

# Unitary Time Changes of Stationary Processes Yield Oscillatory Processes

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## Abstract

A unitary time-change operator  $U_\theta$  is constructed for absolutely continuous, strictly increasing time reparametrizations  $\theta$ , acting on functions that are square-integrable over compact sets. Applying  $U_\theta$  to the Cramér spectral representation of a stationary process yields an oscillatory process in the sense of Priestley with oscillatory function  $\varphi_t(\lambda) = \sqrt{\dot{\theta}(t)} e^{i\lambda\theta(t)}$  and evolutionary spectrum  $dF_t(\lambda) = \dot{\theta}(t)dF(\lambda)$ . It is proved that sample paths of any non-degenerate second-order stationary process almost surely lie in  $L^2_{\text{loc}}(\mathbb{R})$ , making the operator applicable to typical realizations. A zero-localization measure  $d\mu(t) = \delta(Z(t))|\dot{Z}(t)|dt$  induces a Hilbert space  $L^2(\mu)$  on the zero set of an oscillatory process  $Z$ , and the multiplication operator  $(Lf)(t) = t f(t)$  has simple pure point spectrum equal to the zero crossing set of  $Z$ .

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## 1 Gaussian Processes

### 1.1 Definition

**Definition 1. (Gaussian process)** Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be a probability space and  $T$  a nonempty index set. A family  $\{X_t: t \in T\}$  of real-valued random variables on  $(\Omega, \mathcal{F}, \mathbb{P})$  is called a Gaussian process if for every finite subset  $F = \{t_1, \dots, t_n\} \subset T$  the random vector  $(X_{t_1}, \dots, X_{t_n})$  is multivariate normal (possibly degenerate). Equivalently, every finite linear combination  $\sum_{i=1}^n a_i X_{t_i}$  is either almost surely constant or Gaussian. The mean function is  $m(t) := \mathbb{E}[X_t]$  and the covariance kernel is

$$K(s, t) = \text{Cov}(X_s, X_t) \tag{1}$$

For any finite  $(t_i)_{i=1}^n \subset T$ , the matrix  $K_{ij} = K(t_i, t_j)$  is symmetric positive semidefinite, and a Gaussian process is completely determined in law by  $m$  and  $K$ .

**Definition 2.** *The canonical metric associated with a Gaussian process is*

$$\begin{aligned} d(s, t) &= \sqrt{\mathbb{E}[(X_s - X_t)^2]} \\ &= \sqrt{K(s, s) + K(t, t) - 2K(s, t)} \end{aligned} \quad (2)$$

## 1.2 Sample Path Realizations

**Definition 3.** *[Locally square-integrable functions] Define*

$$L^2_{\text{loc}}(\mathbb{R}) := \left\{ f: \mathbb{R} \rightarrow \mathbb{C}: \int_K |f(t)|^2 dt < \infty \text{ for every compact } K \subseteq \mathbb{R} \right\} \quad (3)$$

**Remark 4.** Every bounded measurable set in  $\mathbb{R}$  is compact or contained in a compact set; hence  $L^2_{\text{loc}}(\mathbb{R})$  contains functions that are square-integrable on every bounded interval, including functions with polynomial growth at infinity.

**Theorem 5.** *[Sample paths in  $L^2_{\text{loc}}(\mathbb{R})$ ] Let  $\{X(t)\}_{t \in \mathbb{R}}$  be a second-order stationary process with*

$$\sigma^2 := \mathbb{E}[X(t)^2] < \infty \quad (4)$$

*then, almost surely, every sample path  $t \mapsto X(t) \in L^2_{\text{loc}}(\mathbb{R})$ .*

**Proof.** Fix any bounded interval  $[a, b]$  and consider the random variable

$$Y_{[a, b]} := \int_a^b X(t)^2 dt \quad (5)$$

By stationarity and Fubini's theorem:

$$\begin{aligned} \mathbb{E}[Y_{[a, b]}] &= \mathbb{E}\left[\int_a^b X(t)^2 dt\right] = \int_a^b \mathbb{E}[X(t)^2] dt \\ &= \int_a^b \sigma^2 dt \\ &= \sigma^2(b - a) < \infty \end{aligned} \quad (6)$$

By Markov's inequality, for any  $M > 0$ :

$$P(Y_{[a, b]} > M) \leq \frac{\mathbb{E}[Y_{[a, b]}]}{M} = \frac{\sigma^2(b - a)}{M} \quad (7)$$

Taking  $M \rightarrow \infty$ , the conclusion is

$$P(Y_{[a, b]} < \infty) = 1 \quad (8)$$

i.e., almost surely the sample path is square-integrable on  $[a, b]$ . Since  $\mathbb{R}$  is the countable union of bounded intervals:

$$\mathbb{R} = \bigcup_{n=1}^{\infty} [-n, n] \quad (9)$$

by countable subadditivity of probability:

$$P\left(\bigcap_{n=1}^{\infty} \left\{ \int_{-n}^n X(t)^2 dt < \infty \right\}\right) = 1 \quad (10)$$

Now let  $K$  be any compact set. Then  $K$  is bounded, so

$$K \subseteq [-N, N] \quad (11)$$

for some  $N$ . Therefore:

$$\int_K X(t)^2 dt \leq \int_{-N}^N X(t)^2 dt < \infty \quad (12)$$

almost surely. This holds for every compact set  $K$ , so almost surely every sample path lies in  $L^2_{\text{loc}}(\mathbb{R})$ .  $\square$

### 1.3 Stationary processes

**Definition 6. [Cramér representation]** [1] A zero-mean stationary process  $X$  with spectral measure  $F$  admits the sample path representation

$$X(t) = \int_{\mathbb{R}} e^{i\lambda t} d\Phi(\lambda) \quad (13)$$

which has covariance

$$R_X(t-s) = \int_{\mathbb{R}} e^{i\lambda(t-s)} dF(\lambda) \quad (14)$$

### 1.4 Oscillatory Processes

A particularly tractable class of non-stationary Gaussian processes is that of the oscillatory processes as defined by Maurice Priestley in 1965[2].

**Definition 7. [Oscillatory process]** Let  $F$  be a finite nonnegative Borel measure on  $\mathbb{R}$ . Let

$$A_t \in L^2(F) \forall t \in \mathbb{R} \quad (15)$$

be the gain function and

$$\varphi_t(\lambda) = A_t(\lambda) e^{i\lambda t} \quad (16)$$

be the corresponding oscillatory function then an oscillatory process is a stochastic process which can be represented as

$$\begin{aligned} Z(t) &= \int_{\mathbb{R}} \varphi_t(\lambda) d\Phi(\lambda) \\ &= \int_{\mathbb{R}} A_t(\lambda) e^{i\lambda t} d\Phi(\lambda) \end{aligned} \quad (17)$$

where  $\Phi$  is a complex orthogonal random measure with spectral measure  $F$  which satisfies the relation

$$d\mathbb{E}[\Phi(\lambda)\overline{\Phi(\mu)}] = \delta(\lambda - \mu) dF(\lambda) \quad (18)$$

and has the corresponding covariance kernel

$$\begin{aligned} R_Z(t, s) &= \mathbb{E}[Z(t)\overline{Z(s)}] \\ &= \int_{\mathbb{R}} A_t(\lambda) \overline{A_s(\lambda)} e^{i\lambda(t-s)} dF(\lambda) \\ &= \int_{\mathbb{R}} \phi_t(\lambda) \overline{\phi_s(\lambda)} dF(\lambda) \end{aligned} \quad (19)$$

**Theorem 8. [Real-valuedness criterion for oscillatory processes]** Let  $Z$  be an oscillatory process with oscillatory function

$$\varphi_t(\lambda) = A_t(\lambda) e^{i\lambda t} \quad (20)$$

and spectral measure  $F$ . Then  $Z$  is real-valued if and only if

$$A_t(-\lambda) = \overline{A_t(\lambda)} \quad (21)$$

for  $F$ -almost every  $\lambda \in \mathbb{R}$ , equivalently

$$\varphi_t(-\lambda) = \overline{\varphi_t(\lambda)} \quad (22)$$

for  $F$ -almost every  $\lambda \in \mathbb{R}$ .

**Proof.** Assume  $Z$  is real-valued, i.e.  $Z(t) = \overline{Z(t)} \forall t \in \mathbb{R}$ . Writing its oscillatory representation,  $Z(t) = \int_{\mathbb{R}} A_t(\lambda) e^{i\lambda t} d\Phi(\lambda)$  and taking the complex conjugate gives  $\overline{Z(t)} = \int_{\mathbb{R}} \overline{A_t(\lambda)} e^{-i\lambda t} d\overline{\Phi(\lambda)}$ . For a real-valued process, the orthogonal random measure  $\Phi$  must satisfy  $d\overline{\Phi(\lambda)} = -d\Phi(\lambda)$  which ensures that the spec-

tral representation produces real values. Substituting this identity and using the substitution  $\mu = -\lambda$  it is shown that  $\overline{Z(t)} = \int_{\mathbb{R}} \overline{A_t(-\mu)} e^{i\mu t} d\Phi(\mu)$ . Since  $Z(t) = \overline{Z(t)}$ , comparison of the integrands (which are unique elements of  $L^2(F)$ ) yields  $A_t(\lambda) = \overline{A_t(-\lambda)}$  for  $F$ -a.e.  $\lambda$ . Equivalently, because the oscillatory function (16) is given by  $\varphi_t(\lambda) = A_t(\lambda) e^{i\lambda t}$  we have  $\varphi_t(-\lambda) = \overline{\varphi_t(\lambda)}$  for  $F$ -a.e.  $\lambda$ . Conversely, if  $A_t(-\lambda) = \overline{A_t(\lambda)}$  for  $F$ -a.e.  $\lambda$ , then the same substitution shows that  $\overline{Z(t)} = Z(t) \forall t \in \mathbb{R}$  so  $Z$  is real-valued.  $\square$

**Theorem 9. [Existence]** *Let  $F$  be an absolutely continuous spectral measure and the gain function*

$$A_t(\lambda) \in L^2(F) \forall \mathbb{R} \ni t < \infty \quad (23)$$

*be measurable in both time and frequency then the time-dependent spectral density is defined by*

$$\begin{aligned} S_t(\lambda) &= \int_{\mathbb{R}} |A_t(\lambda)|^2 dF(\lambda) < \infty \\ &= \int_{\mathbb{R}} |A_t(\lambda)|^2 S(\lambda) d\lambda \end{aligned} \quad (24)$$

*and there exists a complex orthogonal random measure  $\Phi$  with spectral measure  $F$  such that for each sample path  $\varpi \in \Theta$  in the space of sample paths having given covariance constituting the ensemble denoted  $\Theta$*

$$Z(t) = \int_{\mathbb{R}} A_t(\lambda) e^{i\lambda t} d\Phi(\lambda) \quad (25)$$

*is well-defined in  $L^2(\Omega)$  and has covariance  $R_Z$  as in (19) above.*

**Proof.** The proof proceeds by constructing the stochastic integral using the standard extension procedure. First, the integral is defined for simple functions of the form

$$g(\lambda) = \lim_{n \rightarrow \infty} \sum_{j=1}^n c_j \mathbf{1}_{E_j}(\lambda) \quad (26)$$

where  $\{E_j\}$  are disjoint Borel sets with  $F(E_j) < \infty$  and  $c_j \in \mathbb{C}$ :

$$\int_{\mathbb{R}} g(\lambda) d\Phi(\lambda) = \lim_{n \rightarrow \infty} \sum_{j=1}^n c_j \Phi(E_j) \quad (27)$$

For simple functions such as this, the isometry property holds:

$$\begin{aligned} \mathbb{E} \left[ \left| \int_{\mathbb{R}} g(\lambda) \Phi d(\lambda) \right|^2 \right] &= \mathbb{E} \left[ \lim_{n \rightarrow \infty} \left| \sum_{j=1}^n c_j \Phi(E_j) \right|^2 \right] \\ &= \lim_{n \rightarrow \infty} \sum_{j=1}^n \sum_{k=1}^n c_j \bar{c}_k \mathbb{E}[\Phi(E_j) \overline{\Phi(E_k)}] \\ &= \lim_{n \rightarrow \infty} \sum_{j=1}^n |c_j|^2 F(E_j) \\ &= \int_{\mathbb{R}} |g(\lambda)|^2 dF(\lambda) \end{aligned} \quad (28)$$

Since simple functions are dense in  $L^2(F)$ , the integral is extended by continuity  $\forall g \in L^2(F)$  since the oscillatory function (16) is defined by

$$\varphi_t(\lambda) = A_t(\lambda) e^{i\lambda t} \in L^2(F) \forall t \in \mathbb{R} \quad (29)$$

and  $A_t \in \cdot$ . Therefore

$$Z(t) = \int_{\mathbb{R}} \varphi_t(\lambda) d\Phi(\lambda) \quad (30)$$

is well-defined in  $L^2(\Omega)$ . The covariance is computed as:

$$\begin{aligned}
R_Z(t, s) &= \mathbb{E}[Z(t)\overline{Z(s)}] \\
&= \mathbb{E}\left[\int_{\mathbb{R}} \varphi_t(\lambda) d\Phi(\lambda) \int_{\mathbb{R}} \overline{\varphi_s(\mu)} d\overline{\Phi(\mu)}\right] \\
&= \int_{\mathbb{R}} \int_{\mathbb{R}} \varphi_t(\lambda) \overline{\varphi_s(\mu)} d\mathbb{E}[\Phi(\lambda)\overline{\Phi(\mu)}] \\
&= \int_{\mathbb{R}} \varphi_t(\lambda) \overline{\varphi_s(\lambda)} dF(\lambda) \\
&= \int_{\mathbb{R}} A_t(\lambda) \overline{A_s(\lambda)} e^{i\lambda(t-s)} dF(\lambda)
\end{aligned} \tag{31} \quad \square$$

## 2 Unitarily Time-Changed Stationary Processes

### 2.1 Unitary time-change operator $U_\theta$

**Definition 10. [Unitary time-change]** Let the time-scaling function  $\theta: \mathbb{R} \rightarrow \mathbb{R}$  be absolutely continuous, strictly increasing, and bijective, with  $\dot{\theta}(t) > 0$  almost everywhere and  $\dot{\theta}(t) = 0$  only on sets of Lebesgue measure zero. Define, for  $f$  measurable,

$$(U_\theta f)(t) = \sqrt{\dot{\theta}(t)} f(\theta(t)) \tag{32}$$

**Proposition 11. [Inverse unitary time-change]** The inverse of the unitary time-change operator  $U$  in Equation (32) is given by

$$(U_\theta^{-1} g)(s) = \frac{g(\theta^{-1}(s))}{\sqrt{\dot{\theta}(\theta^{-1}(s))}} \tag{33}$$

which is well-defined almost everywhere.

**Proof.** Since  $\dot{\theta}(t) = 0$  only on sets of measure zero, and  $\theta^{-1}$  maps sets of measure zero to sets of measure zero because of the fact that absolutely continuous bijective functions preserve measure-zero sets, the denominator  $\sqrt{\dot{\theta}(\theta^{-1}(s))}$  is positive almost everywhere. The expression is therefore well-defined almost everywhere, which suffices for defining an element of  $L^2_{\text{loc}}(\mathbb{R})$ .  $\square$

**Theorem 12. [Local unitarity on compact sets]** For every compact set  $K \subseteq \mathbb{R}$  and  $f \in L^2_{\text{loc}}(\mathbb{R})$ ,

$$\int_K |(U_\theta f)(t)|^2 dt = \int_{\theta(K)} |f(s)|^2 ds \tag{34}$$

Moreover,  $U_\theta^{-1}$  is the inverse of  $U_\theta$  on  $L^2_{\text{loc}}(\mathbb{R})$ .

**Proof.** Let  $f \in L^2_{\text{loc}}(\mathbb{R})$  and let  $K$  be any compact set. The local  $L^2$ -norm of  $U_\theta f$  over  $K$  is:

$$\begin{aligned}
\int_K |(U_\theta f)(t)|^2 dt &= \int_K \left| \sqrt{\dot{\theta}(t)} f(\theta(t)) \right|^2 dt \\
&= \int_K \dot{\theta}(t) |f(\theta(t))|^2 dt
\end{aligned} \tag{35}$$

Since  $\theta$  is absolutely continuous and strictly increasing, applying the change of variables  $s = \theta(t)$  gives

$$ds = \dot{\theta}(t) dt \tag{36}$$

almost everywhere. As  $t$  ranges over the compact set  $K$ ,  $s = \theta(t)$  ranges over  $\theta(K)$ , which is compact. Therefore:

$$\int_K \dot{\theta}(t) |f(\theta(t))|^2 dt = \int_{\theta(K)} |f(s)|^2 ds \tag{37}$$

To verify that  $U_\theta^{-1}$  is indeed the inverse, it is seen that:

$$\begin{aligned} (U_\theta^{-1} U_\theta f)(s) &= \left( U_\theta^{-1} \sqrt{\dot{\theta}(s)} f(\theta(s)) \right)(s) \\ &= \frac{\sqrt{\dot{\theta}(\theta^{-1}(s))}}{\sqrt{\dot{\theta}(\theta^{-1}(s))}} f(\theta(\theta^{-1}(s))) \quad \forall f \in L_{\text{loc}}^2(\mathbb{R}) \\ &= f(s) \end{aligned} \quad (38)$$

since

$$\theta(\theta^{-1}(s)) = s \quad (39)$$

and similarly:

$$\begin{aligned} (U_\theta U_\theta^{-1} g)(t) &= \sqrt{\dot{\theta}(t)} (U_\theta^{-1} g)(\theta(t)) \\ &= \frac{\sqrt{\dot{\theta}(t)}}{\sqrt{\dot{\theta}(\theta^{-1}(\theta(t)))}} g(\theta^{-1}(\theta(t))) \quad \forall g \in L_{\text{loc}}^2(\mathbb{R}) \\ &= \frac{\sqrt{\dot{\theta}(t)}}{\sqrt{\dot{\theta}(t)}} g(t) \\ &= g(t) \end{aligned} \quad (40)$$

since

$$\theta^{-1}(\theta(t)) = t \quad (41)$$

Therefore

$$\begin{aligned} (U_\theta U_\theta^{-1} f)(t) &= (U_\theta^{-1} U_\theta f)(t) \\ &= f(t) \end{aligned} \quad (42)$$

on  $L_{\text{loc}}^2(\mathbb{R})$ . □

### 2.1.1 Inverse Filter for Unitary Time Transformations

**Theorem 13.** *[Inverse Filter for Unitary Time Transformations] Let  $\theta: \mathbb{R} \rightarrow \mathbb{R}$  be absolutely continuous, strictly increasing, and bijective with  $\theta'(t) > 0$  almost everywhere. Let  $Y(u)$  be a stationary process with unit variance, and define*

$$Z(t) = \sqrt{\dot{\theta}(t)} Y(\theta(t)) \quad (43)$$

*as the oscillatory process obtained by the unitary time transformation. Then:*

1. *The forward filter kernel is*

$$h(t, u) = \sqrt{\dot{\theta}(t)} \delta(u - \theta(t)) \quad (44)$$

2. *The inverse filter kernel is*

$$g(t, s) = \frac{\delta(s - \theta^{-1}(t))}{\sqrt{\dot{\theta}(\theta^{-1}(t))}} \quad (45)$$

3. *The composition  $(g \circ h)$  recovers the identity:*

$$\begin{aligned} Y(t) &= \int_{\mathbb{R}} g(t, s) Z(s) ds \\ &= \frac{Z(\theta^{-1}(t))}{\sqrt{\dot{\theta}(\theta^{-1}(t))}} \end{aligned} \quad (46)$$

**Proof.** Parts 1-3 are established in sequence. Let

$$v = \theta(s) \text{ and } s = \theta^{-1}(v) \quad (47)$$

so that

$$ds = \frac{1}{\dot{\theta}(\theta^{-1}(v))} dv \quad (48)$$

and

$$\delta(\theta^{-1}(v) - \theta^{-1}(t)) = \dot{\theta}(\theta^{-1}(t)) \delta(v - t) \quad (49)$$

then substitute the inverse filter in Equation (44)

$$g(t, s) = \frac{\delta(s - \theta^{-1}(t))}{\sqrt{\dot{\theta}(\theta^{-1}(t))}} \quad (50)$$

and unitarily time-changed stationary process operator representation (43)

$$\begin{aligned} Z(t) &= (U_\theta Y)(t) \\ &= \sqrt{\dot{\theta}(t)} Y(\theta(t)) \end{aligned} \quad (51)$$

to verify that each of the equivalent expressions

$$\begin{aligned} Y(t) &= \int_{\mathbb{R}} g(t, s) Z(s) ds \\ &= \int_{\mathbb{R}} \frac{\delta(s - \theta^{-1}(t))}{\sqrt{\dot{\theta}(\theta^{-1}(t))}} Z(s) ds \\ &= \int_{\mathbb{R}} \frac{\delta(s - \theta^{-1}(t))}{\sqrt{\dot{\theta}(\theta^{-1}(t))}} (U_\theta Y)(s) ds \\ &= \int_{\mathbb{R}} \frac{\delta(s - \theta^{-1}(t))}{\sqrt{\dot{\theta}(\theta^{-1}(t))}} \sqrt{\dot{\theta}(s)} Y(\theta(s)) ds \\ &= \int_{\mathbb{R}} \frac{\delta(\theta^{-1}(v) - \theta^{-1}(t))}{\sqrt{\dot{\theta}(\theta^{-1}(t))}} \frac{\sqrt{\dot{\theta}(\theta^{-1}(v))}}{\dot{\theta}(\theta^{-1}(v))} Y(v) dv \\ &= \int_{\mathbb{R}} \frac{\delta(\theta^{-1}(v) - \theta^{-1}(t))}{\sqrt{\dot{\theta}(\theta^{-1}(t))} \sqrt{\dot{\theta}(\theta^{-1}(v))}} Y(v) dv \\ &= \int_{\mathbb{R}} \frac{\dot{\theta}(\theta^{-1}(t)) \delta(v - t)}{\sqrt{\dot{\theta}(\theta^{-1}(t))} \sqrt{\dot{\theta}(\theta^{-1}(v))}} Y(v) dv \\ &= \frac{\dot{\theta}(\theta^{-1}(t))}{\sqrt{\dot{\theta}(\theta^{-1}(t))} \sqrt{\dot{\theta}(\theta^{-1}(t))}} Y(t) \\ &= \frac{\dot{\theta}(\theta^{-1}(t))}{\left(\sqrt{\dot{\theta}(\theta^{-1}(t))}\right)^2} Y(t) \\ &= \frac{\dot{\theta}(\theta^{-1}(t))}{\dot{\theta}(\theta^{-1}(t))} Y(t) \\ &= Y(t) \end{aligned} \quad (52) \quad \square$$

## 2.2 Transformation of Stationary $\rightarrow$ Oscillatory Processes via $U_\theta$

**Theorem 14.** *[Unitary time change yields oscillatory process] Let  $X$  be zero-mean stationary as in Definition 6. For scaling function  $\theta$  as in Definition 10, define*

$$\begin{aligned} Z(t) &= (U_\theta X)(t) \\ &= \sqrt{\dot{\theta}(t)} X(\theta(t)) \end{aligned} \quad (53)$$

Then  $Z$  is a realization of an oscillatory process with oscillatory function

$$\varphi_t(\lambda) = \sqrt{\dot{\theta}(t)} e^{i\lambda\theta(t)} \quad (54)$$

gain function

$$A_t(\lambda) = \sqrt{\dot{\theta}(t)} e^{i\lambda(\theta(t)-t)} \quad (55)$$

and covariance

$$\begin{aligned} R_Z(t, s) &= \mathbb{E}[Z(t)\overline{Z(s)}] \\ &= \mathbb{E}\left[\sqrt{\dot{\theta}(t)} X(\theta(t)) \sqrt{\dot{\theta}(s)} \overline{X(\theta(s))}\right] \\ &= \sqrt{\dot{\theta}(t)\dot{\theta}(s)} \mathbb{E}[X(\theta(t))\overline{X(\theta(s))}] \\ &= \sqrt{\dot{\theta}(t)\dot{\theta}(s)} R_X(\theta(t) - \theta(s)) \\ &= \sqrt{\dot{\theta}(t)\dot{\theta}(s)} \int_{\mathbb{R}} e^{i\lambda(\theta(t)-\theta(s))} dF(\lambda) \end{aligned} \quad (56)$$

**Proof.** Applying the unitary time change operator to the spectral representation of  $X(t)$ :

$$\begin{aligned} Z(t) &= (U_\theta X)(t) \\ &= \sqrt{\dot{\theta}(t)} X(\theta(t)) \\ &= \sqrt{\dot{\theta}(t)} \int_{\mathbb{R}} e^{i\lambda\theta(t)} d\Phi(\lambda) \\ &= \int_{\mathbb{R}} \sqrt{\dot{\theta}(t)} e^{i\lambda\theta(t)} d\Phi(\lambda) \\ &= \int_{\mathbb{R}} \varphi_t(\lambda) d\Phi(\lambda) \end{aligned} \quad (57)$$

where

$$\varphi_t(\lambda) = \sqrt{\dot{\theta}(t)} e^{i\lambda\theta(t)} \quad (58)$$

To verify this constitutes an oscillatory representation according to Definition 7,  $\varphi_t(\lambda)$  has the form  $A_t(\lambda)e^{i\lambda t}$ :

$$\begin{aligned} \varphi_t(\lambda) &= \sqrt{\dot{\theta}(t)} e^{i\lambda\theta(t)} \\ &= \sqrt{\dot{\theta}(t)} e^{i\lambda(\theta(t)-t)} e^{i\lambda t} \\ &= A_t(\lambda) e^{i\lambda t} \end{aligned} \quad (59)$$

where

$$A_t(\lambda) = \sqrt{\dot{\theta}(t)} e^{i\lambda(\theta(t)-t)} \quad (60)$$

Since  $\dot{\theta}(t) \geq 0$  almost everywhere and  $\dot{\theta}(t) = 0$  only on sets of measure zero,  $A_t(\lambda)$  is well defined almost



everywhere. Moreover,  $A_t \in L^2(F)$  for each  $t$  since:

$$\begin{aligned}
 \int_{\mathbb{R}} |A_t(\lambda)|^2 dF(\lambda) &= \int_{\mathbb{R}} \left| \sqrt{\dot{\theta}(t)} e^{i\lambda(\theta(t)-t)} \right|^2 dF(\lambda) \\
 &= \int_{\mathbb{R}} \dot{\theta}(t) |e^{i\lambda(\theta(t)-t)}|^2 dF(\lambda) \\
 &= \dot{\theta}(t) \int_{\mathbb{R}} dF(\lambda) \\
 &= \dot{\theta}(t) F(\mathbb{R}) < \infty
 \end{aligned} \tag{61}$$

where  $|e^{i\alpha}| = 1$  for all real  $\alpha$  is used. The covariance (56) is computed by substituting the spectral representation and applying Fubini's theorem to interchange the order of operations.  $\square$

**Corollary 15.** *[Evolutionary spectrum of unitarily time-changed stationary process][2] The evolutionary spectrum, also called the time-varying spectral density, is*

$$\begin{aligned}
 dF_t(\lambda) &= |A_t(\lambda)|^2 dF(\lambda) \\
 &= \dot{\theta}(t) dF(\lambda)
 \end{aligned} \tag{62}$$

## 2.3 Covariance operator conjugation

**Proposition 16.** *[Operator conjugation] Let*

$$(T_K f)(t) := \int_{\mathbb{R}} K(|t-s|) f(s) ds \tag{63}$$

with stationary kernel

$$K(h) = \int_{\mathbb{R}} e^{i\lambda h} dF(\lambda) \tag{64}$$

Define the transformed kernel

$$K_{\theta}(s, t) := \sqrt{\dot{\theta}(t) \dot{\theta}(s)} K(|\theta(t) - \theta(s)|) \tag{65}$$

then the corresponding integral covariance operator is conjugated  $\forall f \in L^2_{\text{loc}}(\mathbb{R})$  by

$$\begin{aligned}
 (T_{K_{\theta}} f)(t) &= \int_{\mathbb{R}} K_{\theta}(s, t) f(s) ds \\
 &= (U_{\theta} T_K U_{\theta}^{-1} f)(t)
 \end{aligned} \tag{66}$$

**Proof.** For any  $g \in L^2_{\text{loc}}(\mathbb{R})$ , compute:

$$\begin{aligned}
 ((U_{\theta} T_K U_{\theta}^{-1}) g)(t) &= (U_{\theta} (T_K U_{\theta}^{-1} g))(t) \\
 &= \sqrt{\dot{\theta}(t)} (T_K U_{\theta}^{-1} g)(\theta(t)) \\
 &= \sqrt{\dot{\theta}(t)} \int_{\mathbb{R}} K(|\theta(t) - \theta(s)|) (U_{\theta}^{-1} g)(\theta(s)) \dot{\theta}(s) ds \\
 &= \sqrt{\dot{\theta}(t)} \int_{\mathbb{R}} K(|\theta(t) - \theta(s)|) \frac{g(s)}{\sqrt{\dot{\theta}(s)}} \dot{\theta}(s) ds \\
 &= \sqrt{\dot{\theta}(t)} \int_{\mathbb{R}} K(|\theta(t) - \theta(s)|) g(s) \sqrt{\dot{\theta}(s)} ds \\
 &= \int_{\mathbb{R}} \sqrt{\dot{\theta}(t) \dot{\theta}(s)} K(|\theta(t) - \theta(s)|) g(s) ds \\
 &= \int_{\mathbb{R}} K_{\theta}(t, s) g(s) ds \\
 &= (T_{K_{\theta}} g)(t)
 \end{aligned}$$

$\square$

### 3 Zero Localization

The construction

$$\text{stationary } X \xrightarrow{U_\theta} \text{oscillatory } Z \xrightarrow{\mu=\delta(Z)|\dot{Z}| dt} L^2(\mu) \xrightarrow{L:tf(t)} (L, \sigma(L)) \quad (67)$$

produces a self-adjoint operator whose eigenvalues equal the zero set of the realization sample path realization  $Z(t)$  from the ensemble of possible sample path functions having the given covariance structure and whose spectrum equals the closure of the zero set, determined by the choice of time-change  $\theta(t)$ , spectral measure  $F(\lambda)$ , and complex orthogonal random measure  $\Phi(\lambda)$  which uniquely corresponds to a given sample path from the ensemble.

#### 3.1 Zero localization measure

**Definition 17. [Zero localization measure]** Let  $Z$  be real-valued with  $Z \in C^1(\mathbb{R})$  having only simple zeros

$$Z(t_0) = 0 \Rightarrow \dot{Z}(t_0) \neq 0 \quad (68)$$

Define, for Borel  $B \subset \mathbb{R}$ ,

$$\mu(B) = \int_{\mathbb{R}} \mathbf{1}_B(t) \delta(Z(t)) |\dot{Z}(t)| dt \quad (69)$$

**Theorem 18. [Atomicity on the zero set]** For every  $\phi \in C_c^\infty(\mathbb{R})$ ,

$$\int_{\mathbb{R}} \phi(t) \delta(Z(t)) |\dot{Z}(t)| dt = \sum_{t_0: Z(t_0)=0} \phi(t_0) \quad (70)$$

hence

$$\mu(t) = \sum_{t_0: Z(t_0)=0} \delta_{t_0}(t) \quad (71)$$

**Proof.** Since all zeros of  $Z$  are simple and  $Z \in C^1(\mathbb{R})$ , by the inverse function theorem each zero  $t_0$  is isolated. Near each zero  $t_0$ ,  $Z$  is locally monotonic, so the one-dimensional change of variables formula for the Dirac delta can be applied. Specifically, near  $t_0$  where  $Z(t_0) = 0$  and  $\dot{Z}(t_0) \neq 0$ , locally

$$Z(t) = (t - t_0) \dot{Z}(t_0) + O((t - t_0)^2) \quad (72)$$

holds. The distributional identity for the Dirac delta under smooth changes of variables gives:

$$\delta(Z(t)) = \sum_{t_0: Z(t_0)=0} \frac{\delta(t - t_0)}{|\dot{Z}(t_0)|} \quad (73)$$

Therefore:

$$\begin{aligned} \int_{\mathbb{R}} \phi(t) \delta(Z(t)) |\dot{Z}(t)| dt &= \int_{-\infty}^{\infty} \phi(t) |\dot{Z}(t)| \sum_{t_0: Z(t_0)=0} \frac{\delta(t - t_0)}{|\dot{Z}(t_0)|} dt \\ &= \sum_{t_0: Z(t_0)=0} \int_{\mathbb{R}} \phi(t) \frac{|\dot{Z}(t)| \delta(t - t_0)}{|\dot{Z}(t_0)|} dt \\ &= \sum_{t_0: Z(t_0)=0} \frac{|\dot{Z}(t_0)|}{|\dot{Z}(t_0)|} \phi(t_0) \\ &= \sum_{t_0: Z(t_0)=0} \phi(t_0) \end{aligned} \quad (74)$$

This shows that  $\mu$  is the discrete measure

$$\mu(t) = \sum_{t_0: Z(t_0)=0} \delta_{t_0}(t) \quad (75)$$

assigning unit mass to each zero.  $\square$

### 3.2 Hilbert space on zeros and multiplication operator

**Definition 19.** [*Hilbert space on the zero set*] Let  $\mathcal{H} = L^2(\mu)$  with inner product

$$\langle f, g \rangle = \int_{-\infty}^{\infty} f(t) \overline{g(t)} d\mu(t) \quad (76)$$

**Proposition 20.** [*Atomic structure*] Let

$$\mu = \sum_{t_0: Z(t_0)=0} \delta_{t_0} \quad (77)$$

then

$$\mathcal{H} \cong \left\{ f: \{t_0: Z(t_0)=0\} \rightarrow \mathbb{C}: \sum_{t_0: Z(t_0)=0} |f(t_0)|^2 < \infty \right\} \cong \ell^2 \quad (78)$$

with orthonormal basis  $\{e_{t_0}\}_{t_0: Z(t_0)=0}$  where

$$e_{t_0}(t_1) = \delta_{t_0}(t_1) \quad (79)$$

**Proof.** By the atomic form of  $\mu$ , for any  $f \in L^2(\mu)$ :

$$\|f\|_{\mathcal{H}}^2 = \int |f(t)|^2 d\mu(t) \quad (80)$$

$$= \int |f(t)|^2 \sum_{t_0: Z(t_0)=0} \delta_{t_0}(t) \quad (81)$$

$$= \sum_{t_0: Z(t_0)=0} |f(t_0)|^2 \quad (82)$$

This shows the isomorphism with  $\ell^2$  where the functions  $e_{t_0}$  defined by

$$e_{t_0}(t_1) = \delta_{t_0}(t_1) \quad (83)$$

satisfy the relations

$$\begin{aligned} \langle e_{t_0}, e_{t_1} \rangle &= \int e_{t_0}(t) \overline{e_{t_1}(t)} d\mu(t) \\ &= \sum_{t: Z(t)=0} \delta_{t_0}(t) \delta_{t_1}(t) \\ &= \delta_{t_0}(t_1) \\ &= \delta_{t_1}(t_0) \end{aligned} \quad (84)$$

thus forming an orthonormal set. Thus, any  $f(t) \in \mathcal{H}$  can be written as

$$f(t) = \sum_{t_0: Z(t_0)=0} f(t_0) e_{t_0}(t) \quad (85)$$

proving they form a basis.  $\square$

**Definition 21.** [*Multiplication operator*] Define the linear operator

$$L: \mathcal{D}(L) \subset \mathcal{H} \rightarrow \mathcal{H} \quad (86)$$

by

$$(Lf)(t) = t f(t) \quad (87)$$

on the support of  $\mu$  with domain

$$\mathcal{D}(L) := \left\{ f \in \mathcal{H}: \int |t f(t)|^2 d\mu(t) < \infty \right\} \quad (88)$$

**Theorem 22.** *[Self-adjointness and spectrum]  $L$  is self-adjoint on  $\mathcal{H}$  and has pure point, simple spectrum*

$$\sigma(L) = \overline{\{t \in \mathbb{R}: Z(t) = 0\}} \quad (89)$$

*with eigenvalues  $\lambda = t_0$  for each zero  $t_0$  and corresponding eigenvectors  $e_{t_0}$ .*

**Proof.** First, self-adjointness is verified. For  $f, g \in \mathcal{D}(L)$ :

$$\begin{aligned} \langle Lf, g \rangle &= \int (Lf)(t) \overline{g(t)} d\mu(t) \\ &= \int t f(t) \overline{g(t)} d\mu(t) \\ &= \int f(t) \overline{t g(t)} d\mu(t) \\ &= \int f(t) \overline{(Lg)(t)} d\mu(t) \\ &= \langle f, Lg \rangle \end{aligned} \quad (90)$$

Thus  $L$  is symmetric and acts as

$$(Lf)(t_0) = t_0 f(t_0) \quad (91)$$

for each  $t_0$  in the atomic representation where

$$Z(t_0) = 0 \quad (92)$$

This is unitarily equivalent to the diagonal operator on  $\ell^2$  with diagonal entries

$$\{t_0: Z(t_0) = 0\} \quad (93)$$

Such diagonal operators are self-adjoint. For the spectrum calculation:

$$L e_{t_0} = t_0 e_{t_0} \forall \{t_0: Z(t_0) = 0\} \quad (94)$$

holds, so each  $t_0$  is an eigenvalue of  $L$  with eigenvector  $e_{t_0}$  and since  $\{e_{t_0}\}$  forms an orthonormal basis,  $L$  has pure point spectrum. The spectrum of a diagonal operator equals the closure of the set of diagonal entries, hence

$$\sigma(L) = \overline{\{t_0: Z(t_0) = 0\}} \quad (95)$$

The eigenvalues are simple. □

### 3.3 Regularity and Simplicity of Sample Path Zero Crossings

TODO: insert the fundamental theorem on the non-tangency of zero crossings so that it doesn't have to be assumed but is in fact a fundamental theorem of non-degenerate Gaussian processes

**Definition 23.** *[Regularity and simplicity] Assume  $Z \in C^1(\mathbb{R})$  and every zero is simple:*

$$Z(t_0) = 0 \Rightarrow \dot{Z}(t_0) \neq 0 \quad (96)$$

**Lemma 24.** *[Local finiteness and delta decomposition] Under Definition 23, zeros are locally finite and*

$$\delta(Z(t)) = \sum_{t_0: Z(t_0)=0} \frac{\delta(t - t_0)}{|\dot{Z}(t_0)|} \quad (97)$$

*whence*

$$\mu = \sum_{t_0: Z(t_0)=0} \delta_{t_0} \quad (98)$$

**Proof.** Since  $Z \in C^1(\mathbb{R})$  and  $\dot{Z}(t_0) \neq 0$  at each zero  $t_0$ , the inverse function theorem implies that  $Z$  is locally invertible near each zero. Specifically, there exists a neighborhood  $U_{t_0}$  of  $t_0$  such that  $Z|_{U_{t_0}}$  is strictly monotonic and invertible.

This implies zeros are isolated: if  $Z(t_0) = 0$  and  $\dot{Z}(t_0) \neq 0$ , then there exists  $\epsilon > 0$  such that  $Z(t) \neq 0$  for  $0 < |t - t_0| < \epsilon$ . Therefore zeros are locally finite (finitely many in any bounded interval).

For the distributional identity, the one-dimensional change of variables formula for the Dirac delta is considered. If  $g: I \rightarrow \mathbb{R}$  is  $C^1$  on interval  $I$  with  $\dot{g}(x) \neq 0$  for all  $x \in I$ , then

$$\delta(g(x)) = \sum_{x_0: g(x_0)=0} \frac{\delta(x - x_0)}{|\dot{g}(x_0)|} \quad (99)$$

Applying this locally around each zero  $t_0$  of  $Z$ , and since zeros are isolated, the local results can be patched together to obtain the global identity:

$$\delta(Z(t)) = \sum_{t_0: Z(t_0)=0} \frac{\delta(t - t_0)}{|\dot{Z}(t_0)|} \quad (100)$$

Consequently:

$$\begin{aligned} d\mu(t) &= \delta(Z(t)) |\dot{Z}(t)| dt \\ &= \sum_{t_0: Z(t_0)=0} \frac{|\dot{Z}(t)|}{|\dot{Z}(t_0)|} \delta(t - t_0) dt \\ &= \sum_{t_0: Z(t_0)=0} \delta_{t_0}(dt) \end{aligned} \quad (101)$$

where the last equality uses the fact that

$$\frac{|\dot{Z}(t_0)|}{|\dot{Z}(t_0)|} = 1 \quad (102)$$

when evaluating at  $t = t_0$ . □

### 3.4 The Kac-Rice Formula For The Expected Zero Counting Function

**Theorem 25. (Kac-Rice Formula for Zero Crossings)** Let  $Z(t)$  be a centered Gaussian process on  $[a, b]$  with covariance  $K(s, t) = \mathbb{E}[Z(s)Z(t)]$  then the expected number of zeros in  $[a, b]$  is

$$\mathbb{E}[N_{[a,b]}] = \int_a^b \sqrt{\frac{2}{\pi}} \frac{\sqrt{K(t, t) K_{\dot{Z}}(t, t) - K_{Z, \dot{Z}}(t, t)^2}}{K(t, t)} dt \quad (103)$$

where

$$K(t, t) = \mathbb{E}[Z(t)^2] \quad (104)$$

$$K_{\dot{Z}}(t, t) = -\partial_s^2 \partial_t K(s, t)|_{s=t} \quad (105)$$

and

$$K_{Z, \dot{Z}}(t, t) = \partial_s K(s, t)|_{s=t} \quad (106)$$

**Proof.**

The exact zero counting function is

$$N_{[a,b]} = \int_a^b \delta(Z(t)) |\dot{Z}(t)| dt \quad (107)$$

so

$$\begin{aligned} \mathbb{E}[N_{[a,b]}] &= \int_a^b \mathbb{E}[\delta(Z(t)) |\dot{Z}(t)|] dt \\ &= \int_a^b \int_{-\infty}^{\infty} |v| p_{Z, \dot{Z}}(0, v) dv dt \end{aligned} \quad (108)$$

The vector  $(Z(t), \dot{Z}(t))$  is bivariate Gaussian with covariance matrix

$$\Sigma = \begin{pmatrix} K(t, t) & K_{Z, \dot{Z}}(t, t) \\ K_{Z, \dot{Z}}(t, t) & K_{\dot{Z}}(t, t) \end{pmatrix} \quad (109)$$

whose determinant is given by

$$\det \Sigma = K(t, t) K_{\dot{Z}}(t, t) - K_{Z, \dot{Z}}(t, t)^2 \quad (110)$$

the inverse of which satisfies

$$\Sigma_{22}^{-1} = \frac{K(t, t)}{\det \Sigma} \quad (111)$$

yielding

$$p_{Z, \dot{Z}}(0, v) = \frac{1}{\sqrt{2\pi K(t, t)}} \cdot \frac{e^{-\frac{K(t, t)v^2}{2\det \Sigma}}}{\sqrt{2\pi \det \Sigma / K(t, t)}} \quad (112)$$

which factorizes as  $p_Z(0) \cdot p_{\dot{Z}|Z}(v|0)$  where

$$p_Z(0) = \frac{1}{\sqrt{2\pi K(t, t)}} \quad (113)$$

and

$$\dot{Z}|Z=0 \sim \mathcal{N}(0, \det \Sigma / K(t, t)) \quad (114)$$

For zero-mean Gaussian  $Y \sim \mathcal{N}(0, \sigma^2)$ , direct integration gives

$$\begin{aligned} \mathbb{E}[|Y|] &= 2 \int_0^\infty \frac{y}{\sqrt{2\pi\sigma^2}} e^{-y^2/(2\sigma^2)} dy \\ &= \frac{2\sigma}{\sqrt{2\pi}} \int_0^\infty e^{-u} du \\ &= \sqrt{\frac{2}{\pi}} \sigma \end{aligned} \quad (115)$$

so that combining results yields

$$\begin{aligned} \int_{-\infty}^\infty |v| p_{Z, \dot{Z}}(0, v) dv &= \frac{\sqrt{\frac{2}{\pi}} \sqrt{\frac{\det \Sigma}{K(t, t)}}}{\sqrt{2\pi K(t, t)}} \\ &= \sqrt{\frac{2}{\pi}} \frac{\sqrt{\det \Sigma}}{K(t, t)} \end{aligned} \quad (116) \quad \square$$

**Theorem 26. [Expected Zero-Counting Function]** Let  $\theta \in \mathcal{F}$  and let

$$K(t, s) = \text{cov}(Z(t), Z(s)) \quad (117)$$

be twice differentiable at  $s=0$  and  $t=0$  then expected number of zeros of the process  $Z(t)$  in  $[a, b]$  is

$$\mathbb{E}[N_{[a, b]}] = \sqrt{-K(0)} (\theta(b) - \theta(a)) \quad (118)$$

**Proof.** The covariance function of the time-changed process is

$$\begin{aligned} K_\theta(s, t) &= \text{cov}(Z(t), Z(s)) \\ &= \sqrt{\dot{\theta}(s)\dot{\theta}(t)} K(|\theta(t) - \theta(s)|) \end{aligned} \quad (119)$$

For the zero-crossing analysis, consider the normalized process. By the Kac-Rice formula:

$$\mathbb{E}[N_{[a, b]}] = \int_a^b \sqrt{-\lim_{s \rightarrow t} \frac{\partial^2}{\partial s \partial t} K_\theta(s, t)} dt \quad (120)$$

Computing the mixed partial derivative:

$$\frac{\partial}{\partial t} K_{\theta}(s, t) = \frac{1}{2} \frac{\dot{\theta}(t)}{\sqrt{\theta(t)}} \sqrt{\dot{\theta}(s)} K(|\theta(t) - \theta(s)|) + \sqrt{\theta(s)} \sqrt{\dot{\theta}(t)} \dot{K}(|\theta(t) - \theta(s)|) \operatorname{sgn}(\theta(t) - \theta(s)) \dot{\theta}(t) \quad (121)$$

Taking the limit as  $s \rightarrow t$  and using the fact that  $\dot{K}(0) = 0$  for stationary processes:

$$\lim_{s \rightarrow t} \frac{\partial^2}{\partial s \partial t} K_{\theta}(s, t) = \lim_{s \rightarrow t} \dot{\theta}(s) \dot{\theta}(t) \ddot{K}(0) = \dot{\theta}(t)^2 \ddot{K}(0) \quad (122)$$

Substituting into the Kac-Rice formula we have

$$\begin{aligned} \mathbb{E}[N_{[a,b]}] &= \int_a^b \sqrt{-\dot{\theta}(t)^2 \ddot{K}(0)} \, dt \\ &= \sqrt{-\ddot{K}(0)} \int_a^b \dot{\theta}(t) \, dt \\ &= \sqrt{-\ddot{K}(0)} (\theta(b) - \theta(a)) \end{aligned} \quad (123)$$

since  $\dot{\theta}(t) \geq 0$  almost everywhere. □

## Bibliography

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