

## Homework 2 DS-GA 1002

Yuhao Zhao

N17578783

November 19, 2015

### Problem 1.

(a) The position of spider can be modeled By  $F(X,Y)$ .  $X$  is the horizontal position and  $Y$  is the height position.

Since we know that the spider stays twice the time under the painting area, and is uniformly distributed both inside and outside the painting area.

We have  $P((X,Y) \in \text{paint}) = 2P((X,Y) \notin \text{paint})$

$$\text{Let } f_{X,Y}(x,y) = \begin{cases} c & (x,y) \in [4,6] \times [6,8] \\ d & \text{otherwise} \end{cases}$$

Then we have  $c \times 4 = 2 \times d \times 96$  and  $4c = 192d, c = \frac{1}{6}, d = \frac{1}{288}$

Therefore

$$f_{X,Y}(x,y) = \begin{cases} \frac{1}{6} & (x,y) \in [4,6] \times [6,8] \\ \frac{1}{288} & \text{otherwise} \end{cases}$$

(b)

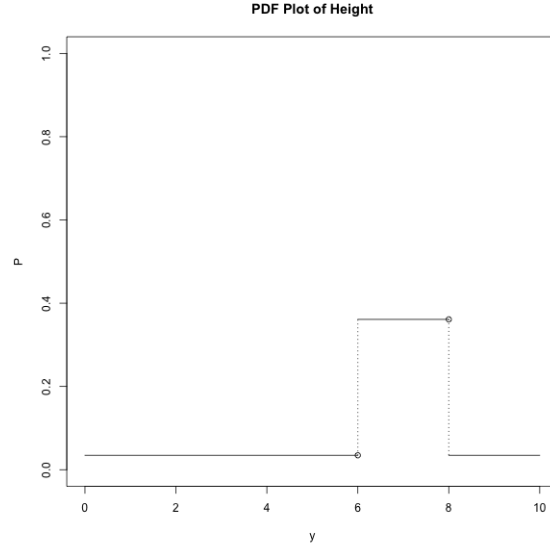
(i) For  $y \in [0,6]$ ,  $F(Y) = P(Y \in [0,y]) = \int_0^y \int_0^{10} f_{X,Y}(u,v) du dv = \int_0^y \int_0^{10} \frac{1}{288} du dv = \frac{10}{288}y$ ,  
the pdf is  $\frac{dF}{dy} = \frac{10}{288}$

(ii) For  $y \in [6,8]$ ,  $F(Y) = P(Y \in [6,y]) + P(Y \in [0,6]) = \int_6^y \int_0^{10} f_{X,Y}(u,v) du dv + \frac{60}{288}$   
 $\int_6^y \int_0^{10} f_{X,Y}(u,v) du dv = \int_6^y \int_0^4 \frac{1}{288} du dv + \int_6^y \int_4^6 \frac{1}{6} du dv + \int_6^y \int_6^{10} \frac{1}{288} du dv = \frac{13}{36}(y-6)$

Therefore, For  $y \in [6,8]$  the pdf is  $\frac{13}{36}$

(iii) For  $y \in [8,10]$ ,  $F(Y) = P(Y \in [0,8]) + P(Y \in [8,y]) = \int_8^y \int_0^{10} f_{X,Y}(u,v) du dv + \frac{67}{72} =$   
 $\frac{67}{72} + \int_8^y \int_0^{10} \frac{1}{288} du dv$

Therefore For  $y \in [8,10]$ , the pdf is  $\frac{10}{288}$



(c) Let event A be that the spider is located outside the painting area.

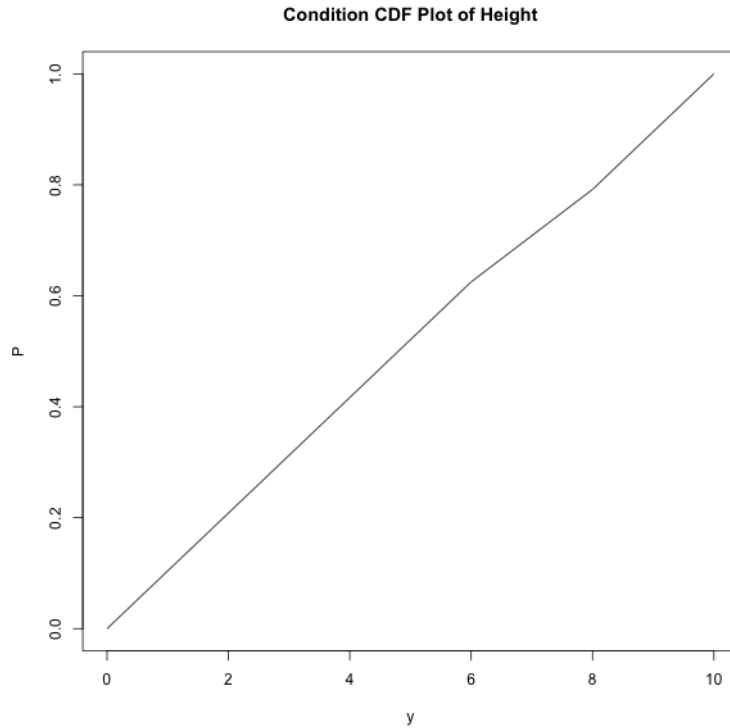
$$F_{X,Y|A} = \frac{p(X \leq x, Y \leq y; A)}{P(A)}, \quad P(A) = \frac{1}{3}$$

$$\text{For } y \in [0, 6], F(Y) = P(Y \in [0, y]) = \int_0^y \int_0^{10} 3f_{X,Y,A}(u, v) du dv = \int_0^y \int_0^{10} \frac{3}{288} du dv = \frac{30}{288}y,$$

$$\text{For } y \in [6, 8], F(Y) = P(Y \in [6, y]) + P(Y \in [0, 6]) = \int_6^y \int_0^{10} f_{X,Y,A}(u, v) du dv + \frac{180}{288}$$

$$\int_6^y \int_0^{10} f_{X,Y,A}(u, v) du dv = \int_6^y \int_0^4 3\frac{1}{288} du dv + \int_6^y \int_6^{10} 3\frac{1}{288} du dv = \frac{24}{288}(y - 6)$$

$$\text{For } y \in [8, 10], F(Y) = P(Y \in [0, 8]) + P(Y \in [8, y]) = \int_8^y \int_0^{10} f_{X,Y,A}(u, v) du dv + \frac{228}{288} = \frac{228}{288} + \int_8^y \int_0^{10} 3\frac{1}{288} du dv = \frac{30}{288}(y - 8) + \frac{228}{288}$$



## Problem 2.

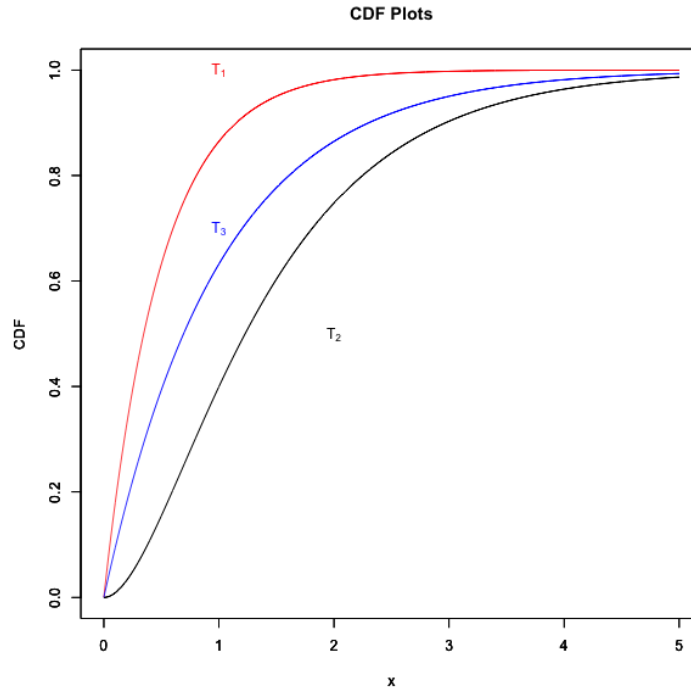
(a) Let  $X$  be the time until Pat receives a call,  $Y$  be the time until Robbie receives a call. If we assume that the customer has no preference between the two restaurant and order pizzas independently, let the time until one of them receives a call be  $T_1$ .  $P(0 < T_1 < t) = 1 - P(X > t, Y > t)$  Since  $X, Y$  are independent by assumption,  $P(X > t, Y > t) = P(X > t) \times P(Y > t) = \int_t^\infty \lambda e^{-\lambda x} dx \int_t^\infty \lambda e^{-\lambda y} dy = (-e^{-\lambda x}|_t^\infty)^2 = e^{-2\lambda t}$   
 $P(0 < T_1 < t) = 1 - P(X > t, Y > t) = 1 - e^{-2\lambda t}$

(b) Making the same assumption as part (a), let  $T_2$  be the time until both of them have received a call. If both of them received the call until  $t$ , the possibility should be  $X$  received at any time within  $[0, t]$ , and  $Y$  received at any time within  $[x, t]$  as well as the symmetric case of  $Y$  received first.

$$P(T_2 < t) = 2 \times \int_0^t \int_x^t \lambda e^{-\lambda x} \lambda e^{-\lambda y} dy dx = 2 \int_0^t \lambda e^{-\lambda x} (-e^{-\lambda y}|_x^t) dx = 2 \int_0^t \lambda e^{-2\lambda x} - \lambda e^{-\lambda t} e^{-\lambda x} dx$$

$$= 2(-\frac{e^{-2\lambda x}}{2}|_0^t + e^{-\lambda t} e^{-\lambda x}|_0^t) = e^{-2\lambda t} - 2e^{-\lambda t} + 1$$

(c) The result is reasonable. Let the individual waiting time distribution be  $T_3$ . At any given time, the probability of at least one of them receive a call should be greater than that of both of them receive a call. The corresponding CDF plots are identical to this fact, in particular  $T_1$  lies above  $T_2$ . The event Pat receives a call at time  $t$  is contained in the event that at least one of them receive a call. Meanwhile, it contain the event that both of them receive call. In the plot we observed that  $T_2(t) < T_3(t) < T_1(t), \forall t > 0$ , this is identical to the fact.



**Problem 3.**

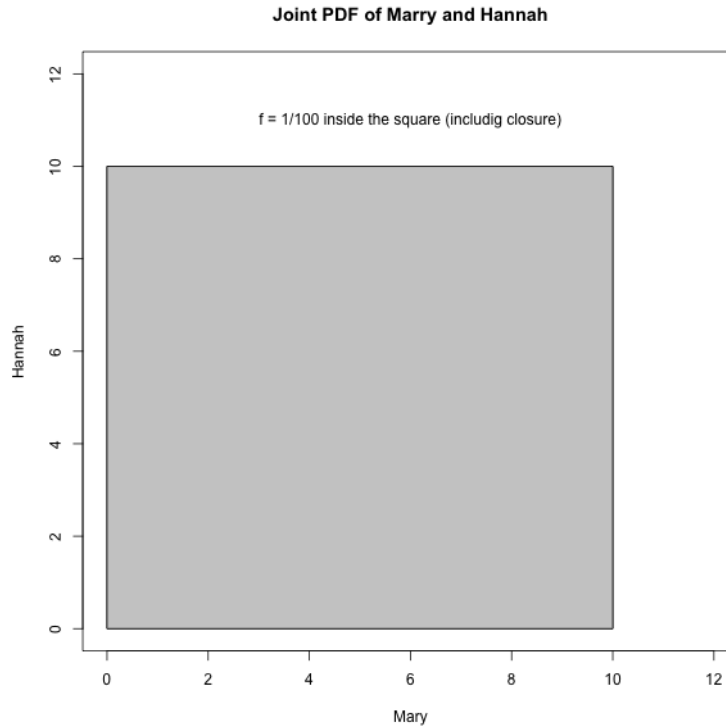
(a) Let  $X$  be a Binomial( $n, \frac{1}{2}$ ), and  $X_k$  be the events that  $X = k$ . Since  $X_k$ 's are disjoint events and  $\bigcup X_k$  is the whole sample space defined by  $X$ . By the axiom of probability:  $P(\bigcup_{k=0}^n X_k) = \sum_{k=0}^n P(X_k) = 1$   
 $\sum_{k=0}^n P(X_k) = \sum_{k=0}^n \binom{n}{k} \left(\frac{1}{2}\right)^k \left(\frac{1}{2}\right)^{n-k} = \frac{\sum_{k=0}^n \binom{n}{k}}{2^n} = 1$   
 Therefore,  $\sum_{k=0}^n \binom{n}{k} = 2^n$

(b) Let  $N$  be the number of mechanical problem that the two people will encounter per month. We assume that the cars encounter problems independently.

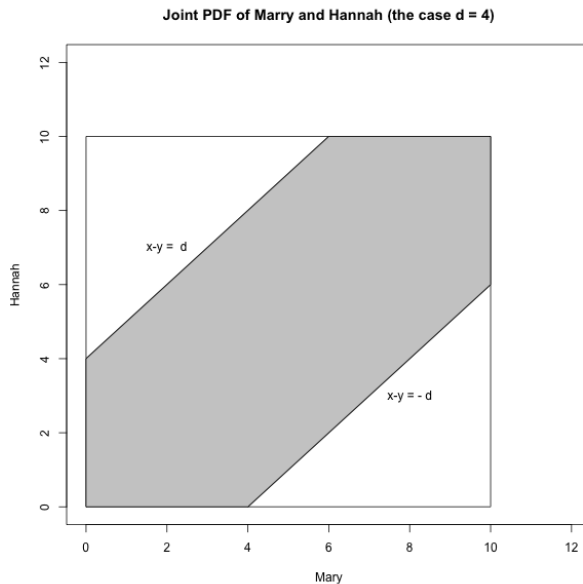
$$P(N = n) = \sum_{k=0}^n \frac{\lambda^k}{k!} e^{-\lambda} \frac{\lambda^{n-k}}{(n-k)!} e^{-\lambda} = e^{-2\lambda} \sum_{k=0}^n \frac{\lambda^n}{k!(n-k)!} = e^{-2\lambda} \frac{\lambda^n}{n!} \sum_{k=0}^n \frac{n!}{k!(n-k)!} = e^{-2\lambda} \frac{\lambda^n}{n!} 2^n = e^{-2\lambda} \frac{(2\lambda)^n}{n!}$$

**Problem 4.**

(a) Let  $X, Y$  be the positions of Mary and Hannah respectively. They are independently and uniformly distributed over  $[0,10] \times [0,10]$ . The joint pdf  $f_{X,Y} = \frac{1}{100}$  for  $(X,Y) \in [0,10] \times [0,10]$  and 0 otherwise.

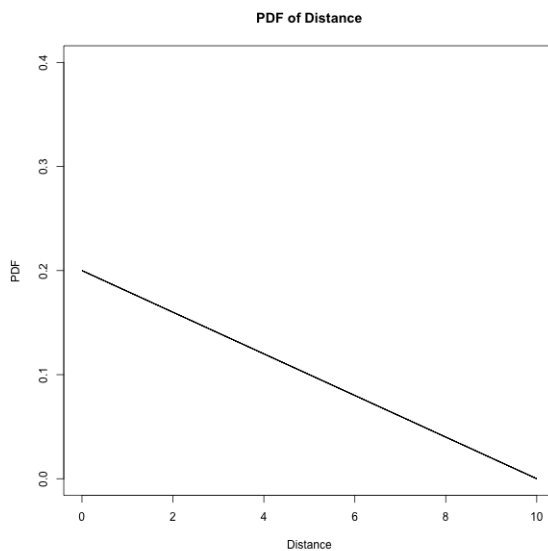


(b)



(c) Let the distance be  $D$ .  $P(D \in [0, d])$  is just the percentage of the shaded region for any give  $d$ .  

$$P(D \in [0, d]) = \frac{100 - (10 - d)^2}{100} = \frac{20d - d^2}{100}$$
the pdf is  $f(d) = \frac{20 - 2d}{100}$



### Problem 5.

(a) Let  $C$  be the bias of the coin,  $C \sim \text{Unif}(\frac{1}{2}, 1)$ , and  $D$  be the outcome of the coin flip. The parameter of the distribution of outcome depends on the bias. In Particular,  $P(D = \text{heads} | C = c) = c$   

$$P(D = \text{heads}) = \int_{0.5}^1 P(D = \text{heads} | C = c) f(c) dc = \int_{0.5}^1 c \times 2 dc = 1 - \frac{1}{4} = \frac{3}{4}$$

$$P(D = \text{tails}) = \frac{1}{4}$$

$$(b) f(c|D = heads) = \frac{f_c(c)P_{D|C}(D=heads|C=c)}{P_D(D=heads)} = \frac{2c}{3/4} = \frac{8c}{3}$$

$$f(c|D = tails) = \frac{f_c(c)P_{D|C}(D=tails|C=c)}{P_D(D=tails)} = \frac{2(1-c)}{1/4} = 8(1-c)$$

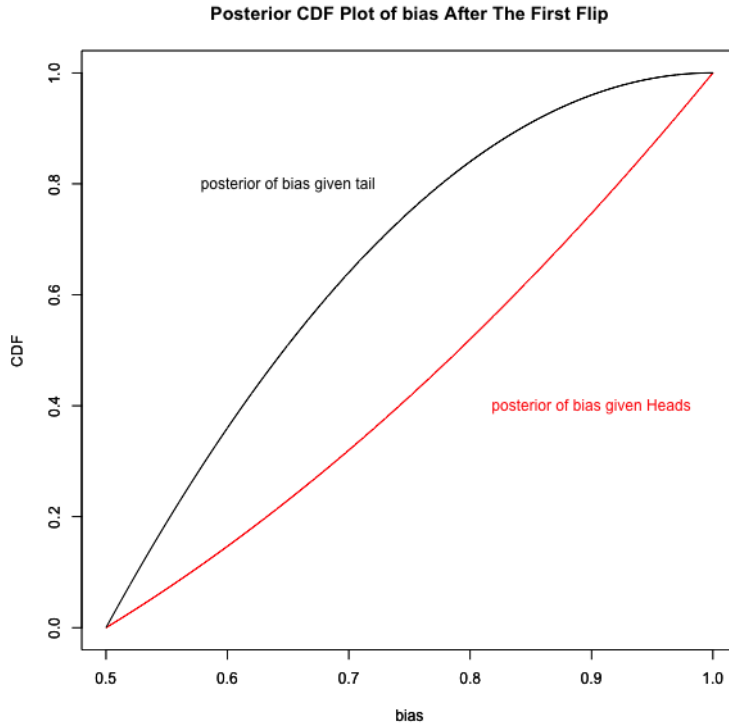
The posterior of bias given getting a head is  $\frac{8c}{3}$ , Thus the cdf is  $\frac{8}{6}c^2 - \frac{1}{3}, 0.5 \leq c \leq 1$

The posterior of bias given getting a tail is  $8 - 8c$ , Thus the cdf is  $8c - 4c^2 - 3, 0.5 \leq c \leq 1$

If we continue the flip:

$$i) \text{ For } D = heads, f_{i+1}(c|D = heads) = \frac{f_{c,i}(c)P_{D|C}(D=heads|C=c)}{P_D(D=heads)} = \frac{f_{c,i}(c)c}{\int_{0.5}^1 cf_{c,i}(c)dc}$$

$$ii) \text{ For } D = tails, f_{i+1}(c|D = tails) = \frac{f_{c,i}(c)P_{D|C}(D=tails|C=c)}{P_D(D=tails)} = \frac{f_{c,i}(c)(1-c)}{\int_{0.5}^1 f_{c,i}(c)(1-c)dc}$$



This is reasonable, because if a head is observed the bias is more likely to be distributed close to 1. Therefore the CDF of the posterior is convex. If a tail is observed, the bias is more likely to be distributed close to 0.5. Thus the CDF is concave.