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Borel-Cantelli Lemmas

Suppose that $\{A_n : n \ge 1\}$ is a sequence of events in a probability space. Then the event $A(i.o.) = \{A_n \text{ ocurrs for infinitely many } n \}$ is given by

$$A(i.o.) = \bigcap_{k=1}^{\infty} \cup_{n=k}^{\infty} A_n,$$

Lemma 1 Suppose that $\{A_n : n \ge 1\}$ is a sequence of events in a probability space. If

$$\sum_{n=1}^{\infty} P(A_n) < \infty, \tag{1}$$

then P(A(i.o.)) = 0; only a finite number of the events occur, wp1.

Proof:

Let $I_n = I\{A_n\}$ denote the indicator rv for the event A_n , and let

$$N = \sum_{n=1}^{\infty} I_n,$$

denote the total number of the events to occur. Then P(A(i.o.)) = 0 if and only if $P(N < \infty) = 1$. But if $E(N) < \infty$, then $P(N < \infty) = 1$ (as is the case with any rv N), and by Tonelli's (Fubini's) theorem,

$$E(N) = \sum_{n=1}^{\infty} P(A_n), \tag{2}$$

which is assumed finite, thus completing the proof.

In general, the converse is not true. Essentially, this is because there exists rvs N such that $P(N < \infty) = 1$ but $E(N) = \infty$. (For example choose an N such that $P(N = n) = c/n^2$, $n \ge 1$, and define $A_n = \{N = n\}$.) But if the events are *independent*, then the converse holds:

Lemma 2 Suppose that $\{A_n : n \ge 1\}$ is a sequence of independent events in a probability space. If

$$\sum_{n=1}^{\infty} P(A_n) = \infty, \tag{3}$$

then P(A(i.o.)) = 1.

Proof: Suppose that (3) holds, and note that if it holds then

$$\sum_{n=k}^{\infty} P(A_n) = \infty, \ k \ge 1.$$
 (4)

Let \overline{A}_n denote the complement of the set A_n .

$$P(A(i.o.)) = \lim_{k \to \infty} P(\bigcup_{n=k}^{\infty} A_n) = 1 - \lim_{k \to \infty} P(\bigcap_{n=k}^{\infty} \overline{A}_n).$$

To complete the proof we will show that

$$P(\bigcap_{n=k}^{\infty} \overline{A}_n) = 0, \ k \ge 1.$$

By independence, and the basic fact that $1 - x \le e^{-x}$, $x \ge 0$,

$$P(\bigcap_{n=k}^{\infty} \overline{A}_n) = \prod_{n=k}^{\infty} P(\overline{A}_n)$$
 (5)

$$= \prod_{n=k}^{\infty} (1 - P(A_n)) \tag{6}$$

$$\leq \prod_{n=k}^{\infty} e^{-P(A_n)}$$

$$= e^{-\sum_{n=k}^{\infty} P(A_n)}$$
(8)

$$= e^{-\sum_{n=k}^{\infty} P(A_n)} \tag{8}$$

$$= e^{-\infty} = 0, (9)$$

where the last equality is from (4).

As an immediate corollary to the two Lemmas, we have a special case of a "0-1" law:

Proposition 0.1 If $\{A_n : n \geq 1\}$ is a sequence of independent events in a probability space, then either P(A(i.o.)) = 0 $(E(N) < \infty$ case) or P(A(i.o.)) = 1 $(E(N) = \infty$ case), where N denotes the total number of A_n to occur;

$$N = \sum_{n=1}^{\infty} I_n,$$

where $I_n = I\{A_n\}$ denote the indicator rv for the event A_n .

0.1Applications

1. Suppose $\{X_n : n \geq 1\}$ are rvs such that $P(X_n = 1) = p_n$, $P(X_n = 0) = 1 - p_n$. Thus these are Bernoulli rvs in which the success probability depends on n and we do not assume independence. Using $A_n = \{X_n = 1\}$ we deduce from Lemma 1 that if

$$\sum_{n=1}^{\infty} p_n < \infty,$$

then X_n only visits 1 a finite number of times, hence eventually must take on value 0, wp1. Thus there exists a random time N such that $X_n = 0$, $n \ge N$ and so $\lim_{n\to\infty} X_n = 0$, wp1. To make this happen, we need that $p_n \to 0$ fast enough; we could, for example, choose $p_n = 1/n^2$ or choose $\{p_n\}$ to be any probability distribution, such as a geometric distribution, $p_n = (1-p)^{n-1}p$, $n \ge 1$ (some 0). (This result remains valid even if the rvs are independent.)

Now assume that the rvs are independent so that by Lemma 2 if

$$\sum_{n=1}^{\infty} p_n = \infty,$$

then $X_n = 1$ for infinitely many n, wp1. To achieve this, we could, of course take $p_n = p$ for some fixed p, but more interesting would be to take, for example, $p_n = 1/n$ in which case $q_n = 1 - p_n = P(X_n = 0)$ also sums to ∞ since $q_n \to 1$; $X_n = 0$ for infinitely many n, wp1 too. Thus in this case X_n continues to visit both 1 and 0 over and over infinitely often.