Shape Constrained Regression for Excel

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Summary

Generalized Additive Models (gam) estimate non-linear but additive relationships between a response and a set of predictors. However, when the training data has a high noise to signal ratio, gam could estimate wiggly (non-monotonic) relationships that are difficult to explain. Shape Constrained Additive Models (scam) impose user-defined monotonic relationships that can be explained by the user.

The mgcv and scam packages are highly effective tools for developing gam and scam models. Both packages use the p-spline basis to construct non-linear features. Although p-splines have desirable properties, they are difficult to implement in other software tools like Excel. There is no closed-form equation that I could export from gam/scam and into Excel for an end-user.

Multivariate Adaptive Regression Splines (mars), from the earth package, allow us to approximate the scam model with linear basis functions – which are Excel-friendly.

Through a data simulation, I demonstrate how to fit gam and scam models using the mgcv and scam packages. Then I approximate the scam model and provide an Excel-friendly regression equation. Reader should be aware that MARS does not guarantee monotonicity. Therefore, after fitting an approximation of the scam model, reader should test the predictions by feeding extreme values of the predictors into the final model.

We also explore the **scar** package which uses monotonic binning. Then I approximate the scar model with non-linear least squares and sigmoid functions. Sigmoid functions guarantee monotonicity. However, sigmoid functions assume that the response tapers off on the tails (i.e., the relationship is S-shaped).

Approximations are achieved by turning the logistic regression problem into a least squares problem. The trick is to use a pseudo response that is continuous (rather than binary).

Data Generating Process (DGP)

The chunk generates 3 data frames: full data, training data, and test data. To demonstrate the wiggliness associated with gam, the sample size of the training data is limited to only 2000 observations.

Each data frame contains a binary response and two predictors. The first predictor has a monotonically increasing relationship with the response, while the second predictor has a monotonically decreasing relationship with the response. Both relationships are sigmoidal and hence highly non-linear.

In addition, the event rate of the response is rare.

library(mgcv)
library(scam)
library(tidyverse)
library(caTools)
library(MLmetrics)

```
library(earth)
library(plotmo)
library(furrr)
library(scar)
library(minpack.lm)
library(car)
library(cgam)
plan(multisession, workers = 6) # set to 2 for most PCs
set.seed(2001)
nobs <- 100000
x1 <- rnorm(nobs)</pre>
x2 <- rnorm(nobs)
x3 <- sample(c('A', 'B', 'C', 'D'), size=nobs, replace=TRUE,
             prob = c(0.8, 0.1, 0.05, 0.05))
# Population log odds is a function of sigmoid features
z \leftarrow -35 + 5 * SSfpl(x1, -2, 2, 0, 0.5) - 5 * SSfpl(x2, -1, 1, 0, 0.1) +
         + 5 * I(x3=='B') + 10 * I(x3=='C') + 10 * I(x3=='D') +
 rnorm(nobs, 0, 20)
y <- rbinom(nobs, 1, prob=boot::inv.logit(z))
full_data <- data.frame(y=y, x1=x1, x2=x2, x3=as.factor(x3))</pre>
summary(full_data)
##
                            x1
                                                  x2
                                                                xЗ
##
  Min.
          :0.00000
                      Min.
                            :-4.469540
                                           Min.
                                                  :-4.349958
                                                                A:79992
##
  1st Qu.:0.00000
                      1st Qu.:-0.672327
                                           1st Qu.:-0.667040
                                                                B:10044
## Median :0.00000
                      Median : 0.006574
                                           Median : 0.002566
                                                                C: 5032
                      Mean : 0.004819
                                                 : 0.000159
                                                                D: 4932
## Mean
           :0.06252
                                           Mean
##
   3rd Qu.:0.00000
                      3rd Qu.: 0.681828
                                           3rd Qu.: 0.676167
           :1.00000
                      Max. : 4.071722
## Max.
                                           Max.
                                                  : 4.590295
train_index <- sample.split(full_data$y, SplitRatio = 2/100)</pre>
train_data <- full_data[train_index,] # small</pre>
test_data <- full_data[!train_index,] # large</pre>
```

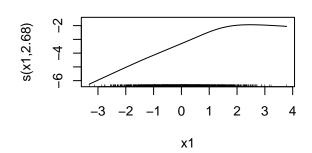
GAM

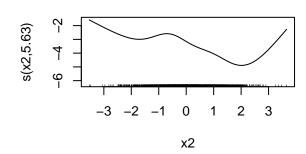
Due to the high noise to signal ratio in the training data, gam estimates very wiggly relationships between the binary response and the predictors. The y-axis is on the log-odds (or logit) scale.

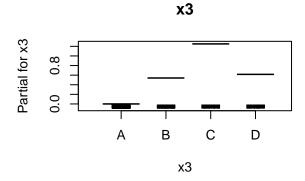
```
## fit a gam
## when population error is high, gam returns non-monotonic relationships
## even when the population relationships are monotonic
mod_gam \leftarrow gam(y \sim s(x1) + s(x2) + x3, data = train_data, family = binomial)
summary(mod_gam)
##
## Family: binomial
## Link function: logit
##
## Formula:
## y \sim s(x1) + s(x2) + x3
##
## Parametric coefficients:
              Estimate Std. Error z value Pr(>|z|)
                          0.1492 -22.299 < 2e-16 ***
## (Intercept) -3.3278
## x3B
                0.5406
                           0.2876 1.879 0.060212 .
## x3C
                1.2506
                           0.3277
                                   3.816 0.000136 ***
                0.6160
                           0.3836
                                   1.606 0.108282
## x3D
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
## Approximate significance of smooth terms:
          edf Ref.df Chi.sq p-value
## s(x1) 2.677 3.421 51.80 < 2e-16 ***
## s(x2) 5.634 6.807 30.41 8.49e-05 ***
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
## R-sq.(adj) = 0.0685 Deviance explained = 12.6\%
## UBRE = -0.57921 Scale est. = 1
plot(mod_gam, pages=1, se=FALSE, main="Unconstrained GAM",
    shift=coef(mod_gam)[1], all.terms=TRUE)
```

Unconstrained GAM

Unconstrained GAM







SCAM

Shape constraints on each predictor reduce variance (wiggliness). Out-of-sample performance is similar between gam and scam.

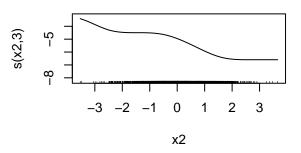
```
##
## Family: binomial
## Link function: logit
##
## Formula:
## y ~ s(x1, bs = "mpi") + s(x2, bs = "mpd") + x3
##
```

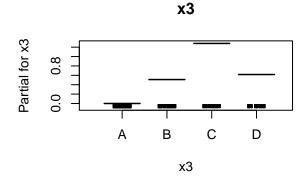
```
## Parametric coefficients:
##
              Estimate Std. Error z value Pr(>|z|)
                                  -1.794
## (Intercept) -5.2813
                           2.9437
                                            0.0763 .
                0.5097
                                    1.773
## x3B
                           0.2875
## x3C
                1.2791
                           0.3255
                                    3.930 8.5e-05 ***
## x3D
                0.6168
                           0.3830
                                    1.610
                                            0.1073
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
##
## Approximate significance of smooth terms:
         edf Ref.df Chi.sq p-value
## s(x1) 1.98 2.466 54.97 1.08e-11 ***
## s(x2) 3.00 3.000 25.99 9.58e-06 ***
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
##
                         Deviance explained = 11.6%
## R-sq.(adj) = 0.0606
## UBRE score = -0.57769 Scale est. = 1
plot(mod_mono_gam, pages=1, se=FALSE, main="Constrained GAM",
    shift=coef(mod_mono_gam)[1], all.terms=TRUE)
```



-3 -2 -1 0 1 2 3 4

Constrained GAM





```
## helper function for test set performance

perf <- function(lst_preds, f_metric=caTools::colAUC, metricname="ROC-AUC"){</pre>
```

```
map_dfr(lst_preds, function(x){
  f_metric(x, test_data$y)
}) %>%
  pivot_longer(everything(), names_to="model", values_to=metricname) %>%
  knitr::kable()
}
```

```
map_dfr(list(mod_gam=mod_gam, mod_mono_gam=mod_mono_gam), AIC) %>%
pivot_longer(everything(), names_to="model", values_to="AIC") %>%
knitr::kable()
```

model	AIC
mod_gam	841.5774
mod_mono_gam	844.6143

model	ROC-AUC
gam	0.7122259
mono_gam	0.7115072

```
perf(myPreds, MLmetrics::LogLoss, "LogLoss")
```

model	LogLoss
gam	0.2180729
$mono_gam$	0.2180345

Generate Pseudo-Training Data

The next step is controversial. I pretend that the scam model is the data generating process. In the actual training data, we observe binary outcomes rather than the log odds. In the pseudo data, I ignore the binary outcomes and "observe" the log odds from the scam model.

The reader should be aware that in the actual training data, y is a binary response. In the pseudo data, y is log odds from the scam model, which is analogous to z in the true data generating process.

```
## generate pseudo-data using Mono GAM
## suppose the scam model is the data generating process (dgp).
## append the log odds for each observation in the training
## data
pseudo_train_data <- train_data
pseudo_train_data$y <- as.numeric(predict(mod_mono_gam, newdata=train_data))</pre>
summary(pseudo_train_data)
##
## Min. :-6.6438 Min. :-3.31420 Min. :-3.53202 A:1609
## 1st Qu.:-3.9143
                   1st Qu.:-0.69239 1st Qu.:-0.64836
                                                         B: 198
## Median :-3.1377 Median : 0.02514 Median :-0.01473
                                                        C: 98
## Mean
         :-3.1751 Mean : 0.01329 Mean : 0.01956 D: 95
## 3rd Qu.:-2.3674
                    3rd Qu.: 0.67314 3rd Qu.: 0.70888
```

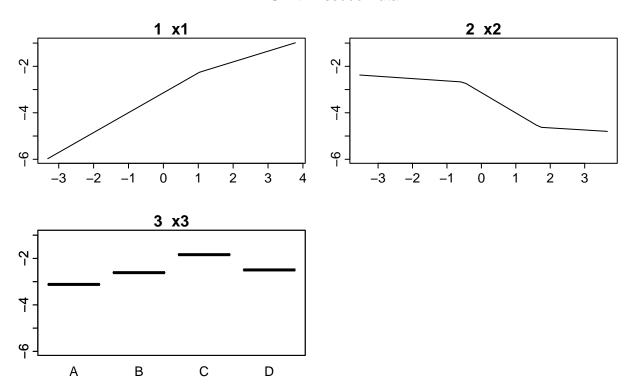
Approximate SCAM with MARS

Max. :-0.2218 Max. : 3.78302 Max. : 3.65809

```
## approximate the Mono GAM model with MARS
## unfortunately, both gam and scam models are very difficult to
## implement outside of R due to how the basis functions are defined.
## an alternative is to approximate the scam relationships
## with simpler basis functions like the linear basis (which can be
## easily implemented in Excel) -- the linear basis is also called the
## reLU (rectified linear unit)
mod_logit_pseudo <- earth(y ~ x1 + x2 + x3, data=pseudo_train_data)</pre>
summary(mod_logit_pseudo)
## Call: earth(formula=y~x1+x2+x3, data=pseudo_train_data)
```

```
##
##
                    coefficients
## (Intercept)
                     -1.8121035
## x3B
                      0.5063307
## x3C
                      1.2805724
## x3D
                      0.6230057
## h(1.03699-x1)
                     -0.8552545
                      0.4611403
## h(x1-1.03699)
## h(-0.518908-x2)
                     0.1005579
## h(x2 - 0.518908)
                    -0.8751410
## h(x2-1.70494)
                     0.7850924
##
## Selected 9 of 9 terms, and 5 of 5 predictors
```

MARS with Pseudo Data



Since the pseudo model is in the log-odds scale, I need to convert the predictions into probabilities to compare against the probability predictions from gam and scam.

```
## test set performance

preds_pseudo <- boot::inv.logit(predict(mod_logit_pseudo, newdata=test_data))

myPreds <- list(gam=preds_gam, mono_gam=preds_mono_gam, pseudo_mono=preds_pseudo)

perf(myPreds, caTools::colAUC, "ROC-AUC")</pre>
```

model	ROC-AUC
gam	0.7122259

model	ROC-AUC
mono_gam pseudo_mono	$0.7115072 \\ 0.7129432$

```
perf(myPreds, MLmetrics::LogLoss, "LogLoss")
```

model	LogLoss
gam mono_gam pseudo_mono	$\begin{array}{c} 0.2180729 \\ 0.2180345 \\ 0.2177323 \end{array}$

Excel-friendly Equation

The equation below predicts the log odds or logit.

```
cat(format(mod_logit_pseudo, style="pmax", use.names=TRUE))
```

```
##
     -1.812103
##
     + 0.5063307 * x3B
    + 1.280572 * x3C
##
##
    + 0.6230057 * x3D
    -0.8552545 * pmax(0, 1.036994 -
##
    + 0.4611403 * pmax(0,
##
                            x1 - 1.036994)
    + 0.1005579 * pmax(0, -0.5189081 -
##
    -0.875141 * pmax(0, x2 - -0.5189081)
                             x2 - 1.704942)
##
    + 0.7850924 * pmax(0,
## Compare predict() against score_function
as.func <- function(object, digits = 20, use.names = TRUE, ...)
  eval(parse(text=paste(
    "function(x)\{\n",
    "if(is.vector(x))\n",
    " x \leftarrow matrix(x, nrow = 1, ncol = length(x))\n",
   "x <- model.matrix(delete.response(object$terms),x) \n",
   "with(as.data.frame(x),\n",
   format(object, digits = digits, use.names = use.names, style = "pmax", ...),
   ")\n",
   "}\n", sep = "")))
score_function <- as.func(mod_logit_pseudo)</pre>
compare_df <- expand.grid(x1 = seq(-10, 10, 0.1),
          x2 = seq(-10, 10, 0.1),
          x3 = factor(c('A', 'B', 'C', 'D')))
earth_preds <- predict(mod_logit_pseudo, newdata=compare_df)</pre>
```

```
score_preds <- score_function(compare_df)
max(abs(earth_preds - score_preds))</pre>
```

[1] 3.552714e-15

Testing monotonicity

MARS does not guarantee monotonic relationships. Reader should always test the model for violations of monotonicity by feeding extreme predictor values into the model.

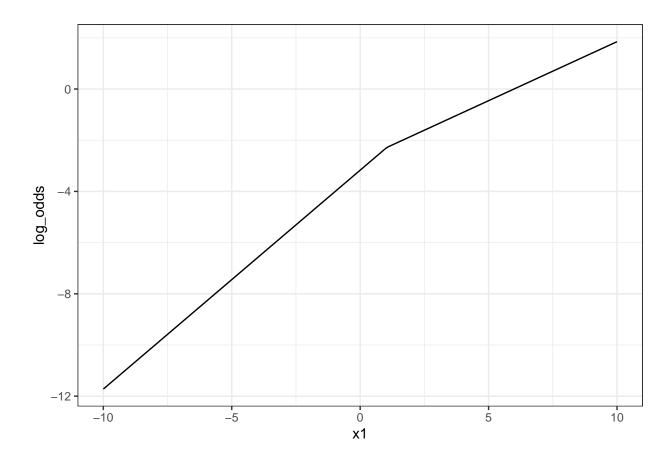
```
mono_check <- function(x, increasing=TRUE){</pre>
  if(increasing==TRUE){
    out \leftarrow all(x == cummax(x))
  } else {
    out <- all(x==cummin(x))</pre>
  return(out)
}
df_x1 \leftarrow data.frame(x1 = seq(-10,10, 0.1),
                      x2 = mean(train_data$x2),
                      x3 = factor('A', levels=c('A','B','C','D'))
df_x2 \leftarrow data.frame(x1 = mean(train_data$x1),
                      x2 = seq(-10,10, 0.1),
                      x3 = factor('A', levels=c('A','B','C','D'))
df_x1$log_odds <- score_function(df_x1)</pre>
df_x2$log_odds <- score_function(df_x2)</pre>
mono_check(df_x1$log_odds)
```

```
## [1] TRUE
```

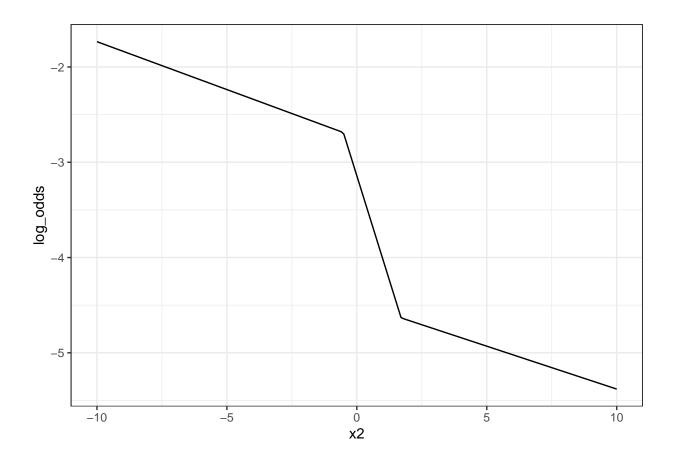
```
mono_check(df_x2$log_odds, FALSE)
```

```
## [1] TRUE
```

```
ggplot(df_x1, aes(x=x1, y=log_odds)) + geom_line() + theme_bw()
```



ggplot(df_x2, aes(x=x2, y=log_odds)) + geom_line() + theme_bw()



How often does monotonicity fail?

Since MARS does not guarantee monotonicity. I repeated, 60 times, the following steps:

- 1. Set a new seed value
- 2. Generate full, train, and test sets
- 3. Fit a scam model to the training data
- 4. Generate pseudo data
- 5. Fit a MARS model to the pseudo data
- 6. Check shape constraints (x1 should increase with y and x2 should decrease with y)

```
# Population log odds is a function of sigmoid features
  z \leftarrow -35 + 5 * SSfpl(x1, -2, 2, 0, 0.5) - 5 * SSfpl(x2, -1, 1, 0, 0.1) +
           + 5 * I(x3=='B') + 10 * I(x3=='C') + 10 * I(x3=='D') +
    rnorm(nobs, 0, 20)
  y <- rbinom(nobs, 1, prob=boot::inv.logit(z))
  # Generate full, train, and test sets
  full_data <- data.frame(y=y, x1=x1, x2=x2, x3=as.factor(x3))
  train_index <- sample.split(full_data$y, SplitRatio = 2/100)</pre>
  train_data <- full_data[train_index,] # small</pre>
  test_data <- full_data[!train_index,] # large</pre>
  # SCAM
  my_mod_mono_gam \leftarrow scam(y \sim s(x1, bs="mpi") + s(x2, bs="mpd") + x3,
                        data = train_data, family = binomial)
  # Pseudo data
  pseudo_train_data <- train_data
  pseudo_train_data$y <- as.numeric(predict(my_mod_mono_gam, newdata=train_data))</pre>
  summary(pseudo_train_data)
  # MARS
  my_mod_pseudo_mars <- earth(y ~ x1 + x2 + x3, data=pseudo_train_data)</pre>
  # Check shape constraints
  df_x1 \leftarrow data.frame(x1 = seq(-10,10, 0.1),
                       x2 = mean(train_data$x2),
                       x3 = factor('A', levels=c('A', 'B', 'C', 'D'))
  df_x2 \leftarrow data.frame(x1 = mean(train_data$x1),
                       x2 = seq(-10, 10, 0.1),
                       x3 = factor('A', levels=c('A','B','C','D'))
                       )
  df_x1$log_odds <- predict(my_mod_pseudo_mars, newdata=df_x1)</pre>
  df_x2$log_odds <- predict(my_mod_pseudo_mars, newdata=df_x2)</pre>
  success <- mono_check(df_x1$log_odds) & mono_check(df_x2$log_odds, FALSE)</pre>
  return(success)
}
```

```
## [1] 0.7333333
```

MARS successfully approximated the scam model about 73.3333333% of the time.

Why Pseudo Data?

The pseudo data filters out the noise in the latent response (z), which drives the binary response (y). Although we could directly pass the actual training data into MARS, the noise in the data generating process could cause earth to return non-monotonic relationships.

Reader should notice that y in the actual training data is a binary outcome.

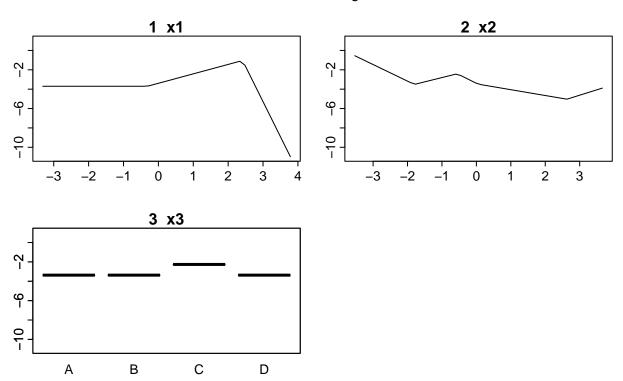
```
mod_earth <- earth(y ~ x1 + x2 + x3, data=train_data, glm=list(family=binomial))
summary(mod_earth)</pre>
```

```
## Call: earth(formula=y~x1+x2+x3, data=train_data, glm=list(family=binomial))
##
## GLM coefficients
##
## (Intercept)
                    -11.4359941
## x3C
                      1.1057666
## h(x1 - 0.315568)
                      0.9740516
## h(x1-2.41403)
                     -8.2227755
## h(x2 - -1.80226)
                      2.5956455
## h(x2 - 0.555908)
                     -2.6511588
## h(x2-0.0490364)
                      1.1672334
## h(2.6295-x2)
                      1.7135309
## GLM (family binomial, link logit):
## nulldev
              df
                       dev
                             df
                                  devratio
                                                AIC iters converged
  935.167 1999
                   825.049 1992
                                      0.118
## Earth selected 8 of 12 terms, and 3 of 5 predictors
## Termination condition: RSq changed by less than 0.001 at 12 terms
## Importance: x1, x2, x3C, x3B-unused, x3D-unused
## Number of terms at each degree of interaction: 1 7 (additive model)
## Earth GCV 0.05595653
                           RSS 110.2407
                                            GRSq 0.04596337
                                                               RSq 0.05927977
```

```
plotmo(mod_earth, type="link", caption="MARS with Training Data")
```

```
## plotmo grid: x1 x2 x3 ## 0.02514023 -0.01473112 A
```

MARS with Training Data



Alternative to SCAM: SCAR

Models developed in scam return smooth relationships between a response and each predictor. However, the smoothness (and monotonicity) is achieved through the p-spline basis, which is difficult to implement. If smoothness is not required, an alternative approach is to use monotonic step functions to approximate the relationships. The scar package achieves this goal.

Unfortunately, the scar package does not take data frames. I wrote two helper functions to convert a data frame into a matrix, which is then passed to the scar function.

Now we can fit a scar model with monotonic binning.

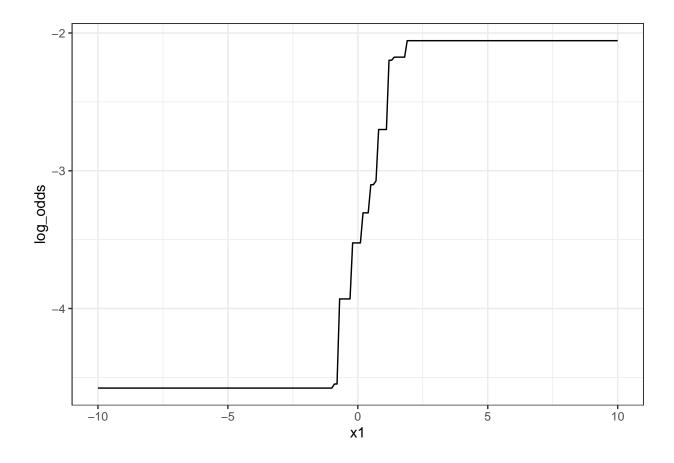
Inside mod_scar is an object called componentfit. It contains partial fitted values by predictor (column) and observation (row). Each observation in the training set corresponds to a row in componentfit. The fitted value for a single observation is the row sum of the partial fitted values.

Given a pair of new predictor values (that are not in the training data), the predict function interpolates the partial fitted values from componentfit. This may be problematic if componentfit is very big. A similar problem arises with generating predictions from k-nearest-neighbors (knn).

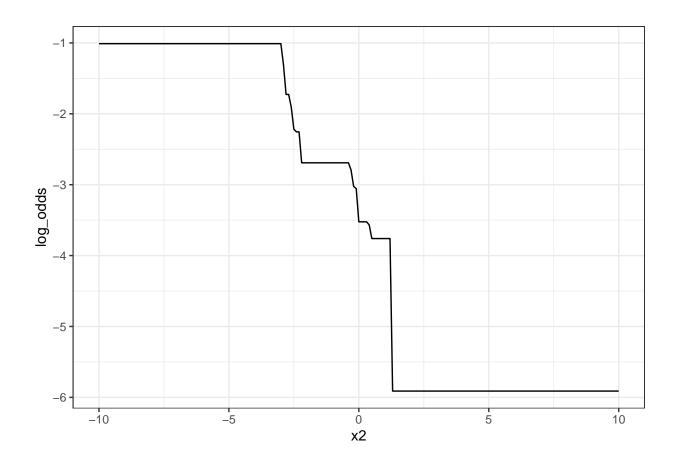
The interpolation of partial fitted values causes the relationship between a response and a predictor to appear "step-like".

Check the partial plots.

```
df_x1$log_odds <- predict_scar(mod_scar, newdata=df_x1)
df_x2$log_odds <- predict_scar(mod_scar, newdata=df_x2)
ggplot(df_x1, aes(x=x1, y=log_odds)) + geom_line() + theme_bw()</pre>
```



ggplot(df_x2, aes(x=x2, y=log_odds)) + geom_line() + theme_bw()



Check test set performance

model	ROC-AUC
gam	0.7122259
mono_gam	0.7115072
pseudo_mono	0.7129432
scar	0.7032823

```
perf(myPreds, MLmetrics::LogLoss, "Log Loss")
```

model	Log Loss
gam	0.2180729

model Log Loss mono_gam 0.2180345 pseudo_mono 0.2177323 scar 0.2226342		
pseudo_mono 0.2177323	model	Log Loss
	pseudo_mono	

Approximating SCAR

Similar to the MARS approximation of scam, we could approximate scar with another method (with the help of pseudo data). The pseudo response would be the log-odds from the scar model. The features would be sigmoid functions that are fed into non-linear least squares. Sigmoid functions guarantee monotonicity.

First, some pseudo data!

```
##
                                                 x2
                                                               хЗ
                             x1
##
   Min.
           :-6.96499
                       Min.
                               :-3.31420
                                           Min.
                                                  :-3.53202
                                                               A:1609
                                                               B: 198
##
   1st Qu.:-3.97509
                       1st Qu.:-0.69239
                                           1st Qu.:-0.64836
  Median :-3.14439
                       Median : 0.02514
                                           Median :-0.01473
                                                                   98
##
   Mean
           :-3.30188
                       Mean
                               : 0.01329
                                           Mean
                                                  : 0.01956
                                                              D:
                                                                   95
##
    3rd Qu.:-2.39797
                       3rd Qu.: 0.67314
                                           3rd Qu.: 0.70888
##
   Max.
           :-0.02539
                       Max.
                               : 3.78302
                                                  : 3.65809
                                           Max.
```

Second, non-linear least squares and sigmoid features. The classical nls function is very sensitive to initial parameter guesses. I have never been successful in using it. I suggest using nlsLM from the minpack.lm package.

```
est <- coef(mod_pseudo_scar)

est_df <- data.frame(Estimate=est)

knitr::kable(est_df)</pre>
```

	Estimate
constant	-1.1143412
a	1.3179421
b	0.9694471
\mathbf{c}	0.4846551
d	1.3170561
e	0.6501136
A1	-1.1071181
B1	1.1311339
xmid1	0.3076759
scal1	0.7142092
A2	-1.2953769
B2	-5.6418496
xmid2	1.1759558
scal2	0.5813312

We could also bootstrap 95% confidence intervals for each estimate.

```
boot_psuedo <- Boot(mod_pseudo_scar, R=1000, ncores=6)</pre>
```

```
CI_nls <- data.frame(confint(boot_psuedo, type="perc"))
names(CI_nls) <- c("p2.5", "p97.5")
CI_nls$signif <- (CI_nls$p2.5 * CI_nls$p97.5) > 0
knitr::kable(CI_nls)
```

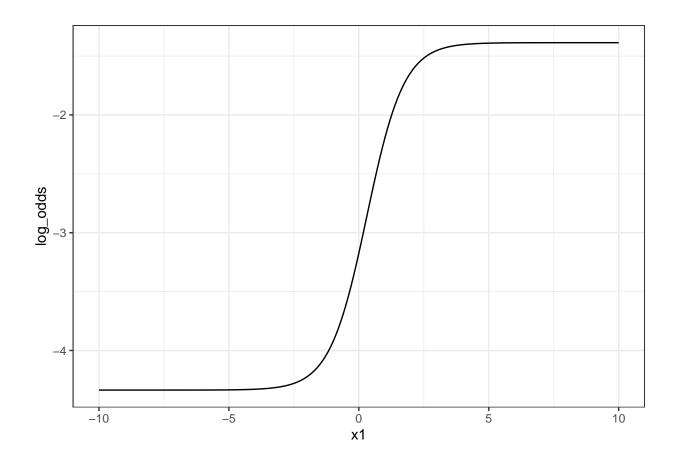
	p2.5	p97.5	signif
constant	-1.5185879	-1.0091381	TRUE
a	1.1410248	1.5684626	TRUE
b	0.9237093	1.1180000	TRUE
\mathbf{c}	0.4273741	0.5429882	TRUE
d	1.2456665	1.3908390	TRUE
e	0.5908989	0.7134469	TRUE
A1	-1.3661630	-0.7545919	TRUE
B1	0.8543728	1.5443743	TRUE
xmid1	0.2458566	0.3724398	TRUE
scal1	0.6595468	0.7744543	TRUE
A2	-1.4793416	-0.7718849	TRUE
B2	-5.8894719	-4.8255832	TRUE
xmid2	1.1158966	1.2595466	TRUE
scal2	0.5386111	0.6301844	TRUE

Some plots.

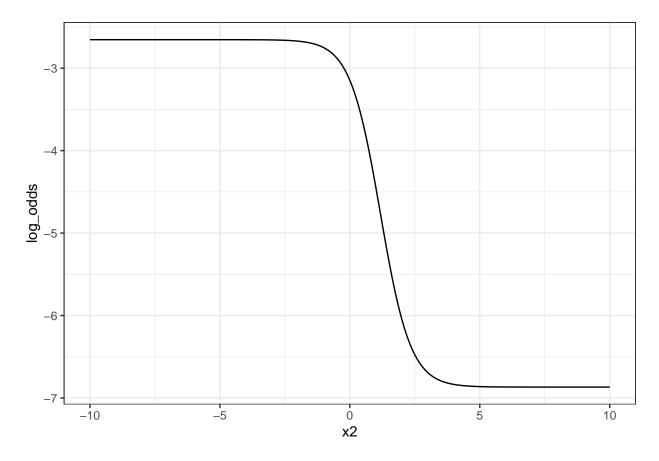
```
df_x1$log_odds <- predict(mod_pseudo_scar, newdata=df_x1)

df_x2$log_odds <- predict(mod_pseudo_scar, newdata=df_x2)

ggplot(df_x1, aes(x=x1, y=log_odds)) + geom_line() + theme_bw()</pre>
```



ggplot(df_x2, aes(x=x2, y=log_odds)) + geom_line() + theme_bw()



Test set performance

model	ROC-AUC
gam	$ \begin{array}{r} 0.7122259 \\ 0.7115072 \end{array} $
mono_gam pseudo_mono	0.7113072 0.7129432
scar pseudo_scar	$\begin{array}{c} 0.7032823 \\ 0.7061607 \end{array}$

perf(myPreds, MLmetrics::LogLoss, "Log Loss")

model	Log Loss
gam	0.2180729
$mono_gam$	0.2180345
pseudo mono	0.2177323

model	Log Loss
scar	0.2226342
$pseudo_scar$	0.2210625

The approximation of scar actually outperforms the original scar model because the approximation extrapolates using the sigmoid functions that were estimated by non-linear least squares.

Sigmoid functions guarantee monotonicity.

CGAM

SCAM and SCAR do not constrain the relationships between the response and categorical predictors. In our simulated data set, the effects of C and D should be similar. However, in our training data set, this does not appear to be the case.

An alternative to the scam package is cgam. One major draw back of this package is that the predict function does not allow for extrapolation. Our test data set requires us to extrapolate, so we will not be able to measure the test set performance of the CGAM model. In this particular case, we have no choice but to fit a pseudo model.

First, we convert x3 from factor to integer.

Second, fit a CGAM with a shape constraint on x3_num.

```
## Call:
## cgam(formula = y ~ s.incr(x1) + s.decr(x2) + incr(x3_num), family = "binomial",
##
      data = train_data)
##
## Coefficients:
##
              Estimate StdErr z.value
                                         p.value
## (Intercept) -3.2362 0.1739 -18.611 < 2.2e-16 ***
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
## Null deviance: 935.1666 on 1999 degrees of freedom
## Residual deviance: 822.5902 on 1980.5 observed degrees of freedom
## Approximate significance of constrained components:
```

```
## edf mixture.of.Beta p.value
## s.incr(x1) 9 0.0330 <2e-16 ***
## s.decr(x2) 6 0.0169 <2e-16 ***
## incr(x3_num) 3 0.0078 2e-04 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## CIC: 0.4174

Third, generate pseudo data.
```

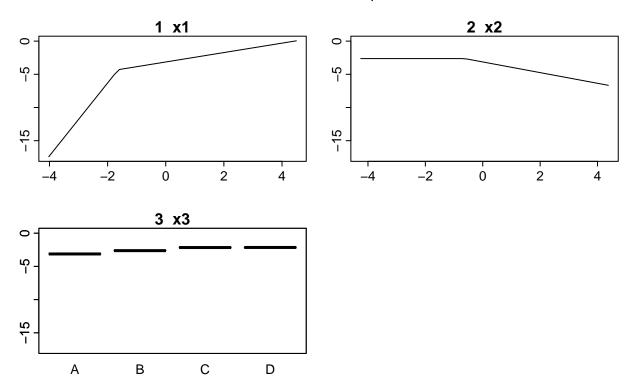
```
##
                           x1
                                             x2
                                                          xЗ
##
          :-16.911
                     Min.
                           :-3.31420
                                      Min.
                                              :-3.53202
                                                          A:1609
  Min.
  1st Qu.: -3.907
                     1st Qu.:-0.69239
                                       1st Qu.:-0.64836
                                                          B: 198
## Median : -3.164
                     Median : 0.02514
                                       Median :-0.01473
                                                          C:
                                                             98
## Mean : -3.236
                     Mean : 0.01329
                                       Mean : 0.01956
                                                             95
## 3rd Qu.: -2.412
                     3rd Qu.: 0.67314
                                       3rd Qu.: 0.70888
  Max.
          : -0.401
                     Max. : 3.78302
                                       Max. : 3.65809
       x3_num
##
## Min.
          :0.0000
## 1st Qu.:0.0000
## Median :0.0000
## Mean
          :0.3395
##
   3rd Qu.:0.0000
## Max.
          :3.0000
```

Next, fit a MARS model on the pseudo data. The threshold was set to 0.01 to get monotonicity. Increasing this threshold raises bias but lowers variance.

```
## Call: earth(formula=y~x1+x2+x3, data=pseudo_train_data_cgam, thresh=0.01)
##
##
                    coefficients
## (Intercept)
                      -3.8123682
## x3B
                       0.4997210
## x3C
                       0.9684266
## x3D
                       0.9780145
## h(-1.63511-x1)
                      -5.4866792
## h(x1--1.63511)
                       0.7094716
## h(x2 - 0.628166)
                      -0.8035952
##
```

```
## plotmo grid: x1 x2 x3
## 0.02514023 -0.01473112 A
```

MARS with CGAM Reponse



Check monotonicity.

```
df_x1$log_odds <- predict(mod_pseudo_cgam, newdata=df_x1)

df_x2$log_odds <- predict(mod_pseudo_cgam, newdata=df_x2)

success <- mono_check(df_x1$log_odds) & mono_check(df_x2$log_odds, FALSE)

success</pre>
```

[1] TRUE

Finally, check test set performance.

model	ROC-AUC
gam	0.7122259
mono_gam	0.7115072
pseudo_mono	0.7129432
scar	0.7032823
pseudo_scar	0.7061607
pseudo_cgam	0.7117960

perf(myPreds, MLmetrics::LogLoss, "Log Loss")

model	Log Loss
gam	0.2180729
mono_gam	0.2180345
pseudo_mono	0.2177323
scar	0.2226342
$pseudo_scar$	0.2210625
$pseudo_cgam$	0.2195729

Conclusion

Imposing shape constraints on data with high noise to signal ratios could greatly reduce variance (wiggliness). The scam package is highly effective in developing models with user-defined monotonicity constraints for each predictor. MARS can approximate the scam model with linear basis functions, which are easy to write down on paper and implement in an Excel formula.

The scar package also develops models with user-defined monotonicity constraints for each predictor. Non-linear least squares can approximate the scar model with sigmoid functions, which are easy to write down on paper (?SSfpl for equation) and implement in an Excel formula.