

TTIC 31230, Fundamentals of Deep Learning

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Pseudo-Likelihood and Contrastive Divergence

Pseudolikelihood

For any distribution Q on assignments of labels to nodes (segmentations), and any assignment \hat{y} , we define $\tilde{Q}(\hat{y})$ as follows.

$$\tilde{Q}(\hat{y}) = \prod_i Q(\hat{y}[i] \mid \hat{y}/i) = \prod_i Q(\hat{y}[i] \mid \hat{y}[N(i)])$$

We then train a graphical model with pseudolikelyhood loss.

$$\Phi^* = \operatorname{argmin}_{\Phi} E_{y \sim \text{Pop}} - \ln \tilde{P}_{\Phi}(y)$$

Pseudolikelihood

$$\mathcal{L}_{\text{PL}} = -\ln \tilde{P}_s(y)$$

We note that by the Markov blanket property for Markov random fields we have

$$\tilde{P}_s(\hat{y}) = \prod_i P_s(\hat{y}[i] \mid \hat{y}[N(i)])$$

Since the loss is directly computed we can directly back-propagate on the loss.

Pseudolikelihood Theorem

$$\operatorname{argmin}_Q E_{y \sim \text{Pop}} - \ln \tilde{Q}(y) = \text{Pop}$$

or equivalently

$$\min_Q E_{y \sim \text{Pop}} - \ln \tilde{Q}(y) = E_{y \sim \text{Pop}} - \ln \widetilde{\text{Pop}}(y)$$

Proof I

We have

$$\min_Q E_{y \sim \text{Pop}} - \ln \tilde{Q}(y) \leq E_{y \sim \text{Pop}} - \ln \widetilde{\text{Pop}}(y)$$

So it suffices to show

$$\min_Q E_{y \sim \text{Pop}} - \ln \tilde{Q}(y) \geq E_{y \sim \text{Pop}} - \ln \widetilde{\text{Pop}}(y)$$

Proof II

We will prove the case of two nodes.

$$\begin{aligned} & \min_Q E_{y \sim \text{Pop}} - \ln Q(y[1]|y[2]) Q(y[2]|y[1]) \\ & \geq \min_{P_1, P_2} E_{y \sim \text{Pop}} - \ln P_1(y[1]|y[2]) P_2(y[2]|y[1]) \\ & = \min_{P_1} E_{y \sim \text{Pop}} - \ln P_1(y[1]|y[2]) + \min_{P_2} E_{y \sim \text{Pop}} - \ln P_2(y[2]|y[1]) \\ & = E_{y \sim \text{Pop}} - \ln \text{Pop}(y[1]|y[2]) + E_{y \sim \text{Pop}} - \ln \text{Pop}(y[2]|y[1]) \\ & = E_{y \sim \text{Pop}} - \ln \widetilde{\text{Pop}}(y) \end{aligned}$$

Contrastive Divergence (CDk)

In contrastive divergence we first construct an MCMC process whose stationary distribution is P_s . This could be Metropolis or Gibbs or something else.

Algorithm CDk: Given a gold segmentation y , start the MCMC process from initial state y and run the process for k steps to get \hat{y} . Then take the loss to be

$$\mathcal{L}_{\text{CD}} = s(\hat{y}) - s(y)$$

If $P_s = \text{Pop}$ then the the distribution on \hat{y} is the same as the distribution on y and the expected loss gradient is zero.

Gibbs CD1

CD1 for the Gibbs MCMC process is a particularly interesting special case.

Algorithm (Gibbs CD1): Given y , select a node i at random and draw $c \sim P(y[i] \mid y[N(i)])$. Define $y[i = c]$ to be the assignment (segmentation) which is the same as y except that node i is assigned label c . Take the loss to be

$$\mathcal{L}_{\text{CD}} = s(y[i = c]) - s(y)$$

Gibbs CD1 Theorem

Gibbs CD1 is equivalent in expectation to pseudolikelihood.

$$\begin{aligned}\mathcal{L}_{\text{PL}} &= E_{y \sim \text{Pop}} \sum_i -\ln P_s(y[i] = c \mid y \setminus i) \\ &= E_{y \sim \text{Pop}} \sum_i -\ln \frac{e^{s(y)}}{Z_i} \quad Z_i = \sum_{c'} e^{s(y[i=c'])} \\ &= E_{y \sim \text{Pop}} \sum_i (\ln Z_i - s(y)) \\ \nabla_{\Phi} \mathcal{L}_{\text{PL}} &= E_{y \sim \text{Pop}} \sum_i \left(\frac{1}{Z_i} \sum_{c'} e^{s(y[i=c'])} \nabla_{\Phi} s(y[i] = c') \right) - \nabla_{\Phi} s(y) \\ &= E_{y \sim \text{Pop}} \sum_i \left(\sum_{c'} P(y[i = c' \mid y \setminus i]) \nabla_{\Phi} s(y[i = c']) \right) - \nabla_{\Phi} s(y)\end{aligned}$$

Gibbs CD1 Theorem

$$\begin{aligned}
\nabla_{\Phi} \mathcal{L}_{\text{PL}} &= E_{y \sim \text{Pop}} \sum_i \left(\sum_{c'} P(y[i = c' \mid y \setminus i]) \nabla_{\Phi} s(y[i] = c') \right) - \nabla_{\Phi} s(y) \\
&= E_{y \sim \text{Pop}} \sum_i \left(E_{c' \sim P(y[i=c' \mid y \setminus i])} \nabla_{\Phi} s(y[i] = c') \right) - \nabla_{\Phi} s(y) \\
&\propto E_{y \sim \text{Pop}} E_i E_{c' \sim P(y[i=c' \mid y \setminus i])} \left(\nabla_{\Phi} s(y[i] = c') - \nabla_{\Phi} s(y) \right) \quad \text{Gibbs CD(1)}
\end{aligned}$$

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