Product rule: $P(X_{1:n}) = P(X_1) \prod_{i=2}^{n} P(X_i | X_{1:i-1}).$

Sum rule: $P(X,Y) = \sum_{y} P(X,Y=y)$.

Bayes rule: P(X|Y) = P(Y|X)P(X)/P(Y). Independence: $P_{XY} = P_X P_Y$.

Conditional independence: $P_{XY|Z} = P_{X|Z}P_{Y|Z}$.

Linearity of expectation: $\mathbb{E}_{x,y}[aX+bY] = a\mathbb{E}_x[X] + b\mathbb{E}_y[Y]$.

Expectation: $\mathbb{E}_p[f(X)] = \sum_{i=0}^n p(x)f(x)$ (don't forget p(x)). **Variance**: $Var[X] = \mathbb{E}[(X - \mathbb{E}[X])^2] = \mathbb{E}[X^2] - \mathbb{E}[X]^2$.

Linearity of variance:

Probability review

 $Var[aX+bY+c] = a^{2}Var[X]+b^{2}Var[Y]+2abCov(X,Y).$

Covariance: $Cov(X,Y) = \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])].$

Cum. dist. function: $\mathbb{P}(x \le t) = F(t)$, where *F* is CDF.

Matrix inversion: $\begin{bmatrix} a & b \\ c & d \end{bmatrix}^{-1} = \frac{1}{ad-bc} \begin{bmatrix} d & -b \\ -c & a \end{bmatrix}.$

Multivariate Gaussian:

$$\mathcal{N}(\mathbf{x};\boldsymbol{\mu},\boldsymbol{\Sigma}) = \frac{1}{\sqrt{(2\pi)^d |\boldsymbol{\Sigma}|}} \exp\left(-\frac{1}{2}(\mathbf{x}-\boldsymbol{\mu})^{\top} \boldsymbol{\Sigma}^{-1}(\mathbf{x}-\boldsymbol{\mu})\right).$$

any linear combination of the RVs is Gaussian.

Properties:

$$egin{aligned} X_A \sim & \mathcal{N}(oldsymbol{\mu}_A, oldsymbol{\Sigma}_{AA}) \ X_A \mid X_B \sim & \mathcal{N}(oldsymbol{\mu}_{A\mid B}, oldsymbol{\Sigma}_{A\mid B}) \ oldsymbol{\mu}_{A\mid B} = & oldsymbol{\mu}_A + oldsymbol{\Sigma}_{AB} oldsymbol{\Sigma}_{BB}^{-1} (oldsymbol{x}_B - oldsymbol{\mu}_B) \ oldsymbol{\Sigma}_{A\mid B} = & oldsymbol{\Sigma}_{AA} - oldsymbol{\Sigma}_{AB} oldsymbol{\Sigma}_{BB}^{-1} oldsymbol{\Sigma}_{BA} \ oldsymbol{MX} \sim & \mathcal{N} igg(oldsymbol{M} oldsymbol{\mu}, oldsymbol{M} oldsymbol{\Sigma} oldsymbol{M}^{ op} igg) \ X + X' \sim & \mathcal{N} ig(oldsymbol{\mu} + oldsymbol{\mu}', oldsymbol{\Sigma} + oldsymbol{\Sigma}' ig) \end{aligned}$$

If asked about conditional distribution of linear functions of Gaussians, notice that the functions can be jointly computed by matrix operation, which results in a Gaussian, or use LoV.

Kalman filters: Motion model updates the state $X_{t+1} = FX_t + \epsilon_t$. Sensor model computes observation $Y_t = HX_t + \eta_t$. $\epsilon_t \sim \mathcal{N}(0, \Sigma_x)$, $\eta_t \sim \mathcal{N}(0,\Sigma_y)$. $\epsilon_t \uparrow \Rightarrow K_{t+1} \uparrow$, $\eta_t \uparrow \Rightarrow K_{t+1} \downarrow$, $\Sigma_t \uparrow \Rightarrow K_{t+1} \uparrow$. Update: $X_{t+1} | y_{1:t+1} \sim \mathcal{N}(\mu_{t+1}, \Sigma_{t+1})$

$$egin{aligned} oldsymbol{\lambda}_{t+1} &| oldsymbol{y}_{1:t+1} \sim \mathcal{N} \left(oldsymbol{\mu}_{t+1} oldsymbol{\omega}_{t+1}
ight) \\ oldsymbol{\mu}_{t+1} &= oldsymbol{F} oldsymbol{\mu}_{t} + oldsymbol{K}_{t+1} - oldsymbol{H} oldsymbol{F} oldsymbol{\mu}_{t}
ight) \\ oldsymbol{\Sigma}_{t+1} &= oldsymbol{I} - oldsymbol{K}_{t+1} oldsymbol{H} oldsymbol{I} oldsymbol{F} oldsymbol{\Sigma}_{t} oldsymbol{F}^{ op} + oldsymbol{\Sigma}_{x}
ight) \end{aligned}$$

 $K_{t+1} = (F\Sigma_t F^\top + \Sigma_x) H^\top (H(F\Sigma_t F^\top + \Sigma_x) H^\top + \Sigma_y)^{-1}$

Entropy: $H[p] = \mathbb{E}_p[-\log p(x)]$.

d-Gaussian: $H[\mathcal{N}(\mu, \Sigma)] = \frac{d}{2}(1 + \log(2\pi)) + \frac{1}{2}\log|\Sigma|$.

1-Gaussian: $H[\mathcal{N}(\mu,\sigma^2)] = \frac{1}{2}\log(2\pi\sigma^2) + \frac{1}{2}$.

 $H[p,q] = H[p] + H[q|p]. H[X|Y] = \mathbb{E}_p \left| \log \frac{p(x,y)}{p(x)} \right|$ **KL-divergence**: $KL(q \parallel p) = \mathbb{E}_q \left[\log \frac{q(x)}{p(x)} \right] = \mathbb{E}_q \left[-\log \frac{p(x)}{q(x)} \right]$. Non-

negative (0 if p = q, ∞ if p(x) = 0 for x with q(x) > 0). Additional expected surprise when observing q samples while assuming p. **Mutual information**: $I(X;Y) = H[X] - H[X \mid Y]$. Symmetric: I(X;Y) = I(Y;X). Information never hurts: $I(X;Y) \ge 0$. **Jensen's inequality**: If f convex: $f(\mathbb{E}[X]) \leq \mathbb{E}[f(X)]$. MLE: $\hat{\theta} = \max_{\theta} p(y | X, \theta) = \max_{\theta} \sum_{i=1}^{n} \log p(y_i | x_i, \theta)$.

MAP: $\hat{\boldsymbol{\theta}} = \operatorname{amax}_{\boldsymbol{\theta}} p(\boldsymbol{\theta} | \boldsymbol{X}, \boldsymbol{y}) = \operatorname{amax}_{\boldsymbol{\theta}} \log p(\boldsymbol{\theta}) + \sum_{i=1}^{n} \log p(y_i | \boldsymbol{x}_i, \boldsymbol{\theta})$ **Bayesian learning:** Prior: $p(\theta)$. Likelihood: $p(y \mid X, \theta) =$ $\prod_{i=1}^n p(y_i \mid x_i, \theta). \text{ Posterior: } p(\theta \mid X, y) = \frac{1}{7} p(\theta) \prod_{i=1}^n p(y_i \mid x_i, \theta).$

 $Z = \int p(\theta) \prod_{i=1}^{n} p(y_i | x_i, \theta) d\theta$. Prediction: $p(y^* | x^*, X, y) = \int p(y^* | x^*, X, y) d\theta$ x^*,θ) $p(\theta|X,y)d\theta$. In general, intractable: GP, VI, and MCMC solve Aleatoric uncertainty: Uncertainty due to irreducible noise in data. Epistemic uncertainty: Uncertainty due to lack of data.

LoTV: $\operatorname{Var}[y^{\star} | x^{\star}] = \mathbb{E}_{\theta}[\operatorname{Var}_{y^{\star}}[y^{\star} | x^{\star}, \theta]] + \operatorname{Var}_{\theta}[\mathbb{E}_{y^{\star}}[y^{\star} | x^{\star}, \theta]].$ Bayesian Linear Regression $f^* = w^\top x^*, y^* = f^* + \epsilon$,

 $\epsilon \sim \mathcal{N}(0, \sigma_n^2)$. Prior: $w \sim \mathcal{N}(0, \sigma_n^2 I)$. Posterior: A random vector is Gaussian if (1) the RVs are Gaussian, and (2) $w | X, y \sim \mathcal{N}(\bar{\mu}, \bar{\Sigma}), \ \bar{\mu} = (X^\top X + \sigma_n^2 \sigma_n^{-2} I)^{-1} X^\top y = \sigma_n^{-2} \Sigma X^\top y,$ $\bar{\boldsymbol{\Sigma}} = (\sigma_n^{-2} \boldsymbol{X}^{\top} \boldsymbol{X} + \sigma_v^{-2} \boldsymbol{I})^{-1}.$

Inference: $y^* = x^{\star \top} w + \epsilon \Rightarrow f^* \mid x^*, X, y \sim \mathcal{N}(x^{\star \top} \bar{\mu}, x^{\star \top} \bar{\Sigma} x^* + \sigma_n^2)$. **Logistic regression**: Bernoulli likelihood $(\pi^y(1-\pi)^{1-y})$.

Recursive update: $p^{(t+1)}(\theta) = p(\theta | y_{1:t+1}) = \frac{1}{Z} p^{(t)}(\theta) p(y_{t+1} | \theta)$.

Online data recursion: $X^{\top}X = \sum_{i=1}^{n} x_i x_i^{\top}$, $X^{\top}y = \sum_{i=1}^{n} y_i x_i$.

Gaussian Processes Problem: BLR can only make linear predictions. Solution: GP describes distributions over (non-linear) functions. In function space, f = Xw, which can be sampled by $f \sim \mathcal{N}(0, X^{\top}X) \Rightarrow$ data points enter as inner products \Rightarrow use kernel function $f \sim \mathcal{N}(0, \hat{k}(X, X))$. $\mathcal{GP}(\mu, k)$ is formally defined as an infinite collection of RVs, of which any finite number are jointly Gaussian. Kernel: Formally, $k(x,x') = \phi(x)^{\top} \phi(x')$ for some feature function $\phi \Rightarrow$ kernel function is more efficient. Intuitively, k(x,x') describes how f(x) and f(x') are related. k(X,X) is symmetric and positive semidefinite ($z^{\top}Mz \ge 0$ for all $z \ne 0$). Must satisfy $k(x,x') \le \sqrt{k(x,x)k(x',x')}$. Stationary: k(x,x') = k(x-x'). **Isotropic**: $k(x,x') = k(||x-x'||_2)$ (same as stat. in 1D). **Linear**: Line. Gaussian (RBF): Smooth, larger ℓ : smoother. Laplace (exponential): Non-smooth, larger ℓ : smoother. **Matern**: $\nu = 1/2$: Laplace, $\nu \to \infty$: Gaussian. Addition, multiplication, scaling, polynomial function of kernel functions are also kernel functions.

Inference: $y^* | x^*, X, y \sim \mathcal{N}(\mu^*, k^* + \sigma_n^2)$ with data X_A : $\mu^* = \mu(x^*) + k_{x^*}^{\top} (K_{AA} + \sigma_n^2 I)^{-1} (y - \mu_A)$ $k^{\star} = k(\mathbf{x}^{\star}, \mathbf{x}^{\star}) - k_{\mathbf{x}^{\star}, A}^{\top} (K_{AA} + \sigma_n^2 \mathbf{I})^{-1} k_{\mathbf{x}^{\star}, A}.$

GP posterior: $\mathcal{GP}(\mu',k')$ such that: $\mu'(x) = \mu(x) + k_{xA}^{\top} (K_{AA} + \sigma_n^2 I)^{-1} (y - \mu_A)$

 $k'(x,x') = k(x,x') - k_{x,A}^{\top} (K_{AA} + \sigma_n^2 I)^{-1} k_{x',A}.$ Forward sampling: Iter sample 1-d Gaussian with prod. rule

 $p(f_1,...,f_n)$. Model selection: Hyperparameters matter a lot \Rightarrow Can be learned by maximizing marginal likelihood, $\hat{\theta} = \min_{\theta} y^{\top} K_{y,\theta}^{-1} y + \operatorname{logdet}(K_{y,\theta}),$

which balances the goodness of the fit (term 1) and model complexity (term 2). Problem: To learn a GP, need to invert matrices, which take

 $\mathcal{O}(n^3)$ (BLR: $\mathcal{O}(dn^2)$). **Local methods**: Stationary kernels de-

pend on distance, so only condition if $|k(x,x')| \ge \tau$. Approximation: Approximate stationary kernel with Random Fourier Transform. **Inducing point (FITC)**: Throw away data where there is a lot (cubic in inducing points, linear in data points). Variational Inference In some cases, not realistic to assume Gaussian \Rightarrow Approximate $p(\theta|y) = \frac{1}{7}p(\theta,y) \approx q_{\lambda}(\theta) \Rightarrow$ Minimize

 $KL(q_{\lambda} \parallel p) \Rightarrow q^{\star} = \operatorname{amax}_{\lambda} \mathbb{E}_{\theta \sim q_{\lambda}} [\log p(\boldsymbol{y} \mid \boldsymbol{\theta})] - KL(q_{\lambda} \parallel p_{\text{prior}})$. I.e., minimizing $KL(q_{\lambda} \parallel p) \equiv$ maximizing expected likelihood, while remaining close to prior. **ELBO**: Lower bounds $\log p(y)$, so it is a good method of model selection.

Forward KL $KL(p || q_{\lambda})$: covers full prob. density, but intractable. **Backward KL** $KL(q_{\lambda} || p)$: greedily covers mode.

Laplace approximation: $q_{\lambda}(\theta) = \mathcal{N}(\hat{\theta}, \Lambda)$, where $\hat{\theta} = \max_{\theta} p(\theta)$ y) and $\Lambda = -H_{\theta} \log p(\theta | y)$. Matches shape of the true posterior around its mode ⇒ Extremely overconfident predictions, because it is greedy.

Problem: Want to compute gradient w.r.t. λ of an expectation w.r.t. λ . Reparameterization trick: Suppose $\epsilon \sim \phi$, $\theta = g(\epsilon, \lambda)$ (diff. and inv.), then $\mathbb{E}_{\theta \sim q_{\lambda}}[f(\theta)] = \mathbb{E}_{\epsilon \sim \phi}[f(g(\epsilon,\lambda))]$. (Can be used for MC obj., unbiased.) **Gaussian**: $\theta = g(\epsilon, \lambda) = \Sigma^{1/2} \epsilon + \mu \sim \mathcal{N}(\mu, \Sigma)$. **Inference**: $p(y^* \mid x^*, y) \approx \int p(y^* \mid f^*) q_{\lambda}(f^* \mid x^*) df^*$ (intractable, but single dimension).

Markov Chain Monte Carlo Markov chain: Seq. of RVs s.t. $\overline{X_{t+1} \perp X_{1:t-1} \mid X_t}$. Stationary dist.: $\pi(x) = \sum_{x'} p(x \mid x') \pi(x')$. (Solve by $\pi = P^{\top}\pi$, $1 \cdot \pi = 1$.) Ergodicity: $\exists t [\forall x, x' [p^{(t)}(x' | x) > 0]]$, where $p^{(t)}$ is prob. to reach x' from x in exactly t steps. (Terminal states \Rightarrow not ergodic.) (Ensure ergodicity by self-loops.) Fundamental theorem of ergodic MCs: Ergodic MC always converges to a unique pos. stat. dist. Detailed balance equation: For an unnormalized dist. q an MC satisfies DBE iff $q(x)p(x'|x) = q(x')p(x|x') \Rightarrow \text{stat. dist.} = \frac{1}{7}q.$ Sampling: Sample MC's stat. dist. by first doing a burn-in t_0 to reach stat. dist.

Idea: Approximate intractable p by drawing m samples from MC with stat. dist. $p(\theta|y) \Rightarrow$ $p(y^* | x^*, y) = \mathbb{E}_{p(\cdot | y)}[p(y^* | x^*, \theta)] = \frac{1}{m} \sum_{i=1}^m p(y^* | x^*, \theta_i).$

Hoeffding's inequality: Compute bound on error. $p(|\mathbb{E}_{p(\cdot|y)}[p(y^{\star}|x^{\star},\theta)] - \frac{1}{m}\sum_{i=1}^{m}p(y^{\star}|x^{\star},\theta_i)| > \epsilon) \le 2\exp(-2m\epsilon^2/C^2), \mathbb{E}_{\theta|x_{1:i},y_{1:i}}[H[y_x|\theta]].$ Want points where the post. is uncertain bewhere C is the upper bound of values (1 for prob. dist.). To get a prob. $\leq \delta$ of error $> \epsilon$, we need $m \geq \log 2 - \log \delta / 2\epsilon^2$ samples. **Metropolis-Hastings**: Arbitrary proposal dist. $r(x' \mid x)$. Fol-

low proposal with prob. $\alpha(x' \mid x) = \min \left\{ 1, \frac{q(x')r(x|x')}{q(x)r(x'|x)} \right\} \Rightarrow$ satisfies DBE to get stat. dist. $\frac{1}{7}q(x)$. Arbitrary proposal influences how fast we converge to stat. dist. Gaussian: Prob. dist. has form $p(x) = \frac{1}{7} \exp(-f(x)) \Rightarrow \alpha(x' \mid x) =$ $\min \left\{ q, \frac{r(x|x')}{r(x'|x)} \exp(f(x) - f(x')) \right\}. \text{ If } r(x' \mid x) = \mathcal{N}(x'; x, \tau \mathbf{I}) \Rightarrow$ r=1 (symmetry). If r proposes low energy (high prob) region, acceptance is 1. **Problem**: Uninformed ⇒ Use gradient information (MALA requires full access to f).

Bayesian Deep Learning Non-linear dependencies.

Prior: $\theta \sim \mathcal{N}(0, \sigma_n^2, I)$. Likelihood: $y \mid x, \theta \sim \mathcal{N}(\mu_{\theta}(x), \sigma_{\theta}^2(x))$.

Homoscedastic: Same noise for all data points, $\sigma_{\theta}^2(x) = c$. Heteroscedastic: Varying noise.

MAP: $\min_{\theta} \frac{1}{2\sigma_n^2} \|\theta\|^2 - \sum_{i=1}^n \log \sigma_{\theta}^2(x_i) + \frac{(y_i - \mu_{\theta}(x_i))^2}{2\sigma_{\theta}^2(x_i)}$. Attenuate loss for certain data points by attributing error to large variance Fails to model epistemic uncertainty ⇒ VI (Gaussian in expectation) and Monte Carlo or MCMC:

$$\mathbb{E}[y^{\star} | x^{\star}, y] \approx \frac{1}{m} \sum_{j=1}^{m} \mu_{\theta_j}(x^{\star}).$$

$$\text{Var}[\mathbf{y}^{\star}\,|\,\mathbf{x}^{\star},\!\mathbf{y}] \approx \frac{1}{m} \textstyle\sum_{j=1}^{m} \sigma_{\theta_{j}}^{2}(\mathbf{x}^{\star}) + \frac{1}{m-1} \textstyle\sum_{j=1}^{m} (\mu_{\theta_{j}}(\mathbf{x}^{\star}) - \bar{\mu}(\mathbf{x}^{\star}))^{2}$$

MCMC: Produce seq. $\theta_1,...,\theta_T$, then $p(y^* \mid x^*,y) \approx \frac{1}{T} \sum_{i=1}^T p(y^* \mid x^*,y)$ x^*, θ_i). **Problem**: Cannot store T times params of network. **Solution**: Approx. with Gaussian and running mean/var. **MC dropout**: Dropout during inference ⇔ VI with Bernoulli.

Prob. ensembles: Train networks on rand. subsets, average.

Calibration: Well-calibrated ⇔ confidence (assigned prob.) \approx frequency. **Reliability diagram**: Bin according to class pred. probs. (assume class 1). Above line: underconfident, below line: overconfident. (No samples \Rightarrow empty bin in diagram). $\operatorname{freq}(B_m) = \frac{1}{|B_m|} \sum_{i \in B_m} \mathbb{1}\{Y_i = 1\},$

 $\operatorname{conf}(B_m) = \frac{1}{|B_m|} \sum_{i \in B_m} p(Y_i = 1 | x_i)$. **ECE**: avg. deviation from

perfect calibration: $\ell_{\text{ECE}} = \sum_{m=1}^{M} \frac{|B_m|}{n} |\text{freq}(B_m) - \text{conf}(B_m)|$.

Active Learning Decide which data to collect: \mathcal{NP} -hard.

Uncertainty sampling: Greedily pick points with maximal mutual information $\Rightarrow x_{t+1} = \max_{x} I(f_x; y_x \mid y_{S_t})$. If Gaussian: $x_{t+1} = \max_{x} \frac{\sigma_{x|S_t}^2}{\sigma_{x}^2} / \sigma_{x}^2(x)$. $\gamma_T = \max_{x_{1:T}} I(f(x_{1:T}); y_{1:T})$. Monotone submodular. Constant factor approx: $I(f(x_{1:T}); y_{1:T}) \ge$ $(1-1/e)\gamma_T$ (near-optimal, $1-1/e\approx 0.63$).

BALD: $x_{t+1} = \max_{x} I(y_x; \theta \mid x_{1:t}, y_{1:t}) = \max_{x} H[y_x \mid x_{1:t}, y_{1:t}]$ cause all θ are certain about their differing pred. Approximate term 2 using VI and MC.

Bayesian Optimization Not only reduce uncertainty, but also maximize objective. $x_{t+1} = \operatorname{amax}_{x} a(x)$.

Regret: $R_T = \sum_{t=1}^T (\max_x f(x) - f(x_t))$. Want algorithm with sublinear regret: $\lim_{T\to\infty} R_T/T = 0$. $f^* = \max_x f(x)$.

GP-UCB: Optimism in the face of uncertainty: pick point where we can hope for best outcome: $a_{\text{UCB}} = \mu_t(x) + \beta_t \sigma_t(x)$. $\mu_t \sigma$ from GP. $R_T \in \mathcal{O}^*(\sqrt{\gamma_T/T})$. **GP bounds**: Linear: $\gamma_T \in \mathcal{O}(d \log T)$, Gaussian: $\mathcal{O}((\log T)^{d+1})$, Matèrn: $\mathcal{O}(T^{d/2\nu+d}(\log T)^{2\nu/2\nu+d})$.

PI: $a_{\text{PI}}(x) = \Phi((\mu_t(x) - f^*)/\sigma_t(x))$ is prob. to improve f^* . Greedy. **EI**: $a_{\text{FI}}(x) = (\mu_t(x) - f^*)\Phi((\mu_t(x) - f^*)/\sigma_t(x)) + \sigma_t(x)\phi((\mu_t(x) - f^*)/\sigma_t)$ is expectation of improvement.

Thompson sampling: Draw sample from GP and select max.

Markov Decision Processes Env. that makes Markov **ass.** (states \mathcal{X} , actions \mathcal{A} , transitions $p(x' \mid x, a)$, rewards r(x,a)). Policy: π maps states to actions (induces MC with $p(x' \mid x) = \sum_{a} \pi(a \mid x) p(x' \mid x, a)$, want to find π that max. long-

term rewards. Horizon T reward: $\mathbb{E}_{\pi}[\sum_{t=0}^{T} r(x_t, a_t)]$. Horizon ∞ reward: $\mathbb{E}_{\pi}\left[\sum_{t=0}^{\infty} \gamma^{t} r(x_{t}, a_{t})\right]$. Geo series: $\sum_{t=0}^{\infty} \gamma^{t} = 1/1 - \gamma$.

 $V^{\pi}(x) = \mathbb{E}_x\left[\sum_{t=0}^{\infty} \gamma^t r(X_t, \pi(X_t)) \mid X_0 = x\right]$ $= r(x,\pi(x)) + \gamma \sum_{x'} p(x'|x,\pi(x)) V^{\pi}(x')$ (Bellman eq).

 $Q(x,a) = r(x,a) + \gamma \sum_{x'} p(x' | x,a) V^{\pi}(x'). V(x) = \max_{a} Q(x,a).$

Bellman theorem: π is optimal \Leftrightarrow greedy w.r.t. V^{π} .

Policy iteration: $\pi \Rightarrow V^{\pi}, V \Rightarrow \pi_V$ (alternate). $\pi_V(x) = \operatorname{amax}_a r(x,a) + \gamma \sum_{x'} p(x'|x,a) V(x')$ (greedy). Converges monotonically, guaranteed to converge in $\mathcal{O}(|\mathcal{X}|^2|\mathcal{A}|/(1-\gamma))$ itera-

tions, expensive (computed efficiently by solving single LSoE). Value iteration: Dynamic programming:

 $V_t(x) = \max_a r(x,a) + \gamma \sum_{x'} p(x'|x,a) V_{t-1}(x')$. Iterates until $\|v_t - v_{t-1}\|_{\infty} \le \epsilon$: ϵ -optimal convergence, in polynomial in iterations, per iteration: $\mathcal{O}(|\mathcal{X}|^2|\mathcal{A}|)$, inexpensive. Then, pick greedy policy.

Reinforcement Learning Learn within unknown MDP. On**policy**: Learn from own data, **Off-policy**: Learn from other policy data, Model-based: Learn MDP and solve, Model-free: Learn value function directly. Data points: $\langle x,a,r,x' \rangle$.

Robbins-Monro: $\sum_{t=0}^{\infty} x_t = \infty$, $\sum_{t=0}^{\infty} x_t^2 < \infty$. E.g. 1/t. ϵ -greedy (based): Pick random action with prob. ϵ_t , or best action according to MDP with prob. $1 - \epsilon_t$. Guaranteed to converge to optimal policy if ϵ_t satisfies RM. Problem: does not quickly eliminate suboptimal actions.

 R_{max} (based): Solves problem. Add fairy tale $p(x^* \mid x^*, a) =$ $1,r(x^{\star},a) = R_{\text{max}}$, assume unexplored go there.

TD-learning (on, free): $V^{\pi}(x) \leftarrow (1-\alpha_t)V^{\pi}(x) + \alpha_t(r+\gamma V^{\pi}(x'))$. Guaranteed to converge if α_t satisfies RM and all states are visited infinitely often. Space: $\mathcal{O}(|\mathcal{X}|)$.

Q-learning (off, free): $Q(x,a) \leftarrow (1-\alpha_t)Q(x,a) + \alpha_t(r+\gamma \max_{a'}Q(x',a))$ Guaranteed to converge if α_t satisfies RM and all state-action pairs are visited infinitely often. Space: $\mathcal{O}(|\mathcal{X}||\mathcal{A}|)$.

DQN (off, free, cont. states): GD on $\frac{1}{2}(Q(x, a; \theta) - (r - \theta))$ $\gamma \max_{a'} Q(x', a'; \theta^{\text{old}}))^2$ (Bellman error). Slow to converge: maintain constant target network. DDQN: Maximization bias makes DQN overestimate. Solution: 2 Q networks where we take minimum to be value.

Policy search (on, free, cont. actions): Parametrize $\pi(x;\theta)$. Maximize expected trajectory reward. $\nabla_{\theta} \mathbb{E}_{\tau \sim \pi_{\theta}}[r(\tau)] =$ $\mathbb{E}_{t \sim \pi_{\theta}}[r(\tau) \nabla_{\theta} \log \pi_{\theta}(\tau)] = \mathbb{E}_{t \sim \pi_{\theta}}[r(\tau) \sum_{t=0}^{T} \nabla_{\theta} \log \pi_{\theta}(a_t \mid x_t)]$ ⇒ Do not need to know MDP to compute gradient. **Baselines**: Large variance \Rightarrow introduce baseline: $\mathbb{E}_{\tau \sim \pi_{\theta}}[r(\tau)\nabla_{\theta}\log \pi_{\theta}] =$ $\mathbb{E}_{\tau \sim \pi_{\theta}}[(r(\tau) - b(\tau))\nabla_{\theta}\log \pi_{\theta}(\tau)].$

REINFORCE (on, free): $\theta \leftarrow \theta + \eta \gamma^t G_t \nabla_{\theta} \log \pi_{\theta}(a_t \mid x_t)$, where $G_t = r(\tau) - b_t = \sum_{t'-t}^{T} \gamma^{t'-t} r_t$

Actor-critic (on, free): REINFORCE gradient = $\mathbb{E}_{(x,a)\sim\pi_{\theta}}[Q(x,a)\nabla_{\theta}\log\pi_{\theta}(a\mid x)]$. So: parametrize actor π_{θ} and critic Q_{θ} . Use in each others' update equations: $\theta_{\pi} \leftarrow \theta_{\pi} + \eta_t Q(x, a \mid \theta_O) \nabla_{\theta} \log \pi_{\theta}(a \mid x).$

 $\theta_O \leftarrow \theta_O - \eta_t(Q(x,a;\theta_O) - r - \gamma Q(x',\pi(x';\theta_\pi);\theta_O)) \nabla_{\theta} Q(x,a;\theta).$

A2C (on, free): Add value network V_{θ} for the baseline: A(x,a) =Q(x,a) - V(x,a) (advantage function). This centers the Q-values.

DDPG (off, free): Replace $\max_{a'} Q(x', a'; \theta^{\text{old}})$ in DQN by $\pi(x'; \theta_{\pi})$, where π should follow the greedy policy w.r.t. Q. **Key idea**: If we use a rich enough parameterization of policies, selecting the greedy policy w.r.t. Q is equivalent to $\theta_{\pi}^{\star} = \operatorname{amax}_{\theta_{\pi}} \mathbb{E}_{x \sim \mu}[Q(x, \pi(x; \theta_{\pi}); \theta_{O})]. \ \mu(x) > 0 \text{ is an exploration}$ distribution with full support. This needs det. π , thus we inject noise for exploration (akin ϵ -greedy).

TD3 (off, free): Add second critic network to address max. bias.

SAC (off, free): Add entropy regularization to loss $\lambda H(\pi_{\theta})$. **Planning:** Det. transition function $x_{t+1} = f(x_t, a_t)$ and reward function $r(x_t, a_t)$. Cannot plan over infinite horizon \Rightarrow **Key idea**: plan over finite horizon *H*, carry out first action, repeat. Optimize $J_H(a_{t:t+H-1}) = \sum_{t'=t}^{t+H-1} \gamma^{t'-t} r(x_t, a_t)$. Local minima, vanishing/exploding gradient ⇒ heuristics ⇒ Random shooting (*m* samples, pick best). Will not work if sparse rewards \Rightarrow Get access to value function to look further: $J_H(a_{t:t+H-1}) = \sum_{t'=t}^{t+H-1} \gamma^{t'-t} r_t + \gamma^H V(x_{t+H}).$

Model-based ML: Estimate f and r off-policy with supervised learning $((x_t,a_t) \mapsto (r_t,x_{t+1})$, regression). **Benefit**: Dramatically decreases sample complexity (need less data). Use MAP estimate \Rightarrow exploited by planning algos \Rightarrow Uncertainty (GP/BNN).