Approximate accelerated stochastic simulation of chemically reacting systems

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Overview

- Stochastic kinetic models
- Issues with the Direct method
- τ -leap method (Gillespie, 2001)
 - Leap conditions
 - Mid-point estimation
- Examples

Source code (R and $\[Mathbb{E}T_EX\]$ of these slides):

https://github.com/csgillespie/talks/

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Stochastic kinetic models

A biochemical network is represented as a set of pseudo-biochemical reactions:

u species & *v* reactions

$$R_i: \quad p_{11}\mathcal{X}_1 + p_{12}\mathcal{X}_2 + \cdots + p_{1u}\mathcal{X}_u \xrightarrow{c_i} q_{11}\mathcal{X}_1 + q_{12}\mathcal{X}_2 + \cdots + q_{1u}\mathcal{X}_u$$

Stochastic rate constant c_i .

Hazard/instantaneous rate: $h_i(X_t, c_i)$ where $X_t = (X_{1,t}, \ldots, X_{u,t})$ is the current state of the system.

Under mass-action stochastic kinetics, the hazard function is proportional to a product of binomial coefficients, with

$$h_i(X_t, c_i) = c_i \prod_{j=1}^u {X_{j,t} \choose p_{ij}}.$$

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Stochastic kinetic model

- Describe the SKM by a Markov jump process (MJP)
- The effect of reaction R_k is to change the value of each species X_i by $q_{ji} p_{ji}$
- The stoichiometry matrix *S* has elements $s_{ij} = q_{ji} p_{ji}$
- It can be shown that the time to the next reaction is

$$t \sim \textit{Exp}(h_0(X_t, c))$$
 where $h_0(X_t, c) = \sum_{i=1}^{v} h_i(X_i, c_i)$

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and the reaction is of type *i* with probability $h_i(X_t, c_i)/h_0(X_t, c)$

The process is easily simulated using the Direct method (Gillespie algorithm)

Example: Lotka-Volterra system

• R₁: Prey reproduction

 $\mathcal{X}_1 \xrightarrow{c_1} 2\mathcal{X}_1$

• R₂: Prey death, predator reproduction

$$\mathcal{X}_1 + \mathcal{X}_2 \xrightarrow{c_2} 2\mathcal{X}_2$$

R₃: Predator death

$$\mathcal{X}_2 \xrightarrow{c_3} \emptyset$$



$$\begin{array}{l} \textbf{X}(\textbf{0}) = (100, 100) \\ \textbf{c} = (0.5, 0.0025, 0.3) \end{array}$$

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- Initialisation: initial conditions, reactions constants, and random number generators
- **Propensities update:** Update each of the v hazard functions, $h_i(x)$
- **9** Propensities total: Calculate the total hazard $h_0 = \sum_{i=1}^{V} h_i(x)$
- Seaction time: $\tau = -\ln[U(0,1)]/h_0$ and $t = t + \tau$
- Seaction selection: A reaction is chosen proportional to it's hazard
- **Reaction execution:** Update species
- Iteration: If the simulation time is exceeded stop, otherwise go back to step 2

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Approximations

Relax some assumptions (e.g. discreteness and stochasticity) in order to make simulation faster and more scalable

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- Diffusion approximation / chemical Langevin equation (CLE)
- Linear noise approximation (LNA) / moment closure (2MA)
- ODE
- Hybrid discrete-continuous models

Poisson leap

- If all reactions are zeroth-order, then the model is a homogeneous Poisson process
- Hence the number of reactions in (t_0, t_1) follows a Poisson distribution
- For a more general model, if we consider a *small* time interval, (t, t + Δt), then:
 - the hazard rates should be approximately constant
 - the number of reactions (of a given type) can be sampled from a Poisson distribution

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• A balance between speed and accuracy

Poisson leap method

Set t = 0. Initialise the rate constants and the initial molecule numbers x

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- Calculate $h_i(x, c_i)$, for i = 1, ..., v, and simulate the *v*-dimensional reaction vector *r*, with *i*th entry a $Po(h_i(x, c_i)\Delta t)$ random quantity
- Opdate the state according
- Update $t := t + \Delta t$
- Output t and x. If $t < T_{max}$ return to step 2

Poisson leap method

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How do you chose Δt ?

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- Suppose we are interested in parameter inference
- A possible prior could be independent Uniform priors over U(-8,8) for each log(c_i)
- Three samples from this prior yield very different realisations
 - Probability of extinction by time t = 30, is around 0.86
- Each simulation would require a very different Δt



Basic τ -leap method

- Suppose a temporal leap τ will result in a state change λ
- Choose a value of τ that satisfies the *leap condition*. For each reaction, R_i , we want

$$|h_j(\mathbf{x}+\lambda)-h_j(\mathbf{x})|$$

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to be small

- Sample $k_j \sim P(h_j(\mathbf{x})\tau)$
- Compute λ
- Set $t := t + \tau$ and $\mathbf{x} := \mathbf{x} + \lambda$

Choosing τ

- If the reactions don't depend on x, the leap condition will be satisfied exactly for any τ (and so exact)
- If population numbers are large, then it would take a large number of reactions to "noticeably" change the hazard functions
- If satisfying the leap condition requires a very small value τ << 1/h₀(**x**), then we may as well use the exact SSA

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A procedure for selecting τ (Gillespie 2001)

• The expected net change in $(t, t + \tau)$ will be:

$$ar{\lambda} = \sum_{j=1}^{\nu} [h_j(\mathbf{x})\tau] \mathbf{s}_j = \tau \xi(\mathbf{x})$$

 So we require that the *expected* changes in the propensity functions in time τ, are bounded by some fraction of all propensity functions, i.e.

$$|h_j(\mathbf{x} + \lambda) - h_j(\mathbf{x})| < \epsilon h_0(\mathbf{x}) \quad \text{for } j = 1, \dots, v.$$

• Estimate the difference using a Taylor expansion:

$$h_j(\mathbf{x} + \lambda) - h_j(\mathbf{x}) \simeq \sum_{i=1}^u \tau \xi_i(\mathbf{x}) \frac{\partial}{\partial x_i} h_j(\mathbf{x})$$

Defining

$$b_{ji}(\mathbf{x}) \equiv \frac{\partial h_j(\mathbf{x})}{\partial x_i}$$
 $(j = 1, \dots, v; i = 1, \dots, u)$

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A procedure for selecting au

• The requirement becomes

$$\tau \left| \sum_{i=1}^{u} \xi_i(\mathbf{x}) b_{ji}(\mathbf{x}) \right| \leq \epsilon h_0(\mathbf{x}) \quad (j = 1, \dots, v)$$

• The largest value of τ consistent with this condition (and hence optimal) is

$$\tau = \min_{j \in [1, v]} \left\{ \frac{\epsilon h_0(\mathbf{x})}{\left| \sum_{i=1}^{u} \xi_i(\mathbf{x}) b_{ji}(\mathbf{x}) \right|} \right\}$$

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Typical values of ϵ are around 0.05









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The estimated-midpoint technique

- The leap condition requires that the hazards functions do not "appreciably" change in the course of a leap
- But we want to take large leaps, so we will inevitably get computational errors
- This is similar to solving the ODE

$$\frac{dX(t)}{dt} = f(X)$$

using an Euler scheme

• A standard technique is to use a *second-order Runge-Kutta* or *modified Euler* method

$$X(t + \Delta t) = X(t) + f[X(t) + 0.5f(X(t))\Delta t]\Delta t$$

i.e. we use an Euler method to estimate the midpoint during $[t, t + \Delta t]$, then calculate the increment in X by evaluating the slope function f at that estimated midpoint

Example: Death model

• The death model contains a single reaction

$$\mathcal{X} \xrightarrow{\mu} \emptyset$$

and has hazard function $h_1(x, \mu) = \mu x$ and state change vector s = -1. • The solution to the CME is:

$$\Pr(X = x; t) = \binom{x_0}{x} e^{-\mu \tau (x_0 - x)} (1 - e^{-\mu \tau})^x$$

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Death model



$$\Pr(X=x;\tau) = \binom{x_0}{x} e^{-\mu\tau(x_0-x)} (1-e^{-\mu\tau})^x$$

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Death model: p-leap



If we perform a single leap of length τ , the number of executed reactions are

$$P_{\rho}(k; \mu x_0, \tau) = rac{e^{\mu x_0 \tau} (\mu x_0 \tau)^k}{k!}$$
 for $k = 0, ...$

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Death model: τ -leap + mid-point



We estimate the mid-point to be:

$$x' \equiv x_0 - \lfloor 0.5\tau \mu x_0 \rfloor$$

so the number of reactions executed is

$$P_{\rho}(k; \mu x_0, \tau) = rac{e^{\mu x' \tau} (\mu x' \tau)^k}{k!}$$
 for $k = 0, ...$

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For these parameter values and this model, we have the approximate relationship (obtained via simulation)

 $\epsilon \simeq$ 0.556 $imes \Delta t$

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Summary

- Similar issues arise when solving SDEs with the Euler-Maruyama scheme
 - In fact, if we substitute Poisson with Gaussian random numbers in the p-leap scheme, we get the Euler-Maruyama algorithm
- Choosing a fixed Δt for a wide range of parameter combinations doesn't make sense

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• Is it possible to these ideas when constructing bridges for SDEs?

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