

Shock Persistence and Shock Frequency in VIX

A Quantitative Analysis of Volatility Dynamics

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Outline

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2 Data

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What is VIX?

- **VIX** = CBOE Volatility Index, derived from S&P 500 option prices.
- Often called the “fear gauge” — rises when markets expect turbulence.
- Understanding VIX dynamics is crucial for:
 - Risk management and hedging
 - Derivatives pricing
 - Portfolio allocation

Research Questions

① How persistent is volatility?

How long does a VIX shock take to decay?

② How frequently do large spikes occur?

Can we model extreme events as a point process?

③ Do shocks cluster?

Is there self-excitation in shock arrivals?

④ How do regimes affect shock dynamics?

Do crisis periods show different behavior?

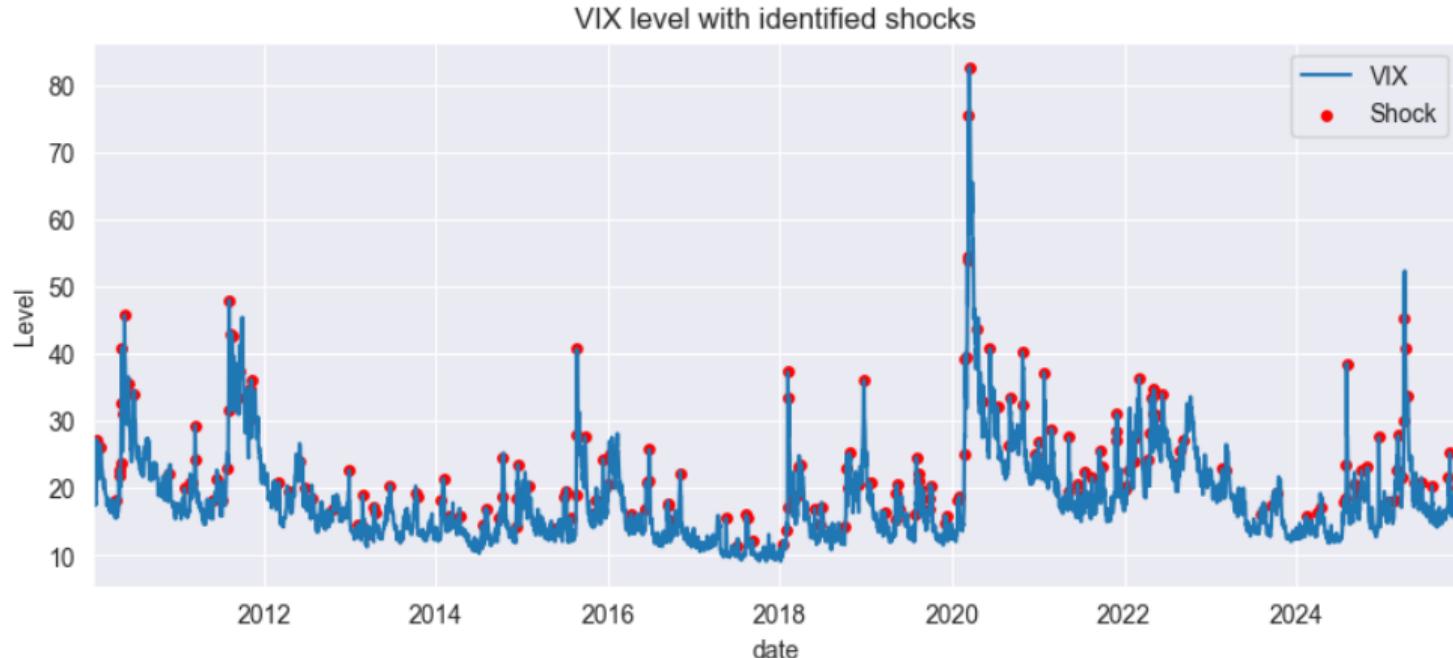
⑤ Can we forecast VIX volatility?

Do GARCH-type models beat simple baselines out-of-sample?

Data Overview

- **Source:** Yahoo Finance (ticker ^VIX)
- **Period:** January 2010 – November 2025
- **Observations:** 4,148 business days
- **Pre-processing:**
 - Forward-fill missing dates
 - 0.1% winsorization to limit outlier influence
 - Compute $\log(\text{VIX})$ and daily log-changes $\Delta \log(\text{VIX})$

VIX Time Series



- Red markers indicate identified shock days (top 5% of $\Delta \log VIX$).

Volatility Models: GARCH Family

GARCH(1,1):

$$\sigma_t^2 = \omega + \alpha \varepsilon_{t-1}^2 + \beta \sigma_{t-1}^2$$

EGARCH(1,1):

$$\ln \sigma_t^2 = \omega + \beta \ln \sigma_{t-1}^2 + \alpha(|z_{t-1}| - \mathbb{E}|z|) + \gamma z_{t-1}$$

GJR-GARCH(1,1):

$$\sigma_t^2 = \omega + (\alpha + \gamma \mathbf{1}_{\varepsilon_{t-1} < 0}) \varepsilon_{t-1}^2 + \beta \sigma_{t-1}^2$$

- GARCH: Symmetric; EGARCH/GJR: Asymmetric (leverage effect via γ).
- Distribution: Auto-selected (Normal, Student-t, GED) via PIT diagnostics.

HAR-RV: Realized Volatility Model

Heterogeneous Autoregressive Realized Volatility (Corsi, 2009):

$$RV_{t+1} = \beta_0 + \beta_d RV_t + \beta_w \overline{RV}_t^{(w)} + \beta_m \overline{RV}_t^{(m)} + \varepsilon_{t+1}$$

where:

- $RV_t = r_t^2$ (proxy for realized variance)
- $\overline{RV}_t^{(w)} = \frac{1}{5} \sum_{i=0}^4 RV_{t-i}$ (weekly average)
- $\overline{RV}_t^{(m)} = \frac{1}{22} \sum_{i=0}^{21} RV_{t-i}$ (monthly average)

Use: Baseline model capturing heterogeneous investor horizons.

Shock Identification: Two Approaches

Method 1: Quantile-Based (Fixed Threshold)

$$\text{Shock}_t = \mathbf{1}\{\Delta \log(\text{VIX})_t \geq Q_{0.95}\}$$

- Simple, interpretable: top 5% of daily changes.

Method 2: Volatility-Relative (Surprise-Based)

$$\text{Shock}_t = \mathbf{1}\{|r_t| > k \cdot \sigma_t\}, \quad k = 2$$

- Captures “surprises” relative to expected volatility.
- More meaningful in high-volatility regimes.

Point Process Models for Shock Arrivals

Homogeneous Poisson Process (HPP):

$$\lambda(t) = \lambda \quad (\text{constant rate})$$

Non-Homogeneous Poisson Process (NHPP):

$$\lambda(t) = \exp(\beta_0 + \beta_1 \cdot \text{lagged_log_VIX}_t)$$

Hawkes Self-Exciting Process:

$$\lambda(t) = \mu + \sum_{t_i < t} \alpha \cdot e^{-\beta(t-t_i)}$$

- Each shock temporarily increases intensity \Rightarrow **clustering**.
- **Branching ratio** $= \alpha/\beta < 1 \Rightarrow$ stationary.

- **Out-of-Sample Design:**

- Train on first 75% of data (through Dec 2021).
- Monthly rolling re-estimation of GARCH.
- Forecast 1-step-ahead variance into the remaining 25%.

- **Baselines:**

- EWMA ($\lambda = 0.94$)
- 63-day rolling variance
- HAR-RV model

- **Metrics:**

- Log-score (predictive density evaluation)
- 95% coverage rate
- PIT histogram (calibration diagnostic)
- Diebold–Mariano test (statistical significance)

Volatility Model Comparison

| Model | Distribution | AIC | Persistence | Half-life | Leverage (γ) |
|--------------------|--------------|---------------|-------------|-----------|-----------------------|
| GARCH(1,1) | GED | 27,531 | 0.852 | 4.3 days | – |
| EGARCH(1,1) | GED | 27,395 | 0.934 | 10.2 days | Yes |
| GJR-GARCH(1,1) | GED | 27,447 | 0.866 | 4.8 days | -0.27 |

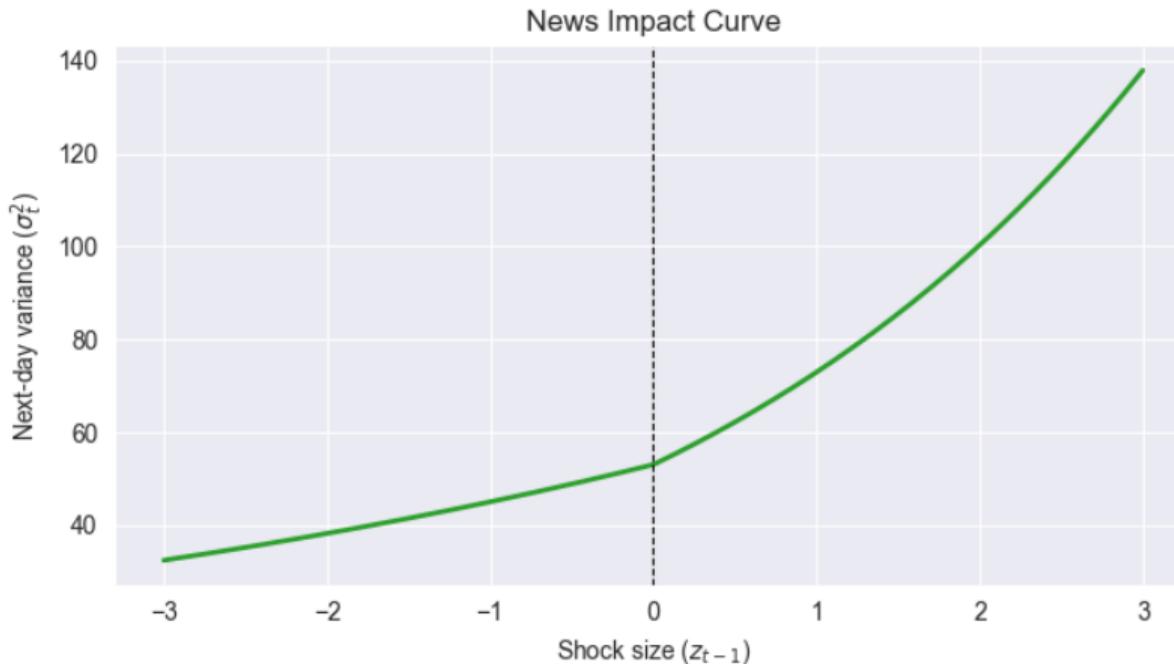
- EGARCH achieves lowest AIC \Rightarrow best in-sample fit.
- Higher persistence in EGARCH \Rightarrow shocks decay more slowly (≈ 10 days half-life).
- GJR-GARCH $\gamma < 0$: negative returns increase volatility more.

HAR-RV Results

| Component | Coefficient | Interpretation |
|--------------------------|-------------|----------------------------|
| β_{daily} | 0.117 | Short-term (1-day) effect |
| β_{weekly} | 0.231 | Medium-term (5-day) effect |
| β_{monthly} | 0.042 | Long-term (22-day) effect |
| R^2 | 0.049 | Explanatory power |

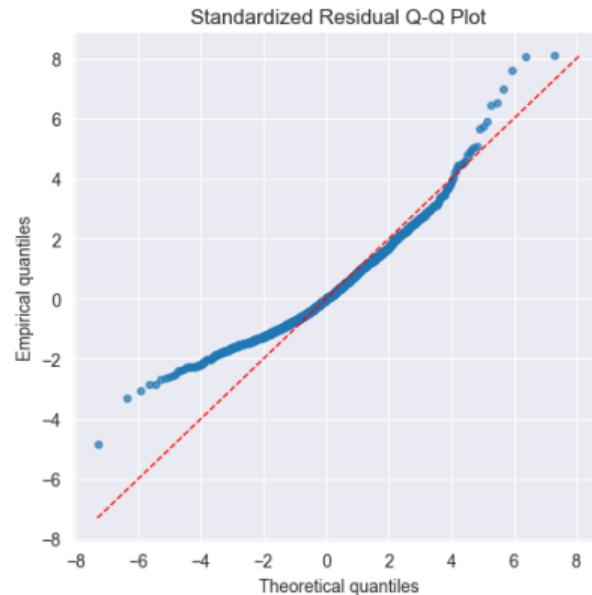
- Weekly component dominates \Rightarrow swing traders' horizon matters most.
- Low R^2 expected: squared returns are noisy proxies for true variance.
- HAR-RV provides parsimonious baseline for comparison.

News Impact Curve (Asymmetry)



- Positive shocks (VIX spikes) increase future variance more than negative shocks of equal magnitude.

Q-Q Plot of Standardized Residuals



- Points hug the 45° line in the tails \Rightarrow GED captures fat tails well.

Shock Statistics: Two Methods Compared

| Metric | Quantile (95%) | Vol-Relative (2σ) |
|--------------------|----------------|----------------------------|
| Threshold (avg) | 0.127 | 0.147 |
| Total shocks | 208 | 207 |
| Rate (shocks/year) | 9.0 | 9.0 |

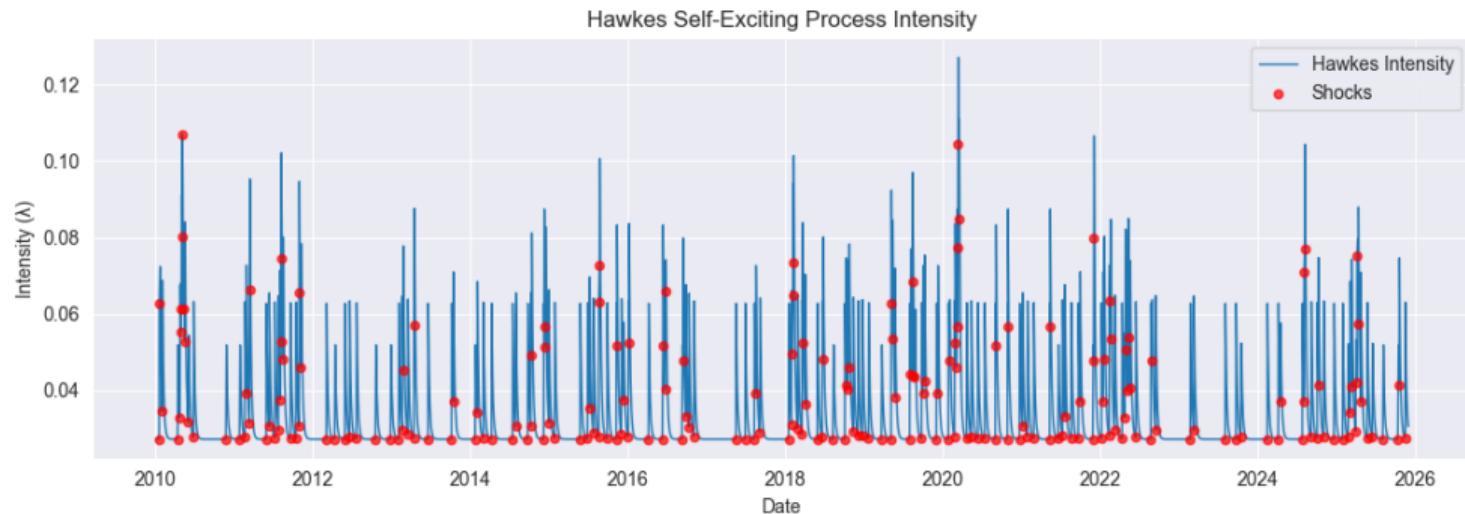
- Both methods identify ~ 9 shocks/year.
- Vol-relative threshold adapts to market conditions.
- NHPP: Lagged log VIX coefficient = $-0.16 \Rightarrow$ low VIX predicts fewer shocks.

Hawkes Self-Exciting Process

| Parameter | Value | Interpretation |
|---------------------|-----------|-----------------------------------|
| Baseline μ | 0.027/day | Background intensity |
| Excitation α | 0.043 | Jump after each shock |
| Decay β | 0.183 | Rate of decay |
| Branching ratio | 0.23 | Fraction triggered by past shocks |
| Half-life | 3.8 days | Excitation halving time |

- Branching ratio $< 1 \Rightarrow$ process is **stationary**.
- $\approx 23\%$ of shocks are “triggered” by previous shocks.
- Excitation decays with ~ 4 -day half-life.

Hawkes Intensity Over Time



- Intensity spikes after each shock, then decays exponentially.
- Clustering visible during crisis periods.

Compound Poisson Process: Motivation

Problem: HPP/NHPP/Hawkes model *when* shocks occur, but not *how big*.

Compound Poisson Process (CPP):

$$S(T) = \sum_{i=1}^{N(T)} J_i$$

where:

- $N(T) \sim \text{Poisson}(\lambda T)$: number of shocks by time T
- $J_i \sim F$: random jump sizes (iid)
- $S(T)$: cumulative shock impact over horizon T

Key Insight: Models *both* timing and magnitude \Rightarrow risk quantification.

CPP: Jump Size Distribution Selection

Candidates for jump size distribution F :

- Exponential: $f(x) = \lambda e^{-\lambda x}$ (memoryless)
- Gamma: $f(x) \propto x^{k-1} e^{-x/\theta}$ (flexible shape)
- Lognormal: $\ln(J) \sim N(\mu, \sigma^2)$ (multiplicative)
- **Pareto**: $f(x) = \frac{\alpha x_m^\alpha}{x^{\alpha+1}}$ (heavy tails)
- Weibull: $f(x) \propto x^{k-1} e^{-(x/\lambda)^k}$ (hazard rate)

Selection via AIC + KS test:

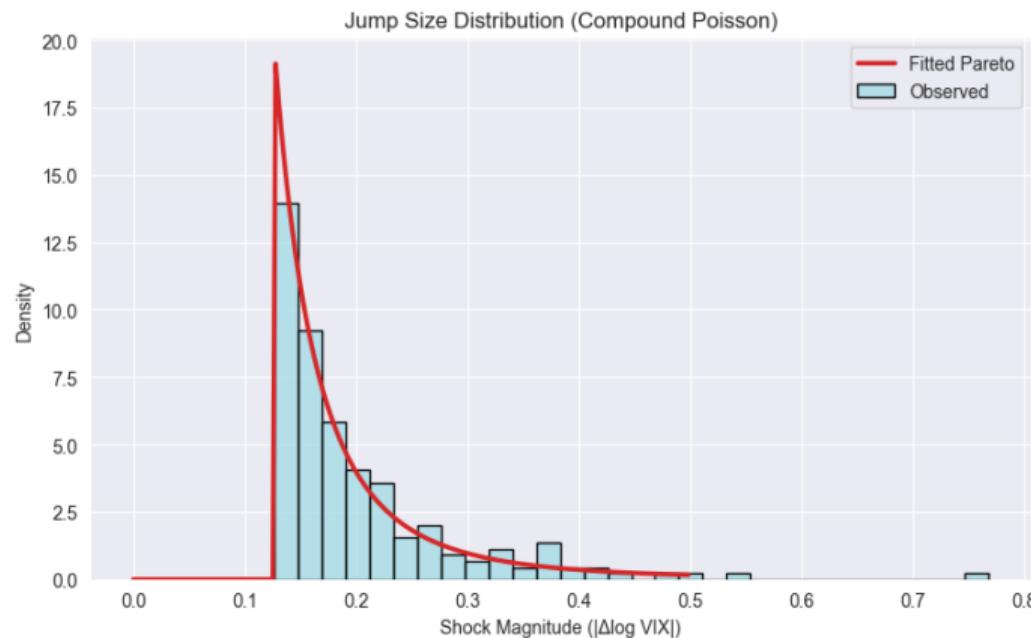
- Best fit: **Pareto** ($\alpha = 2.50$, $x_{\min} = 0.127$)
- KS p-value = 0.42 \Rightarrow cannot reject fit

CPP: Fitted Parameters

| Parameter | Value | Interpretation |
|-----------------|------------|-------------------------------|
| λ | 12.64/year | Shock arrival rate |
| $\mathbb{E}[J]$ | 0.211 | Mean jump size (21% log-move) |
| $\text{Std}[J]$ | 0.189 | Jump size volatility |
| $\mathbb{E}[S]$ | 2.67/year | Expected annual impact |
| VaR (95%) | 4.24 | 95th percentile annual impact |
| CVaR (95%) | 5.01 | Expected Shortfall (tail avg) |

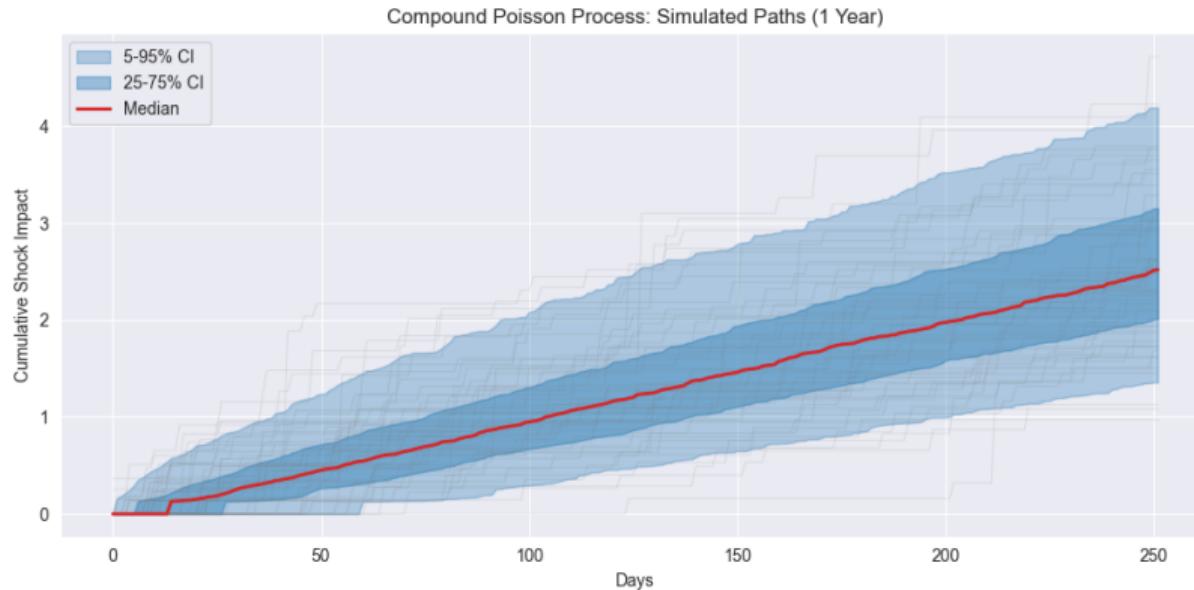
- $\mathbb{E}[S] = \lambda \cdot \mathbb{E}[J] = 12.64 \times 0.211 = 2.67$
- VaR/CVaR computed via 10,000 Monte Carlo simulations.

CPP: Jump Size Distribution



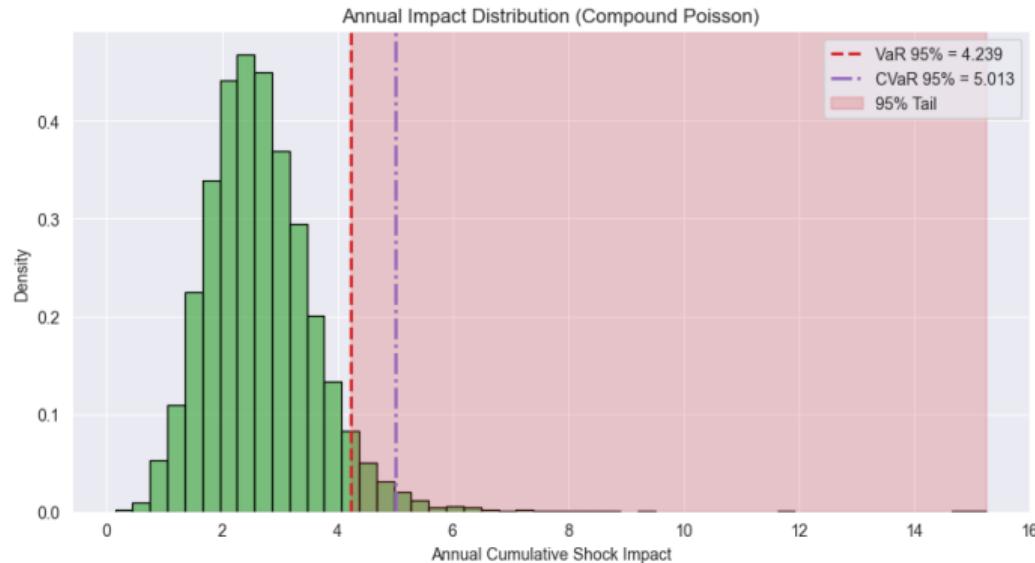
- Pareto distribution captures heavy right tail of shock magnitudes.
- Most shocks are moderate; a few are extreme.

CPP: Simulated Paths



- Gray: 50 sample paths of cumulative annual shock impact.
- Shaded: 5–95% and 25–75% confidence bands.
- Red: Median trajectory.

CPP: VaR and CVaR Distribution



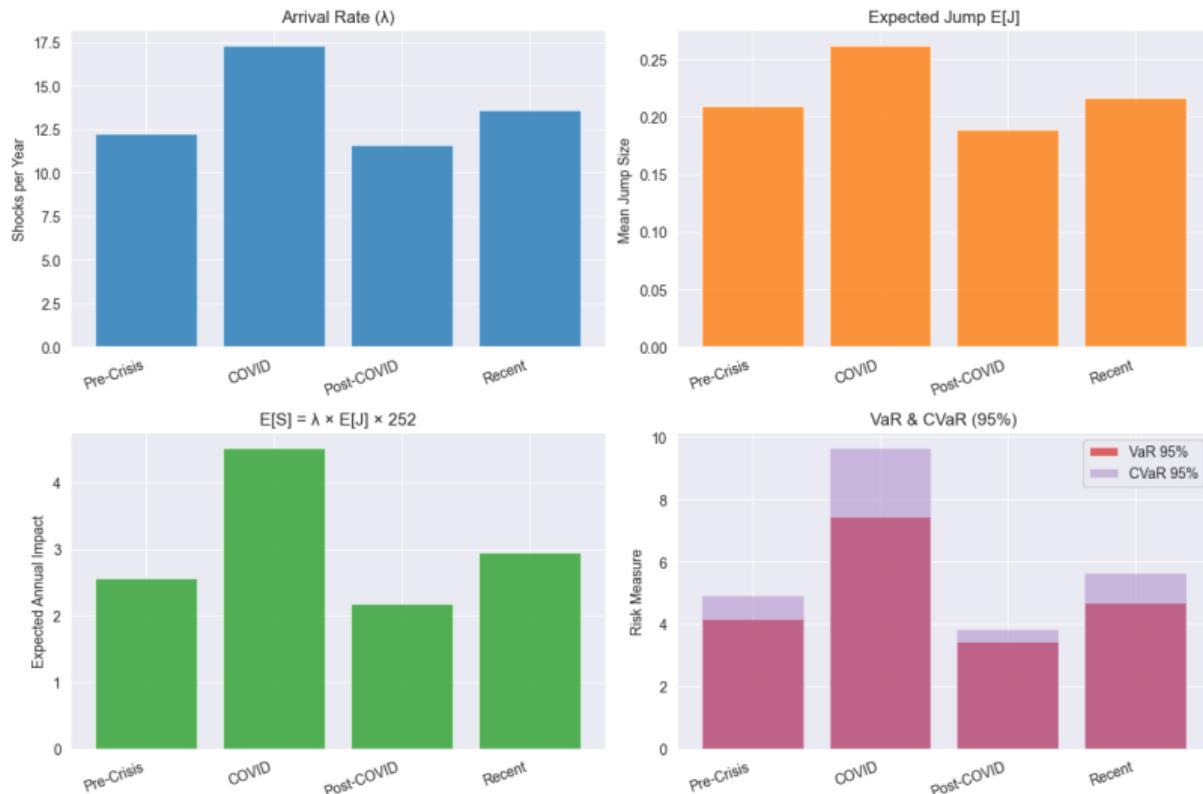
- Distribution of annual cumulative shock impact.
- $\text{VaR}(95\%) = 4.24$: “In 95% of years, total impact ≤ 4.24 .”
- $\text{CVaR}(95\%) = 5.01$: “Average impact in worst 5% of years.”

CPP: Regime Comparison

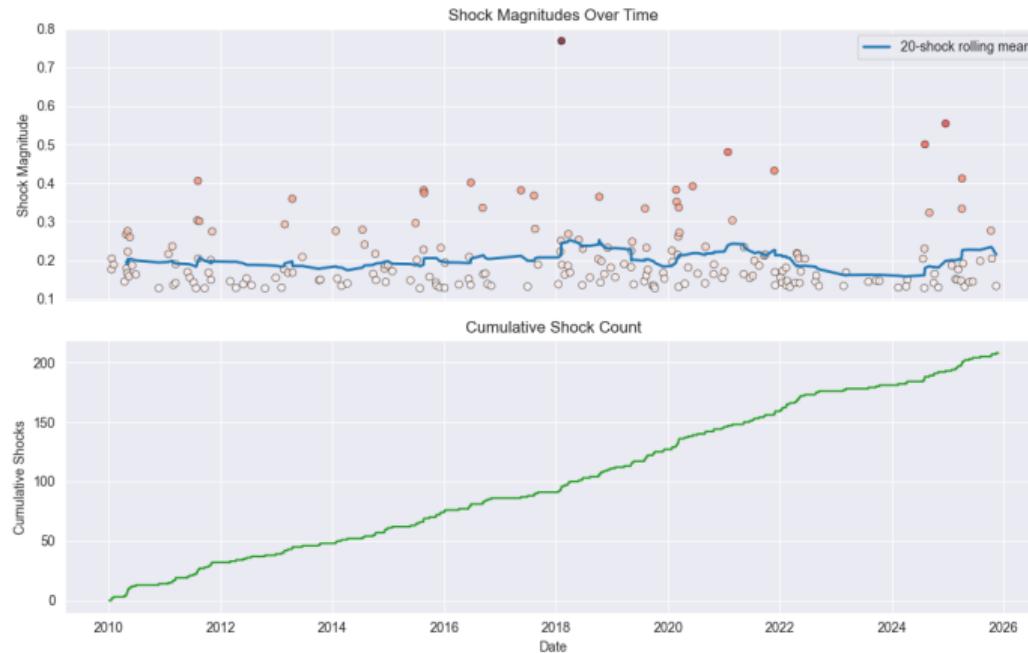
| Regime | λ/Year | $\mathbb{E}[J]$ | $\mathbb{E}[S]/\text{Year}$ | VaR 95% | CVaR 95% |
|--------------|-----------------------|-----------------|-----------------------------|-------------|-------------|
| Pre-Crisis | 12.3 | 0.209 | 2.57 | 4.15 | 4.92 |
| COVID | 17.3 | 0.262 | 4.53 | 7.44 | 9.65 |
| Post-COVID | 11.6 | 0.188 | 2.19 | 3.44 | 3.85 |
| Recent | 13.6 | 0.216 | 2.95 | 4.70 | 5.63 |

- COVID: Both higher arrival rate AND larger jumps \Rightarrow 76% higher $\mathbb{E}[S]$.
- VaR nearly doubles during crisis periods.

CPP: Regime Risk Comparison



Shock Magnitudes Over Time



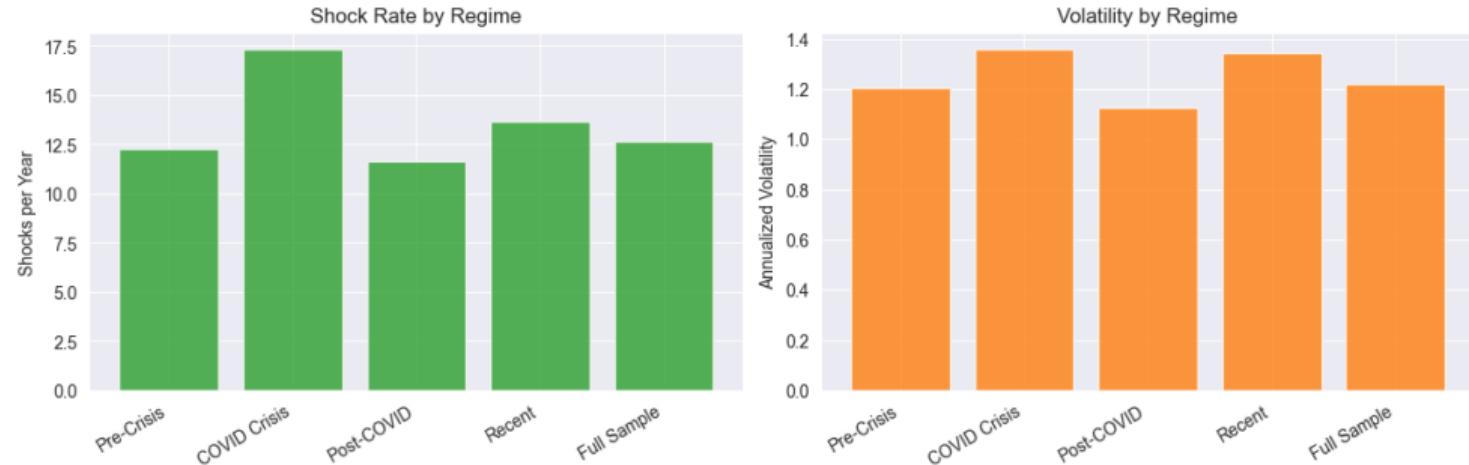
- Top: Individual shock magnitudes with 20-shock rolling mean.
- Bottom: Cumulative shock count shows acceleration during crises.

Regime Analysis: Shock Rates Across Periods

| Regime | Period | Obs | Shocks | Rate/Year | Ann. Vol |
|--------------|-----------|-------|--------|-------------|-------------|
| Pre-Crisis | 2010–2019 | 2,606 | 127 | 12.3 | 1.21 |
| COVID | 2020 | 262 | 18 | 17.3 | 1.36 |
| Post-COVID | 2021–2023 | 781 | 36 | 11.6 | 1.12 |
| Recent | 2024–2025 | 499 | 27 | 13.6 | 1.34 |
| Full Sample | 2010–2025 | 4,148 | 208 | 12.6 | 1.22 |

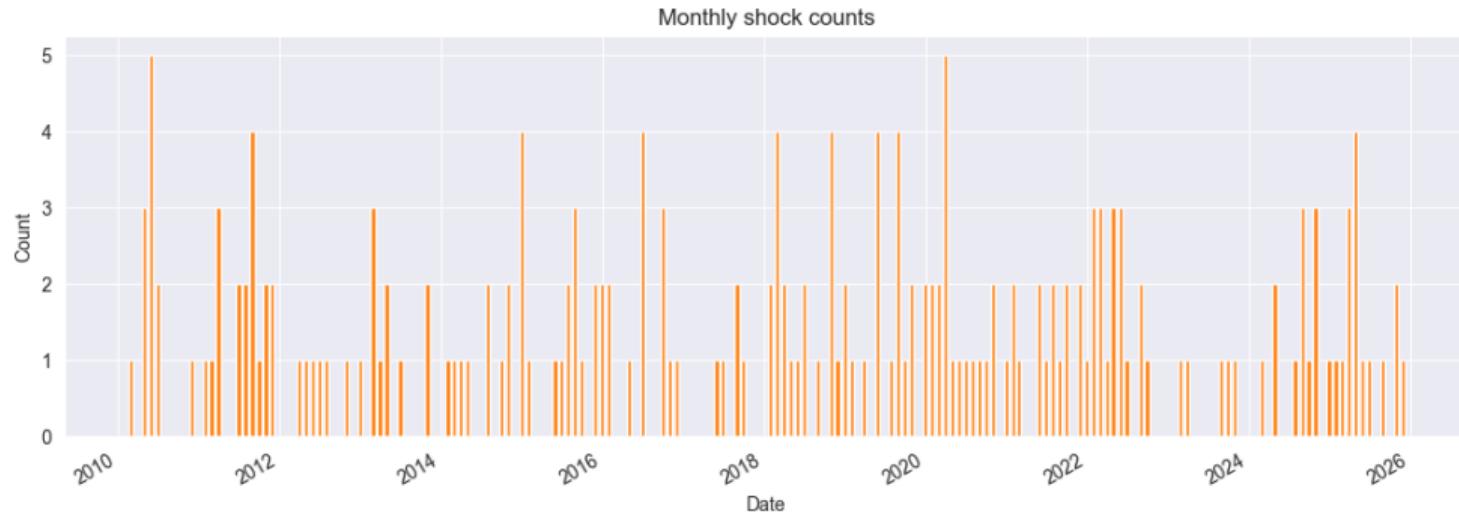
- COVID period: **41% higher** shock rate (17.3 vs 12.3/year).
- Volatility elevated across all crisis periods.

Regime Comparison Visualization



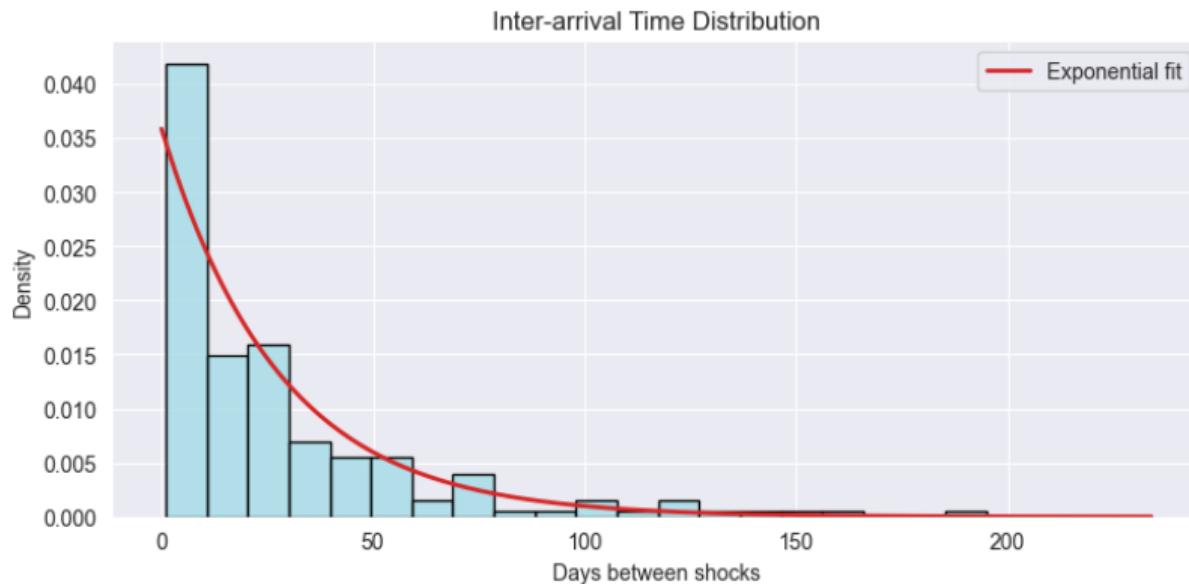
- Bar chart shows clear elevation during COVID and recent volatility.

Monthly Shock Counts



- Notable clustering: 2011–12 (Euro crisis), 2018 (Volmageddon), 2020 (COVID), 2022 (rate hikes).

Inter-Arrival Time Distribution



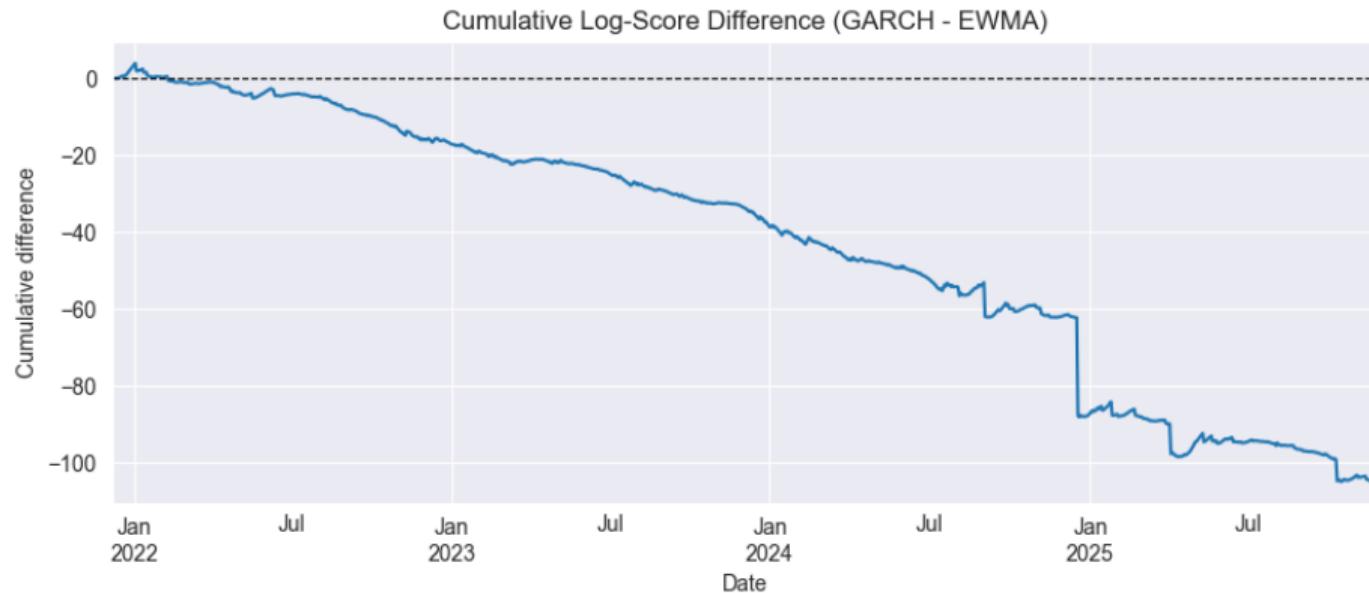
- Histogram vs. exponential PDF: reasonable fit validates Poisson assumption.

Forecast Evaluation: Log-Scores

| Model | Log-Score (higher = better) |
|-------------|-----------------------------|
| GARCH | 1.275 |
| EWMA | 1.376 |
| Rolling Var | 1.276 |
| HAR-RV | 1.271 |

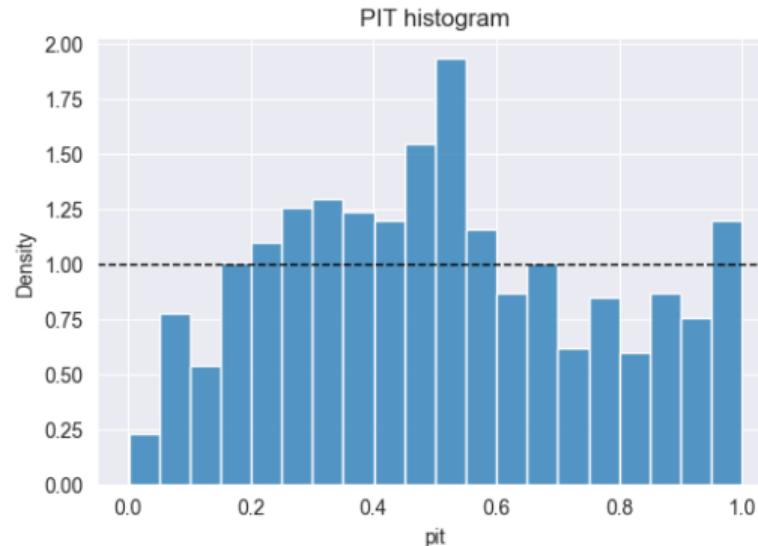
- EWMA slightly outperforms GARCH out-of-sample.
- Diebold–Mariano $p < 0.001 \Rightarrow$ difference is statistically significant.
- HAR-RV comparable to rolling variance baseline.

Cumulative Log-Score Difference



- Downward trend: EWMA consistently accumulates better scores.
- GARCH rarely catches up, even briefly.

PIT Histogram (Calibration)



- Near-uniform distribution indicates well-calibrated density forecasts.
- Mean ≈ 0.51 , std ≈ 0.26 (ideal: 0.5, 0.29).
- 95% coverage: 95.0% (at nominal).

Key Findings

- ① **Persistence:** VIX volatility shocks have a half-life of 4–10 days; EGARCH captures longer memory.
- ② **Leverage Effect:** GJR-GARCH $\gamma = -0.27$ confirms asymmetric response.
- ③ **Self-Excitation:** Hawkes branching ratio = 0.23 \Rightarrow shocks trigger more shocks.
- ④ **Compound Poisson:** Pareto-distributed jumps; VaR 95% = 4.24/year.
- ⑤ **Regime Dependence:** COVID showed 76% higher expected annual impact.
- ⑥ **Forecasting:** EWMA beats GARCH on log-score; both well-calibrated.

Practical Implications

- **Risk Management:**
 - Use EGARCH/GJR-GARCH for asymmetric VaR calculations.
 - Hawkes intensity provides real-time “shock alert” signals.
 - CPP gives VaR/CVaR for aggregate annual shock impact.
- **Option Pricing:**
 - Leverage effect matters: adjust for skewed vol-of-vol.
 - Pareto jumps justify heavy-tail adjustments in pricing.
- **Portfolio Allocation:**
 - Regime detection enables adaptive hedging strategies.
 - CPP regime analysis shows crisis periods need 2× risk budget.

Limitations

- GARCH re-estimation is computationally expensive; rolling used here.
- EWMA's superiority may reflect VIX's strong mean-reversion.
- NHPP covariates limited to lagged VIX; macro indicators could improve.
- Hawkes estimation assumes exponential decay; other kernels possible.
- Daily data only; intraday dynamics not captured.

Future Work

- Incorporate regime-switching GARCH (Markov-switching) models.
- Test realized volatility from high-frequency data.
- Extend NHPP/Hawkes with external regressors (credit spreads, VIX term structure).
- Multivariate Hawkes for cross-asset shock contagion.
- Deploy real-time monitoring dashboard with Hawkes intensity tracking.

- VIX exhibits **persistent, asymmetric volatility clustering** well captured by EGARCH.
- Large spikes arrive at $\sim 13/\text{year}$ and exhibit **self-excitation** (Hawkes branching ratio 0.23).
- **Compound Poisson** quantifies aggregate shock risk: VaR 95% = 4.24/year.
- **COVID regime** showed 76% higher expected annual impact than baseline.
- Out-of-sample, **EWMA remains a tough benchmark** to beat for density forecasting.
- The reproducible pipeline (`runall.py`) with **28 unit tests** enables transparent research.

Thank you!

Questions?

Appendix: Project Structure

```
Shock-Persistence-and-Shock-Frequency-in-VIX/
+-- runall.py                      # Main pipeline
+-- src/
|   +-- config.py                  # Parameters
|   +-- data_pipeline.py          # Data loading
|   +-- volatility_models.py    # GARCH/EGARCH/GJR/HAR-RV
|   +-- shock_modeling.py       # HPP/NHPP/Hawkes/CPP
|   +-- forecast_evaluation.py # OOS evaluation
|   +-- visualization.py        # Plotting
+-- tests/
|   +-- test_models.py           # 28 unit tests
+-- figures/                      # Generated plots
```

Appendix: Key Equations

Log-Score:

$$S_t = \log f(r_t | \mu_t, \sigma_t^2)$$

Hawkes Intensity:

$$\lambda(t) = \mu + \sum_{t_i < t} \alpha \cdot e^{-\beta(t-t_i)}$$

Compound Poisson Process:

$$S(T) = \sum_{i=1}^{N(T)} J_i, \quad N(T) \sim \text{Poisson}(\lambda T), \quad J_i \sim F$$

Expected Annual Impact:

$$\mathbb{E}[S] = \lambda \cdot \mathbb{E}[J] \cdot T$$

Appendix: References

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