

# Homework 5

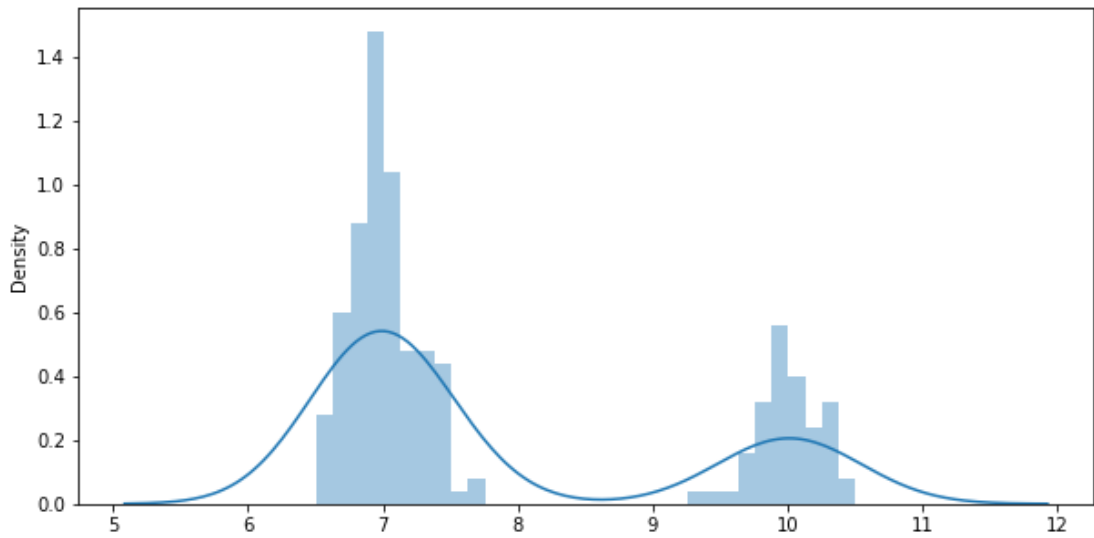
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注：因为我不懂 R 语言，所以代码都是用 Python 写的。代码文件见 homework5-code.ipynb

## 7.1

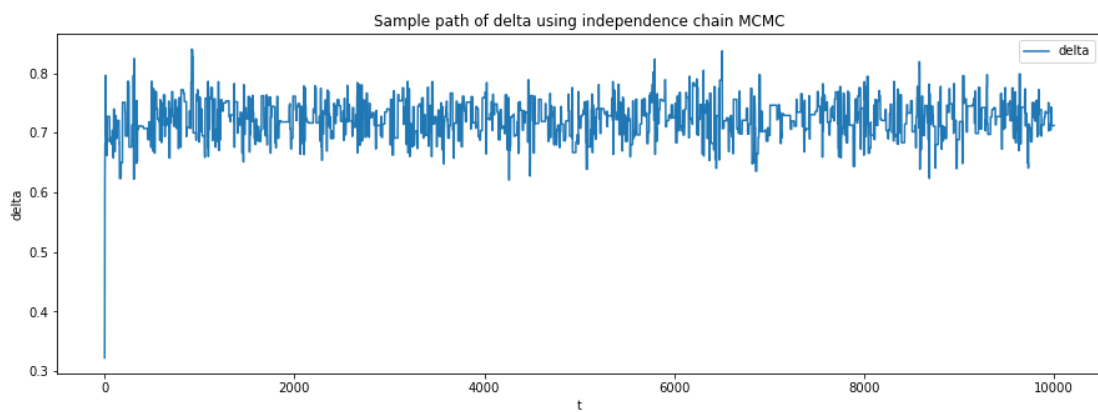
解：

a.

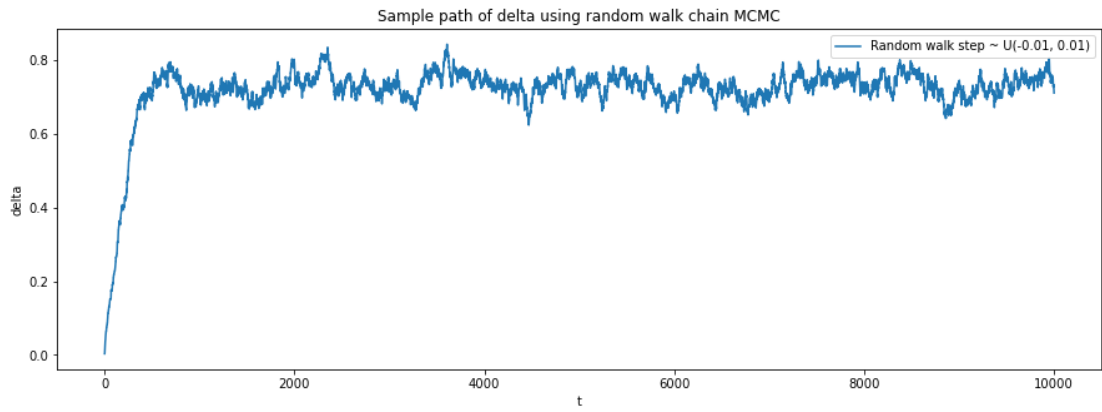
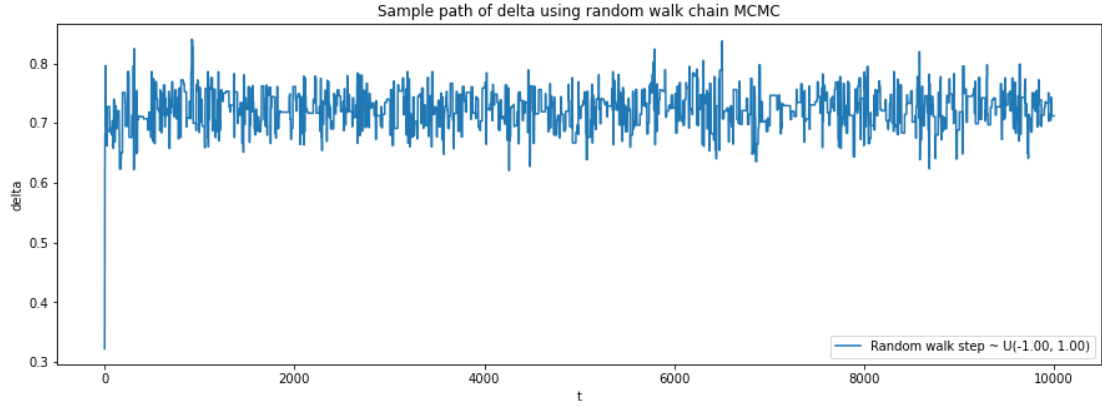


b.

$$P(X, \delta) = \prod_{i=1}^n \left( \delta \frac{1}{0.5\sqrt{2\pi}} e^{-\frac{1}{2} \frac{(x_i-7)^2}{0.5^2}} + (1-\delta) \frac{1}{0.5\sqrt{2\pi}} e^{-\frac{1}{2} \frac{(x_i-10)^2}{0.5^2}} \right)$$



c.



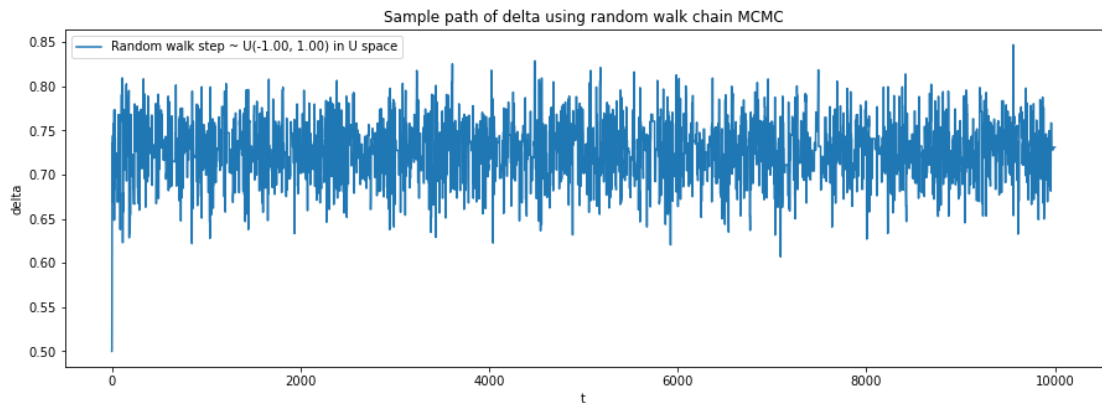
d.

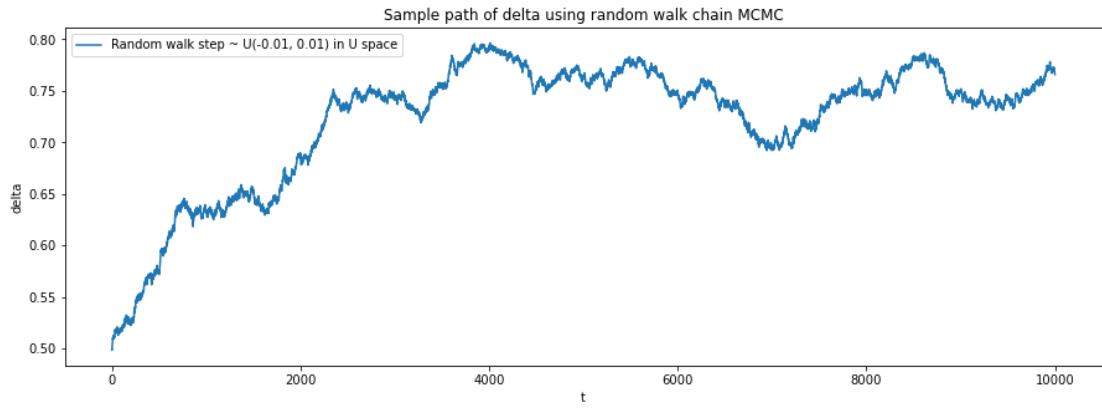
$$U = \ln \frac{\delta}{1-\delta}, \delta = \frac{1}{(e^{-U} + 1)}$$

在 U 空间进行  $\epsilon \sim U(-b, b)$  的随机游走，则

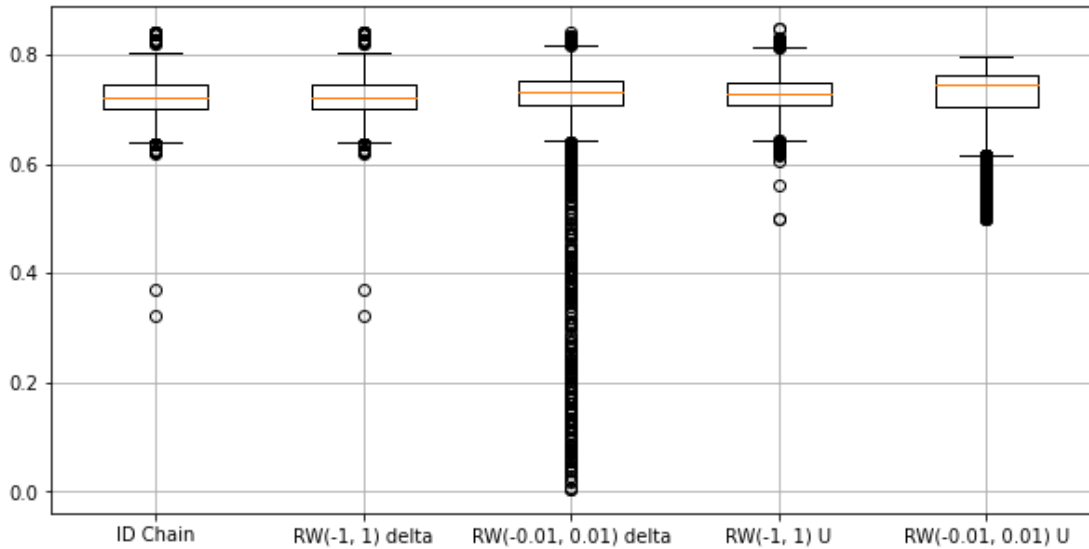
$$g(U^t | U^*) = g(U^* | U^t) = \frac{1}{2b}$$

$$|Jacob(U)| = \frac{e^{-U}}{(e^{-U} + 1)^2} = e^{-U} \delta^2$$





e.

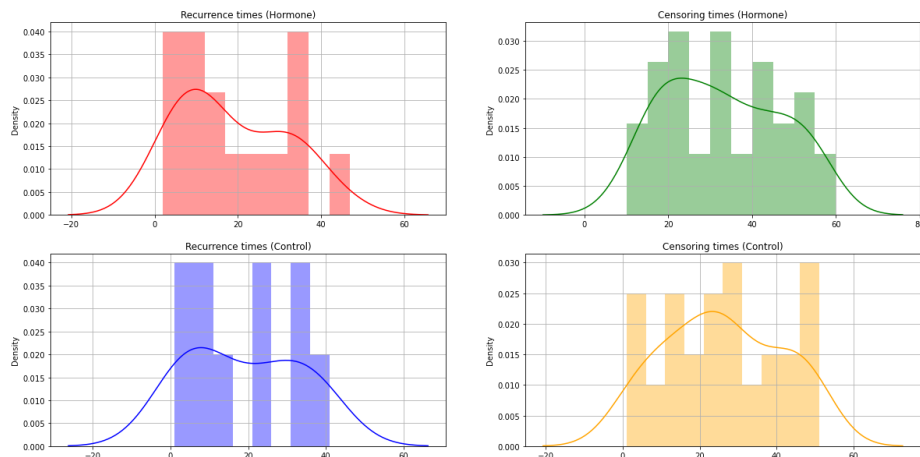


可见，当随机游走步数为 $U(-1,1)$ 时，在 $\delta$ 空间的随机游走和独立链的结果是等价的。当随机游走的步长越小，收敛速度越慢。在U空间中的随机游走不论是方差还是收敛速度都要弱于在 $\delta$ 空间的随机游走。这是因为非线性变换造成的步长扭曲。

## 7.5

解：

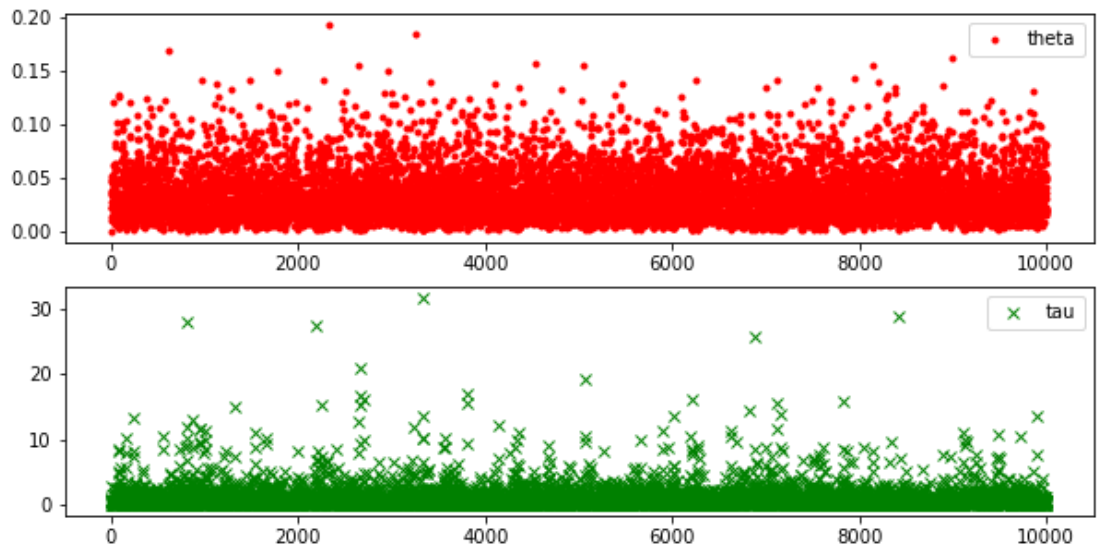
a.



b.

$$\begin{aligned}
 P(\theta|\tau) &= \frac{\theta^a \tau^b e^{(-c\theta - d\theta\tau)}}{\int_0^{+\infty} \theta^a \tau^b e^{(-c\theta - d\theta\tau)} d\theta} \\
 &= \frac{\theta^a e^{(-c\theta - d\theta\tau)}}{(d\tau + c)^{-a-1} \Gamma(a+1)} \\
 &\sim \Gamma(a+1, d\tau + c) \\
 P(\tau|\theta) &= \frac{\theta^a \tau^b e^{(-c\theta - d\theta\tau)}}{\int_0^{+\infty} \theta^a \tau^b e^{(-c\theta - d\theta\tau)} d\tau} \\
 &= \frac{\tau^b e^{-d\theta\tau}}{(d\theta)^{-b-1} \Gamma(b+1)} \\
 &\sim \Gamma(b+1, d\theta)
 \end{aligned}$$

c. 采样结果为:

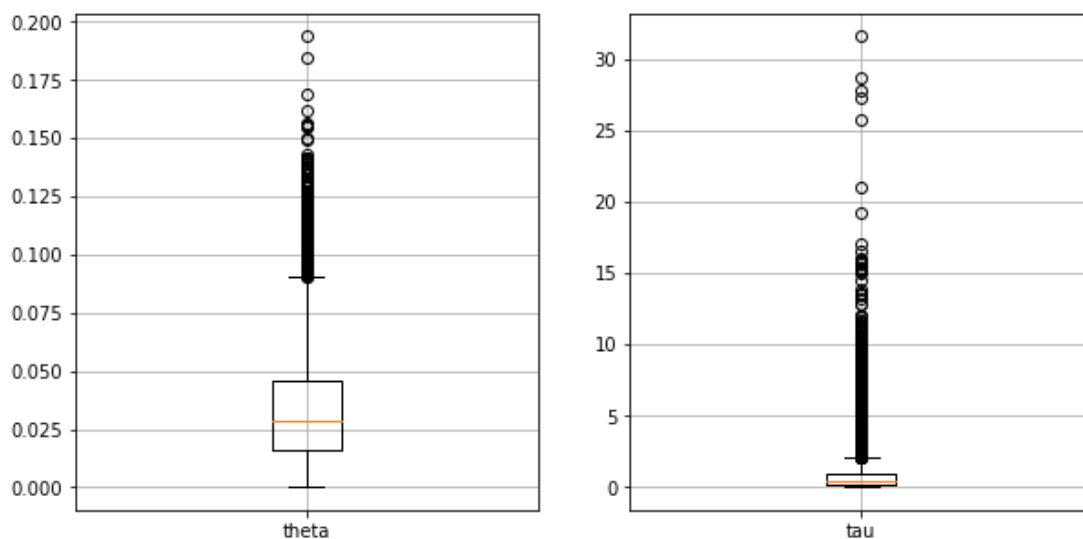


Using Burn-in and Run Length method, get  $R_{\text{theta}} = 1.000192$ ,  $R_{\text{tau}} = 1.000344$

可见，拟合得相当好。

d.

	marginal mean	std	95% interval LB	95% interval UB
theta	0.034032	0.023688	0.004239	0.094301
tau	0.916754	1.523507	0.052322	4.537479

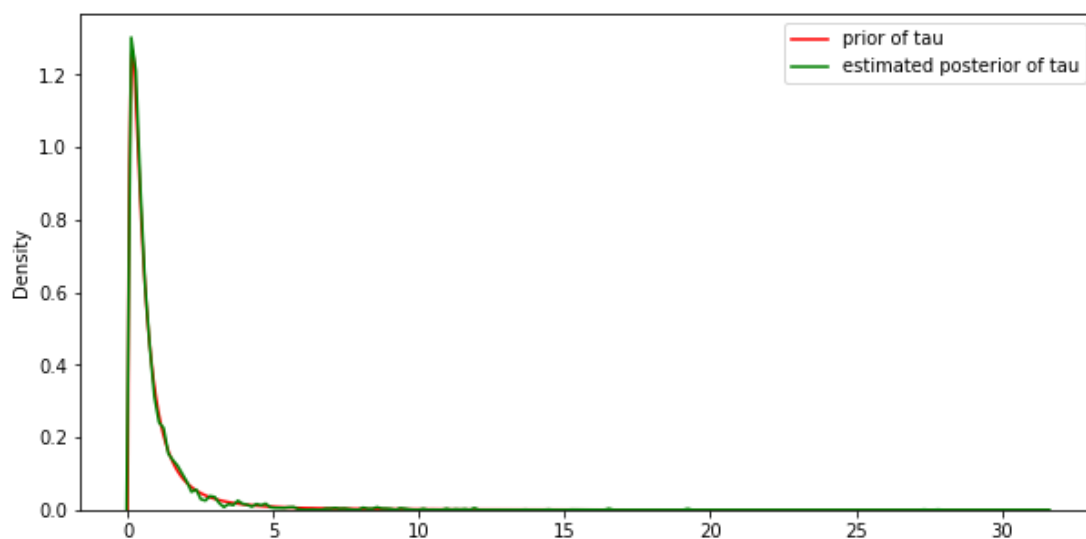


e.  $\tau$ 的先验分布为

$$\begin{aligned}
 f(\tau) & \\
 &\propto \int_0^{+\infty} \theta^a \tau^b e^{(-c\theta - d\theta\tau)} d\theta \\
 &= (d\tau + c)^{-a-1} \Gamma(a+1) \tau^b
 \end{aligned}$$

归一化后，得

$$f(\tau) = \frac{c^{a-b} d^{b+1} \Gamma(a+1)}{\Gamma(a-b) \Gamma(b+1)} (d\tau + c)^{-a-1} \tau^b$$



f.

$\tau$ 是一个比例系数， $\tau < 1$ 时才意味着该激素对病人有效。而根据吉布斯采样得到的均值和置信区间，我们倾向于认为该激素对病人无效。

g.

The statics of origin hyperparameters:

	marginal mean	std	95% interval LB	95% interval UB
theta	0.034032	0.023688	0.004239	0.094301
tau	0.916754	1.523507	0.052322	4.537479

The statics of origin hyperparameters divide by 2:

	marginal mean	std	95% interval LB	95% interval UB
theta	0.033225	0.033424	0.000729	0.122396
tau	5.244312	33.200553	0.046139	32.349776

The statics of origin hyperparameters times 2:

	marginal mean	std	95% interval LB	95% interval UB
theta	0.033540	0.016717	0.008956	0.073503
tau	0.496958	0.478717	0.067289	1.824140

可见， $\tau$ 对超参数非常敏感。建议先对超参数进行更准确的估计。

## 7.7

解：

a. 证明：

$$\begin{aligned}
 & \mu^{(t+1)} | (\alpha^{(t)}, \beta^{(t)}, y) \\
 &= E[y_{ij} - \alpha_i^{(t)} - \beta_{ij}^{(t)} - \epsilon_{ij}] \\
 &= \frac{1}{n} \sum_{i=1}^I \sum_{j=1}^{J_i} (y_{ij} - \alpha_i^{(t)} - \beta_{ij}^{(t)} - \epsilon_{ij}) \\
 &= \frac{1}{n} y_{ij} - \frac{1}{n} \sum_{i=1}^I J_i \alpha_i^{(t)} - \frac{1}{n} \sum_{i=1}^I \sum_{j=1}^{J_i} \beta_{ij}^{(t)} - \frac{1}{n} \sum_{i=1}^I \sum_{j=1}^{J_i} \epsilon_{ij}
 \end{aligned}$$

而 $y_{ij}, \alpha_i^{(t)}, \beta_{ij}^{(t)}$ 都是已知量，且 $\epsilon_{ij} \sim N(0, \sigma_\epsilon^2)$ ，故

$$\begin{aligned}
 & \mu^{(t+1)} | (\alpha^{(t)}, \beta^{(t)}, y) \sim N \left( y_{..} - \frac{1}{n} \sum_i J_i \alpha_i^{(t)} - \frac{1}{n} \sum_{j(i)} \beta_{j(i)}^{(t)}, \frac{\sigma_\epsilon^2}{n} \right) \\
 & P(\alpha_i^{(t+1)} | (\mu^{(t+1)}, \beta^{(t)}, y)) \\
 & \propto P(\alpha_i^{(t+1)}) \prod_{j=1}^{J_i} P(\epsilon_{ij} = y_i - \alpha_i^{(t+1)} - \mu^{(t+1)} - \beta_{ij}^{(t)}) \\
 & \propto e^{-\frac{1}{2} \frac{\alpha_i^{(t+1)^2}}{\sigma_\alpha^2}} \prod_{j=1}^{J_i} e^{-\frac{1}{2} \sum_{j=1}^{J_i} \frac{(y_i - \alpha_i^{(t+1)} - \mu^{(t+1)} - \beta_{ij}^{(t)})^2}{\sigma_\epsilon^2}}
 \end{aligned}$$

$$\begin{aligned}
& \propto e^{-\frac{1}{2} \left( \frac{\alpha_i^{(t+1)^2}}{\sigma_\alpha^2} + \sum_{j=1}^{J_i} \frac{\alpha_i^{(t+1)^2} - 2\alpha_i^{(t+1)} y_{ij} + 2\alpha_i^{(t+1)} \mu^{(t+1)} + 2\alpha_i^{(t+1)} \beta_{ij}^{(t)}}{\sigma_\epsilon^2} \right)} \\
& = e^{-\frac{1}{2} \left( \frac{\alpha_i^{(t+1)^2} \left( \frac{\sigma_\epsilon^2}{J_i} + \sigma_\alpha^2 \right) - \alpha_i^{(t+1)} \frac{1}{J_i} \sum_{j=1}^{J_i} (y_{ij} - \mu^{(t+1)} - \beta_{ij}^{(t)})}{\frac{\sigma_\alpha^2 \sigma_\epsilon^2}{J_i}} \right)} \\
& \propto \exp \left( -\frac{1}{2} \frac{\left( \alpha_i^{(t+1)} - \frac{1}{\left( \frac{\sigma_\epsilon^2}{J_i} + \sigma_\alpha^2 \right)} \frac{1}{J_i} \sum_{j=1}^{J_i} (y_{ij} - \mu^{(t+1)} - \beta_{ij}^{(t)}) \right)^2}{\frac{\sigma_\alpha^2 \sigma_\epsilon^2}{J_i} \cdot \frac{1}{\left( \frac{\sigma_\epsilon^2}{J_i} + \sigma_\alpha^2 \right)}} \right)
\end{aligned}$$

故,  $\alpha_i^{(t+1)} \sim N \left( \frac{J_i V_1}{\sigma_\epsilon^2} \left( y_{i \cdot} - \mu^{(t+1)} - \frac{1}{J_i} \beta_{j(i)}^{(t)} \right), V_1 \right)$ , 其中,  $V_1 = \left( \frac{J_i}{\sigma_\epsilon^2} + \frac{1}{\sigma_\alpha^2} \right)^{-1}$

同理,  $\beta_{j(i)}^{(t)} \sim N \left( \frac{V_2}{\sigma_\epsilon^2} \left( y_{ij} - \mu^{(t+1)} - \alpha_i^{(t+1)} \right), V_2 \right)$ , 其中  $V_2 = \left( \frac{1}{\sigma_\epsilon^2} + \frac{1}{\sigma_\beta^2} \right)^{-1}$

证毕.

b. 证明:

$$\begin{aligned}
& P \left( \mu^{(t+1)} | (\gamma^{(t)}, \eta^{(t)}, y) \right) \\
& = \prod_{i=1}^I P \left( \alpha_i = \gamma_i^{(t)} - \mu^{(t+1)} \right) \\
& \propto e^{-\frac{1}{2} \left( \sum_{i=1}^I \frac{(\gamma_i^{(t)} - \mu^{(t+1)})^2}{\sigma_\alpha^2} \right)} \\
& \propto e^{-\frac{1}{2} \left( \frac{I \mu^{(t+1)^2} - 2 \mu^{(t+1)} \sum_{i=1}^I \gamma_i^{(t)}}{\sigma_\alpha^2} \right)} \\
& \propto \exp \left( -\frac{1}{2} \frac{\left( \mu^{(t+1)} - \frac{1}{I} \sum_{i=1}^I \gamma_i^{(t)} \right)^2}{\frac{\sigma_\alpha^2}{I}} \right)
\end{aligned}$$

故  $\mu^{(t+1)} | (\gamma^{(t)}, \eta^{(t)}, y) \sim N \left( \frac{1}{I} \sum_{i=1}^I \gamma_i^{(t)}, \frac{\sigma_\alpha^2}{I} \right)$

同理,

$$\begin{aligned}
& P \left( \gamma_i^{(t+1)} | (\mu^{(t+1)}, \eta^{(t)}, y) \right) \\
& = P \left( \alpha_i = \gamma_i^{(t+1)} - \mu^{(t+1)} \right) \cdot \prod_{j=1}^{J_i} P \left( \beta_{ij} = \eta_{ij}^{(t)} - \gamma_i^{(t+1)} \right)
\end{aligned}$$

$$\begin{aligned}
& \propto e^{-\frac{1}{2} \frac{(\gamma_i^{(t)} - \mu^{(t+1)})^2}{\sigma_\alpha^2}} \cdot e^{-\frac{1}{2} \left( \sum_{j=1}^{J_i} \frac{(\eta_{ij}^{(t)} - \gamma_i^{(t+1)})^2}{\sigma_\beta^2} \right)} \\
& \propto e^{-\frac{1}{2} \left( \frac{\gamma_i^{(t)2} - 2\mu^{(t+1)}\gamma_i^{(t)} + \frac{1}{J_i} \sum_{j=1}^{J_i} \gamma_i^{(t+1)2}}{\sigma_\alpha^2} + \frac{\sigma_\beta^2}{J_i} - 2 \frac{\gamma_i^{(t+1)} \sum_{j=1}^{J_i} \eta_{ij}^{(t)}}{\sigma_\beta^2} \right)} \\
& = e^{-\frac{1}{2} \left( \left( \frac{J_i}{\sigma_\beta^2} + \frac{1}{\sigma_\alpha^2} \right) \gamma_i^{(t+1)2} - 2 \left( \frac{1}{\sigma_\beta^2} \sum_{j=1}^{J_i} \eta_{ij}^{(t)} + \frac{\mu^{(t+1)}}{\sigma_\alpha^2} \right) \gamma_i^{(t+1)} \right)} \\
& \propto \exp \left( -\frac{1}{2} \frac{\left( \gamma_i^{(t+1)} - \left( \frac{J_i}{\sigma_\beta^2} + \frac{1}{\sigma_\alpha^2} \right)^{-1} \left( \frac{1}{\sigma_\beta^2} \sum_{j=1}^{J_i} \eta_{ij}^{(t)} + \frac{\mu^{(t+1)}}{\sigma_\alpha^2} \right) \right)^2}{\left( \frac{J_i}{\sigma_\beta^2} + \frac{1}{\sigma_\alpha^2} \right)^{-1}} \right)
\end{aligned}$$

故  $\gamma_i^{(t+1)} | (\mu^{(t+1)}, \eta^{(t)}, y) \sim N \left( V_3 \left( \frac{1}{\sigma_\beta^2} \sum_{j=1}^{J_i} \eta_{ij}^{(t)} + \frac{\mu^{(t+1)}}{\sigma_\alpha^2} \right), V_3 \right)$ , 其中  $V_3 = \left( \frac{J_i}{\sigma_\beta^2} + \frac{1}{\sigma_\alpha^2} \right)^{-1}$

同理, 有  $\eta_{ij}^{(t+1)} | (\mu^{(t+1)}, \gamma^{(t+1)}, y) \sim N \left( V_2 \left( \frac{y_{ij}}{\sigma_\epsilon^2} + \frac{\gamma_i^{(t+1)}}{\sigma_\beta^2} \right), V_2 \right)$