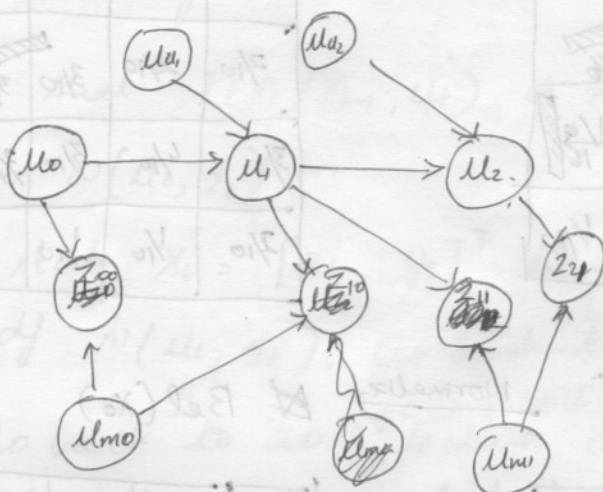


# SAM:



Consider a 3 state  
vehicle/robot state,  
2 landmark situations.  
We derive SAM  
equations for this.

$z_{00} \rightarrow$  landmark member.  
 $\hookrightarrow$  time instant

Find  $\arg\max_{u, u_m}$

$$\text{Find } u^*, u_m^* = \arg\max_{u, u_m} P(u, u_m / u_0, z).$$

$$= \arg\max_{u, u_m} P(u, u_m, u_0, z)$$

$$\propto P(u_2 / u_1, u_{01}) \cdot P(z_{24} / u_{m1}, u_2).$$

$$P(u_1 / u_0, u_{01}) \cdot P(z_{10} / u_{m1}) \cdot P(z_{00} / u_{m0})$$

$$P(u_0) \cdot P(z_{00} / u_{m0})$$

Actually you know how to compute all of  
the above term as they are just motion  
models/sensor model. Given those distributions  
it is merely a search in the space of

$u$ 's &  $u_m$ 's to get the maximum value

$$u = \begin{bmatrix} u_0 \\ u_1 \\ u_2 \end{bmatrix}^T = \begin{bmatrix} u_{x0}, u_{y0}, u_{z0}, \dots, u_{02} \end{bmatrix}^T$$

$$u_m = [u_{m0x} \ u_{m0y} \ u_{m1x} \ u_{m1y}]^T$$

$$P(u_t / u_{t-1}, u_{ut}) = P(u_t / \hat{u}_t)$$

$$\propto \exp \{ (u_t - \hat{u}_t)^T \Sigma_t^{-1} (u_t - \hat{u}_t) \}$$

~~Now we follow Kaess derivation closely by generalizing for a time  $t$~~

$$\text{Consider } P(u_2 / u_1, u_{u2}) =$$

$$\text{Consider } P(u_2 / u_1, u_{u2}) = P(u_2 / \hat{u}_2)$$

$$\propto \exp \left\{ -\frac{1}{2} [u_2 - \hat{u}_2]^T \Sigma_{u_2}^{-1} [u_2 - \hat{u}_2] \right\} \rightarrow (1)$$

where  $\hat{u}_2$  = predicted odometry reading  
 $= f(u_1, u_{u2})$ ;  $u_2$  = measured/obtained odometry

$$\Sigma_2 = F \Sigma_1 F^T + G \Sigma_{u_2} G^T$$

Kaess write  $[u_2 - \hat{u}_2]^T \Sigma_{u_2}^{-1} [u_2 - \hat{u}_2]$  as  $\|u_2 - \hat{u}_2\|_{\Sigma_{u_2}}^2$

Hence (1) is  $\propto \exp \left\{ -\frac{1}{2} \|u_2 - \hat{u}_2\|_{\Sigma_{u_2}}^2 \right\} \rightarrow (2)$ .

$$\text{Consider } P(z_{2f} / u_m, u_2) = P(z_{2f} / \hat{z}_{2f})$$

$$\propto \exp \left\{ -\frac{1}{2} \|z_{2f} - \hat{z}_{2f}\|_H^2 \right\}$$

where  $H = \frac{\partial h}{\partial z}$

Why are we considering only  $\hat{\Sigma}_{u_2}$  & not  $\Sigma_z$ .

1) Because  $P(\mu_2/\mu_1, u_2)$  is a function of only control noise  $\hat{\Sigma}_{u_2}$  and it is only the belief  $\hat{Bel}(x_t) = N(\hat{u}_t, \hat{\Sigma}_t)$   
 $= \int P(x_t/x_{t-1}, u_t) \cdot Bel(x_{t-1})$  is a function of the state covariance  $\hat{\Sigma}_t$

2) Another way of looking at it is that the optimization routine searches over all possible  $\mu_1$  and for every such  $\mu_1$  it considers it only needs to consider control noise for that ~~the~~  $\mu_1$ 's evolution to  $\mu_2$

Consider  $P(z_{21}/\mu_2, \mu_{m1}) = P(z_{21}/\hat{z}_{21})$

$$\propto \exp\left(-\frac{1}{2} \|z_{21} - \hat{z}_{21}\|_{Q_{21}}^2\right) \quad (3)$$

where  $Q_{21}$  is the measurement noise covariance of the 1st landmark observed from second state / or second time sample.

$$\hat{z}_{21} = h(\hat{\mu}_2, \hat{\mu}_{m1}) \quad (4)$$

From (5) take Taylor series of  $\mu_i - f(\mu_{i-1}, u_i)$  to get

$$\mu_i^0 + \frac{\partial \mu_i}{\partial \mu_i} \delta \mu_i \bigg|_{\mu_i} - \left[ f(\mu_{i-1}^0) + \frac{\partial f(\mu_{i-1}, u_i)}{\partial \mu_{i-1}} \delta \mu_{i-1} \right] \bigg|_{\mu_{i-1}^0}$$

$$= \mu_i^0 + \delta \mu_i - f(\mu_{i-1}^0) - F \delta \mu_{i-1} \quad (6)$$




$$\text{Now } \mu^*, \mu_m^* = \underset{\mu, \mu_m}{\operatorname{argmax}} P(\mu, \mu_m | \mu_u, Z)$$

$$= \underset{\mu, \mu_m}{\operatorname{argmin}} (-\log P(\mu, \mu_m, \mu_u, Z))$$

$$= \sum_{i=1}^2 \frac{1}{2}$$

$$= \underset{\mu, \mu_m}{\operatorname{argmin}} \left\{ \sum_{i=1}^2 \|\mu_i - \hat{\mu}_i\|_{\Sigma_i}^2 + \sum_{i,k} \|\mathbf{z}_{ik} - \hat{\mathbf{z}}_{ik}\|_{\Sigma_{ik}}^2 \right\}$$

Only for those  $\mathbf{z}_k$ 's seen from  $\mu_i$

We now write  $\mu_i = \mu_i^0 + \delta \mu_i$  as a perturbation around the obtained odometry. 

$\hat{\mu}_i = f(\mu_{i-1}, \mu_i)$  is written as

$$f(\mu_{i-1}^0) + \frac{\partial f(\mu_{i-1}^0, \mu_i)}{\partial \mu_i} \delta \mu_i = f(\mu_{i-1}^0) + F \delta \mu_{i-1}$$

Only we find Taylor series for  $h(\mu_i, \mu_m) - \mathbf{z}_k$ .

$$\mathbf{z}_k \approx h(\mu_i, \mu_m) \approx \frac{\partial h}{\partial \mu_i}$$

$$= h(\mu_i^0, \mu_m^0) + \frac{\partial h}{\partial \mu_i} \delta \mu_i \Big|_{\mu_i^0} + \frac{\partial h}{\partial \mu_m} \delta \mu_m \Big|_{\mu_m^0} - \mathbf{z}_k$$

$$= h(\mu_i^0, \mu_m^0) + H_{ik} \delta \mu_i + J_{ik} \delta \mu_m - \mathbf{z}_k \rightarrow (7)$$

Now  $\mu_i = f(\mu_{i-1}, \mu_i)$

$$= \mu_i^0 + \delta \mu_i - f(\mu_{i-1}^0, \mu_i) - F \delta \mu_{i-1}$$

$$= \delta \mu_i - F \delta \mu_{i-1}$$

~~Now  $f(u_{i-1}, u_i) = u_i$~~

Now  $f(u_{i-1}, u_i) = u_i$

$$= f(u_{i-1}^o, u_i) + F \delta u_{i-1} - u_i - \delta u_i$$

$$= -a_i + F \delta u_{i-1} + G \delta u_i, \text{ where } G = -I_{3 \times 3}$$

(odometry prediction error)  $\rightarrow$  (8)

and  $h(u_i, u_{mk}) = z_{ik}$

$$= h(u_i^o, u_{mk}) - z_{ik} + H_{ik} \delta u_i + J_k \delta u_{mk}$$

$$= H_{ik} \delta u_i + J_k \delta u_{mk} - c_k \rightarrow 9$$

$\rightarrow$  measurement error.

$\therefore$  We want to find

$$\arg \min_{u, u_m} \left\{ \sum_{i=1}^2 \| F \delta u_{i-1} + G \delta u_i - a_i \|^2_{\Sigma_{u_i}} + \sum_{i,k} \right.$$

$$\left. \sum_{i,k} \| H_{ik} \delta u_i + J_k \delta u_{mk} - c_k \|^2_{\Sigma_{ik}} \right\}$$

We write the  $F \delta u_{i-1} + G \delta u_i - a_i$  as (10)

$$\begin{bmatrix} F & G \end{bmatrix} \begin{bmatrix} \delta u_{i-1} \\ \delta u_i \end{bmatrix} - a_i$$

$\begin{matrix} 3 \times 3 & 3 \times 3 \\ 3 \times 6 & 6 \times 1 \end{matrix}$ 
 $\begin{matrix} 3 \times 1 \\ 3 \times 1 \end{matrix}$ 
 $\begin{matrix} 6 \times 1 \\ 3 \times 1 \end{matrix}$

$$= A x - b$$

$\therefore$  First term in (10) is of the form

$$[A x - b]^T [\Sigma_u]^{-1} [A x - b]$$

Let  $\Sigma_{ii}^{-1} = \gamma$  and since  $\gamma$  is a symmetric matrix then  $[\gamma] = [\gamma]^{1/2} [\gamma]^{1/2}$ , where  $\gamma^{1/2}$  is the square root matrix of  $\gamma$ .

~~Square root matrix:  $R$  is said to be square root matrix of  $A$ , where  $A = VDV^T$ , then~~

~~$$R = VD^{1/2}V^T$$~~

~~$$RR = [VD^{1/2}V^T][VD^{1/2}V^T] = VD^{1/2}D^{1/2}V^T = VDV^T = A,$$~~

How do you find the square root of a matrix?

The covariance matrix and its inverse are symmetric. A symmetric matrix, ~~is also~~  $B$  is diagonalizable as

→  $B = VDV^{-1}$ , where  $D$  is a diagonal matrix

→ The diagonal entries of  $D$  are ~~the~~ eigenvalues of  $B$ .

→ Column vectors of  $V$  are the <sup>vector</sup> eigenvalues of  $B$ .

→ The column vectors are orthonormal

Let  $S$  be the sqrt matrix of  $D$ , <sup>obtained</sup> then by taking square root of diagonal entries of  $D$

Then let  $R = VS V^T$ , then

$$RR = VS V^T VS V^T = VSS V^T = VDV^T = B$$

(Since  $V$  is an orthonormal matrix

$$VV^T = VV^T = I \text{ \& } V^T = V^{-1}).$$

Then  $R$  is the sqrt matrix of  $B$  or  $R = B^{1/2}$



$$\begin{aligned}
 \text{Then } & [Ax-b]^T \Sigma^{-1} [Ax-b] \\
 &= [Ax-b]^T \gamma [Ax-b] \\
 &= [Ax-b]^T \gamma^{1/2}
 \end{aligned}$$

$$\text{Also } RR^T = VSV^T[VSV^T]^T = VDV^T = B$$

Not only  $RR = B$ ,  $RR^T = B$ , when  $B$  is symmetric and diagonal.

$$\begin{aligned}
 \text{Now } & [Ax-b]^T \Sigma^{-1} [Ax-b] \\
 &= [Ax-b]^T \gamma [Ax-b]
 \end{aligned}$$

$\gamma$  is symmetric &  $\gamma = RR^T$

$$= [Ax-b]^T RR^T [Ax-b] =$$

$$\begin{aligned}
 &= \|R^T [Ax-b]\|^2 = \|\gamma^{1/2} [Ax-b]\|^2 \quad \text{--- (11)} \\
 &= \|\gamma^{1/2} e\|^2
 \end{aligned}$$

Second term on (10) is written

$$\text{similarly as } \|n^{1/2} [Cy-d]\|^2 \quad \text{--- (12)}$$

$$= \|n^{1/2} f\|^2$$

(11) + (12) is of the form

$$[\gamma^{1/2} e]^T [\gamma^{1/2} e] + [n^{1/2} f]^T [n^{1/2} f]$$

$$\text{or } [Re]^T [Re] + [Sf]^T [Sf]$$

$$= \begin{bmatrix} R & 0 \\ 0 & S \end{bmatrix} \begin{bmatrix} e \\ f \end{bmatrix}^T \begin{bmatrix} R & 0 \\ 0 & S \end{bmatrix} \begin{bmatrix} e \\ f \end{bmatrix}$$

--- (13)

hence all the computations can be lumped into a minimization

$$= [Ax-b]^T \gamma [Ax-b]$$

$$= [Ax-b]^T \gamma [Ax-b]$$

$$\theta^* = \arg \min_{\theta} \|A\theta - b\|^2$$

$$\text{Also } RR^T = V \Lambda V^T R \theta^T = V \Lambda V^T = B$$

Next  $RR^T = B$ ,  $\theta^T$  is solved by  $RR^T \theta = B \theta$ , when  $B$

is symmetric  $\theta$  of value at each

$$\text{New } [Ax-b]^T \gamma [Ax-b]$$

$$= [Ax-b]^T \gamma [Ax-b]$$

$\gamma$  is symmetric &  $\gamma = RR^T$

$$= [Ax-b]^T R R^T [Ax-b] = \|R^T [Ax-b]\|^2$$

$$= \|R^T [Ax-b]\|^2 = \|\gamma^{1/2} [Ax-b]\|^2 \quad (11)$$

$$= \|\gamma^{1/2} e\|^2$$

Second term in (10) is written

$$\text{similarly as } \|n^{1/2} [Cy-d]\|^2 \quad (12)$$

$$= \|n^{1/2} f\|^2$$

(11) + (12) is of the form

$$[\gamma^{1/2} e]^T [\gamma^{1/2} e] + [n^{1/2} f]^T [n^{1/2} f]$$

$$= [R e]^T [R e] + [S f]^T [S f]$$

$$= \begin{bmatrix} R & 0 \\ 0 & S \end{bmatrix} \begin{bmatrix} e \\ f \end{bmatrix}^T \begin{bmatrix} R & 0 \\ 0 & S \end{bmatrix} \begin{bmatrix} e \\ f \end{bmatrix}$$

(13)



Hence all the summation terms in (10)  
can be lumped into a minimization  
of the form

$$\theta^* = \arg \min_{\theta} \|A\theta - b\|^2$$

$\theta$  is solved by SVD or QR etc.

How to solve for  $\theta$  incrementally  
- ISAM?