

Nonparametric estimation using wavelet methods. L1

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Examples of nonparametrics

Estimate a density of probability

We observe X_1, \dots, X_n i.i.d. with probability P having a density f w.r. to Lebesgue measure. Our aim is to estimate f .

Regression framework

We observe $(X_1, Y_1), \dots, (X_n, Y_n)$ i.i.d.
 X_i Uniform on $[0,1]$,

$$Y_i = f(X_i) + \epsilon_i, \quad \epsilon_i \sim N(0, 1).$$

Our aim is to estimate f .

Examples of nonparametrics

White noise model

$$dY_t^\epsilon = f(t)dt + \epsilon dW_t, \quad t \in [0, 1], \epsilon = 1/\sqrt{n}.$$

We observe : $\forall \phi \in L_2([0, 1]), \quad Y_\phi = \int_0^1 \phi(t)f(t)dt + \epsilon \xi_\phi,$

$(\xi_\phi, \xi_\eta) \sim N \left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} \|\phi\|^2 & \langle \phi, \eta \rangle \\ \langle \phi, \eta \rangle & \|\eta\|^2 \end{pmatrix} \right)$ Our aim is to estimate f .

Examples of nonparametrics (more involved)

EDS

$$dX_t = b(t)dt + f(t)dW_t.$$

We observe $X_i = X_{i\Delta}$, $i = 1, \dots, n$.

Our aim is to estimate f .

Examples of nonparametrics (more involved)

Inverse models

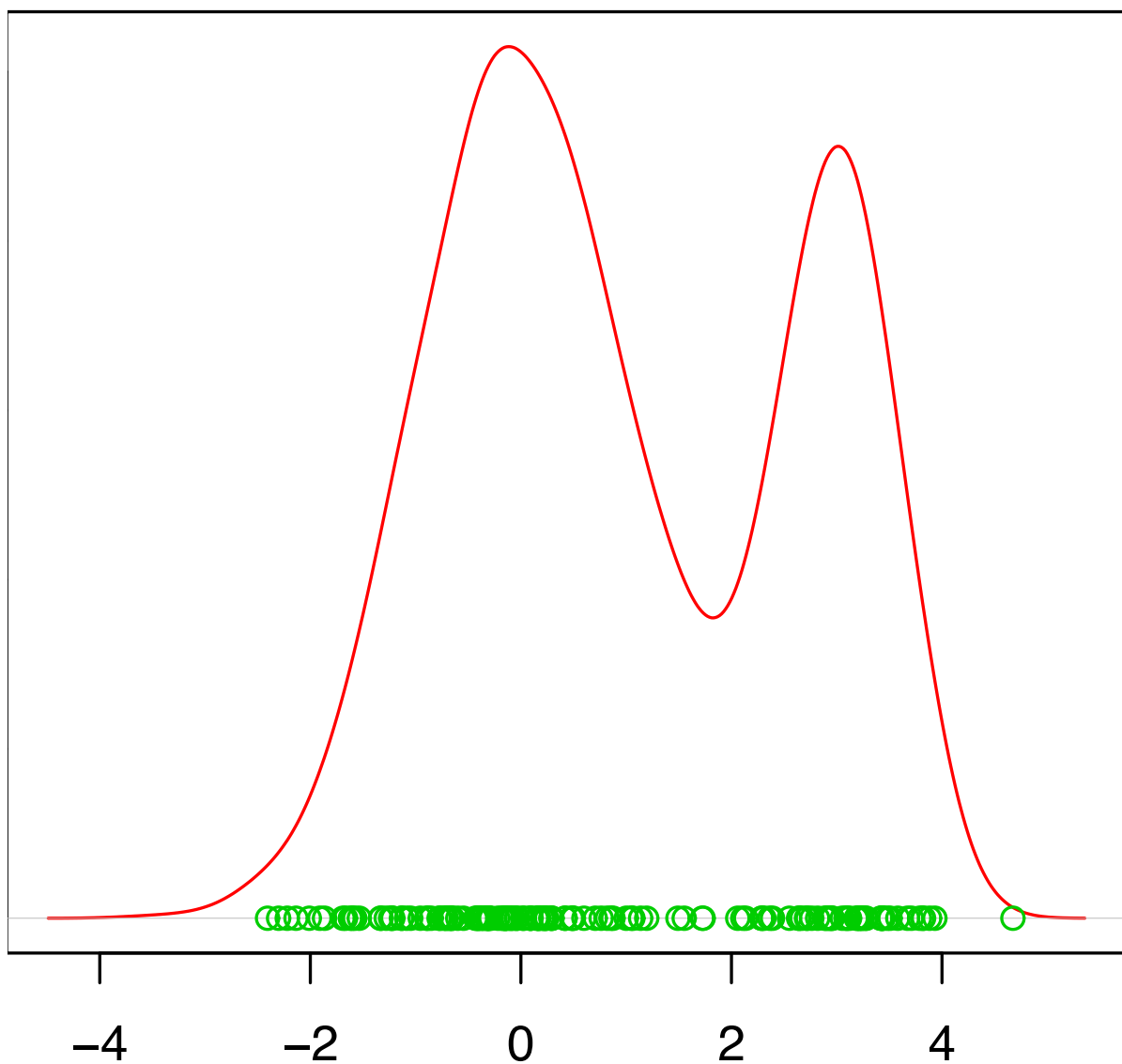
$$dY_t^\epsilon = Af(t)dt + \epsilon dW_t, \quad t \in [0, 1], \epsilon = 1/\sqrt{n}.$$

We observe : $\forall \phi \in L_2([0, 1])$, $Y_\phi = \int_0^1 \phi(t)Af(t)dt + \epsilon\xi_\phi$, A known linear operator, for instance

$$Af(s) = \int g(s - t)f(t)dt, \quad A \text{ Radon transform...}$$

Our aim is to estimate f .

density.default(x = z1)



N = 3000 Bandwidth = 0.3006

Why is it difficult ?

Estimate a density of probability

We observe X_1, \dots, X_n i.i.d. with probability P having a density f w.r. to Lebesgue measure. Our aim is to estimate f .

Easy : Estimate $F(x) = P(X_i \leq x)$:

$$\hat{F}_n(x) = \frac{1}{n} \sum_{i=1}^n 1_{]-\infty, x]}(X_i), \quad \hat{F}_n(x) - F(x) = \frac{\xi_n(x)}{\sqrt{n}}$$

$$\{\xi_n(x), x \in \cdot\} \xrightarrow{\text{weakly}} \{B_0(F(x)), x \in \cdot\}$$

Kolmogorov Smirnov.

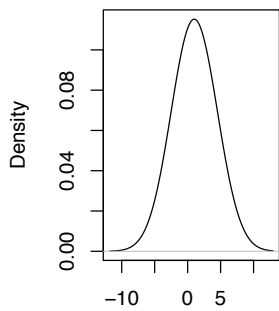
2 obstructions to differentiation :

- $\hat{F}_n(x)$ is not differentiable.
- $B_0(F(x))$ is not differentiable either.

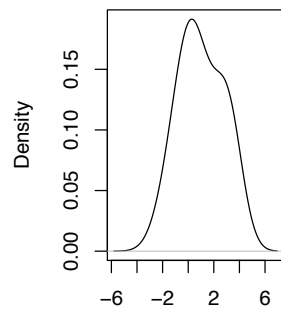
Parzen kernel method

$$\begin{aligned}\hat{K}_{h_n}(x) &= \int \frac{1}{h_n} K\left(\frac{x-y}{h_n}\right) dF_n(y) \\ &= \frac{1}{nh_n} \sum_{i=1}^n K\left(\frac{x-X_i}{h_n}\right)\end{aligned}$$

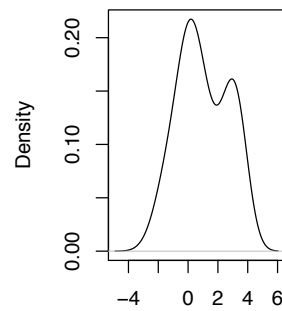
`density.default(x = z, bw =` `density.default(x = z, bw =` `density.default(x = z, bw = 0` `density.default(x = z, bw = 0`



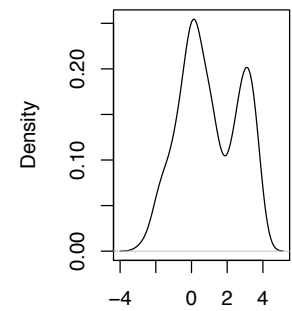
N = 300 Bandwidth = 3



N = 300 Bandwidth = 1

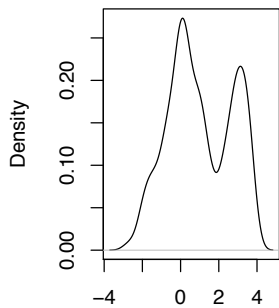


N = 300 Bandwidth = 0.7

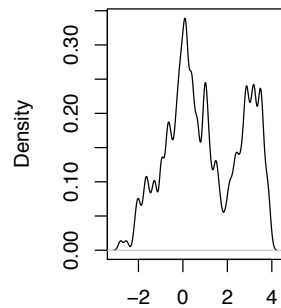


N = 300 Bandwidth = 0.4

`density.default(x = z, bw = 0` `density.default(x = z, bw = 0`

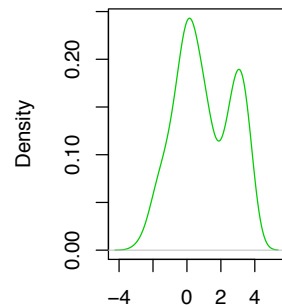


N = 300 Bandwidth = 0.3



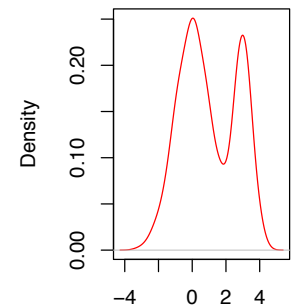
N = 300 Bandwidth = 0.1

`density.default(x = z)`



N = 300 Bandwidth = 0.4793

`density.default(x = z)`



N = 3000 Bandwidth = 0.3041

Different bandwidths

Minimax framework

- Our aim is to estimate $f \in V$
- We have a loss function $l(\hat{f}, f)$
(for instance $l(\hat{f}, f) = \|\hat{f} - f\|_p^p, \|\hat{f} - f\|_\infty$)

f^* is minimax (exactly) if

$$\sup_V \inf_{\hat{f}} l(\hat{f}, f) = \inf_{\hat{f}} \sup_V l(\hat{f}, f).$$

f_n^* is minimax (up to constants) if

$$c \inf_{\hat{f}_n} \sup_V l(\hat{f}_n, f) \leq \sup_V l(f_n^*, f) \leq C \inf_{\hat{f}_n} \sup_V l(\hat{f}_n, f) \quad \forall n.$$

Minimax framework

Two steps :

Find a rate $r(n)$

1. Lower bound $\sup_V \mathbb{E} \frac{n}{f} l(\hat{f}_n, f) \geq c r(n)$
2. Upper bound : construct an estimation method f_n^* with $\sup_V \mathbb{E} \frac{n}{f} l(f_n^*, f) \leq C r(n)$.

Functional constraints and losses

Models : density-1, regression-2, white noise model-3

$$V = V_\alpha(L) = \{f : [0, 1] \mapsto \mathbb{R}, \sup_{|x-y| \leq \delta} |f(x) - f(y)| \leq L\delta^\alpha, \forall \delta, \\ |f(0)| \leq L\}$$

$$l(\hat{f}, f) = \|\hat{f} - f\|_p^p$$

Upper bound

$$\hat{K}_{h_n}(x) = \frac{1}{nh_n} \sum_{i=0}^n K\left(\frac{x - X_i}{h_n}\right) \quad \text{density (Rosenblatt 1956)}$$

$$\hat{K}_{h_n}(x) = \frac{1}{nh_n} \sum_{i=0}^n K\left(\frac{x - X_i}{h_n}\right) Y_i \quad \text{regression (Nadaraya-Watson 1964)}$$

$$\hat{K}_{h_n}(x) = \int_0^1 \frac{1}{h_n} K\left(\frac{x - t}{h_n}\right) dY_t^\epsilon \quad \text{white noise}$$

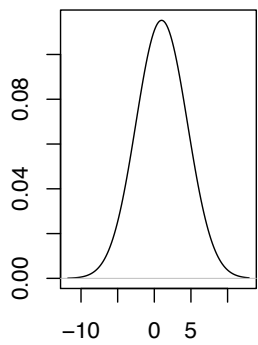
Theorem 2

For $0 \leq \alpha \leq 1$, in models 1, 2 or 3, $\int K(x) dx = 1$, K compactly supported $[-M, M]$, $\|K\|_\infty < \infty$.

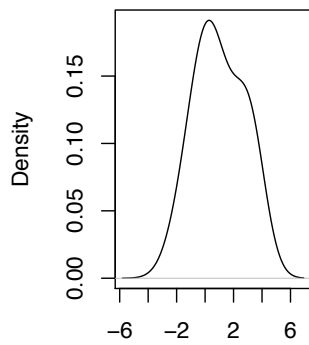
$$f_n^* = \hat{K}_{h_n}, \quad h_n = n^{\frac{-1}{1+2\alpha}}$$

$$\sup_{V_\alpha(L)} \|f_n^* - f\|_2^2 \leq C n^{\frac{-2\alpha}{1+2\alpha}} = Cr(n)$$

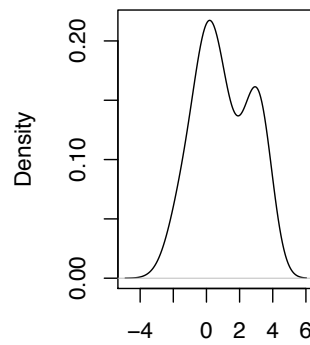
`density.default(x = z, bw = 3)` `density.default(x = z, bw = 1)` `density.default(x = z, bw = 0.7)` `density.default(x = z, bw = 0.4)`



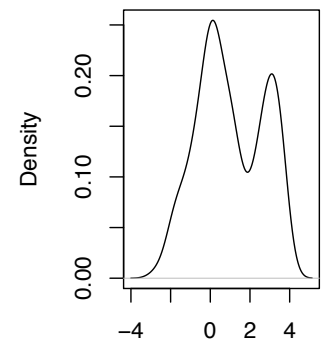
N = 300 Bandwidth = 3



N = 300 Bandwidth = 1

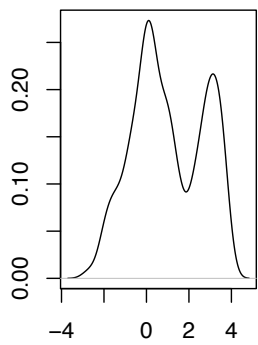


N = 300 Bandwidth = 0.7

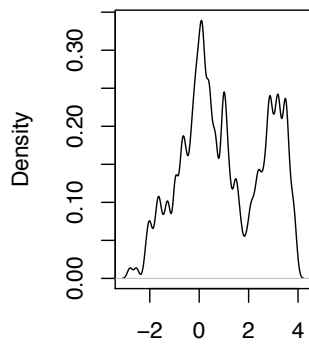


N = 300 Bandwidth = 0.4

`density.default(x = z, bw = 0.3)` `density.default(x = z, bw = 0.1)`

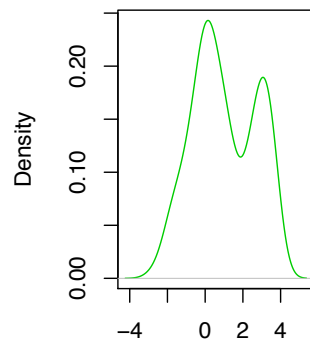


N = 300 Bandwidth = 0.3



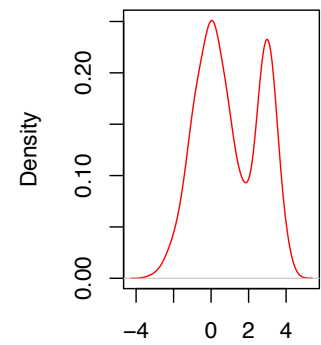
N = 300 Bandwidth = 0.1

`density.default(x = z)`



N = 300 Bandwidth = 0.4793

`density.default(x = z)`



N = 3000 Bandwidth = 0.3041

Upper bound, density model

$$\mathbb{E}_f^n(\hat{K}_{h_n}(x) - f(x))^2 = \mathbb{E}_f^n(\hat{K}_{h_n}(x) - K_{h_n}f(x))^2 + (K_{h_n}f(x) - f(x))^2$$

Balance bias, variance

Variance

$$\begin{aligned}\mathbb{E}_f^n(\hat{K}_{h_n}(x) - K_{h_n}f(x))^2 &\leq \frac{1}{n} \int K_{h_n}(x)^2 f(x) dx \\ &\leq \frac{L}{nh_n} \int K(x)^2 dx \\ &\leq \frac{L}{nh_n} 2M \|K\|_\infty^2\end{aligned}$$

Balance bias, variance

Bias

$$\begin{aligned} |K_{h_n}f(x) - f(x)| &= \left| \int K_{h_n}(u)[f(x-u) - f(x)]du \right| \\ &\leq \int |K_{h_n}(u)| \sup_{|u/h_n| \leq M} |f(x-u) - f(x)| du \\ &\leq 2M \|K\|_{\infty} L(2Mh_n)^{\alpha} \end{aligned}$$

Balance bias, variance

$$\begin{aligned} \mathbb{E}_f^n (\hat{K}_{h_n}(x) - f(x))^2 &= \mathbb{E}_f^n (\hat{K}_{h_n}(x) - K_{h_n}f(x))^2 + (K_{h_n}f(x) - f(x))^2 \\ &\leq \frac{L}{nh_n} 2M \|K\|_\infty^2 + 4M^2 \|K\|_\infty^2 L^2 (2Mh_n)^{2\alpha} \end{aligned}$$

$$\mathbb{E}_f^n \|\hat{K}_{h_n} - f\|_2^2 \leq c \left[\frac{L}{nh_n} 2M \|K\|_\infty^2 + 4M^2 \|K\|_\infty^2 L^2 (2Mh_n)^{2\alpha} \right]$$

$$c \leq 2M + 1$$

Optimized for

$$h_n^* \sim n^{\frac{-1}{2\alpha+1}}$$

$$\mathbb{E}_f^n \|\hat{K}_{h_n^*} - f\|_2^2 \leq C n^{\frac{-2\alpha}{2\alpha+1}}$$

Questions

- What about the lower bound? (Is $n^{\frac{-2\alpha}{2\alpha+1}}$ the minimax rate?)
- What is the behavior for a general L_p loss?
- How to choose h_n in practice?

Orthogonal series methods

$f \in L^2([0, 1])$, $\mathcal{E} = \{\psi_i, i \in \mathbb{N}\}$ orthonormal basis of $L^2([0, 1], dt)$,

$$f = \sum \theta_i \psi_i, \quad x_i = \int \psi_i dY, \quad i \in \mathbb{N},$$

General estimator

$$\hat{f} = \sum_{i \in A} \hat{\theta}_i \psi_i.$$

Two choices : A , $\hat{\theta}_i$.

Orthogonal series methods

$$A = (\text{generally}) \{0, \dots, K\}$$

$$\hat{\theta}_i = \frac{1}{n} \sum_{i=0}^n \psi_i(X_i) \quad \text{density}$$

$$\hat{\theta}_i = \frac{1}{n} \sum_{i=0}^n \psi_i(X_i) Y_i \quad \text{regression}$$

$$\hat{\theta}_i = \int_0^1 \psi_i(t) dY_t^\epsilon \quad \text{white noise}$$

$$\rightarrow \hat{f}_K$$

Upper bounds

- If we assume f belongs to a polynomially tail compact domain :

For $s > 0$, fixed,

$$V = \{f = \sum \theta_k \psi_k, \sum_{k>K} \theta_k^2 \leq LK^{-2s}, \forall K, \|f\|_\infty \leq L\}$$

Upper bounds

-

$$f \in V(s, M) = \{f = \sum \theta_k \psi_k, \sum_{k>K} \theta_k^2 \leq L^2 K^{-2s}, \forall K, \|f\|_\infty \leq L\}$$

-

$$\begin{aligned} \overset{n}{f} \|\hat{f}_K - f\|^2 &= \sum_{k \leq K} \overset{n}{f} (\hat{\theta}_k - \theta_k)^2 + \sum_{k > K} \theta_k^2 \\ &\leq (K+1) \frac{L}{n} + \sum_{k > K} \theta_k^2 \\ &\leq (K+1) \frac{L}{n} + L^2 K^{-2s} \end{aligned}$$

Optimized for $K_s^* = c[n]^{\frac{1}{1+2s}} \leq K_0^* = cn$
(decreasing in s)

$$\sup_{f \in V(s, M)} E \|\hat{f}_{K^*} - f\|^2 \leq c' n^{\frac{-2s}{1+2s}}.$$

Kernels versus series

Easier calculation for series (for proof and computation)

Tuning parameters $K \sim h^{-1}$ gives an interpretation of the bandwidth parameter and the 'dimension' of the problem.

Space V depends on the basis, on the numbering in the basis, only allows a L_2 loss function.

Bases and functional spaces

Trigonometric basis and Sobolev spaces

- : $L_2([0, 1])$ of periodic functions

$$\psi_0 = 1$$

$$\psi_{2k}(x) = \sqrt{2} \cos 2k\pi x$$

$$\psi_{2k+1}(x) = \sqrt{2} \sin 2k\pi x$$

Let $\beta \in \mathbb{R}_*$, the following Sobolev space,

$$W(\beta, L) = \{f : [0, 1] \mapsto \mathbb{R} : f^{\beta-1} \text{ absolut. continuous } \int (f^\beta)^2(x) dx \leq L\}$$

$$W^{per}(\beta, L) = \{f \in W(\beta, L), \text{ periodic}\}$$

Trigonometric basis and Sobolev spaces

Let

$$\Theta((a_j), Q) = \{\theta \in l^2 : \sum_j a_j^2 \theta_j^2 \leq Q^2\}$$

We have,

$$W^{per}(\beta, L) = \Theta((a_j), Q) := \Theta(\beta, Q),$$

$$a_j = j^\beta, \text{ } j \text{ even}$$

$$a_j = (j-1)^\beta, \text{ } j \text{ odd}$$

$$Q = \frac{L}{\pi^\beta}$$

Trigonometric basis and Sobolev spaces

$$\Theta(\beta, Q) = \{\theta \in l^2 : \sum_j j^{2\beta} \theta_j^2 \leq Q^2\}$$

$$\Theta(\beta, Q) \subset V(\beta, Q)$$

$$\sum_{j \geq K} \theta_j^2 \leq \sum_{j \geq K} \left[\frac{j}{K}\right]^{2\beta} \theta_j^2 \leq K^{-2\beta} \sum_j j^{2\beta} \theta_j^2$$