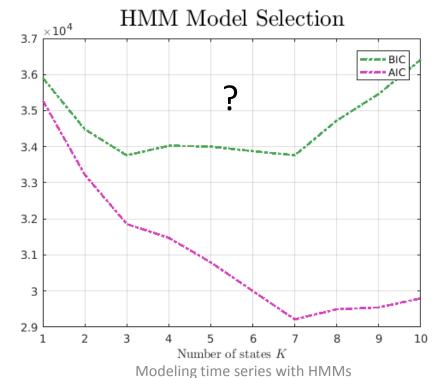




Model Selection for HMMs









Limitations of classical finite HMMs for Segmentation Undefined

Cardinality: Choice of hidden states is based on Model Selection heuristics, there is little understanding of the strengths and weaknesses of such methods in the Fixed Transition

xed Transilic Matrix

hidden states

Topology: We assume that all time series share the same set of emission models and switch among them in exactly the same manner [2].

[1] Emily Fox et al., An HDP-HMM for Systems with State Persistence, ICML, 2008

[2] Emily Fox et al., Sharing Features among Dynamical Systems with Beta Processes, NIPS, 2009

Solution: Bayesian Non-Parametrics



Bayesian Non-Parametrics

 Bayesian: Use Bayesian inference to estimate the parameters; i.e. priors on model parameters!

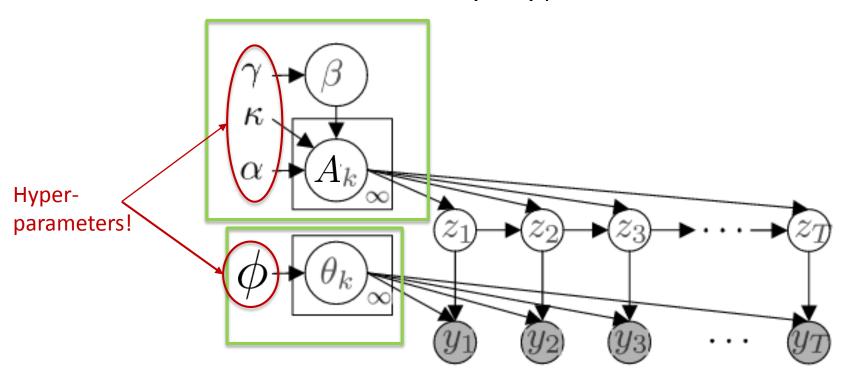
$$\lambda = \{\pi, A, \Theta\}$$

- **Non-parametric:** Does NOT mean methods with "no parameters", rather models whose complexity (# of states, # Gaussians) is inferred from the data.
 - 1. Number of parameters grows with sample size.
 - 2. Infinite-dimensional parameter space!



BNP for HMMs: <u>HDP-HMM</u>

Hierarchical Dirichlet Process (HDP) prior on the transition Matrix!



Normal Inverse Wishart (NIW) prior on emission parameters!

Emily Fox et al., An HDP-HMM for Systems with State Persistence, ICML, 2008



BNP for HMMs: HDP-HMM

© Cardinality

The Dirichlet Process (DP) is a prior distribution over distributions.

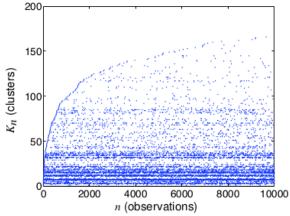
$$M(.) = \sum_{k=1}^{\infty} C_k \delta_{\Theta_k}(.)$$
 where $\sum_{k=1}^{\infty} C_k = 1$

• Used for clustering with infinite mixture models; i.e. instead of setting K in a GMM, the K is learned from data.

$$K_n = \#$$
 clusters in sample of size n

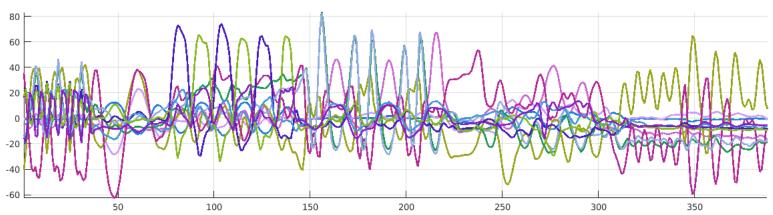
This only gives us an estimate of the K clusters! Cannot be use directly on transition matrix:

$$A \in \mathbb{R}^{K \times K}$$

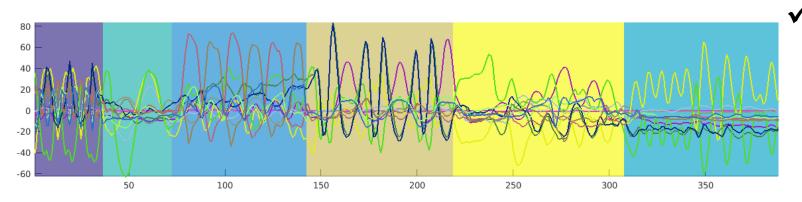


The Hierarchical Dirichlet Process (HDP) is a hierarchy of DPs!

Segmentation with HDP-HMM



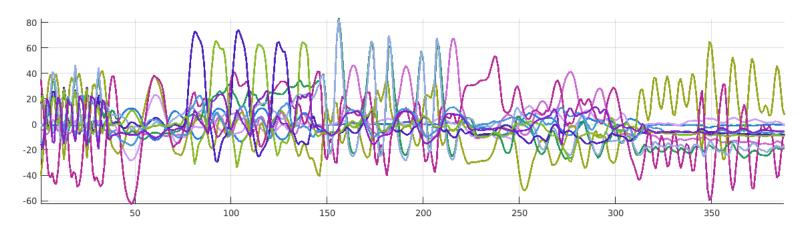
 $\mathbf{y}_i \in \mathbb{R}^{12}$



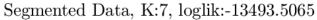
Emily Fox et al., An HDP-HMM for Systems with State Persistence, ICML, 2008

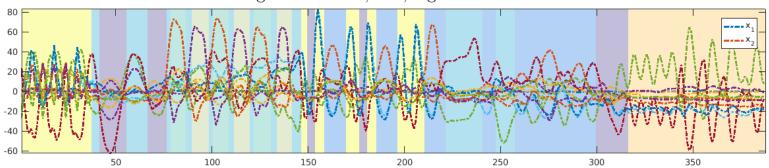


Compare to Model Selection with Classical HMMs



$$\mathbf{y}_i \in \mathbb{R}^{12}$$

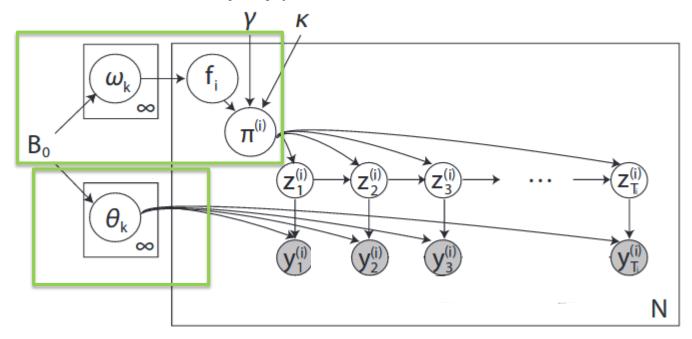






BNP for HMMs: BP-HMM

- © Cardinality
- © Topology
- The Beta Process (BP) prior on the transition Matrix!



Normal Inverse Wishart (NIW) prior on emission parameters!

Emily Fox et al., Sharing Features among Dynamical Systems with Beta Processes, NIPS, 2009



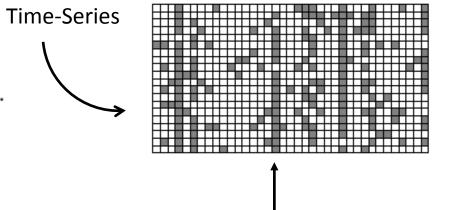
BNP for HMMs: BP-HMM

- © Cardinality
- © Topology
- The Beta Process (BP) prior on the transition Matrix!

Beta Process (BP)

Distribution on objects of the form

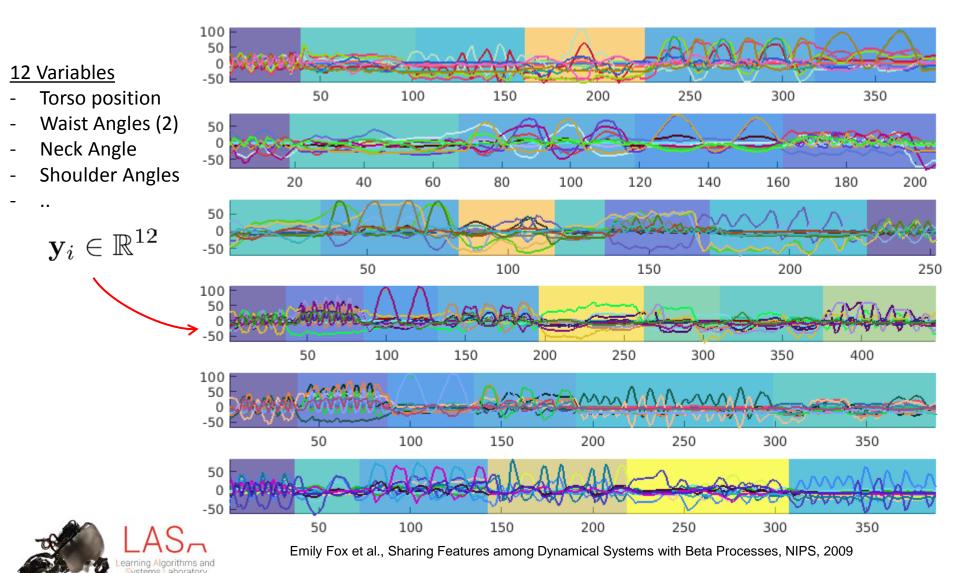
$$\theta = \sum_{k=1}^{\infty} w_k \delta_{\phi_k}$$
 with $w_k \in [0, 1]$.



Features (i.e. shared HMM States)



Segmentation in Human Motion Data



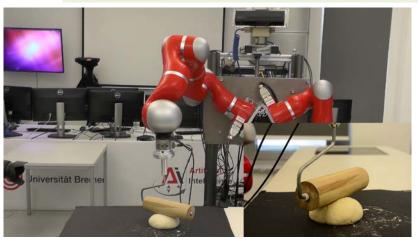
ÉCOLE POLYTECHNIQUE FÉDÉRALE DE LAUSANNE

Applications in Robotics





Learning Complex Sequential Tasks from Demonstration





Learning Complex Sequential Tasks from Demonstrations: Pizza Dough Rolling

Nadia Figueroa, Lucia Pais and Aude Billard





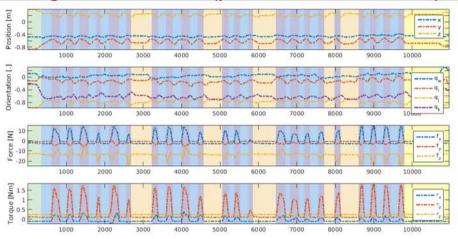




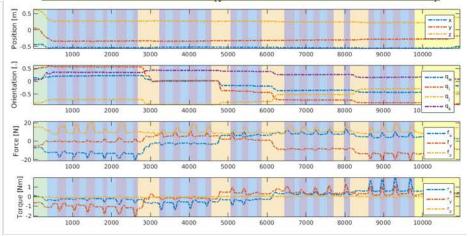


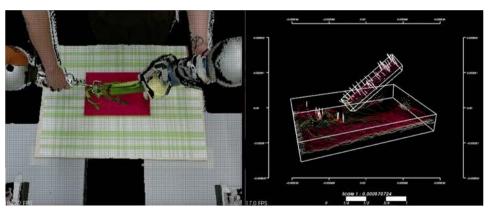


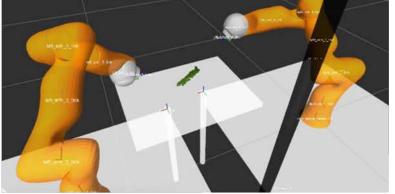
Right Arm EE data (pos, orientation, wrench)



Left Arm EE data (pos, orientation, wrench)







Object Features (mean, std of RGB Channels)

