

# Global Financial Stability Randomization Review

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## INTRODUCTION

Since the beginning of the recent global financial crisis financial stability has been shifted into the center of attention. “To be able to detect potential threats to financial stability and take appropriate macro prudential measures early on, policymakers do not only need to monitor and assess financial stability but also to project its likely future development. One of the lessons to be learned from the recent financial and economic crisis is that a very broad range of indicators must be monitored to be able to assess overall financial stability in a reliable manner.”(**Petr Jakubík, Tomáš Slacík 2013**).

The most common methods of measuring financial stability are: early warning systems, macro-stress testing, financial stability indices. In this report I will show the examples of financial stability indices.

The financial stability indexes for nine countries, from February 2006 to February 2014, were studied and divided in three groups.

A Prediction of USA as a linear function of all countries sampled, by means of a multiple regression model, was conducted. Each country’s significant level, expect for USA, was tested using an approximate permutation test. Finally, each divided group’s model was tested and result were compared.

## FINANCIAL STABILITY

Financial Stability (FS) is a condition, when a financial system is able to withstand economic shocks, managing financial risks, processing payment transactions and intermediating financial resources.

After the global financial crisis the topic of financial stability became very important. It became clear that for detecting financial stability risks and crisis and taking appropriate decisions one should be able to project future financial stability development. Globalization, financial innovations and STR (Scientific-technological revolution) started many financial processes bringing new transmission channels and making the process of measuring financial stability more complicated. To assess financial stability a big number of economic and financial indicators must be monitored.

There are different methods of measuring financial stability and each has its advantages and disadvantages. The most common methods are:

- Early warning systems,
- Macro-stress testing
- Financial stability indices

Early warning systems are created from potential indicators which are aiming to predict the probability of financial crisis. But this systems doesn’t provide a deep and detailed financial stability analysis.

Macro-stress testing provides more detailed analysis, which can measure financial resistance to risks and threats. Stress tests can reveal not only the source of risk but also vulnerabilities of a financial system.

Financial stability indices provide information on the measure of changes in a representative group of individual indicators. There are different indices which provide the information about financial stability. Let’s have a look at some of them:

Cleveland Financial Stress Index tracks stress in six types of markets: credit markets, equity markets, foreign exchange markets, funding markets, real estate markets, and securitization markets. It is a coincident indicator of systemic stress, where a high value of CFSI indicates high systemic financial stress. Units of

CFSI are expressed as standardized differences from the mean. The CFSI provides a continuous measure of stress. For interpretation of index it is divided into four levels, which called grades. The grade thresholds are dynamic and move slowly over time. The four grades from the Probability of Systemic Stress Episode are:

CFSI Grade	Threshold Description	Range	Probability
Grade 1	Low stress period	$\text{CFSI} \leq -0.50$	1.9
Grade 2	Normal stress period	$-0.50 < \text{CFSI} < 0.59$	8.7
Grade 3	Moderate stress period	$0.59 < \text{CFSI} < 1.68$	26.3
Grade 4	Significant stress period	$\text{CFSI} \geq 1.68$	56.3

Another example is Financial Instability Index (FII) created by *Petr Jakubík and Tomáš Slacík*. FII gauges the level of financial market stress in selected Central, Eastern and Southeastern European (CESEE) countries. Index observes eleven variables relative to financial stability:

- Money market year-on-year change,
- Money market volatility,
- Spread between domestic and German interbank offered rates,
- Exchange rate year-on-year change,
- Exchange rate volatility,
- Stock index year-on-year change,
- Stock index volatility,
- Ten-year government bond yield,
- Ten-year government bond yield -year-on-year change,
- Ten-year government bond yield -volatility,
- Composite EMBI Global.

In the output each indicator is transformed into the number between 0 and 1, and the closer output is to the 0, the more financially stable the economy is in a certain period of time (**Petr Jakubík, Tomáš Slacík 2013**).

## GLOBAL DATA DESCRIPTION

To conduct the empirical study, an analysis of the data globally using the USA as a benchmark was used. The data was collected from Eurostat and World Bank Group. The time window was from January 2006 to January 2014.

The research was complied into a initial data frame called “GData”

GData has 97 rows and 10 columns of data was collected from Eurostat and World Bank Group.

GData has comprised of nine country’s FS indexes.

They are as followed:

- Euro zone countries: Germany, France, Italy, Spain
- Countries from EU but not from Euro zone: Poland, United Kingdom
- None EU or Euro zone countries: Brazil, Russia

## Prediction of Multiple Regression Model

A Prediction of USA as a linear function of all countries sampled, by means of a multiple regression model is displayed below:

```

#install.packages("readxl")
suppressMessages(suppressWarnings(library("readxl")))

Gdata<-read_excel("WorldData.xlsx")

#   USA, BRA,   FR, GR, IT, PO, RUS,   SPA,   UK
US <- Gdata$USA
BRA <- Gdata$BRA
FR <- Gdata$FR
GR <- Gdata$GR
IT <- Gdata$IT
PO <- Gdata$PO
RUS <- Gdata$RUS
SPA <- Gdata$SPA
UK <- Gdata$UK

# Fit a linear model and run a summary of its results.
mod1<-lm(US ~ BRA + FR + GR + IT + PO + RUS + SPA + UK)
summary(mod1)

##
## Call:
## lm(formula = US ~ BRA + FR + GR + IT + PO + RUS + SPA + UK)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -0.117604 -0.032565  0.000571  0.028090  0.221158
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   0.26652    0.07369   3.617 0.000497 ***
## BRA           0.05495    0.08513   0.645 0.520295
## FR            0.44318    0.12165   3.643 0.000455 ***
## GR           -0.28504    0.11179  -2.550 0.012511 *
## IT            0.24164    0.15398   1.569 0.120179
## PO           -0.21295    0.10220  -2.084 0.040099 *
## RUS          -0.27930    0.15221  -1.835 0.069897 .
## SPA           0.23137    0.10890   2.125 0.036426 *
## UK            0.19408    0.11415   1.700 0.092620 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.05825 on 88 degrees of freedom
## Multiple R-squared:  0.5622, Adjusted R-squared:  0.5224
## F-statistic: 14.13 on 8 and 88 DF,  p-value: 4.986e-13

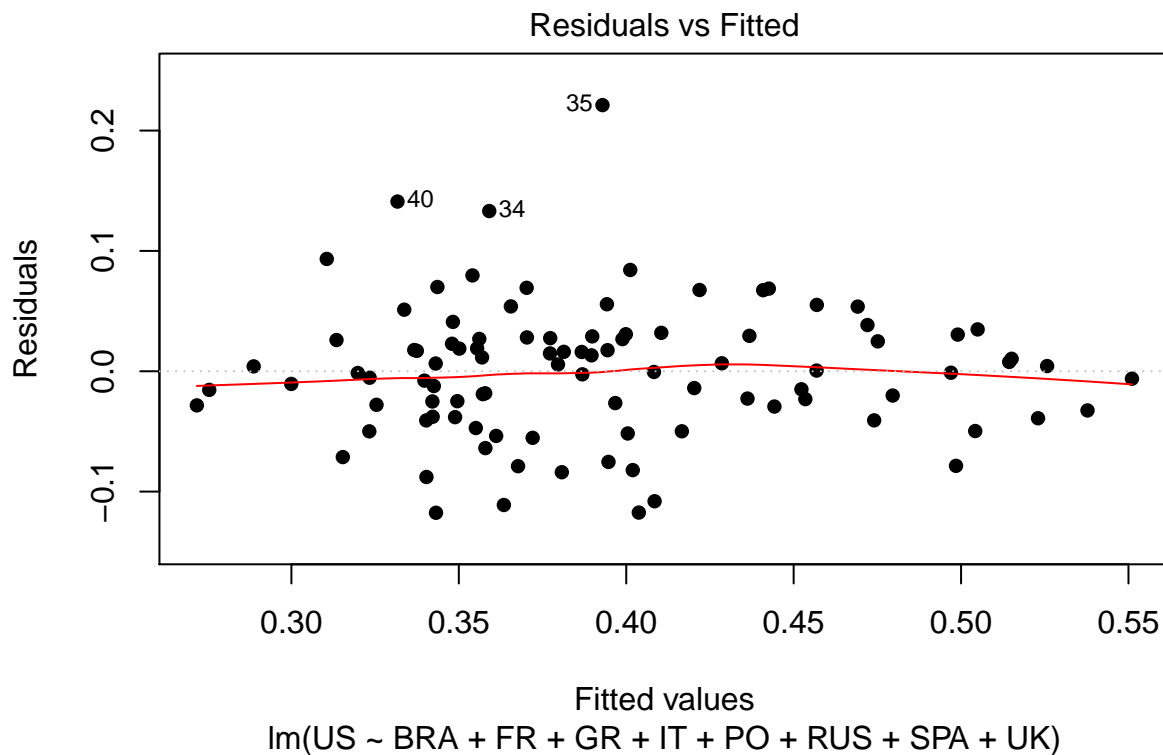
# CIs for model parameters
confint(mod1, level=0.95)

##              2.5 %      97.5 %
## (Intercept) 0.12007192 0.412965333
## BRA         -0.11423313 0.224137445
## FR           0.20143197 0.684931478
## GR          -0.50720967 -0.062879194

```

```
## IT          -0.06437181  0.547649510
## PO          -0.41606233 -0.009840704
## RUS         -0.58178399  0.023193091
## SPA         0.01495113  0.447780266
## UK          -0.03276827  0.420932339
```

```
# Plot model residuals.
par(mfrow=c(1,1))
plot(mod1, pch=16, which=1)
```



### Significance of the all countries, except USA

The testing began by testing whether a set of independent variables has no partial effect on the dependent variable, “US”.

The model is:

$$US = B_0 + B_1BRA + B_2FR + B_3GR + B_4IT + B_5PO + B_6RUS + B_7SPA + B_8UK + e$$

**Null Hypothesis:** The initial assumption is that there is no relation, which is expressed as:

$$H_0: B_1 = B_2 = B_3 = B_4 = B_5 = B_6 = B_7 = B_8 = 0.$$

**Alternative Hypothesis:** At least one of the independent variables IS useful in explaining/predicting US, expressed as:

**H1:** At least one  $B_i$  is “not equal to” 0.

## Approximate Permuatition Test

All country's significant level, expect for USA, using an approximate permutation test was tested below:

```
#install.packages("combinat")
suppressMessages(suppressWarnings(library("combinat")))
#install.packages("readxl")
suppressMessages(suppressWarnings(library("readxl")))

## **Prediction of Multiple Regression Model**

Gdata<-read_excel("WorldData.xlsx")

#   USA, BRA,   FR, GR, IT, PO, RUS,   SPA,   UK
US <- Gdata$USA
BRA <- Gdata$BRA
FR <- Gdata$FR
GR <- Gdata$GR
IT <- Gdata$IT
PO <- Gdata$PO
RUS <- Gdata$RUS
SPA <- Gdata$SPA
UK <- Gdata$UK

# Fit a linear model and run a summary of its results.
mod2<-glm(US ~ BRA + FR + GR + IT + PO + RUS + SPA + UK, family = gaussian)
a<-summary(mod2)
a

##
## Call:
## glm(formula = US ~ BRA + FR + GR + IT + PO + RUS + SPA + UK,
##      family = gaussian)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.117604  -0.032565   0.000571   0.028090   0.221158
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.26652    0.07369   3.617 0.000497 ***
## BRA          0.05495    0.08513   0.645 0.520295
## FR           0.44318    0.12165   3.643 0.000455 ***
## GR          -0.28504    0.11179  -2.550 0.012511 *
## IT           0.24164    0.15398   1.569 0.120179
## PO          -0.21295    0.10220  -2.084 0.040099 *
## RUS          -0.27930    0.15221  -1.835 0.069897 .
## SPA          0.23137    0.10890   2.125 0.036426 *
## UK           0.19408    0.11415   1.700 0.092620 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for gaussian family taken to be 0.003392862)
##
##      Null deviance: 0.68201  on 96  degrees of freedom
```

```

## Residual deviance: 0.29857  on 88  degrees of freedom
## AIC: -265.72
##
## Number of Fisher Scoring iterations: 2

ZUSinttrue<-abs(a$coefficients[1,3])
ZBRAtrue<-abs(a$coefficients[2,3])
ZFRtrue<-abs(a$coefficients[3,3])
ZGRtrue<-abs(a$coefficients[4,3])
ZITtrue<-abs(a$coefficients[5,3])
ZPOtrue<-abs(a$coefficients[6,3])
ZRUStrue<-abs(a$coefficients[7,3])
ZSPAtrue<-abs(a$coefficients[8,3])
ZUKtrue<-abs(a$coefficients[9,3])

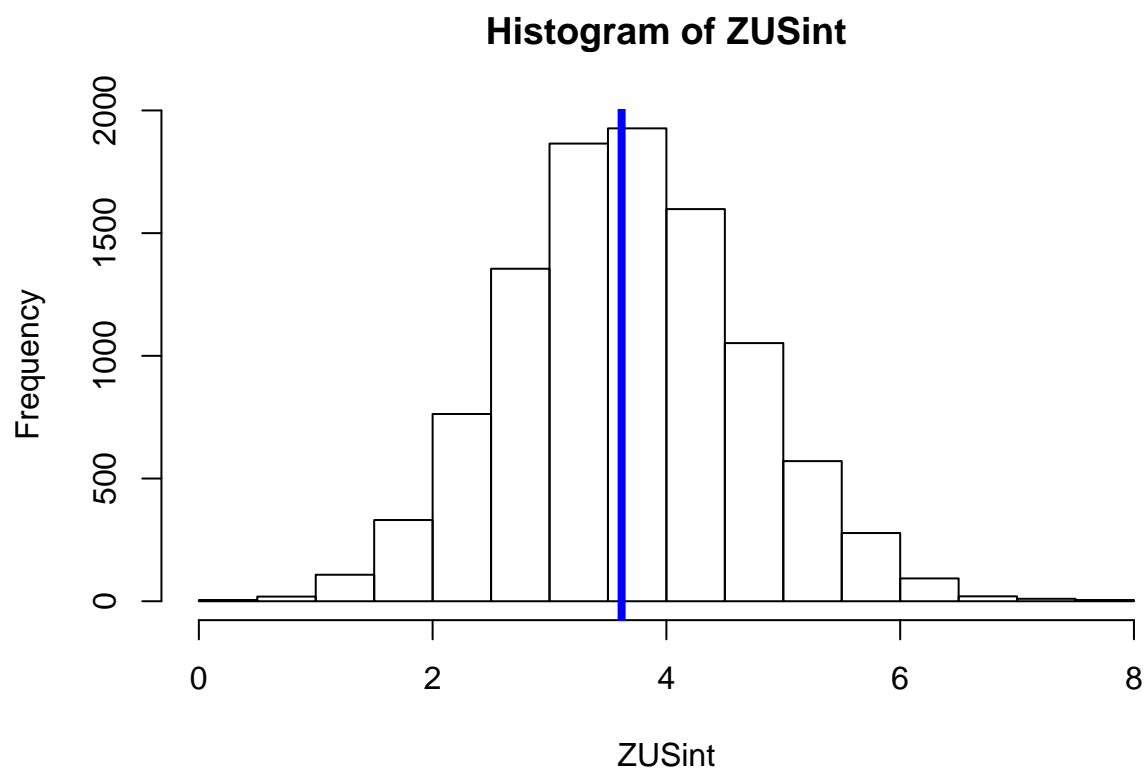
#number of rearrangements to be examined
nr=10000 #number of rearrangements to be examined

ZUSint=numeric(nr);ZBRA=numeric(nr);ZFR=numeric(nr);
ZGR=numeric(nr);ZIT=numeric(nr);ZPO=numeric(nr);
ZRUS=numeric(nr);ZSPA=numeric(nr);ZUK=numeric(nr)

for (i in 1:nr){
  newUS<-sample(US, 97)
  mod2<-glm(newUS ~ BRA + FR + GR + IT + PO + RUS + SPA + UK, family = gaussian)
  a<-summary(mod2)
  ZUSint[i]<-abs(a$coefficients[1,3])
  ZBRA[i]<-abs(a$coefficients[2,3])
  ZFR[i]<-abs(a$coefficients[3,3])
  ZGR[i]<-abs(a$coefficients[4,3])
  ZIT[i]<-abs(a$coefficients[5,3])
  ZPO[i]<-abs(a$coefficients[6,3])
  ZRUS[i]<-abs(a$coefficients[7,3])
  ZSPA[i]<-abs(a$coefficients[8,3])
  ZUK[i]<-abs(a$coefficients[9,3])}

#US
par(mfrow=c(1,1))
hist(ZUSint)
abline(v=ZUSinttrue, lwd=4, col="blue")

```



```
#True t-value of intercept "B0"
ZUSinttrue

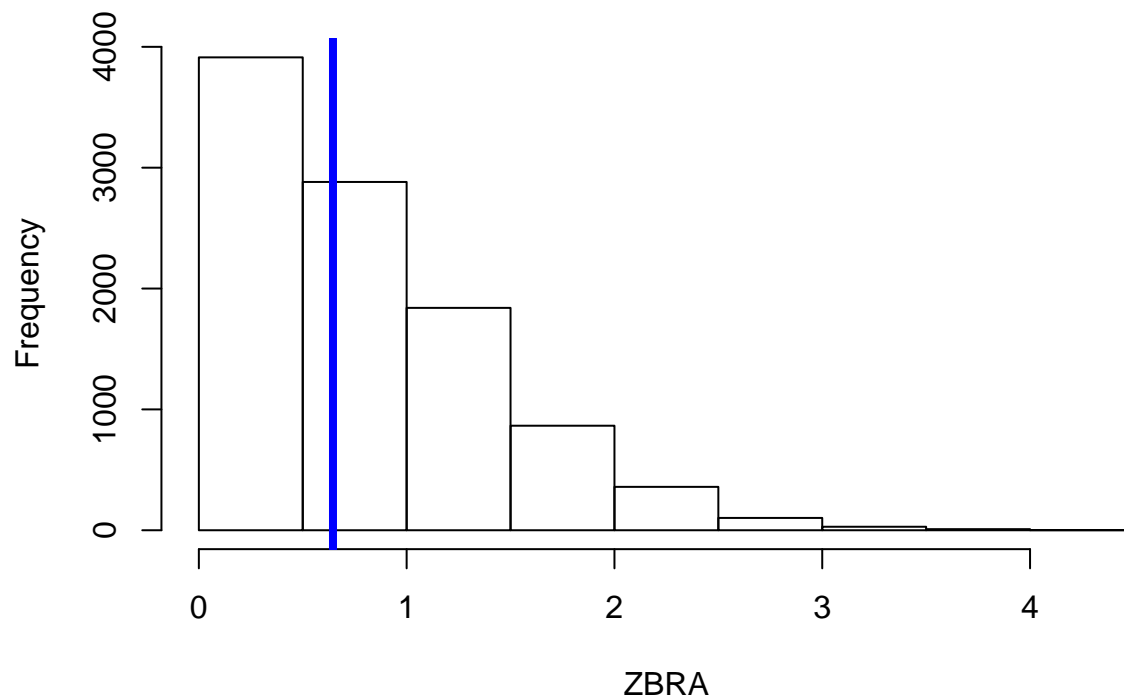
## [1] 3.616672

#P-Value of intercept "B0"
length(ZUSint[ZUSint>= ZUSinttrue])/nr

## [1] 0.5136

#BRA
par(mfrow=c(1,1))
hist(ZBRA)
abline(v=ZBRAttrue, lwd=4, col="blue")
```

## Histogram of ZBRA



```
#True t-Value of Time (T) "B1"  
ZBRAtrue
```

```
## [1] 0.6454808
```

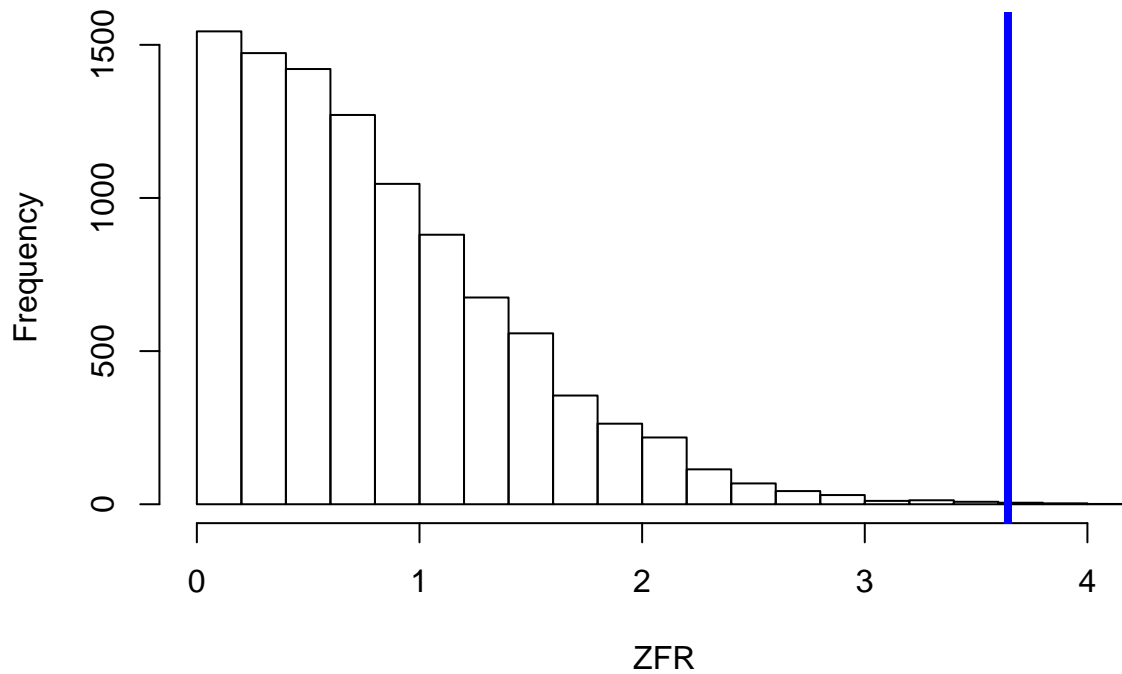
```
#P-value of Time (T) "B1"  
length(ZBRA[ZBRA>= ZBRAtrue])/nr
```

```
## [1] 0.5155
```

```
#FR  
par(mfrow=c(1,1))  
hist(ZFR)  
abline(v=ZFRtrue, lwd=4, col="blue")
```



## Histogram of ZFR



```
#True t-Value of Time moving the grass (TMG) "B2"  
ZFRtrue
```

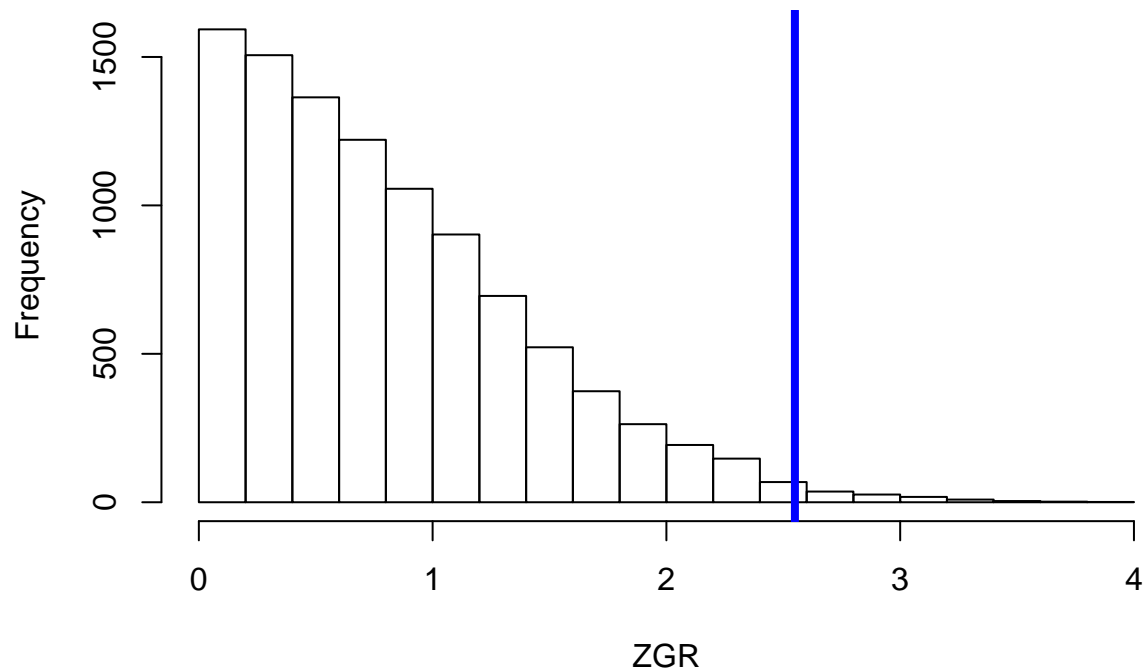
```
## [1] 3.64315
```

```
#P-Value of Time moving the grass (TMG) "B2"  
length(ZFR[ZFR>= ZFRtrue])/nr
```

```
## [1] 8e-04
```

```
#GR  
par(mfrow=c(1,1))  
hist(ZGR)  
abline(v=ZGRtrue, lwd=4, col="blue")
```

## Histogram of ZGR



```
#True t-Value of Time moving the grass (TMG) "B3"  
ZGRtrue
```

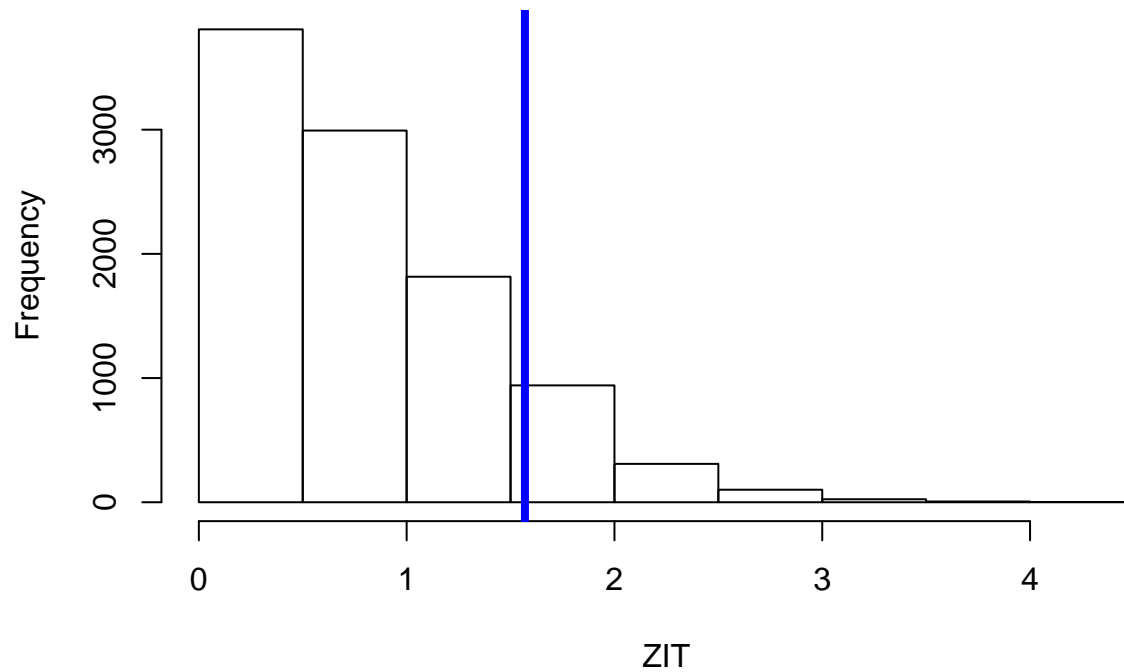
```
## [1] 2.54975
```

```
#P-Value of Time moving the grass (TMG) "B3"  
length(ZGR[ZGR>= ZGRtrue])/nr
```

```
## [1] 0.011
```

```
#IT  
par(mfrow=c(1,1))  
hist(ZIT)  
abline(v=ZITtrue, lwd=4, col="blue")
```

## Histogram of ZIT



```
#True t-Value of Time moving the grass (TMG) "B4"  
ZITtrue
```

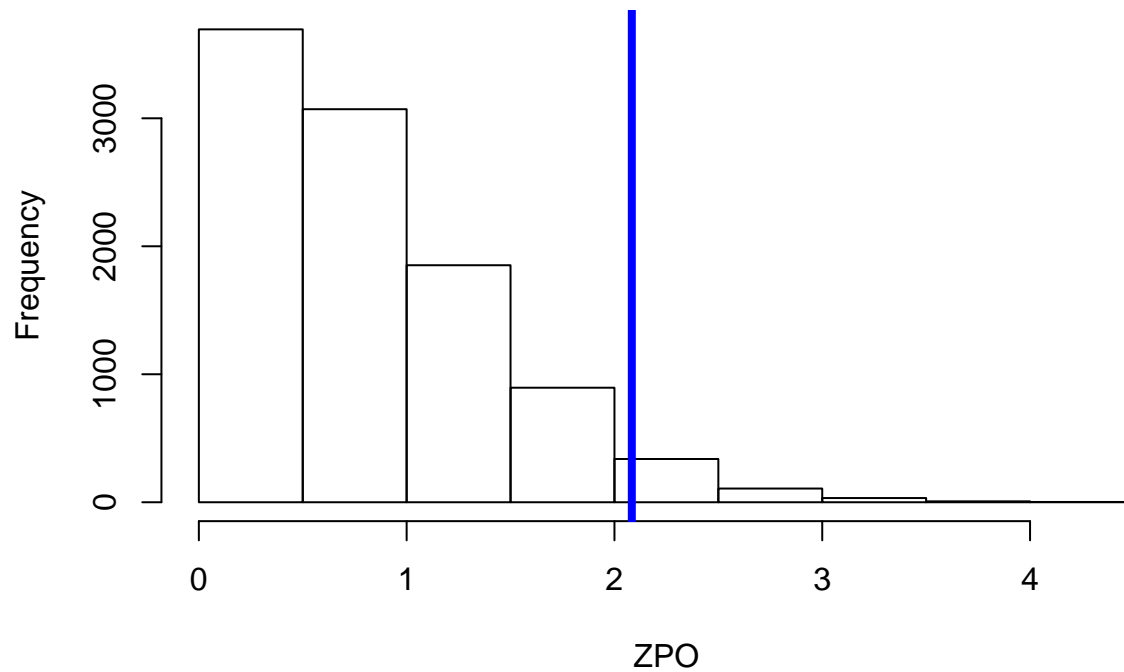
```
## [1] 1.569247
```

```
#P-Value of Time moving the grass (TMG) "B4"  
length(ZIT[ZIT>= ZITtrue])/nr
```

```
## [1] 0.1197
```

```
#P0  
par(mfrow=c(1,1))  
hist(ZP0)  
abline(v=ZP0true, lwd=4, col="blue")
```

## Histogram of ZPO



```
#True t-Value of Time moving the grass (TMG) "B5"  
ZP0true
```

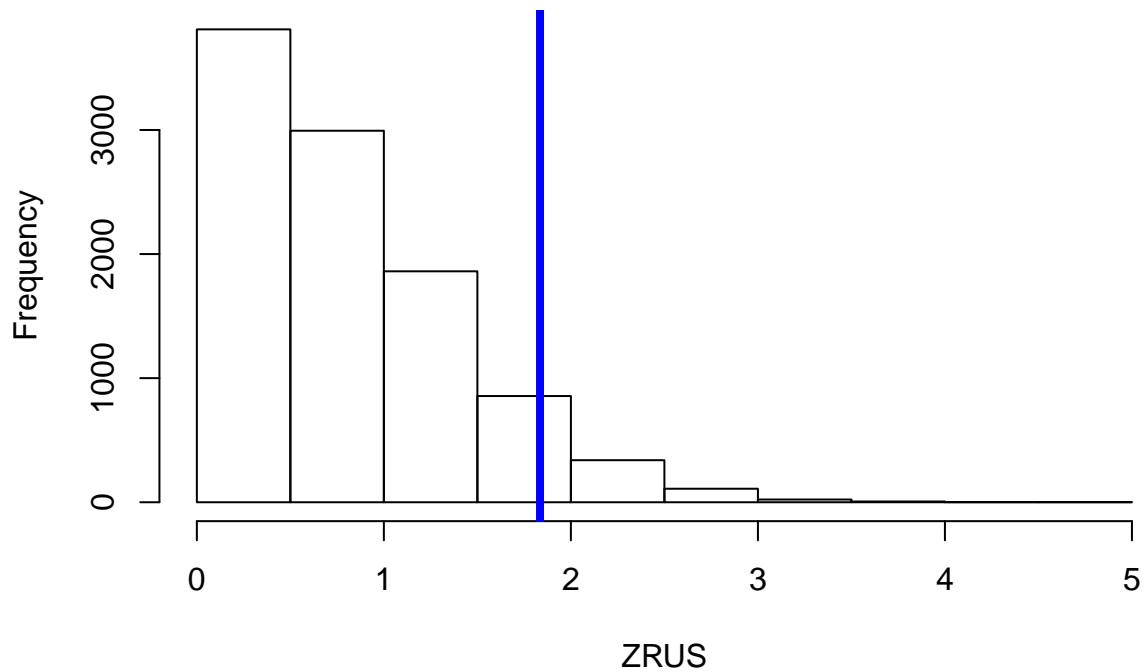
```
## [1] 2.083574
```

```
#P-Value of Time moving the grass (TMG) "B5"  
length(ZP0[ZP0>= ZP0true])/nr
```

```
## [1] 0.0399
```

```
#RUS  
par(mfrow=c(1,1))  
hist(ZRUS)  
abline(v=ZRUStrue, lwd=4, col="blue")
```

## Histogram of ZRUS



```
#True t-Value of Time moving the grass (TMG) "B6"  
ZRUStrue
```

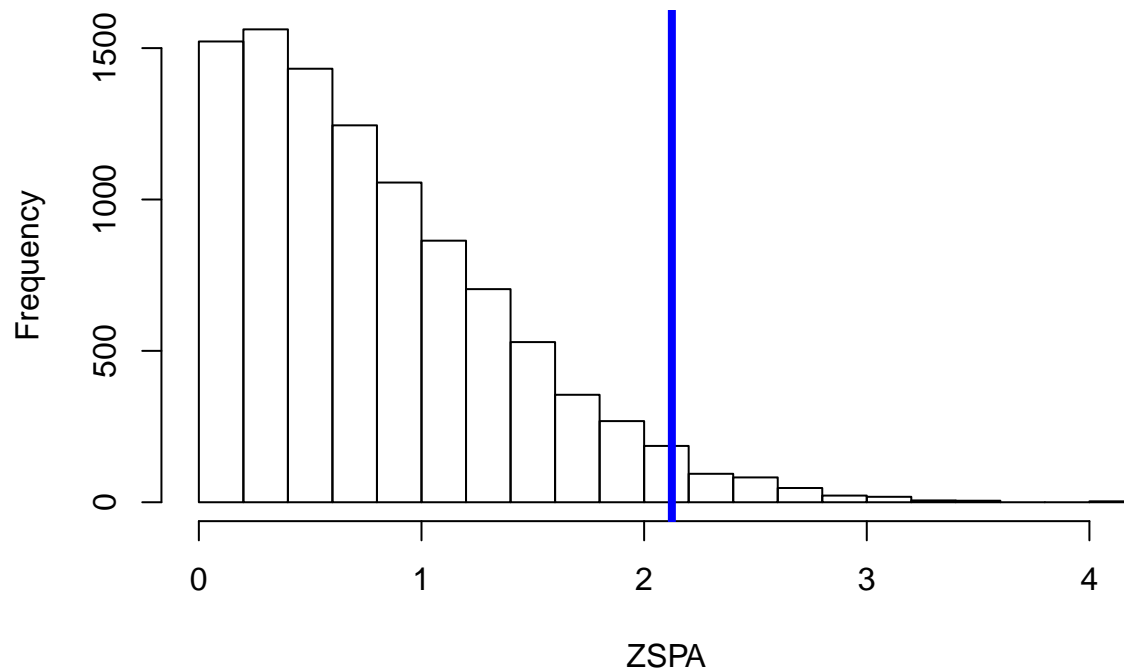
```
## [1] 1.834916
```

```
#P-Value of Time moving the grass (TMG) "B6"  
length(ZFR[ZRUS>= ZRUStrue])/nr
```

```
## [1] 0.0677
```

```
#SPA  
par(mfrow=c(1,1))  
hist(ZSPA)  
abline(v=ZSPAtrue, lwd=4, col="blue")
```

## Histogram of ZSPA



```
#True t-Value of Time moving the grass (TMG) "B7"  
ZSPAtrue
```

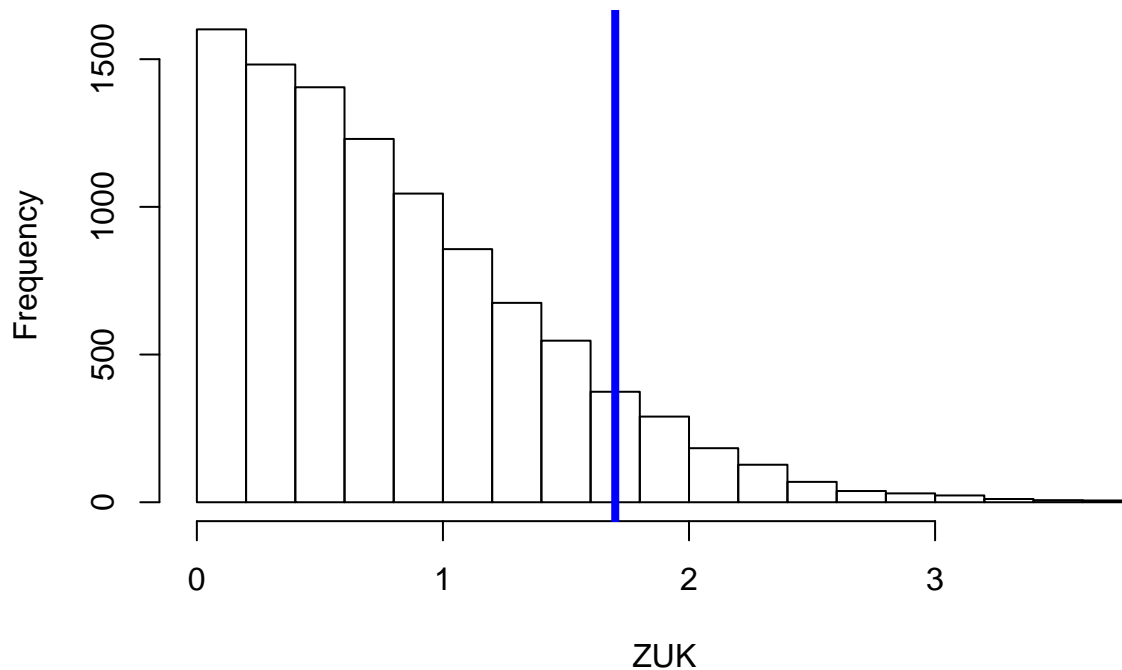
```
## [1] 2.124583
```

```
#P-Value of Time moving the grass (TMG) "B7"  
length(ZSPA[ZSPA>= ZSPAtrue])/nr
```

```
## [1] 0.0334
```

```
#UK  
par(mfrow=c(1,1))  
hist(ZUK)  
abline(v=ZUKtrue, lwd=4, col="blue")
```

## Histogram of ZUK



```
#True t-Value of Time moving the grass (TMG) "B8"
ZUKtrue
```

```
## [1] 1.700228
```

```
#P-Value of Time moving the grass (TMG) "B8"
length(ZUK[ZUK>= ZUKtrue])/nr
```

```
## [1] 0.0965
```

Some countries large p-values (p “greater than” 0.05) indicates weak evidence against the null hypothesis so, you fail to reject the null hypothesis (Ho is not rejected).

## Multiple Regression through the Origin

```
#install.packages("combinat")
suppressMessages(suppressWarnings(library("combinat")))
#install.packages("ape")
suppressMessages(suppressWarnings(library("ape")))
#install.packages("readxl")
suppressMessages(suppressWarnings(library("readxl")))

Gdata<-read_excel("WorldData.xlsx")

#   USA, BRA,   FR, GR, IT, PO, RUS,   SPA,   UK
US <- Gdata$USA
BRA <- Gdata$BRA
```

```

FR <- Gdata$FR
GR <- Gdata$GR
IT <- Gdata$IT
PO <- Gdata$PO
RUS <- Gdata$RUS
SPA <- Gdata$SPA
UK <- Gdata$UK

Ga<-data.frame(US,BRA,FR,GR,IT,PO,RUS,SPA,UK)

#number of rearrangements to be examined
nr<-10000

#Permutation method using number of rearrangements
lmorigin(US ~ BRA + FR + GR + IT + PO + RUS + SPA + UK, data =Ga, nperm=nr)

## Regression through the origin
## Permutation method = raw data
## Computation time = 58.920000 sec

##
## Regression through the origin
##
## Call:
## lmorigin(formula = US ~ BRA + FR + GR + IT + PO + RUS + SPA +      UK, data = Ga, nperm = nr)
##
## Coefficients and parametric test results
##
##      Coefficient Std_error t-value Pr(>|t|)
## BRA      0.157980  0.085498  1.8478  0.06796 .
## FR       0.240237  0.115025  2.0886  0.03960 *
## GR      -0.110315  0.107439 -1.0268  0.30731
## IT       0.357773  0.160494  2.2292  0.02832 *
## PO      -0.224204  0.108870 -2.0594  0.04238 *
## RUS      0.088970  0.120572  0.7379  0.46252
## SPA      0.074966  0.106512  0.7038  0.48338
## UK       0.317847  0.116056  2.7387  0.00745 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Two-tailed tests of regression coefficients
##
##      Coefficient p-param  p-perm
## BRA      0.157980  0.0680 0.066793 .
## FR       0.240237  0.0396 0.036396 *
## GR      -0.110315  0.3073 0.306069
## IT       0.357773  0.0283 0.028297 *
## PO      -0.224204  0.0424 0.041196 *
## RUS      0.088970  0.4625 0.459854
## SPA      0.074966  0.4834 0.488251
## UK       0.317847  0.0074 0.007299 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##

```



```

## One-tailed tests of regression coefficients:
## test in the direction of the sign of the coefficient
##
##      Coefficient p-param p-perm
## BRA      0.157980  0.0340 0.0358 *
## FR       0.240237  0.0198 0.0178 *
## GR      -0.110315  0.1537 0.1527
## IT       0.357773  0.0142 0.0151 *
## PO      -0.224204  0.0212 0.0201 *
## RUS      0.088970  0.2313 0.2297
## SPA      0.074966  0.2417 0.2487
## UK       0.317847  0.0037 0.0039 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.06207567 on 89 degrees of freedom
## Multiple R-square: 0.9780958    Adjusted R-square: 0.9761269
##
## F-statistic: 496.7687 on 8 and 89 DF:
##      parametric p-value      : 2.234866e-70
##      permutational p-value: 9.999e-05
## after 10000 permutations of raw data
##

```

Both (parametric and permutational) with some large p-values (p “greater than” 0.05) indicates weak evidence against the null hypothesis so, you fail to reject the null hypothesis (Ho is not rejected).

## Significance of the all countries in A1, A2 and A3

The Gdata was divided in three selected groups of countries using the USA as benchmark (i.e. dependent country); with the other countries as independent countries.

The three groups of countries are as followed:

- A1 - Euro zone countries: Germany, France, Italy, Spain
- A2 - Countries from EU but not from Euro zone: Poland, United Kingdom
- A3 - Not EU or Euro zone countries: Brazil, Russia

It has been decided to pick these countries for A1 to compare the countries with strong economy (Germany, France) with countries with lower economy (Italy, Spain) within Euro currency zone. For the same reason, the United Kingdom with Poland for A2. For A3, USA as a representative of a strong and stable economy and Russia, Brazil as countries with growing economy.

The testing began whether a set of independent variables has no partial effect on the dependent variable, “US”.

The models are:

$US\_A1 = B0 + B1FR + B2GR + B3IT + B4SPA + e$ ;  $US\_A2 = B0 + B1PO + B2UK + e$ ;  $US\_A3 = B0 + B1BRA + B2RUS + e$

**Null Hypothesis:** The initial assumption is that there is no relation, which is expressed as:

**Ho\_A1:**  $B1 = B2 = B3 = B4 = 0$ ; **Ho\_A2:**  $B1 = B2 = 0$ ; **Ho\_A3:**  $B1 = B2 = 0$

**Alternative Hypothesis:** At least one of the independent variables IS useful in explaining/predicting US, expressed as:

H1\_A1: At least one Bi is “not equal to” 0; H1\_A2: At least one Bi is “not equal to” 0;  
H1\_A3: At least one Bi is “not equal to” 0

## LM and GLMs Comparison

The results is displayed below:

```
# Model A1, A2 and A3

#install.packages("texreg")
suppressMessages(suppressWarnings(library("texreg")))
#install.packages("readxl")
suppressMessages(suppressWarnings(library("readxl")))

Gdata<-read_excel("WorldData.xlsx")

#   USA, BRA,   FR, GR, IT, PO, RUS,   SPA,   UK
US <- Gdata$USA
BRA <- Gdata$BRA
FR <- Gdata$FR
GR <- Gdata$GR
IT <- Gdata$IT
PO <- Gdata$PO
RUS <- Gdata$RUS
SPA <- Gdata$SPA
UK <- Gdata$UK

# Estimate with OLS (Model A1, A2 and A3):
regA1<-lm(US ~ FR + GR + IT + SPA)
regA2<-lm(US ~ PO + UK)
regA3<-lm(US ~ BRA + RUS)

# Estimate with GLS (MModel A1, A2 and A3):
regA11<-glm(US ~ FR + GR + IT + SPA, family = gaussian)
regA22<-glm(US ~ PO + UK, family = gaussian)
regA33<-glm(US ~ BRA + RUS, family = gaussian)

# Compare:
screenreg(l = list(regA1,regA11,regA2,regA22,regA3,regA33))
```

```
##
## =====
##           Model 1      Model 2      Model 3      Model 4      Model 5      Model 6
## -----
## (Intercept)      0.16 ***      0.16 ***      0.09          0.09          0.16 *      0.16 *
##                  (0.04)      (0.04)      (0.05)      (0.05)      (0.07)      (0.07)
## FR                0.39 **      0.39 **
##                  (0.12)      (0.12)
## GR               -0.14          -0.14
##                  (0.11)      (0.11)
## IT                0.25          0.25
##                  (0.16)      (0.16)
## SPA              0.02          0.02
##                  (0.09)      (0.09)
```

```

## PO                0.12      0.12
##                (0.07)    (0.07)
## UK                0.58 ***   0.58 ***
##                (0.10)    (0.10)
## BRA              0.11      0.11
##                (0.10)    (0.10)
## RUS              0.35 **    0.35 **
##                (0.13)    (0.13)
## -----
## R^2              0.48      0.33      0.11
## Adj. R^2        0.46      0.32      0.09
## Num. obs.       97      97      97      97      97      97
## RMSE            0.06      0.07      0.08
## AIC              -257.42    -237.16    -209.00
## BIC              -241.98    -226.86    -198.70
## Log Likelihood   134.71     122.58     108.50
## Deviance         0.35      0.45      0.61
## =====
## *** p < 0.001, ** p < 0.01, * p < 0.05

```

Both (LM and GLM models) with some large p-values (p “greater than” 0.05) indicates weak evidence against the null hypothesis so, you fail to reject the null hypothesis (Ho is not rejected).

## CONCLUSION

Completing the testing process of the significance of the Global Model, Model\_A1, Model\_A2 and Model\_A3, it can be concluded that at least one of the countries stability index is not significant.

After the global financial crisis, the topic of financial stability became very important. It became clear that for detecting financial stability risks, crisis and taking appropriate decisions one should be able to project future financial stability development. Globalization, financial innovations and STR(Scientific-technological revolution) started many financial processes bringing new transmission channels and making the process of measuring financial stability more complicated. To assess financial stability, a big number of economic and financial indicators must be monitored; to protect a stable country’s financial stability index globally against high risk countries.

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