

## **General ARMA(p,q) models**

Shumway and Stoffer: 3.1-3.2

## Introduction to autoregressive models

- *Autoregression* - our earlier example of  $x_t = x_{t-1} - 0.9x_{t-2} + w_t$
- Intuitively, makes forecasting possible - an exciting perspective
- More formally, a stationary process

$$x_t = \phi_1 x_{t-1} + \dots + \phi_p x_{t-p} + \omega_t$$

is an autoregressive process of order  $p$  with  $\phi_p \neq 0$ .

- For simplicity,  $\omega_t$  is Gaussian with mean zero and variance  $\sigma^2$

### Some remarks about an AR(p) model

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- Let  $\Phi(B) = 1 - \phi_1 B - \phi_2 B^2 - \dots - \phi_p B^p$  be an **autoregressive operator**. It can be viewed as a polynomial in  $B$  of order  $p$ . Then the autoregressive model of order  $p$  can be written concisely as

$$\Phi(B)x_t = w_t$$

- If the mean  $\mu$  of  $x_t$  is not zero, it is useful to center the series and consider the process in terms of  $x_t - \mu$ , e.g.

$$x_t = \alpha + \phi_1 x_{t-1} + \dots + \phi_p x_{t-p} + \omega_t$$

where  $\alpha = \mu(1 - \phi_1 - \dots - \phi_p)$

### AR(1) process as easy example

- Assume  $p = 1$  and consider AR(1) process first for simplicity

$$x_t = \phi x_{t-1} + w_t \quad (1)$$

- If  $|\phi| < 1$ , (1) allows the infinite representation  $x_t = \sum_{j=0}^{\infty} \phi^j w_{t-j}$
- Clearly, the mean of the process is  $E x_t = \sum_{j=0}^{\infty} \phi^j E w_{t-j} = 0$ .
- Its autocovariance function is

$$\gamma(h) = \text{cov}(x_{t+h}, x_t) = \sigma_w^2 \sum_{j=0}^{\infty} \phi^j \phi^{j+h} = \frac{\sigma_w^2 \phi^h}{1 - \phi^2}$$

- As a corollary, the autocorrelation function of the AR(1) process is

$$\rho(h) = \frac{\gamma(h)}{\gamma(0)} = \phi^h$$

for all  $h \geq 0$

- Note also that  $\rho(h)$  satisfies the recursive relationship  $\rho(h) = \phi\rho(h - 1)$ ,  
 $h = 1, 2, \dots$

### Causality of an autoregressive process

- Representation of the AR(1) process as linear process  $\sum_{j=0}^{\infty} \psi_j w_{t-j}$  is only possible when  $|\phi| < 1$
- If  $|\phi| > 1$  we have the process that depends on the **future** and is thus non-causal; more specifically, the solution is

$$x_t = - \sum_{j=1}^{\infty} \phi^{-j} \omega_{t+j}$$

and its autocovariance is  $\gamma_x(h) = \sigma_{\omega}^2 \phi^{-2} \phi^{-h} / (1 - \phi^{-2})$

- The above makes it clear that we can define a process  $y_t = \phi^{-1} y_{t-1} + v_t$  with  $v_t \text{ iid } N(0, \phi^{-2} \sigma_{\omega}^2)$  that is stochastically equal to the process  $x_t$ ; this suggests there's no need to consider non-causal autoregressive stationary processes

### General stationary solution of the AR(p) model

- Represent  $x_t$  as a linear process with unknown coefficients:

$$x_t = \sum_{j=0}^{\infty} \psi_j w_{t-j} = \Psi(B)w_t$$

where the polynomial  $\Psi(B) = \sum_{j=0}^{\infty} \psi_j w_{t-j}$

- From the identity

$$\Phi(B)\Psi(B)w_t = w_t$$

determine the coefficients  $\psi_j$  recursively

- As an example, for AR(1) the sequence of equations is

$$\psi_1 - \phi_1 = 0$$

$$\psi_2 - \psi_1\phi = 0$$

...

$$\psi_j - \psi_{j-1}\phi = 0$$

- Clearly,  $\psi_j = \psi_{j-1}\phi$  and, thus,  $\psi_j = \phi^j$



General moving average models MA(q) (moving average of order  $q$ )

- The general MA(q) process is defined as

$$x_t = w_t + \theta_1 w_{t-1} + \theta_2 w_{t-2} + \dots + \theta_q w_{t-q} \quad (2)$$

where  $\theta_q \neq 0$  and  $w_t$  is a white noise

- Again, in general  $w_t \sim N(0, \sigma_w^2)$
- Introduce the **moving average operator**  $\Theta(B) = 1 + \theta_1 B + \theta_2 B^2 + \dots + \theta_q B^q$
- Then, the concise form of (2) is

$$x_t = \Theta(B)w_t$$

### An example: MA(1)

- Consider

$$x_t = w_t + \theta w_{t-1}$$

- Clearly,  $E x_t = 0$ ; moreover, it is easy to check that the autocovariance function is

$$\gamma(h) = \begin{cases} (1 + \theta^2)\sigma_w^2, & h = 0 \\ \theta\sigma_w^2, & h = 1 \\ 0, & |h| \geq 1 \end{cases}$$

- Note that the autocovariance function of MA(1) cuts off at the lag 1

### Non-uniqueness and invertibility of MA models

- It is easy to find out that the MA(1) processes with  $\sigma_w^2 = 1$  and  $\theta = 5$ , on one hand, and  $\sigma_w^2 = 25$  and  $\theta = \frac{1}{5}$  possess the same autocovariance function
- Let us assume these processes are driven by the Gaussian white noise; then, they are  $x_t = w_t + \frac{1}{5}w_{t-1}$ ,  $w_t \sim iid N(0, 25)$  and  $x_t = v_t + 5v_{t-1}$ ,  $v_t \sim iid N(0, 1)$ .
- All of their finite-dimensional distributions are identical  $\rightarrow$  the processes themselves are identical!!!
- To avoid this, it is necessary to assume that  $|\theta| < 1$ . If this is the case, there is a unique representation of the MA(1) process as

$$w_t = \sum_{j=0}^{\infty} (-\theta)^j x_{t-j}$$

- Note that  $MA(1)$  process is **Always** stationary.

### General invertibility condition

- For general  $MA(q)$  processes, it is not easy to state the invertibility condition in terms of the moving average polynomial coefficients
- However, it can be easily done in terms of the *roots* of the polynomial  $\Theta(z)$  viewed as a complex variable function
- The  $MA(q)$  process is causal if and only if  $\Theta(z) \neq 0$  for all  $|z| \leq 1$ .

### Example: a general MA(q) model

- Let  $x_t = \Theta(B)\omega_t$  where  $\Theta(B) = 1 + \theta_1 B + \dots + \theta_q B^q$ . Clearly,  $E x_t = 0$
- Moreover,

$$\gamma(h) = \text{cov}(x_{t+h}, x_t) = \sigma_w^2 \sum_{j=0}^{q-h} \theta_j \theta_{j+h}$$

for any  $0 \leq h \leq q$  while it is equal to zero for any  $h > q$

- The autocorrelation is

$$\rho(h) = \frac{\sigma_w^2 \sum_{j=0}^{q-h} \theta_j \theta_{j+h}}{1 + \theta_1^2 + \dots + \theta_q^2}$$

for any  $0 \leq h \leq q$

- The autocovariance (autocorrelation) that cuts off after  $q$  lags is the signature of an MA( $q$ ) model

### General ARMA (autoregressive moving average) models

- A general ARMA(p,q) model is defined as

$$x_t = \phi_1 x_{t-1} + \dots + \phi_p x_{t-p} + w_t + \theta_1 w_{t-1} + \dots + \theta_q w_{t-q}$$

where  $\phi_p \neq 0$  and  $\theta_q \neq 0$ ;  $w_t$  is a white noise with  $\sigma_w^2 > 0$ .

- In concise form it is

$$\Phi(B)x_t = \Theta(B)w_t$$

### Parameter redundancy in ARMA(p,q) models

- It is **always** assumed that the polynomials  $\Phi(B)$  and  $\Theta(B)$  do not have common factors
- To ensure that the model is *causal* we require that the autoregressive polynomial  $\Phi(B)$  does not have any roots inside the unit circle
- To ensure that the model is *invertible* we require, likewise, that the moving average polynomial  $\Theta(B)$  does not have any roots inside the unit circle

## Causality of ARMA(p,q) models

- The ARMA (p,q) process  $x_t$  that satisfies

$$\Phi(B)x_t = \Theta(B)\omega_t$$

is said to be **causal** when it can be represented as

$$x_t = \sum_{j=0}^{\infty} \psi_j \omega_{t-j} = \Psi(B)\omega_t$$

where  $\sum_{j=0}^{\infty} |\psi_j| < \infty$  and  $\Psi(B) = \sum_{j=0}^{\infty} \psi_j B^j$

- **Property** The ARMA(p,q) process  $x_t$  is causal if and only if all of the roots of its autoregressive polynomial  $\phi(z)$  are outside the unit circle:  $\phi(z) = 0$  only if  $|z| > 1$ . In that case,  $\psi(z) = \frac{\theta(z)}{\phi(z)}$  for any  $|z| \leq 1$



## Invertibility of ARMA(p,q) processes

- The ARMA (p,q)  $x_t$  process that satisfies

$$\Phi(B)x_t = \Theta(B)\omega_t$$

is said to be **invertible** when there exists it can be represented as

$$\pi(B)x_t = \omega_t$$

where  $\pi(B) = \sum_{j=0}^{\infty} \pi_j B^j$  and  $\sum_{j=0}^{\infty} |\pi_j| < \infty$  and

- **Property** The ARMA(p,q) process  $x_t$  is invertible if and only if all of the roots of its moving average polynomial  $\theta(z)$  are outside the unit circle:  $\theta(z) = 0$  only if  $|z| > 1$ . In that case,  $\pi(z) = \frac{\phi(z)}{\pi(z)}$  for any  $|z| \leq 1$