

# Theoretical Statistics

# Overview

Statistics 101

Bayesian statistics

# Setup

## Definition

Let  $(\mathcal{X}, \mathcal{F})$  be a measurable space and  $\mathcal{P} = \{P_\theta \mid \theta \in \Omega\}$  a set of probability measures on  $(\mathcal{X}, \mathcal{F})$ . We call  $(\mathcal{X}, \mathcal{F}, \mathcal{P})$  a *statistical model*.

## Definition

A function  $T : \mathcal{X} \rightarrow \mathbb{R}^n$  is called a *statistic*.

## Definition

Given a sample  $X$ , we want to find a statistic  $\delta$  such that  $\delta(X)$  would be close to  $g(\theta)$ . In this context  $\delta$  is called an *estimator*.

# Loss and Risk

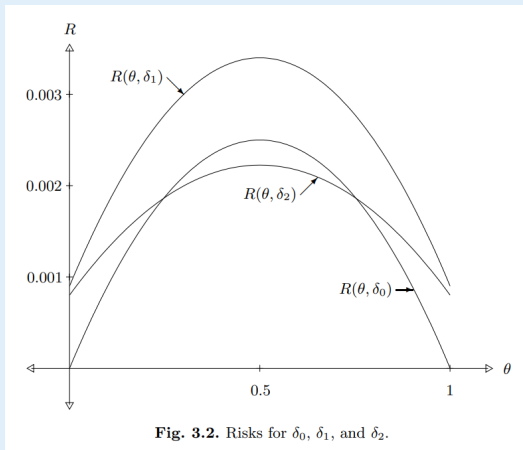
## Definition

To be able to compare different estimators we define the *loss* function  $L(\theta, d)$  as the cost of estimating  $g(\theta)$  with  $d$ .

We can then define the *risk* of a particular estimator  $\delta$  as

$$R(\theta, \delta) = \mathbb{E}_{X \sim P_\theta} [L(\theta, \delta(X))]$$

# Loss and Risk



# Loss and Risk (Decision Theory)

## Definition

Alternatively we could consider the *action space*  $\mathcal{A}$ , define the loss function and the *nonrandomized* decision rule as

$$L : \Omega \times \mathcal{A} \rightarrow \mathbb{R}$$

$$\delta : \mathcal{X} \rightarrow \mathcal{A}$$

Then the risk is defined in the same way

$$R(\theta, \delta) = \mathbb{E}_{\theta}[L(\theta, \delta(X))]$$

# Hypothesis Testing

## Hypothesis testing as a Decision Problem

Given hypotheses  $H_0 : \theta \in \Omega_0$  and  $H_1 : \theta \in \Omega_1$  we can associate

1. an action space  $\mathcal{A} = \{0, 1\}$ .
2. a loss function

$$L(\theta, a) = I\{a = 1, \theta \in \Omega_0\} + I\{a = 0, \theta \in \Omega_1\}$$

# Sufficient statistics

## Definition

Given a statistical model  $(\mathcal{X}, \mathcal{F}, \mathcal{P})$ , a statistic  $T$  is called a *sufficient* if  $X|T = t$  under  $P_\theta$  does not depend on  $\theta$ .

Equivalently, if  $I(\Theta, X) = I(\Theta, T(X))$  for any distribution on  $\Theta$ .

## Simulating samples using sufficient statistics

If  $T$  is a sufficient statistic, then define

$$Q_t(B) = \mathbb{P}(X \in B \mid T = t)$$

We can now sample  $X \sim P_\theta$ ,  $T = T(X)$  and  $\tilde{X} \sim Q_T$ .



# Neyman's Factorization Theorem

## Theorem

Let  $(\mathcal{X}, \mathcal{F}, \mathcal{P})$  be a statistical model where each  $P_\theta$  is dominated by  $\mu$  and has density  $p_\theta$ .  $T(X)$  is a sufficient statistic if and only if there exist functions  $g_\theta(x)$  and  $h(x)$  such that for all  $\theta$

$$p_\theta(x) = g_\theta(T(x))h(x)$$

# Rao-Blackwell theorem

## Theorem

Let  $T$  be a sufficient statistic for  $(\mathcal{X}, \mathcal{F}, \mathcal{P})$  and  $\delta$  an estimator of  $g(\theta)$  and define

$$\nu(T) = \mathbb{E}[\delta(X) \mid T]$$

If  $L(\theta, \cdot)$  is convex and  $R(\theta, \delta) < \infty$ , then

$$R(\theta, \nu) \leq R(\theta, \delta)$$

# Bayesian statistics

## Definition

In Bayesian statistics, one assumes an additional prior distribution  $\Lambda(\cdot)$  on  $\Omega$ . Then we might want to optimize

$$\int_{\Omega} R(\theta, \delta) \Lambda(d\theta)$$

Estimators  $\delta$  which optimize this criterion are called Bayes estimators.

# Bayesian statistics

## Theorem

Suppose that  $(\mathcal{X}, \mathcal{F}, \mathcal{P})$  is a statistical model together with a prior distribution  $\Lambda$  on  $\Omega$  and that  $L(\theta, d) \geq 0$  for all relevant arguments. If

1.  $\mathbb{E}[L(\Theta, \delta_0)] < \infty$  for some  $\delta_0$
2. for a.e.  $x$  there exists a value  $\delta_\Lambda(x)$  that minimizes

$$\mathbb{E}[L(\Theta, d)|X = x]$$

over  $d$

then  $\delta_\Lambda$  is a Bayes estimator.

# Bayesian sufficient statistics

## Definition

If  $(\mathcal{X}, \mathcal{F}, \mathcal{P})$  is a statistical model and  $\Lambda(\cdot)$  is a Bayesian prior, then  $T$  is a *Bayesian sufficient statistic* if for almost every  $x$

$$P(\theta|X = x) = P(\theta|T(X) = t(x))$$

## Definition

If  $(\mathcal{X}, \mathcal{F}, \mathcal{P})$  is a statistical model and  $\Lambda(\cdot)$  is a Bayesian prior, then  $T$  is a *predictive sufficient statistic* if for almost every  $x$

$$P(X^* = x^*|X = x) = P(X^* = x^*|T(X) = t(x))$$

# Admissibility

## Definition

Let  $(\mathcal{X}, \mathcal{F}, \mathcal{P})$  be a statistical model and let  $\delta$  be an estimator. If there is a  $\delta_*$  such that

$$R(\theta, \delta_*) \leq R(\theta, \delta)$$

with strict inequality for at least one  $\theta$ , then  $\delta$  is called inadmissible. Otherwise it is admissible.

## Theorem

If  $\delta_\Lambda$  is an essentially unique Bayes estimator, it is admissible.

# Complete Class Theorems

## Definition

A family of estimators is called a complete class if any  $\delta$  **not** in the class is inadmissible.

## Theorem

Let  $(\mathcal{X}, \mathcal{F}, \mathcal{P})$  be a statistical model and let  $\mathcal{B}_0$  be the class of all Bayes estimators w.r.t. different priors  $\Lambda(\cdot)$  on **finite** subsets of  $\Omega$ . Suppose that

- ▶  $\mathcal{P}$  is dominated with densities  $p_\theta(x) > 0$  everywhere
- ▶  $L(\theta, \cdot)$  is nonnegative and strictly convex
- ▶  $L(\theta, a) \rightarrow \infty$  as  $\|a\| \rightarrow \infty$

Then the set of pointwise limits of  $\mathcal{B}_0$  is a complete class.

# Bayesian vs Frequentist

$$I(\theta, \delta(x))$$

- ▶ Frequentist -  $R(\theta, \delta) = \mathbb{E}_{\theta}[I(\theta, \delta(X))]$
- ▶ Bayesian -  $\rho(x, \delta) = \mathbb{E}[I(\Theta, \delta(X)) | X = x]$
- ▶ Coherency – decisions/beliefs are consistent with Bayes theorem.  
Frequentist statistics are incoherent – Dutch Book argument.



# Objective vs Subjective Bayes

- ▶ Subjective Bayesian – allows any priors
- ▶ Objective Bayesian – restricts to diffuse priors
- ▶ Empirical Bayes – choose prior from data (James-Stein estimator)
- ▶ <https://stats.stackexchange.com/questions/381825/objective-vs-subjective-bayesian-paradigms>

# James-Stein estimator

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Suppose that  $X_i \sim N(\theta_i, 1)$  and consider the loss function  $l(\theta, \delta(X)) = \|\theta - \delta(X)\|^2$ . Then the estimator

$$\delta(x) = x$$

is admissible for  $n = 1, 2$ , but inadmissible for  $n > 2$ .  
It is dominated by the James-Stein estimator:

$$\delta_{JS}(x) = \left(1 - \frac{n-2}{\|x\|^2}\right)x$$

# References I

-  José M Bernardo and Adrian FM Smith, *Bayesian theory*, vol. 405, John Wiley & Sons, 2009.
-  Robert W Keener, *Theoretical statistics: Topics for a core course*, Springer Science & Business Media, 2010.