

# Statistics for Data Science

Lecture 3

Transforms and Expectations

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# Recall

- Random variable  $X$  has a cdf  $F_X(x)$
- Cumulative distribution function.
- We also covered
  - Mass density functions
  - Probability density functions

# Content

- Expected Values
- Estimating a random variable
- Functions of a random variable
- Support Sets
- Another theorem

# Expected Value

- One more basic definition...
- The expected value of a random variable is the long-run average value of repetitions
- More specifically, it is the probability-weighted average of all possible values
- What is the expected value of rolling a fair six-sided die?
- What about if the die isn't fair?

# Expected Value

- The expected value of a random variable is denoted by  $E$ .

- For a discrete random variable  $x$ , the expected value

$$E[x] = \sum_x x p(x)$$

- We have already seen one example of this: the arithmetic mean, which is the expected value when every value of  $x$  is equally likely
- Fair dice roll:

- $E[x] = 1 \frac{1}{6} + 2 \frac{1}{6} + 3 \frac{1}{6} + 4 \frac{1}{6} + 5 \frac{1}{6} + 6 \frac{1}{6} = 3.5$

# Expected Value

- The expected value is more powerful, for example, an unfair dice roll:
  - $1 \frac{1}{12} + 2 \frac{1}{12} + 3 \frac{1}{12} + 4 \frac{1}{12} + 5 \frac{1}{12} + 6 \frac{7}{12} = 4.75$
- Or the amount you should expect to win if you play roulette:
  - $E[\text{profit on } £1] = -£1 \frac{37}{38} + £35 \frac{1}{38} = -£0.0526$

# Expected values

- So  $E g(X)$  is

$$\int_{-\infty}^{\infty} g(x) f_x(x) dx$$

if  $X$  is continuous

$$\sum_{x \in X} g(x) f_x(x) = \sum_{x \in X} g(x) P(X = x) \quad \text{if } X \text{ is discrete}$$

If the sum or integral exists. If  $E |g(X)| = \infty$  then  $E g(X)$  does not exist.

# Minimising Distance

- Suppose we measure the distance between a random variable  $X$  and a constant  $b$  by  $(X - b)^2$ .
- The closer to  $X$  that  $b$  is, the smaller the squared value above.
- If we can determine the value of  $b$  that minimises  $E(X-b)^2$  we will have a good predictor of  $X$ .



# Without using calculus

- $E(X - b)^2 = E(X - EX + EX - b)^2$  Add +/- EX which does nothing.  
 $= E((X - EX) + (EX - b))^2$  Group terms.  
 $= E(X - EX)^2 + (EX - b)^2 + 2E((X - EX)(EX - b))$  Expanded sq

Note,

$$E((X - EX)(EX - b)) = (EX - b)E(X - EX) = 0$$

- $(EX - b)$  is a constant and comes out of the expectation

$$E(X - EX) = EX - EX = 0$$

$$E(X - b)^2 = E(X - EX)^2 + (EX - b)^2$$

- Choose  $b = EX$

$$\min_b E(X - b)^2 = E(X - EX)^2$$

# Functions of a Random Variable

- Any function of  $X$ , call it  $g(X)$  is also a random variable.
- We can introduce  $Y$ , a new random variable to represent this:

$$Y = g(X)$$

# Sample Spaces

- As  $Y$  is a function of  $X$  we can describe the probabilistic behaviour of  $Y$  in terms of  $X$ . So for any set  $A$ :

$$P(Y \in A) = P(g(X) \in A)$$

- So if  $y = g(x)$ ,  $g$  defines a mapping from the sample space of  $X$   $\mathcal{X}$  to a new sample space  $\mathcal{Y}$  of  $Y$ .

$$g(x): \mathcal{X} \rightarrow \mathcal{Y}$$

# Inverse mapping

- We can also associate an inverse mapping  $g^{-1}$
- A mapping from subsets of  $Y$  to subsets of  $X$ , defined by

$$g^{-1}(A) = \{x \in X: g(x) \in A\}$$

# Sets into Sets

- $g^{-1}(A)$  is the set of points in  $\mathcal{X}$  that  $g(x)$  takes into the set  $A$ .
- $A$  can be a point set  $\{y\}$ .

$$g^{-1}(\{y\}) = \{x \in \mathcal{X}: g(x) \in y\}$$

- We can write  $g^{-1}(\{y\})$  as  $g^{-1}(y)$  in this case.
- $g^{-1}(y)$  can still be a set when there are multiple  $x$  for which  $g(x) = y$ .
- Now for any set  $A \subset \mathcal{Y}$ ,

$$P(Y \in A) = P(g(X) \in A) = P(\{x \in \mathcal{X}: g(x) \in A\}) = P(X \in g^{-1}(A))$$

Defines the pdf of  $Y$ . Showing the pdf satisfies the axioms left as an exercise.

If  $X$  is discrete, then  $\mathcal{X}$  is countable.

- $Y = g(X)$  is  $\mathcal{Y} = \{y: y = g(x), x \in \mathcal{X}\}$  is also a countable set.
- Therefore  $Y$  is also a discrete random variable.
- The pmf  $f_Y(y)$  can be found by identifying  $g^{-1}(y)$  for each  $y \in \mathcal{Y}$  and summing the appropriate probabilities ( $x \in g^{-1}(y)$ )



# An Example

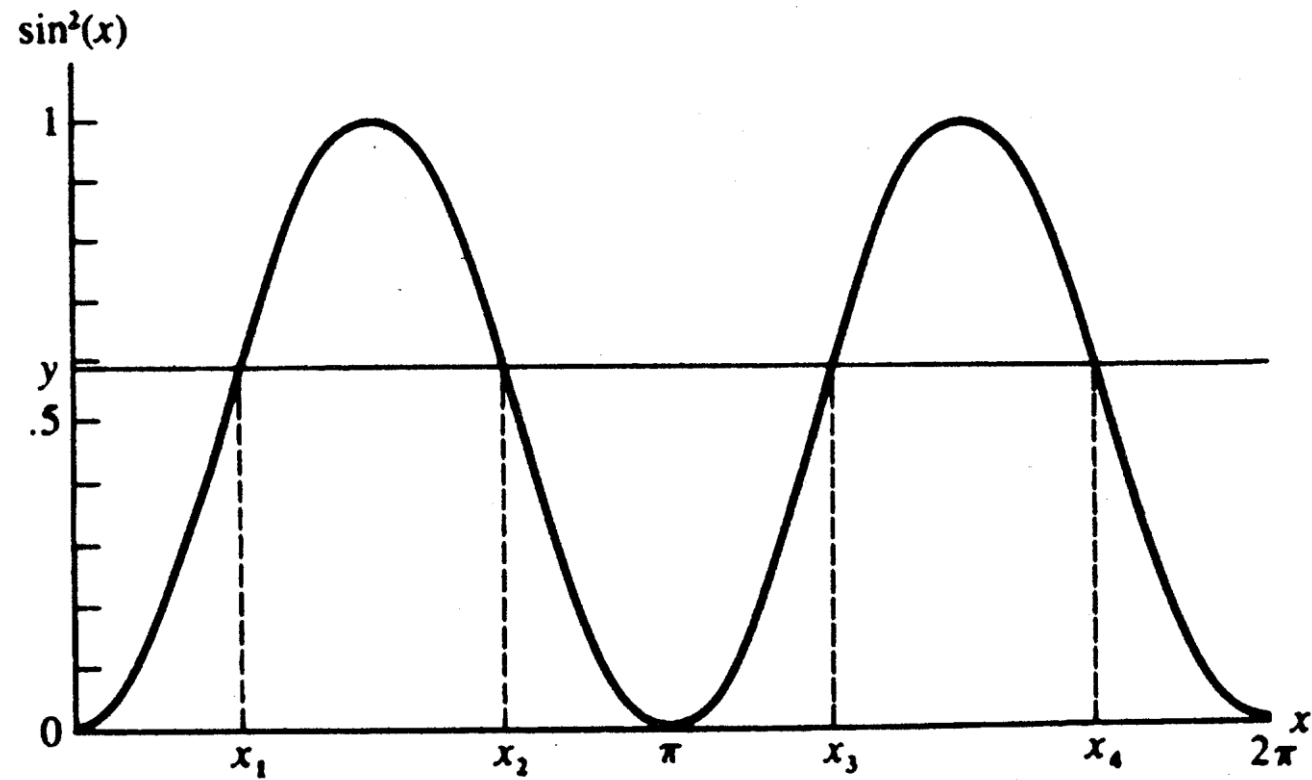
- Random Variable  $X$  with uniform distribution on interval  $(0, 2\pi)$

$$f_x(x) = \begin{cases} 1/(2\Omega) & 0 < x < 2\Omega \\ 0 & \textit{Otherwise} \end{cases}$$

- Now consider

$$Y = \sin^2(X)$$

# Example



# An Example

$$P(Y \leq y) = P(X \leq x_1) + P(x_2 \leq X \leq x_3) + P(X \geq x_4)$$

- From symmetry of  $\sin^2(x)$  and that  $X$  has uniform distribution

$$P(X \leq x_1) = P(X \geq x_4) \text{ and } P(x_2 \leq X \leq x_3) = 2P(x_2 \leq X \leq \pi)$$

- So

$$P(Y \leq y) = 2P(X \leq x_1) + 2P(x_2 \leq X \leq \pi)$$

# An Example

- Where  $x_1$  and  $x_2$  are the two solutions to:

$$\sin^2(x) = y \quad 0 < y < \pi$$

- The resulting cdf expression was not simple, even though it looked like it should be when we started.
- It's important to keep track of the sample spaces of the random variable to avoid confusion.

# Support Sets

- When the transform is from  $X$  to  $Y = g(X)$  it is convenient to use:
- $\mathcal{X} = \{x: f_x(x) > 0\}$  and  $\mathcal{Y} = \{y: y = g(x) \text{ for some } x \in \mathcal{X}\}$  [1]
- The pdf of  $X$  is positive only on the set  $\mathcal{X}$  and zero elsewhere.
- This known as the support set of a distribution.

# Monotone functions

- It's easier if the  $g(x)$  are monotone.
- Functions are monotone if:

$$u > v \Rightarrow g(u) > g(v)$$

Or,

$$u < v \Rightarrow g(u) < g(v)$$

- If  $x \rightarrow g(x)$  is monotone then it is one-to-one and onto from  $\mathcal{X} \rightarrow \mathcal{Y}$ 
  - Each  $x$  goes to only one  $y$ .
  - And each  $y$  comes from at most one  $x$
- And for  $\mathcal{Y}$  defined as in [1] for each  $y \in \mathcal{Y}$  there is an  $x \in \mathcal{X}$  such that  $g(x) = y$ .

- If  $g$  is monotone then  $g^{-1}$  is single-valued.

$$g^{-1}(y) = x \text{ if and only if } y = g(x)$$

- If  $g$  is increasing then:

$$\{x \in X : g(x) \leq y\} = \{x \in X : g^{-1}(g(x)) \leq g^{-1}(y)\} = \{x \in X : x \leq g^{-1}(y)\}$$

- If  $g$  is decreasing then:

$$\{x \in X : g(x) \leq y\} = \{x \in X : g^{-1}(g(x)) \geq g^{-1}(y)\} = \{x \in X : x > g^{-1}(y)\}$$



- If  $g(x)$  is increasing then:

$$F_Y(y) = \int_{-\infty}^{g^{-1}(y)} f(x) dx = F_X(g^{-1}(y))$$

- If  $g(x)$  is decreasing then:

$$F_Y(y) = \int_{g^{-1}(y)}^{-\infty} f(x) dx = F_X(g^{-1}(y))$$

# Theorem

- These results allow us to state the following theorem:
- Let  $X$  have cdf  $F_x(x)$ , let  $Y = g(X)$  and let  $\mathcal{X}$  and  $\mathcal{Y}$  be defined as in [1]
  - a. If  $g$  is an increasing function on  $\mathcal{X}$ ,  $F_y(y) = F_x(g^{-1}(y))$  for  $y \in \mathcal{Y}$ .
  - b. If  $g$  is a decreasing function on  $\mathcal{X}$  and  $X$  is a continuous random variable,  $F_y(y) = 1 - F_x(g^{-1}(y))$  for  $y \in \mathcal{Y}$ .