

Homework 5 DS-GA 1002

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Problem 1.

(a) Assuming that each bet is independent of others, Let X_i be the Bernoulli distribution with $X_i = 1$ if getting the right number. Then $X(x) = \sum_{i=1}^{100} X_i \sim \text{Binomial}(100, x)$ with $p = \frac{1}{38}$

By CLT, $X \rightarrow^d N(\frac{100}{38}, 100 \frac{37}{38^2}) \approx N(2.632, 2.562)$

If Bob is making money, let the number of winning be x , x must satisfy $36x - 100 > 0, x > \frac{100}{36}$

$$P(X \geq \frac{100}{36}) = 1 - P(X \leq \frac{100}{36}) = 1 - P(\frac{X - \frac{100}{38}}{\sqrt{100 \frac{37}{38^2}}} \leq \frac{\frac{100}{36} - \frac{100}{38}}{\sqrt{100 \frac{37}{38^2}}}) = 1 - P(Z < 0.2302) \approx 0.4636$$

(b) Let Y_i be the Bernoulli distribution with $Y_i = 1$ if getting the right color. Then $Y(y) = \sum_{i=1}^{100} Y_i \sim \text{Binomial}(100, y)$ with $p = \frac{18}{38}$

By CLT, $Y \rightarrow^d N(\frac{100 \times 18}{38}, 100 \frac{18 \times 20}{38^2}) \approx N(47.368, 24.931)$

If Bob is making money, let the number of winning be y , y must satisfy $2y - 100 > 0, y > 50$

$$P(Y > 50) = 1 - P(Y \leq 50) = 1 - P(\frac{Y - \frac{100 \times 18}{38}}{\sqrt{100 \frac{18 \times 20}{38^2}}} \leq \frac{50 - \frac{100 \times 18}{38}}{\sqrt{100 \frac{18 \times 20}{38^2}}}) = 1 - P(Z < 0.7273) \approx 0.2991$$

(c) By CLT, for n goes to infinity, for betting on numbers, the expectation is $\frac{100}{38} \times 36 - 100 = -\frac{100}{19}$
for betting on color, the expectation is $\frac{100 \times 18}{38} \times 2 - 100 = -\frac{100}{19}$

Therefore, he is expected to lose money finally.

However, $sd(36X) \approx 78 > sd(2Y) \approx 78 \approx 9$

betting on colors is asymptotically better

$$(d) P(X = x | x \geq \frac{100}{36}) = \frac{P(X=x)}{0.4636}, E(X | x \geq \frac{100}{36}) = \int_{\frac{100}{36}}^{\infty} \frac{xf(x)}{0.4636} dx$$

$$= \int_{\frac{100}{36}}^{\infty} \frac{xe^{-\frac{(x-2.632)^2}{2 \times 2.562}}}{\sqrt{2 \times 2.562 \times \pi}} \times \frac{1}{0.4636} dx \approx 4.0043$$

The expected gain is $4.0043 \times 36 - 100 \approx 44.1548$

$$P(Y = y | y \geq 50) = \frac{P(Y=y)}{0.2991}, E(Y | y \geq 50) = \int_{50}^{\infty} \frac{yf(y)}{0.2991} dy = \int_{50}^{\infty} \frac{ye^{-\frac{(y-47.368)^2}{2 \times 24.931}}}{\sqrt{2 \times 24.931 \times \pi}} \times \frac{1}{0.2991} dy \approx 53.156$$

The expected gain is $53.156 \times 2 - 100 \approx 6.312$

Therefore, he will choose betting on the numbers. The expected gain is 44.1548

Problem 2.

(a) Since we know that the maximum of weight is 1400lbs. the max sd is 1400.

By the confidence interval for variables with bounded variance, we know that:

$$P(\mu \in [\bar{X}_n - \frac{b}{\sqrt{\alpha n}}, \bar{X}_n + \frac{b}{\sqrt{\alpha n}}]) \geq 1 - \alpha, \sigma x_i \leq b$$

we set $b = 1400$, and $\alpha = 0.05$

$$\frac{1400}{\sqrt{0.05 \times n}} \leq 2.5, n \geq 6272000$$

So we need to sample 1568000 individuals.

(b) Assuming $n = 1000$ is enough for the sample variance to be closed to the true variance, from the asymptotic confidence interval from the sample mean:

$$P(\mu \in [\bar{X}_n - \frac{\hat{\sigma}}{\sqrt{n}} Q^{-1}(\frac{\alpha}{2}), \bar{X}_n + \frac{\hat{\sigma}}{\sqrt{n}} Q^{-1}(\frac{\alpha}{2})]) \approx 1 - \alpha$$

$$\frac{\sqrt{144}}{\sqrt{n}} \times 1.96 \leq 2.5, n \geq 88$$

So we need at least 88 individuals.

(c) For $n = 20$, 0 intervals out of 10000 don't contain the true mean. For $n = 1000$, 0 intervals out of 10000 don't contain the true mean. This means the confidence interval is precise for both $n = 20$ and $n = 1000$

The true variance of the weights are 135.97. For $n = 20$, the distribution of sample variance is biased and heavy tailed. This means that the variance calculated from 20 sample is not enough to estimate the true sample variance. For $n = 1000$, the distribution of variance is unbiased and light tailed. Therefore using sample variance of 1000 data is a close to the true variance.

Question 3.

(a) We want to show: $P(A \leq a - \epsilon) - P(|A_n - A| > \epsilon) \leq P(A_n \leq a) \leq P(A \leq a + \epsilon) + P(|A_n - A| > \epsilon)$

On the one hand, $P(A_n \leq a) = P(A_n \leq a, A \leq a + \epsilon) + P(A_n \leq a, A \geq a + \epsilon)$

$$\leq P(A \leq a + \epsilon) + P(A_n - A \leq a - A, a - A \leq -\epsilon) \leq P(A \leq a + \epsilon) + P(A_n - A \leq -\epsilon)$$

$$\leq P(A \leq a + \epsilon) + P(A_n - A \leq -\epsilon) + P(A_n - A \geq \epsilon) = P(A \leq a + \epsilon) + P(|A_n - A| > \epsilon)$$

On the other hand, $P(A \leq a - \epsilon) = P(A \leq a - \epsilon, A_n \leq \epsilon + A) + P(A \leq a - \epsilon, A_n \geq \epsilon + A)$

$$= P(A_n \leq \epsilon + A, A + \epsilon \leq a) + P(A \leq a - \epsilon, A_n \geq \epsilon + A) \leq P(A_n \leq a) + P(A_n \leq \epsilon)$$

$$\leq P(A_n \leq a) + P(A_n - A \geq \epsilon) + P(A_n - A \leq -\epsilon) = P(A_n \leq a) + P(|A_n - A| \geq \epsilon)$$

Therefore, $P(A_n \leq a) \geq P(A \leq a - \epsilon) - P(|A_n - A| \geq \epsilon)$

These two inequalities shows $P(A \leq a - \epsilon) - P(|A_n - A| > \epsilon) \leq P(A_n \leq a) \leq P(A \leq a + \epsilon) + P(|A_n - A| > \epsilon)$

(b) Since $A_n \rightarrow A$ in probability, $\forall \epsilon > 0, \lim_{n \rightarrow \infty} P(|A_n - A| > \epsilon) = 0$

From (a), $P(A \leq a - \epsilon) - P(|A_n - A| > \epsilon) \leq P(A_n \leq a) \leq P(A \leq a + \epsilon) + P(|A_n - A| > \epsilon)$

This yields, $P(A \leq a - \epsilon) \leq P(A_n \leq a) \leq P(A \leq a + \epsilon) \iff F_A(a - \epsilon) \leq P(A_n \leq a) \leq F_A(a + \epsilon)$

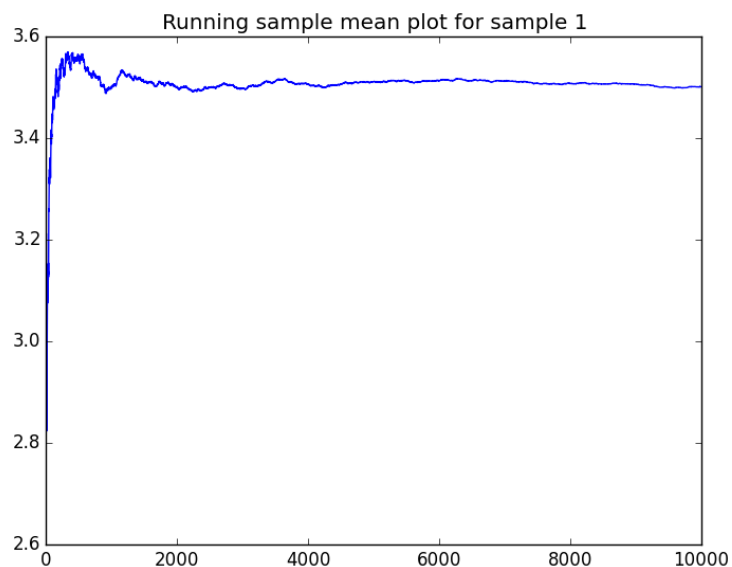
Since ϵ is an arbitrary number, this means $F_A(a) = F_{A_n}(a), \lim_{n \rightarrow \infty} P(A_n \leq a) = P(A \leq a)$

By definition A_n converge to A in distribution.

(c) Converge in distribution does not imply converge in probability. A counterexample is $X \sim N(0, 1)$, $X_n = -X$, X_n and X has the same cdf but they don't converge in probability.

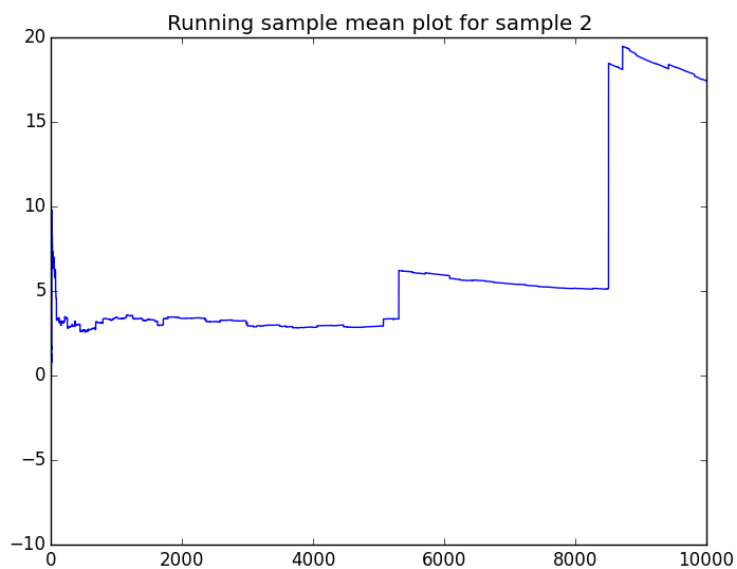
Problem 4.

(a)



A reasonable estimator for c can be 3.5. This method will work only if the mean and variance of the sample is well defined and finite.

(b)



The method doesn't work. There are extremely large data point such that the running mean does not seem to converge.

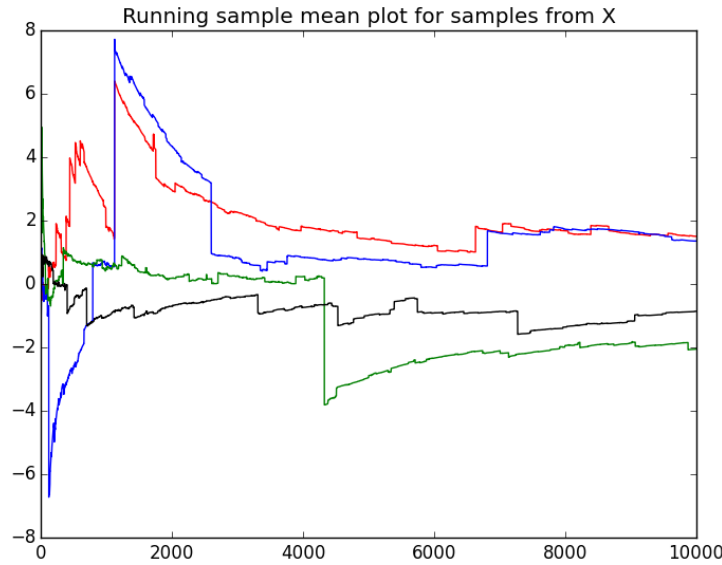
(c) α is uniformly distributed from $[-\pi/2, \pi/2]$, $P(X \in [0, x]) = \int_0^{\arctan(x)} \frac{1}{\pi} d\alpha = \frac{\arctan(x)}{\pi}$

$$f(x) = d \frac{\arctan(x)}{\pi} / dx = \frac{1}{\pi(1+x^2)}$$

$$E(x) = \int_{-\infty}^{\infty} \frac{x}{\pi(1+x^2)} dx = \int_{-\infty}^{\infty} \frac{1}{2\pi(1+x^2)} dx^2 = \int_0^{\infty} \frac{1}{2\pi(1+t)} dt = \frac{\log(t+1)}{2\pi} \Big|_0^{\infty} = \infty$$

Therefore the mean of X is not defined. Therefore, for this random variable, Law of large number doesn't work.

(d) We can sample points from $\alpha \sim \text{unif}[-\pi/2, \pi/2]$, and transform to x by $x = \tan(\alpha)$



From the plots, it's clear that the sample means do not converge to the same value. LLN fails.

Problem 5.

(a) If the sample mean is greater than 0.5 then the Republican is more likely to win the election. Otherwise, the Democrat is more likely to win. This will work when the sample is randomly selected and are iid, and the number of sample is sufficiently large.

(b) Since R_i is Bernoulli distribution, $\text{Var}(R_i) = p(1-p) \leq \frac{1}{4}$

From the sample, $\bar{R}_n = 45/94 < \frac{1}{2}$, the Democrat will win the election.

By Chebyshev's inequality: $P(|\bar{R}_n - p| \geq \alpha) \leq \frac{\text{Var}(\bar{R}_n)}{\alpha^2} = \frac{1}{4n\alpha^2}$

$$P(|\bar{R}_n - p| \geq \alpha) \geq P(p \geq \bar{R}_n + \alpha) = P(p \geq \frac{45}{94} + \alpha)$$

$$\text{Let } \alpha = \frac{1}{2} - \frac{45}{94}, \text{ we have } P(p > \frac{1}{2}) \leq \frac{1}{4 \times 94000 \times (\frac{1}{2} - \frac{45}{94})^2} \approx 0.005875$$

It does not mean that it's an upper bound of the probability that the prediction is right.

(c) Let $n_{Y,R}$ be the number of young people vote for Republican, $n_{Y,D}$ the number of young people vote for Democrat, $n_{O,R}$ be the number of old people vote for Republican, $n_{O,D}$ the number of young people vote for Democrat.

$$E(X_{n_1, n_2}) = aE(\sum_{n_1} Y_i) + bE(\sum_{n_2} O_i) = \frac{n_{Y,R} + n_{Y,D}}{n_{Y,R} + n_{Y,D} + n_{O,R} + n_{O,D}} \times \mu_Y + \frac{n_{O,R} + n_{O,D}}{n_{Y,R} + n_{Y,D} + n_{O,R} + n_{O,D}} \times \mu_O$$

$$E(\sum_{n_1} Y_i) = n_1 \mu_Y, E(\sum_{n_2} O_i) = n_2 \mu_O$$

$$\text{Therefore, } a = \frac{n_{Y,R} + n_{Y,D}}{n_{Y,R} + n_{Y,D} + n_{O,R} + n_{O,D}} \times \frac{1}{n_1}, b = \frac{n_{O,R} + n_{O,D}}{n_{Y,R} + n_{Y,D} + n_{O,R} + n_{O,D}} \times \frac{1}{n_2}$$

$$(d) \text{ From (c) } X_{59000, 35000} = \frac{0.2}{0.2+0.55} \times \frac{24000}{59000} + \frac{0.55}{0.2+0.55} \times \frac{21000}{35000} = \frac{809}{1475} \approx 0.5799031 > 0.5$$

We predict that Republican will win the election

We assume that the sample for both groups are selected randomly. Within each group the data are i.i.d

$$Var(X_{n_1, n_2}) \leq a^2 \times n_1 \times \frac{1}{4} + b^2 \times n_2 \times \frac{1}{4} = \frac{1}{4} \left(\frac{4}{15}^2 \frac{1}{59000} + \frac{11}{15}^2 \frac{1}{35000} \right)$$

$$\text{By Chebyshev's inequality: } P(|\bar{X}_{n_1, n_2} - p| \geq \alpha) \leq \frac{Var(\bar{X}_{n_1, n_2})}{\alpha^2}$$

$$P(|\bar{X}_{n_1, n_2} - p| \geq \alpha) \geq P(p \leq \bar{X}_{n_1, n_2} - \alpha), \text{ let } \alpha = \frac{809}{1475} - \frac{1}{2}$$

$$P(p \leq \frac{1}{2}) \leq \frac{1}{\alpha^2} \times \frac{1}{4} \left(\frac{4}{15}^2 \frac{1}{59000} + \frac{11}{15}^2 \frac{1}{35000} \right) \approx 0.001763$$