1. Problem

A polling firm dials phone numbers in the London area at random. Among the 100 people interviewed, **32** were over 65 years old. This is not surprising because in the most recent census, **25**% of London's residents were over 65.

Are the numbers in bold parameters or statistics?

- (a) 32 is a parameter, 25% is a statistic
- (b) 32 is a statistic, 25% is a parameter
- (c) Both are parameters
- (d) Both are statistics

Solution

A **parameter** is a number associated with the population. A **statistic** is a number associated with the sample. Therefore 32 is a statistic and 25% is a parameter.

2. Problem

In a meat factory, a machine is packaging minced meat in boxes. We are interested in the average weight of meat in each box. What is the parameter space?

- (a) [0,1]
- (b) $[0, \infty)$
- (c) $\{0, 1, 2, \ldots\}$
- (d) $(-\infty, \infty)$

Solution

The **parameter space** is the set of all possible values of the parameter. In our case the parameter is the average weight of meat in the package which can be any number from 0 to infinity.

3. Problem

An airline is monitoring its flight schedule for a period of one month. Suppose there are 100 flights in a month. What is the sample space?

- (a) [0, 100]
- (b) $[0, \infty)$
- (c) $\{0, 1, 2, \ldots\}$
- (d) $\{0, 1, 2, \dots, 100\}$

Solution

The **sample space** is the set of all possible values we can observe. This cannot be less than 0 and no more than 100 and it must be whole numbers because we are counting so $\{0, 1, 2, ..., 100\}$.

4. Problem

Let X_1, \ldots, X_n be a random sample from the Exponential (μ) distribution. The parameter μ denotes the mean of the distribution. The asymptotic distribution of \bar{X} according to the central limit theorem is

- (a) N(0,1)
- (b) $N(\mu, \sigma^2/n)$
- (c) $N(\mu, \mu^2/n)$
- (d) $N(\mu,\mu/n)$

Solution

The mean of the exponential distribution is μ and the variance is μ^2 . Therefore, according to the CLT, the asymptotic distribution of \bar{X} is $N(\mu,\mu^2/n)$.

5. Problem

Let X_1, \ldots, X_n be a random sample from a distribution with moment generating function (mgf) M(t). What is the correct formula for the mgf of the statistic $T = \sum X_i$?

- (a) M(t)
- (b) M(t/n)
- (c) $M(t)^n$
- (d) $M(t/n)^n$

Solution

By the definition of mgf, $M(t) = E[\exp(tX)]$ so $E[\exp(tT)] = E[\exp(t(X_1 + \ldots + X_n))] = E[\exp(tX_1) \ldots \exp(tX_n)] = E[\exp(tX_1)] \ldots E[\exp(tX_n)] = M(t) \ldots M(t) = M(t)^n$.

6. Problem

The following is a random sample of size 4 from the Pareto distribution with parameter θ

The pdf is

$$f(x|\theta) = \frac{\theta}{x^{\theta+1}}, \ \theta > 1, \ x > 1$$

Use these data to estimate θ using the method of moments.

Solution

The mean of the distribution is

$$\mu = \int_{1}^{\infty} x \frac{\theta}{x^{\theta+1}} dx$$

$$= \int_{1}^{\infty} \frac{\theta}{x^{\theta}} dx$$

$$= \int_{1}^{\infty} \theta \times x^{-\theta} dx$$

$$= -\frac{\theta}{\theta - 1} \times x^{-\theta + 1} \Big]_{1}^{\infty}$$

$$= \frac{\theta}{\theta - 1}.$$

We set

$$\begin{array}{rcl} \displaystyle \frac{\theta}{\theta-1} & = & \bar{x} \\ \Rightarrow \theta & = & \bar{x}(\theta-1) \\ \Rightarrow \bar{x} & = & (\bar{x}-1)\theta \\ \Rightarrow \theta & = & \displaystyle \frac{\bar{x}}{\bar{x}-1} \end{array}$$

Substituting the data we have

$$\bar{x} = (3.28 + 2.33 + 1.14 + 3.97)/4 = 2.68$$
 $\hat{\theta} = \frac{2.68}{2.68 - 1} = 1.595$

7. Problem

The following is a random sample of size 5 from a $N(\mu, \sigma^2)$ population

The pdf is

$$f(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{1}{2\sigma^2}(x-\mu)^2\right\}.$$

Use these data to estimate the mean μ and the variance σ^2 using the maximum likelihood method.

- (a) What is your estimate for the mean μ ?
- (b) What is your estimate for the variance σ^2 ?

Solution

The logarithm of the pdf is

$$\log f(x|\mu, \sigma^2) = -\frac{1}{2}\log(2\pi\sigma^2) - \frac{1}{2\sigma^2}(x-\mu)^2,$$

so the log-likelihood is

$$\ell(\mu, \sigma^2) = -\frac{n}{2}\log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^{n} (x_i - \mu)^2.$$

To estimate μ and σ^2 , we differentiate with respect to both parameters and set the derivative equal to 0. This gives us a system of 2 equations to solve for μ and σ^2 .

From the derivative with respect to μ we have

$$\frac{1}{\sigma^2} \sum_{i=1}^n (x_i - \mu) = 0 \tag{1}$$

From the derivative with respect to σ^2 we have

$$-\frac{n}{2\sigma^2} + \frac{1}{2(\sigma^2)^2} \sum_{i=1}^{n} (x_i - \mu)^2 = 0$$
 (2)

From equation (1) we have

$$\sum_{i=1}^{n} (x_i - \mu) = 0$$

$$\Rightarrow \sum_{i=1}^{n} x_i - n\mu = 0$$

$$\Rightarrow \mu = \sum_{i=1}^{n} x_i / n.$$

So the MLE for μ is \bar{x} . This is unbiased as \bar{x} is always unbiased for the mean of the population.

From equation (2), by substituting $\mu = \bar{x}$, we have

$$-\frac{n}{2\sigma^2} + \frac{1}{2(\sigma^2)^2} \sum_{i=1}^n (x_i - \bar{x})^2 = 0$$

$$\Rightarrow -n + \frac{1}{\sigma^2} \sum_{i=1}^n (x_i - \bar{x})^2 = 0$$

$$\Rightarrow \frac{1}{\sigma^2} \sum_{i=1}^n (x_i - \bar{x})^2 = n$$

$$\Rightarrow \sigma^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2$$

This MLE is not exactly S^2 because S^2 is divided by (n-1), not by n. Because S^2 is unbiased, then the MLE must be biased.

Substituting the data to these formulae, we have

$$\hat{\mu} = (3.2 + 6.7 + 3.2 + 3.9 + 3)/5 = 4$$

$$\hat{\sigma}^2 = ((-0.8)^2 + (2.7)^2 + (-0.8)^2 + (-0.1)^2 + (-1)^2)/5 = 1.916$$

- (a) The estimate for the mean is 4.
- (b) The estimate for the variance is 1.916.

8. Problem

The following is a random sample of size 5 from the geometric distribution with parameter θ

The pmf is

$$f(x|\theta) = \theta \times (1-\theta)^x, \ \theta \in (0,1), \ x = 0,1,2,...$$

Use these data to estimate θ using the maximum likelihood method.

Solution

The logarithm of the pdf is

$$\log f(x|\theta) = \log \theta + x \log(1-\theta),$$

so the log-likelihood is

$$\ell(\theta) = n \log \theta + \log(1 - \theta) \sum_{i=1}^{n} x_i.$$

To estimate θ , we differentiate the log-likelihood with respect to θ and set the derivative equal to 0. This gives us an equation which we need to solve to find θ .

The derivative is

$$\frac{n}{\theta} - \frac{1}{1 - \theta} \sum_{i=1}^{n} x_i = 0$$

$$\Rightarrow \frac{n}{\theta} = \frac{1}{1 - \theta} \sum_{i=1}^{n} x_i$$

$$\Rightarrow n(1 - \theta) = \theta \sum_{i=1}^{n} x_i$$

$$\Rightarrow n = \theta (\sum_{i=1}^{n} x_i + n)$$

$$\Rightarrow \frac{n}{\sum_{i=1}^{n} x_i + n} = \theta$$

$$\Rightarrow \theta = \frac{1}{\overline{x} + 1}$$

Substituting the data to these formulae, we have

$$\bar{x} = (3+0+1+1+2)/5 = 1.4$$
 $\hat{\theta} = \frac{1}{1.4+1} = 0.417$

9. Problem

Let $X \sim N(6.3, 4.41)$. Compute the probability $P(X \leq 7)$ using the normal distribution table. (Give 4 decimal points in your answer.)

Solution

If
$$X \sim N(\mu, \sigma^2)$$
, then $Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$.

Therefore
$$P(X < 7) = P(Z < \frac{7-6.3}{2.1}) = P(Z < 0.33)$$
.

We look for the number 0.33 in the margin of the normal distribution table and read the corresponding probability: 0.6293.

10. Problem

Let $X \sim N(9.7, 18.49)$. Compute the probability $P(X \le 6.4)$ using the normal distribution table. (Give 4 decimal points in your answer.)

Solution

If
$$X \sim N(\mu, \sigma^2)$$
, then $Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$.

Therefore
$$P(X < 6.4) = P(Z < \frac{6.4 - 9.7}{4.3}) = P(Z < -0.77).$$

Because the number -0.77 is negative, it can't be found in the margin of the normal distribution table.

Because of the symmetry of N(0,1) around 0, we have the property $\Phi(-z) = 1 - \Phi(z)$ so instead we look for 0.77 and read the corresponding probability 0.7794. Then, the probability $P(X \le 6.4) = 0.2206$.

11. Problem

Let $X \sim N(5.7, 3.24)$. Compute the probability P(3.702 < X < 7.698) using the normal distribution table. (Give 4 decimal points in your answer.)

Solution

If
$$X \sim N(\mu, \sigma^2)$$
, then $Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$.

Therefore
$$P(X < 7.698) = P(Z < \frac{7.698 - 5.7}{1.8}) = P(Z < 1.11)$$
.

We look for the number 1.11 in the margin of the normal distribution table and read the corresponding probability: 0.8665.

We do the same for the lower bound: $P(X < 3.702) = P(Z < \frac{3.702 - 5.7}{1.8}) = P(Z < -1.11)$.

We look for the number -1.11 in the margin of the normal distribution table and read the corresponding probability: 0.1335.

Therefore, the probability within is 0.8665 - 0.1335 = 0.733.

12. Problem

Let $X \sim N(2.9, 0.81)$. Find x corresponding to probability $P(X \le x) = 0.7389$ using the normal distribution table. (Give 2 decimal points in your answer.)

Solution

If
$$X \sim N(\mu, \sigma^2)$$
, then $Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$.

Therefore
$$P(X < x) = P(Z < \frac{x-2.9}{0.9}) = P(Z < z)$$
, where $z = \frac{x-2.9}{0.9}$.

To have P(Z < z) = 0.7389 we must choose z appropriately.

We look for the number 0.7389 in the body of the normal distribution table and read the corresponding quantile from the margin: z=0.64. Then, rearranging $\frac{x-2.9}{0.9}=0.64$ we get $x=2.9+0.64\times0.9=3.48$

13. Problem

Let $X \sim N(2.9, 0.81)$. Find x corresponding to probability $P(X \le x) = 0.3409$ using the normal distribution table. (Give 2 decimal points in your answer.)

Solution

If
$$X \sim N(\mu, \sigma^2)$$
, then $Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$.

Therefore
$$P(X < x) = P(Z < \frac{x - 2.9}{0.9}) = P(Z < z)$$
, where $z = \frac{x - 2.9}{0.9}$.

To have P(Z < z) = 0.3409 we must choose z appropriately.

Because 0.3409 < 0.5 the number 0.3409 cannot be found in the normal distribution table.

Because of the symmetry of N(0,1) around 0, we have the property $\Phi(-z) = 1 - \Phi(z)$ so if $\Phi(z) = 0.3409 \Rightarrow 1 - \Phi(-z) = 1 - 0.3409 = 0.6591$.

We seek for 0.6591 in the body of the table and read the corresponding quantile from the margin: $-z=0.41 \Rightarrow z=-0.41$. Then, rearranging $\frac{x-2.9}{0.9}=-0.41$ we get $x=2.9+-0.41\times0.9=2.53$

14. Problem

Let $X \sim \text{Gamma}(a, \theta)$ with shape a = 3 and scale $\theta = 2$. Use Python to find the number x such that $P(X \le x) = 0.43$. (Round your answer to 3 decimal places.)

Solution

Use the following code

from scipy.stats import gamma

 $\operatorname{gamma.ppf}(0.43, a = 3, \operatorname{scale} = 2)$

15. Problem

The daily expenses of summer tourists in Bath are analysed. A survey with 60 tourists is conducted. This shows that the tourists spend on average 132.3 pounds. The variance σ^2 is equal to 136.8.

Determine a 95% confidence interval for the average daily expenses of a tourist.

- (a) What is the lower confidence bound?
- (b) What is the upper confidence bound?

Solution

For a 95% confidence interval we find that $\alpha = 0.05$ and $z_{1-\alpha/2} = 1.96$.

Then, the 95% confidence interval for the average expenses μ is given by:

$$\begin{bmatrix} \bar{x} - 1.96\sqrt{\frac{\sigma^2}{n}}, \ \bar{x} + 1.96\sqrt{\frac{\sigma^2}{n}} \end{bmatrix}$$

$$= \begin{bmatrix} 132.3 - 1.96\sqrt{\frac{136.8}{60}}, \ 132.3 + 1.96\sqrt{\frac{136.8}{60}} \end{bmatrix}$$

$$= [129.34, \ 135.26].$$

- (a) The lower confidence bound is 129.34.
- (b) The upper confidence bound is 135.26.

16. Problem

The weekly expenses for groceries for a family of four were monitored for 11 weeks. On average the family spent 132.7 pounds. The sample variance S^2 was equal to 191.3.

Determine a 95% confidence interval for the average weekly expenses of the family.

- (a) What is the lower confidence bound?
- (b) What is the upper confidence bound?

Solution

We are not told what the variance is so we have to use quantiles from the t distribution.

For a 95% confidence interval we find that $\alpha = 0.05$ and $t_{1-\alpha/2} = 2.228$ at n-1 = 10 degrees of freedom.

Then, the 95% confidence interval for the average expenses μ is given by:

$$\begin{bmatrix} \bar{x} - t_{1-\alpha/2} \sqrt{\frac{S^2}{n}}, \ \bar{x} + t_{1-\alpha/2} \sqrt{\frac{S^2}{n}} \end{bmatrix}$$

$$= \begin{bmatrix} 132.7 - 2.228 \sqrt{\frac{191.3}{11}}, \ 132.7 + 2.228 \sqrt{\frac{191.3}{11}} \end{bmatrix}$$

$$= [123.409, 141.991].$$

- (a) The lower confidence bound is 123.409.
- (b) The upper confidence bound is 141.991.