COVARIANCE, LLSE, CONDITIONAL EXPECTATION, MARKOV CHAINS 10

COMPUTER SCIENCE MENTORS 70

November 14 to November 18, 2016

1 Covariance

1.1 Introduction

The **covariance** of two random variables *X* and *Y* is defined as:

$$Cov(X,Y) := E((X - E(X)) \cdot (Y - E(Y)))$$

1.2 Warm Up

1. Prove that Cov(X, X) = Var(X):

Solution:

$$\mathrm{Cov}(X,X) = \mathrm{E}(X\cdot X) - \mathrm{E}(X)\cdot \mathrm{E}(X) = \mathrm{E}(X^2) - \mathrm{E}(X)^2$$

2. Prove that if X and Y are independent, then Cov(X,Y) = 0:

Solution:

$$\mathrm{Cov}(X,Y) = \mathrm{E}(X \cdot Y) - \mathrm{E}(X) \cdot \mathrm{E}(Y)$$

Remember that a property of expectation is that if X and Y are independent, then $E(XY) = E(X) \cdot E(Y)$, so we get 0 when we subtract

3. Prove that Cov(X + Y, Z) = Cov(X, Z) + Cov(Y, Z):

$$\begin{aligned} \operatorname{Cov}(X+Y,Z) &= \operatorname{E}((X+Y) \cdot Z) - \operatorname{E}(X+Y) \cdot \operatorname{E}(Z) \\ &= \operatorname{E}(X \cdot Z) + \operatorname{E}(Y \cdot Z) - (\operatorname{E}(X) \cdot \operatorname{E}(Z) + \operatorname{E}(Y) \cdot \operatorname{E}(Z)) \\ &= \operatorname{E}(X \cdot Z) - \operatorname{E}(X) \cdot \operatorname{E}(Z) + \operatorname{E}(Y \cdot Z) - \operatorname{E}(Y) \cdot \operatorname{E}(Z) \\ &= \operatorname{Cov}(X,Z) + \operatorname{Cov}(Y,Z) \end{aligned}$$

1.3 Questions

1. Roll 2 dice. Let A be the number of 6's you get, and B be the number of 5's, find Cov(A,B)

Solution: $E(A) = \frac{1}{6}$ for one die, by linearity of expectation, two dice make $\frac{1}{3}$, same for E(B) $E(A) = \frac{1}{3}$, $E(B) = \frac{1}{3}$

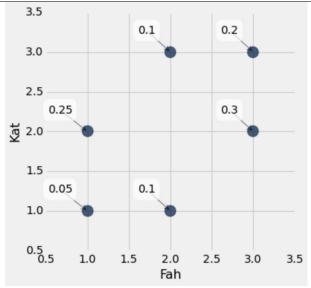
AB can be either 0 (if no 5's or 6's show up) or 1 (get a 5 and a 6).

$$\begin{split} E(AB) &= 1 \cdot P[\text{get a 5 and a 6}] \\ &= P[\text{first die} = 5 \text{ and second die} = 6] + P[\text{first die} = 6 \text{ and second die} = 5] \\ &= \frac{1}{36} + \frac{1}{36} \end{split}$$

$$Cov(AB) = E(AB) - E(A) \cdot E(B)$$
$$= \frac{1}{18} - \frac{1}{9}$$
$$= -\frac{1}{18}$$

2. Consider the following distribution with random variables Fah and Kat:

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Find the covariance of Fah and Kat.

Solution:
$$E(Fah) = 1 \cdot .3 + 2 \cdot .2 + 3 \cdot .5 = 2.2$$

$$E(Kat) = 1 \cdot .15 + 2 \cdot .55 + 3 \cdot .3 = 2.15$$

$$E(KatFah) = 1 \cdot 1 \cdot .05 + 1 \cdot 2 \cdot .25 + 2 \cdot 1 \cdot .1 + 2 \cdot 3 \cdot .1 + 3 \cdot 2 \cdot .3 + 3 \cdot 3 \cdot .2 = 4.95$$

$$cov(Kat, Fah) = 4.95 - 2.2 \cdot 2.15 = 0.22$$

2 LLSE

2.1 Introduction

Theorem: Consider two random variables, X, Y with a given distribution P[X=x,Y=y]. Then

$$\mathrm{L}[Y|X] = \mathrm{E}(Y) + \frac{\mathrm{Cov}(X,Y)}{\mathrm{Var}(X)}(X - \mathrm{E}(X))$$

2.2 Questions

1. Assume that

$$Y = \alpha X + Z$$

where *X* and *Z* are independent and E(X) = E(Z) = 0. Find L[X|Y].

Solution:

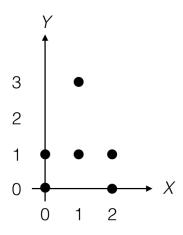
$$Cov(X,Y) = E(XY) - E(X)E(Y)$$
$$= E(X \cdot (\alpha X + Z)) = \alpha E(X^{2})$$

$$Var(Y) = \alpha^{2}Var(X) + Var(Z)$$
$$= \alpha^{2}E(X^{2}) + E(Z^{2})$$

Therefore,

$$L[X|Y] = \frac{\alpha E(X^2)}{\alpha^2 E(X^2) + E(Z^2)} \cdot Y$$

- 2. The figure below shows the six equally likely values of the random pair (X, Y). Specify the functions of:
 - $L[Y \mid X]$
 - $E(X \mid Y)$
 - *L*[*X* | *Y*]
 - $E(Y \mid X)$



Solution: Let's calculate some useful properties of the distribution first and then see how we can use them to calculate the estimates.

$$\begin{aligned} |\Omega| &= 6 \implies P[\text{one point}] = \frac{1}{6} \\ E(X) &= 0 \left(\frac{2}{6}\right) + 1 \left(\frac{2}{6}\right) + 2 \left(\frac{2}{6}\right) \\ &= 1 \\ E(Y) &= 0 \left(\frac{2}{6}\right) + 1 \left(\frac{3}{6}\right) + 30 \left(\frac{1}{6}\right) \\ &= 1 \\ E(XY) &= 0 \left(\frac{3}{6}\right) + 1 \left(\frac{1}{6}\right) + 2 \left(\frac{1}{6}\right) + 3 \left(\frac{1}{6}\right) \\ &= 1 \\ \text{Cov}(X, Y) &= 0 \implies \text{L}[Y|X] = E(Y) \end{aligned}$$

- $L[Y \mid X]$: Using the LLSE formula: $L[Y \mid X] = E[Y] + \frac{Cov(X,Y)}{Var(Y)}(Y E[Y]) = E[Y]$. Therefore $L[Y \mid X] = 1$
- $E[X \mid Y]$: Notice the symmetry across X = 1. For all values of y, $E[X \mid Y = y]$ is the same; therefore $E[X \mid Y] = E[X] = 1$.
- $L[X \mid Y]$: The MMSE estimator for X given Y is a linear function, therefore $\boxed{L[X \mid Y] = E[X \mid Y] = 1}$
- $E[Y \mid X]$ For this one we can't make use of symmetry or directly apply what we calculated above. We must go back to the definition of conditional expectation. We can calculate $E[Y \mid X = x]$ for every point x, and that entirely defines the expression:

$$E(Y \mid X = x) = \begin{cases} \frac{1}{2} & \text{if } x = 0\\ 2 & \text{if } x = 1\\ \frac{1}{2} & \text{if } x = 2 \end{cases}$$

The above equation is sufficient, but we can go further by realizing that these points are part of a flipped absolute value function centered around x = 1:

$$E[Y \mid X] = \frac{-3}{2}|X - 1| + 2.$$
 Indeed, this is not linear, which is why L[Y | $X \mid \neq E[Y \mid X]$.

3.1 Introduction

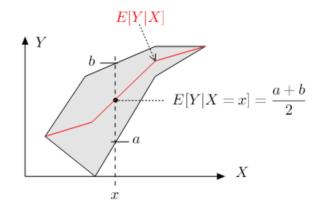
The **conditional expectation** of *Y* given *X* is defined by

$$E[Y|X = x] = \sum_{y} y \cdot P[Y = y|X = x] = \sum_{y} y \cdot \frac{P[X = x, Y = y]}{P[X = x]}$$

Properties of Conditional Expectation

$$\begin{split} \mathbf{E}(a|Y)) &= a \\ \mathbf{E}(aX + bZ|Y) &= a \cdot \mathbf{E}(X|Y) + b \cdot \mathbf{E}(Z|Y) \\ \mathbf{E}(X|Y) &\geq 0 \text{ if } X \geq 0 \\ \mathbf{E}(X|Y) &= \mathbf{E}(X) \text{ if } X,Y \text{ independent} \\ \mathbf{E}(\mathbf{E}(X|Y)) &= \mathbf{E}(X) \end{split}$$

Solution: Here is a picture that shows that conditioning creates a new random variable with a new distribution. Figure 9 of note 26 does so.



3.2 Questions

1. Prove E(E(Y|X)) = E(Y)

Solution:

$$\mathrm{E}(\mathrm{E}(Y|X)) = \sum_x \mathrm{E}(Y|X=x) \cdot \mathrm{P}[X=x]$$

$$\begin{split} &= \sum_{x} (\sum_{y} y \cdot \mathbf{P}[Y = y | X = x]) \cdot \mathbf{P}[X = x] \\ &= \sum_{y} y \cdot \sum_{x} y \cdot \mathbf{P}[X = x | Y = y]) \cdot \mathbf{P}[Y = y] \\ &= \sum_{y} y \cdot \mathbf{P}[Y = y] \cdot \sum_{x} \mathbf{P}[X = x | Y = y]) \\ &= \sum_{y} y \cdot \mathbf{P}[Y = y] = E[Y] \end{split}$$

2. Prove $E(h(X) \cdot Y|X) = h(X) \cdot E(Y|X)$

Solution:

$$\begin{split} \mathbf{E}(h(X) \cdot Y | X) &= \sum_{y} h(X) \cdot y \cdot \mathbf{P}[Y = y | X] \\ &= h(X) \sum_{y} y \cdot \mathbf{P}[Y = y | X] \\ &= h(X) \cdot \mathbf{E}[Y | X] \end{split}$$