

STAT 231

November 2, 2016

Roadmap

- Wrap up Interval estimation using Confidence Intervals.
- Relationship between likelihood Intervals and C.I for large samples
- Hypothesis Testing

Case I : C.I for μ ; Gaussian (μ, σ)
 σ known.

Case II : Binomial problem. $Y \sim \text{Bin}(n, \theta)$
C.I for θ

Case III : $Y_1, \dots, Y_n \sim \mathcal{G}(\mu, \sigma)$
indep. μ, σ unknown.

C.I for μ ?

PIVOTAL QUANTITY : $\frac{\bar{Y} - \mu}{\frac{S}{\sqrt{n}}}$ ✓

PIVOTAL DISTRIBUTION

Student's t distribution:

with $n-1$ df

COVERAGE INTERVAL

$$\bar{Y} \pm t^* \frac{S}{\sqrt{n}}$$

where t^* = from the t -table

Row = $n-1$

Column = ~~the~~ depends
on the level of
confidence.

$$\text{CONFIDENCE INTERVAL: } \left[\bar{y} \pm t^* \frac{s}{\sqrt{n}} \right]$$

\bar{y} : sample mean

$$s = \sqrt{\frac{1}{n-1} \sum (y_i - \bar{y})^2}$$

Case IV $Y_1, \dots, Y_n \sim G_e(\mu, \sigma)$

μ, σ unknown.

Objective: To find the C.I for σ

$$\text{PIVOTAL QUANTITY} = \frac{(n-1)S^2}{\sigma^2}$$

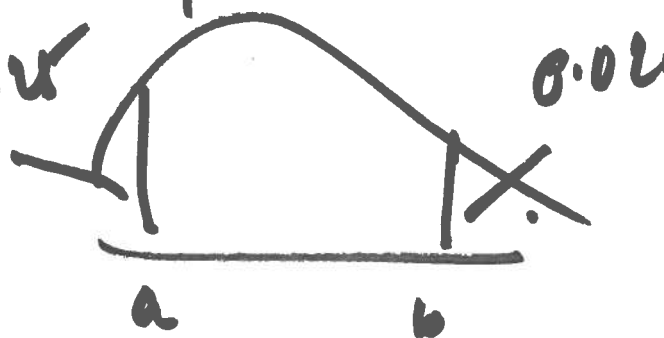
$$\text{PIVOTAL DISTRIBUTION} = \chi^2_{n-1}$$

$$df = n - 1$$

COVERAGE INTERVAL

$$\left[\frac{(n-1)S^2}{b}, \frac{(n-1)S^2}{a} \right]$$

where b, a are computed from the χ^2 table



$$\text{Confidence Interval for } \sigma^2 = \left[\frac{(n-1)s^2}{b}, \frac{(n-1)s^2}{a} \right]$$

$$\text{C.I for } \sigma : \left[\sqrt{\frac{(n-1)s^2}{b}}, \sqrt{\frac{(n-1)s^2}{a}} \right]$$

Equal-tailed Confidence Interval
but that need not be the case.

Case I Other distributions
with large sample sizes

(a) Poisson : $Y_1, \dots, Y_n \sim \text{Poi}(\mu)$

P.Q :
$$\frac{\bar{Y} - \mu}{\sqrt{\bar{Y}/n}}$$

PIVOTAL DISTRIBUTION = Z

Coverage : $\bar{Y} \pm z^* \sqrt{\bar{Y}/n}$

C.I =
$$\boxed{\bar{y} \pm z^* \sqrt{\bar{y}/n}}$$

(b) Exponential distribution.

$$Y_1, \dots, Y_n \sim \text{Exp}(\mu)$$

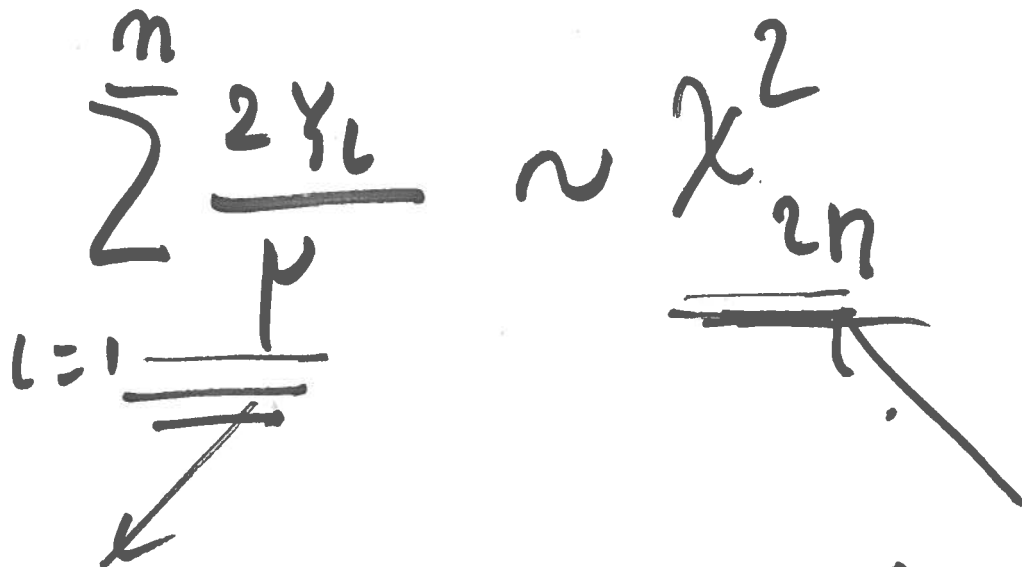
Objective: To find the C.I for μ .

(The following result is true for all n)

Result: If $Y \sim \text{Exp}(\mu)$, then

$$\frac{2Y}{\mu} \sim \text{Exp}(2) \text{ --- } \textcircled{1}$$
$$= \chi^2(2)$$

From (1),

$$\sum_{l=1}^n \frac{2Y_l}{\mu} \sim \chi^2_{2n}$$


PIVOTAL QUANTITY

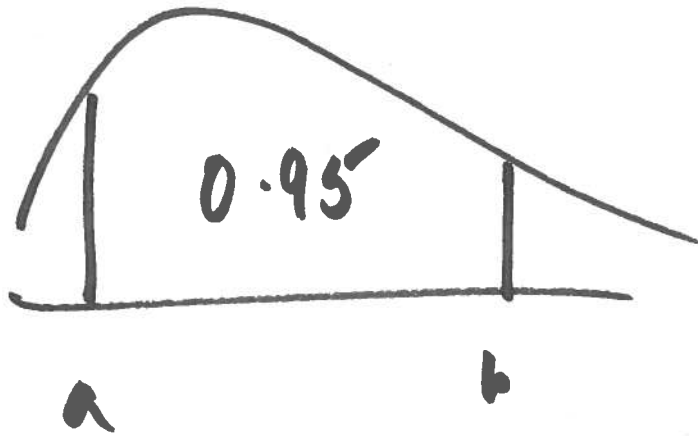
PIVOTAL

DISTRIBUTION

Example: $n = 10$

Go to the χ^2_{20} table and find the two end points a and b

95% C.I



$$P(a \leq \chi^2_{20} \leq b) = 0.95$$

$$P\left(a \leq \underbrace{\sum_{i=1}^n \frac{2Y_i}{\mu}}_{\chi^2_{2n}} \leq b\right) = 0.95$$

$$\mu \leq \frac{2 \sum Y_i}{a} \quad \mu \geq \frac{2 \sum Y_i}{b}$$

Coverage Interval is

$$\left(\frac{2 \sum Y_i}{b}, \frac{2 \sum Y_i}{a} \right)$$

Confidence Interval

$$\left(\frac{2 \sum y_i}{b}, \frac{2 \sum y_i}{a} \right)$$

$$\left(\frac{2 \bar{y} n}{b}, \frac{2 n \bar{y}}{a} \right)$$

There are two approaches to interval estimation

- Through Likelihood function.
- Through Sampling distributions.

~~Also~~

Question: How are these two intervals related?

For large samples, we call the following theorem.

Theorem: If $L(\theta)$ is based on (Y_1, \dots, Y_n)
 θ = unknown parameter, then

$$\cancel{\Lambda(\theta)} = \frac{\cancel{L(\theta)}}{\cancel{L(\hat{\theta})}} \sim \cancel{\chi^2_i}$$

if n is large.

$\tilde{\theta}$ = estimator corresponds

$$\chi^2(\theta) = -2 \log \frac{L(\theta)}{L(\hat{\theta})} \sim \chi^2_1$$

$\Lambda(\theta) = \text{LIKELIHOOD RATIO}$
 TEST STATISTIC.

Theorem: $p = 0.1$

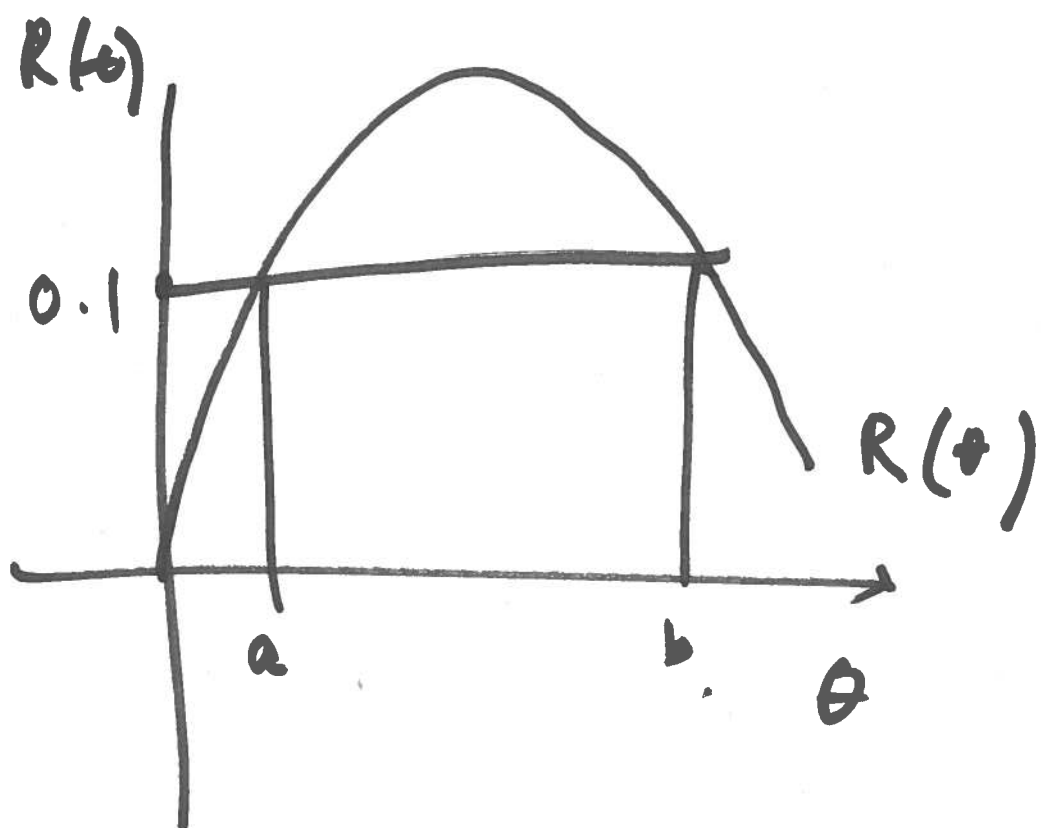
A $100 p\%$ likelihood interval.
is an approximate $100 q\%$ C.I.

where $q = 2 P(Z \leq \sqrt{-2 \log p})$

Application

10% Likelihood Interval.

$$= 96.8\% \quad 2 P(Z \leq \sqrt{-2 \log 0.1})$$



10%

Outline of the proof

100 p% l.i

$$= \{ \theta : R(\theta) \geq p \}$$

$$= \{ \theta : \underline{-2 \log R(\theta)} \leq -2 \log p \}$$

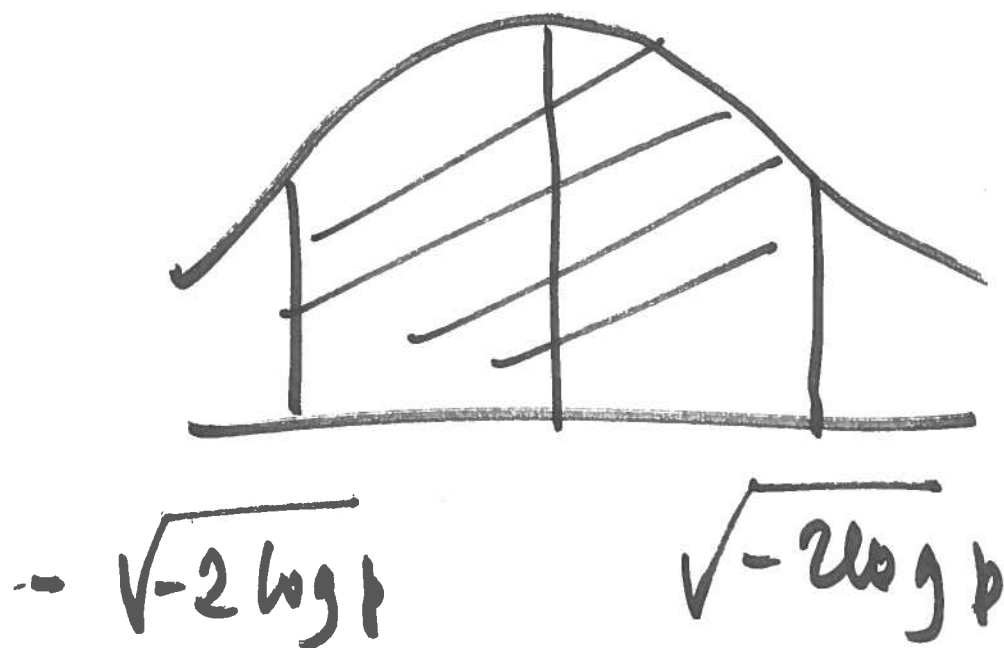
$$-2 \log \frac{L(\theta)}{L(\hat{\theta})}$$

Corresponding Coverage probability

$$P\left(-2 \log \underbrace{\frac{L(\theta)}{L(\hat{\theta})}} \leq -2 \log p\right)$$

$$P\left(\Lambda(\theta) \leq -2 \log p\right)$$

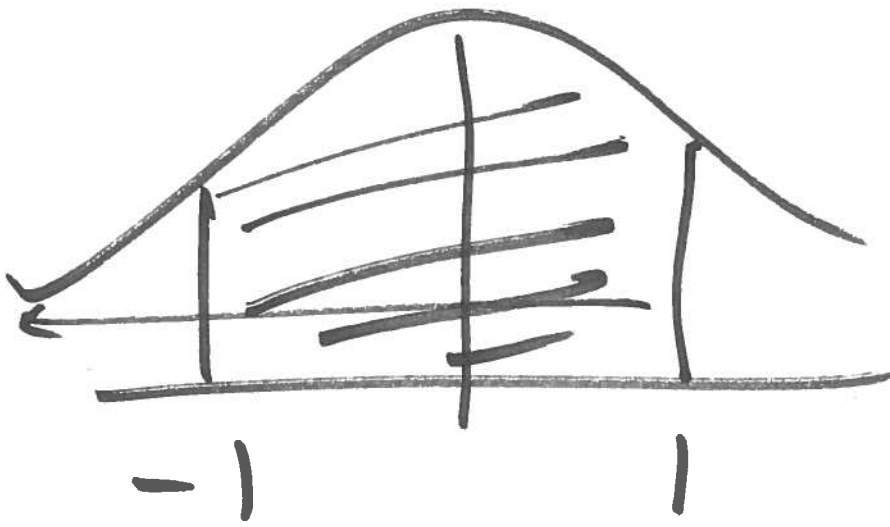
$$\begin{aligned} &P\left(Z^2 \leq -2 \log p\right) \\ &= P\left(|Z| \leq \sqrt{-2 \log p}\right) \end{aligned}$$



$$P(Z \leq \sqrt{-2\log p}) = (1 -$$

$$P(Z \leq \sqrt{-2\log p})$$

$$2 P(Z \leq \sqrt{-2\log p}) - 1$$



$$P(Z \leq 1) - (P(Z \leq -1))$$

$$= P(Z \geq 1)$$

$$= 1 - P(Z \leq 1)$$

$p = 0.1 \Rightarrow$ Confidence Interval
96.8%

$p = 0.01 \Rightarrow$

Example: Suppose we have a 95% C.I. What likelihood interval would this correspond to?

$$\left\{ \theta : R(\theta) \geq e^{-\frac{z^*{}^2}{2}} \right\}$$

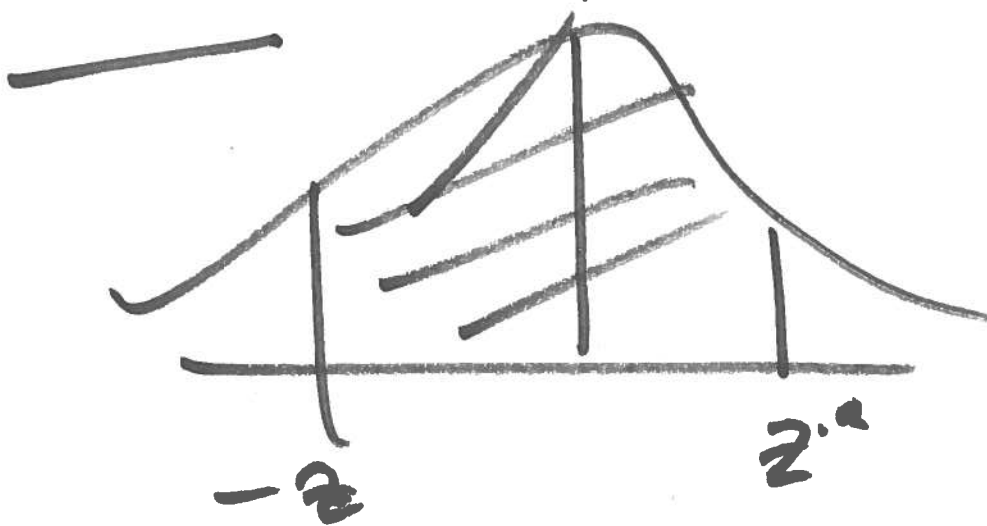
where z^* = value from the Z -table

for a 95% C.I., the p would

$$be \quad e^{-\frac{1.96^2}{2}} = 14.65\%$$

For large samples, a 95% C.I
is equivalent to a 15% l.i
approximately

90% C.I $0.9 Z^* = 1.64$



95% C.I

$$P(-1.96 \leq z \leq 1.96) = 0.95$$

$$\Rightarrow P(z^2 \leq 1.96^2) = 0.95$$

$$P(\chi^2_1 \leq 1.96^2) = 0.95$$

$$P(\underline{\Lambda}(\theta) \leq 1.96^2) = 0.95$$

$$P\left(-2 \log \frac{L(\theta)}{L(\hat{\theta})} \leq 1.96^2\right)$$

$$P\left(\frac{L(\theta)}{L(\hat{\theta})} \geq \underline{\underline{e^{-1.96^2/2}}}\right)$$

$R(\theta) \geq p$

C.I

90%

→

26% l.i

