

Deep Learning for Visual Recognition

Assignment 1: Machine Learning Basics

Bilge Ulusay
Ewald Bindereif
Ulvi Shukurzade
Cüneyt Erem

November 23, 2020

1 Theoretical Exercises

1.1 Bias of an estimator

$$\hat{\sigma}_m^2 = \frac{1}{m} \sum_{i=1}^m (x_i - \hat{\mu}_m)^2$$

$$\text{Bias}(\hat{\sigma}_m^2) = \mathbb{E}[\hat{\sigma}_m^2] - \sigma^2 = \mathbb{E}\left[\frac{1}{m} \sum_{i=1}^m (x_i - \hat{\mu}_m)^2\right] - \sigma^2 \stackrel{(1)}{=} \frac{(m-1) \cdot \sigma^2}{m} - \sigma^2 = \frac{-\sigma^2}{m} \stackrel{(2)}{\neq} 0$$

Step (1) holds due to the hint on the problem sheet and (2) since $\sigma > 0$. It follows that $\hat{\sigma}_m^2$ is a **biased** estimator. ✓

$$\tilde{\sigma}_m^2 = \frac{1}{m-1} \sum_{i=1}^m (x_i - \hat{\mu}_m)^2$$

$$\text{Bias}(\tilde{\sigma}_m^2) = \mathbb{E}[\tilde{\sigma}_m^2] - \sigma^2 = \mathbb{E}\left[\frac{1}{m-1} \sum_{i=1}^m (x_i - \hat{\mu}_m)^2\right] - \sigma^2 \stackrel{(1)}{=} \frac{(m-1) \cdot \sigma^2}{m-1} - \sigma^2 = 0$$

It follows that $\tilde{\sigma}_m^2$ is an **unbiased** estimator. ✓

Points 5/5

1.2 Bias Variance Trade-off

$$\text{Bias}(\hat{\theta})^2 = (\mathbb{E}[\hat{\theta}] - \theta)^2 = \mathbb{E}[\hat{\theta}]^2 - 2\mathbb{E}[\hat{\theta}]\theta + \theta^2$$

$$\text{Var}(\hat{\theta}) = \mathbb{E}[\hat{\theta}^2] - \mathbb{E}[\hat{\theta}]^2$$

$$\text{Bias}(\hat{\theta})^2 + \text{Var}(\hat{\theta}) = \mathbb{E}[\hat{\theta}^2] - 2\mathbb{E}[\hat{\theta}]\theta + \theta^2$$

$$\text{MSE}(\hat{\theta}) = \mathbb{E}[(\hat{\theta} - \theta)^2] = \mathbb{E}[\hat{\theta}^2 - 2\hat{\theta}\theta + \theta^2] = \mathbb{E}[\hat{\theta}^2] - 2\mathbb{E}[\hat{\theta}]\theta + \theta^2 = \text{Bias}(\hat{\theta})^2 + \text{Var}(\hat{\theta}) \quad \checkmark$$

A nice solution, fewer boring computation steps compared to mine. Points 5/5


1.3 MAP for a conditional likelihood

$$\begin{aligned}
 \theta_{\text{MAP}} &= \underset{\theta}{\operatorname{argmax}} p(\theta|X, Y) \\
 &\stackrel{(1)}{=} \underset{\theta}{\operatorname{argmax}} p(X, Y|\theta)p(\theta) \\
 &= \underset{\theta}{\operatorname{argmax}} p(Y|X, \theta)p(X|\theta)p(\theta) \\
 &\stackrel{(2)}{=} \underset{\theta}{\operatorname{argmax}} p(Y|X, \theta)p(X)p(\theta) \\
 &\stackrel{(3)}{=} \underset{\theta}{\operatorname{argmax}} p(Y|X, \theta)p(\theta) \\
 &\stackrel{(4)}{=} \underset{\theta}{\operatorname{argmax}} (\log p(Y|X, \theta) + \log p(\theta)) \\
 &= \underset{\theta}{\operatorname{argmax}} \left(\sum_{i=1}^m \log p(y^{(i)}|x^{(i)}, \theta) + \log p(\theta) \right)
 \end{aligned}$$

(1) holds by Bayes' rule: $p(\theta|X, Y) = \frac{p(X, Y|\theta)}{p(X, Y)}$; thus the denominator can be neglected for argmax concerning θ .

(2) holds as θ and X are independent.

(3) holds as $p(X)$ represents a constant and thus can be neglected for argmax concerning θ .

(4) holds after applying the logarithm to the product.  Points 5/5


1.4 Derivatives and the chain rule

We derive the following relationship between the sigmoid function σ and its derivative for $t \in \mathbb{R}$:

$$\begin{aligned}
 \sigma'(t) &= [(1 + e^{-t})^{-1}]' = -(1 + e^{-t})^{-2} \cdot (-e^{-t}) \\
 &= \frac{e^{-t} + 1 - 1}{(1 + e^{-t})^2} \\
 &= \frac{1 + e^{-t}}{(1 + e^{-t})^2} - \frac{1}{1 + e^{-t}} \cdot \frac{1}{1 + e^{-t}} \\
 &= \frac{1}{1 + e^{-t}} \cdot \left(1 - \frac{1}{1 + e^{-t}} \right) \\
 &= \sigma(t)(1 - \sigma(t)) \quad \text{✓}
 \end{aligned}$$

Now we calculate the partial derivative of $L(W, b)$ with respect to W_j :

$$\begin{aligned}
 \frac{\partial L(W, b)}{\partial W_j} &= \frac{\partial}{\partial W_j} \left(\sum_{i=1}^N (y_i - f_{W, b}(x_i))^2 \right) \\
 &= \sum_{i=1}^N \frac{\partial}{\partial W_j} (y_i - f_{W, b}(x_i))^2 \\
 &= \sum_{i=1}^N 2(y_i - f_{W, b}(x_i)) \frac{\partial (y_i - f_{W, b}(x_i))}{\partial W_j} \quad (\text{by chain rule}) \\
 &= -2 \sum_{i=1}^N (y_i - f_{W, b}(x_i)) \frac{\partial f_{W, b}(x_i)}{\partial W_j} \\
 &= -2 \sum_{i=1}^N (y_i - f_{W, b}(x_i)) \cdot f_{W, b}(x_i) \cdot (1 - f_{W, b}(x_i)) \cdot x_{ij}, \quad \text{✓}
 \end{aligned}$$

where x_{ij} is the j -th component of x_i . The last step follows from the derivation above, because $f_{W, b}$ is a sigmoid function, and the chain rule $(\frac{\partial}{\partial W_j} (\langle W, x_i \rangle + b) = \frac{\partial}{\partial W_j} (W_1 x_{i1} + \dots + W_d x_{id} + b) = x_{ij})$. 

Now we calculate the partial derivative of $L(W, b)$ with respect to b :

$$\begin{aligned}
\frac{\partial L(W, b)}{\partial b} &= \frac{\partial}{\partial b} \left(\sum_{i=1}^N (y_i - f_{W, b}(x_i))^2 \right) \\
&= \sum_{i=1}^N \frac{\partial}{\partial b} (y_i - f_{W, b}(x_i))^2 \\
&= \sum_{i=1}^N 2(y_i - f_{W, b}(x_i)) \frac{\partial (y_i - f_{W, b}(x_i))}{\partial b} \quad (\text{by chain rule}) \\
&= -2 \sum_{i=1}^N (y_i - f_{W, b}(x_i)) \frac{\partial f_{W, b}(x_i)}{\partial b} \\
&= -2 \sum_{i=1}^N (y_i - f_{W, b}(x_i)) \cdot f_{W, b}(x_i) \cdot (1 - f_{W, b}(x_i)), \quad \checkmark
\end{aligned}$$

The last step follows from the derivation above, because $f_{W, b}$ is a sigmoid function, and the chain rule ($\frac{\partial}{\partial b} (\langle W, x_i \rangle + b) = \frac{\partial}{\partial b} (W_1 x_{i1} + \dots + W_d x_{id} + b) = 1$).

The gradient of $L(W, b)$ with respect to the vector W is equal to the vector, which contains all the partial derivatives of $L(W, b)$ with respect to a single component of W :


$$\nabla_W L(W, b) = \left(\frac{\partial L(W, b)}{\partial W_1}, \dots, \frac{\partial L(W, b)}{\partial W_d} \right)^T$$

We can update W_j by the equation

$$W_j = W_j - \eta \frac{\partial L(W, b)}{\partial W_j}, \quad (1)$$

and b by

$$b = b - \eta \frac{\partial L(W, b)}{\partial b}, \quad (2)$$

where η denotes the learning rate.  Points 5/5

Theory - Points 20/20