Conditional Diffusion Models for Uncertainty Estimation in Super Resolution Microscopy

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Abstract

Deep learning has recently attracted considerable attention from researchers in the natural sciences, particularly microscopists, for fast extraction of physically relevant information from images. However, simple and interpretable uncertainty quantification is lacking in these applications, and remains a necessary modeling component in high-risk research. In order to quantify uncertainty in otherwise deterministic image translation architectures, we propose a hybrid generative modeling framework based on denoising diffusion probabilistic models (DDPMs). Specifically, our model combines a deterministic neural network with a DDPM, which can improve conditional synthesis speed and fidelity of the DDPM, while providing a natural mechanism for uncertainty estimation via Langevin dynamics. We apply our model to the task of single molecule localization in fluorescence microscopy, and demonstrate that blending the DeepSTORM architecture with a DDPM permits simultaneous high-fidelity super-resolution with uncertainty estimation of kernel density estimates (KDEs) regressed by DeepSTORM. Our results suggest the proposed solution is an interesting addition to the modeling toolkit for fluorescence microscopists and the field of deep image translation in general.

18 1 Introduction

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Deep learning has attracted tremendous attention from researchers in the natural sciences, with several foundational applications arising in microscopy, e.g., (Weigert 2018; Falk 2019). Recently, the application of deep image translation in single-molecule localization microscopy (SMLM) has received considerable interest (Ouyang 2018; Nehme 2020; Speiser 2021). SMLM techniques are a mainstay of fluorescence microscopy and can be used to produce a pointillist representation of biomolecules in the cell at diffraction-unlimited precision (Rust 2006; Betzig 2006). As this technology enables increasingly precise measurements of the cellular environment, there is an increasing need for machine learning methods to report uncertainty for quality control.

In previous applications of deep models to localization microscopy, super-resolution images can be 27 recovered from a sparse set of localizations with conditional generative adversarial networks (Ouyang 28 2018) or kernel density estimation can be performed using convolutional networks (Nehme 2020; Speiser 2021). Here, we focus on the latter class of models which perform single molecule localization 30 using neural networks. In this approach, one estimates molecular coordinates by predicting kernel 31 density estimates (KDEs) y, which are latent in the raw data x, using a convolutional neural network. Importantly, inferences in SMLM are often necessarily made on a single measurement, thus common measures of model performance are based on localization errors computed over ensembles of 34 simulated images. However, this choice precludes computation of aleatoric uncertainty at test time 35 under a fixed model, and may result in the application of models to out of distribution datasets.

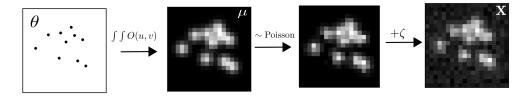


Figure 1: Generative model of single molecule localization microscopy images

Bayesian probability theory offers us mathematically grounded tools to reason about model uncertainty, but these usually come with a prohibitive computational cost (Gal 2022). A few approaches 38 to avoiding this intractibility in deep models have been deterministic uncertainty quantification 39 (Amersfoort 2020), ensembling (Lakshminarayanan et al., 2017) or Monte Carlo dropout (Gal and 40 Ghahramani, 2016). Here, we report a method which models estimates uncertainty in KDE predic-41 tions by combining deterministic deep learning with deep generative modeling in a hybrid algorithm. 42 Our approach produces pixel-wise uncertainties in model predictions with no modification to the 43 existing architecture, and can be used for downstream filtering of erroneous image regions. We choose to model a distribution on high-resolution KDE predictions conditioned on a low-resolution 45 input using a denoising diffusion probabilistic model (DDPM) (Ho 2020; Song 2021), referred to 46 here as simply "diffusion model". Such models are one class of score based generative models which 47 implicitly compute the score of the data distribution at each noise scale (Song 2021) and are well 48 suited conditional image generation tasks (Saharia 2021). Most importantly, score-based models 49 provide a natural mechanism for uncertainty quantification. Our approach could be readily integrated 50 with existing localization performance measures to address both model accuracy on training data and 51 precision on datasets produced by experiments.

2 Background

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2.1 Image Likelihood and Localization Error

The central objective of single molecule localization microscopy is to infer a set of molecular coordinates θ from measured low resolution images \mathbf{x} . The likelihood on a particular pixel k, i.e., $p(\mathbf{x}_k|\theta)$ is taken to be a convolution of Poisson and Gaussian distributions, due to shot noise $p(s_k) = \operatorname{Poisson}(\omega_k)$ and sensor readout noise $p(\zeta_k) = \mathcal{N}(o_k, \sigma_k^2)$

$$p(\mathbf{x}_k|\theta) = A \sum_{q=0}^{\infty} \frac{1}{q!} e^{-\omega_k} \omega_k^q \frac{1}{\sqrt{2\pi}\sigma_k} e^{-\frac{(\mathbf{x}_k - g_k q - o_k)}{2\sigma_k^2}} \approx \text{Poisson}(\omega_k')$$
 (1)

where A is some normalization constant and $\omega_k' = \omega_k + \sigma_k^2$. For the sake of generality, we include a per-pixel gain factor g_k , which is often unity. In practice, the summation in (1) can be difficult to work with, and it is commond to instead use a Poisson-Normal approximation for simplification, valid under a range of experimental conditions (Huang 2013). This result can be seen from the fact the the convolution of two Poisson distributions is also Poisson. The expectation of the Poisson process at each pixel of the image is computed from the optical transfer function O(u,v), which is often a two-dimensional isotropic Gaussian.

$$\omega = i_0 \int O(u) du \int O(v) dv \tag{2}$$

The above integration can be carried out by computing differences of error functions, as detailed in Appendix A. The complete generative process is depicted in Figure 1.
Reliable estimation of θ from x, for example by maximum likelihood estimation or with a deep model, requires performance matrices for model selection. We use the Fisher information as an

model, requires performance metrics for model selection. We use the Fisher information as an information theoretic criteria to assess the quality of the model tested here, with respect to the root mean squared error (RMSE) of our predictions of θ (Chao 2016). The Poisson log-likelihood

 $\ell(\mathbf{x}|\theta)$ is also convenient for computing the Fisher information matrix (Smith 2010) and thus the

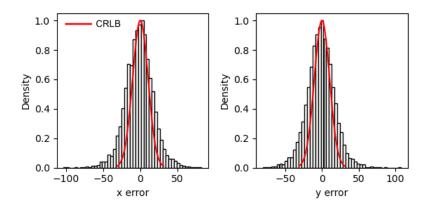


Figure 2: Localization errors of the trained model

Cramer-Rao lower bound, which bounds the variance of a statistical estimator of θ , from below i.e., $var(\hat{\theta}) \geq I^{-1}(\theta)$. The Fisher information is straightforward to compute under the Poisson log-likelihood, which is detailed in the Appendix

$$\mathcal{I}_{ij}(\theta) = \mathbb{E}\left(\frac{\partial \ell}{\partial \theta_i} \frac{\partial \ell}{\partial \theta_j}\right) = \sum_k \frac{1}{\omega_k'} \frac{\partial \omega_k'}{\partial \theta_i} \frac{\partial \omega_k'}{\partial \theta_j}$$
(3)

2.2 Kernel density estimation with deep networks

Direct optimization of the likelihood in (1) from observations \mathbf{x} alone is challenging when fluorescent emitters are dense within the field of view and fluorescent signals significantly overlap. However, convolutional neural networks (CNN) have recently proven to be powerful tools fluorescence microscopy to extract parameters describing fluorescent emitters such as color, emitter orientation, z-coordinate, and background signal (Zhang 2018; Kim 2019; Zelger 2018). For localization tasks, CNNs typically employ upsampling layers to reconstruct Bernoulli probabilities of emitter occupancy (Speiser 2021) or kernel density estimates with higher resolution than experimental measurements (Nehme 2020). Kernel density estimates, denoted by \mathbf{y} , are the most common data structure used in SMLM, and can be easily generated from molecular coordinates, alongside observations \mathbf{x} , using well-understood models of the optical impulse response (Zhang 2007).

87 3 Conditional Diffusion for Uncertainty-Aware Super Resolution

We consider datasets $(\mathbf{x}_i, \mathbf{y}_i, \hat{\mathbf{y}}_i)_{i=1}^N$ of observed images \mathbf{x}_i true kernel density estimate (KDE) images \mathbf{y}_i , and KDE estimates $\hat{\mathbf{y}}_i = \phi(\mathbf{x}_i)$. Observations \mathbf{x}_i are simulated under the Poisson likelihood (1) and KDEs are generated using (2) alone, followed by appropriate normalization.

91 3.1 Problem Statement

Point estimates \hat{y}_i produced by the traditional deep architectures for super resolution microscopy produce strong results, but lack uncertainty quantification. Recent advances in generative modeling, particularly DDPMs, therefore present a unique opportunity to integrate uncertainty awareness into the super-resolution microscopy toolkit. However, sampling from DDPMs is computationally expensive, given that generation amounts to solving a complex stochastic differential equation, effectively mapping a simple base distribution to the complex data distribution. The solution of such equations requires numerical integration with very small step sizes, resulting in thousands of neural network evaluations (Saharia 2021; Vahdat 2021). Furthermore, for conditional generation tasks in high-risk applications, generation complexity is further exacerbated by the need for the highest level of detail in generated samples.

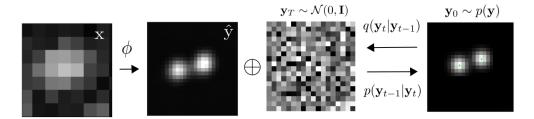


Figure 3: Conditional diffusion model for sampling kernel density estimates

Since we are primarily interested in the conditional distribution $p(\mathbf{y}|\mathbf{x})$, we propose that DDPM sampling is preceded by a deterministic transformation ϕ , trained to predict \mathbf{y} from \mathbf{x} . Reasoning for this choice in the current application is two-fold:

Synthesis Speed. By training a preprocessor ϕ to obtain an approximate estimate of y, we can reduce the number of iterations, since the DDPM only needs to model the remaining mismatch, resulting in a less complex model from which sampling becomes easier. Speed is critical in SMLM applications, which can produce large volumes of image data in a single experiment. Moreover, we note that this approach is analagous to preconditioned stochastic gradient langevin dynamics (Li 2016), wherein ϕ identifies the posterior mode followed by Langevin dynamics to sample from the posterior.

Sample Fidelity. Since Langevin dynamics will often be initialized in low-density regions of the data distribution, inaccurate score estimation in these regions will negatively affect the sampling process (Song 2019). Moreover, mixing can be difficult because of the need of traversing low density regions to transition between modes of the distribution. Preprocessing with a deterministic mapping ϕ can ameliorate this issue, by eliminating the need for score estimation in low density regions.

The preprocessor ϕ is realized by a CNN with upsampling layers. Consider the Markov chain wherein the KDE ${\bf y}$ is latent in and inferred from a noisy measurement ${\bf x}$, i.e., ${\bf x} \to \phi({\bf x}) \to \hat{{\bf y}}$. By the data processing inequality the function ϕ can only destroy information in ${\bf x}$ pertaining to ${\bf y}$ i.e., $I({\bf x};{\bf y}) \ge I(\phi({\bf x});{\bf y})$ or $h({\bf y}|\phi({\bf x})) \ge h({\bf y}|{\bf x})$ where I is the mutual information and h is the entropy. In other words, the function ϕ , while deterministic, can introduce additional uncertainty about ${\bf y}$ in downstream stochastic models by destroying information. Here, we are interested in measuring the upper bound $h({\bf y}|\phi({\bf x}))$, as this is the relevant quantity when a deterministic transformation ϕ is an unavoidable first step.

In practice, a DDPM Ψ can be trained on pairs $(\mathbf{y}_i, \hat{\mathbf{y}}_i)_{i=1}^N$. The conditional DDPM generates a target KDE \mathbf{y}_0 in T refinement steps. Starting with a pure noise image $\mathbf{y}_T \sim \mathcal{N}(0, \mathbf{I})$, the model iteratively refines the KDE through successive iterations according to learned conditional transition distributions $p(\mathbf{y}_{t-1}|\mathbf{y}_t)$ such that $\mathbf{y}_0 \sim p(\mathbf{y}|\hat{\mathbf{y}})$

3.2 Gaussian Diffusion

Diffusion models (Sohl-Dickstein 2015; Ho 2020) are a class of generative models inspired by nonequilibrium statistical physics, which slowly destroy structure in a data distribution $p(\mathbf{y}_0|\mathbf{x})$ via a fixed Markov chain referred to as the *forward process*. In the present context, we apply leverage recent results from (Ho 2020; Song 2021; Saharia 2021) for applying this framework to sampling from $p(\mathbf{y}|\mathbf{x},\hat{\mathbf{y}})$. We note that this approach is analagous to preconditioned stochastic gradient langevin dynamics (Li 2016), wherein ϕ identifies the posterior mode followed by sampling with Langevin dynamics. The forward process gradually adds Gaussian noise to the KDE \mathbf{y} according to a variance schedule $\beta_{0:T}$

$$q(\mathbf{y}_t|\mathbf{y}_0) = \prod_{t=1}^{T} q(\mathbf{y}_t|\mathbf{y}_{t-1}) \quad q(\mathbf{y}_t|\mathbf{y}_{t-1}) = \mathcal{N}\left(\sqrt{1-\beta_t}\mathbf{y}_{t-1}, \beta_t I\right)$$
(4)

The usual procedure is then to learn a parametric representation of the *reverse process*, and therefore generate samples from $p(\mathbf{y}_0)$, starting from noise. Formally, $p_{\theta}(\mathbf{y}_0|\hat{\mathbf{y}}) = \int p_{\theta}(\mathbf{y}_{0:T}|\hat{\mathbf{y}})d\hat{\mathbf{y}}_{1:T}$ where

140 \mathbf{y}_t is a latent representation with the same dimensionality of the data. $p_{\theta}(\mathbf{y}_{0:T}|\hat{\mathbf{y}})$ is a Markov process, starting from a noise sample $p_{\theta}(\mathbf{y}_T) = \mathcal{N}(0, \mathbf{I})$.

$$p_{\theta}(\mathbf{y}_{0:T}) = p_{\theta}(\mathbf{y}_T) \prod_{t=1}^{T} p_{\theta}(\mathbf{y}_{t-1}|\mathbf{y}_t) \quad p_{\theta}(\mathbf{y}_{t-1}|\mathbf{y}_t) = \mathcal{N}\left(s_{\theta}(\mathbf{y}_t), \beta_t I\right)$$
(5)

where we reuse the variance schedule of the forward process (Ho 2020). We omit conditioning on $\hat{\mathbf{y}}$ for each transition density $p_{\theta}(\mathbf{y}_{t-1}|\mathbf{y}_{t})$, as this is only considered at t=0 i.e., $p_{\theta}(\mathbf{y}_{1}|\mathbf{y}_{0},\hat{\mathbf{y}})$. An important property of the forward process is that it admits sampling \mathbf{y}_{t} at an arbitrary timestep t in closed form (Ho 2020). Using the notation $\alpha_{t}:=1-\beta_{t}$ and $\gamma_{t}:=\prod_{s=1}^{t}\alpha_{s}$, we have $q(\mathbf{y}_{t}|\mathbf{y}_{0})=\mathcal{N}\left(\sqrt{\gamma_{t}}\mathbf{y}_{0},(1-\gamma_{t})I\right)$. Training is performed by optimizing the usual variational bound on negative log likelihood:

$$\mathcal{L}(\theta) = \mathbb{E}\left[-\log p_{\theta}(\mathbf{y}_{0}|\mathbf{x})\right] \le \mathbb{E}\left[-\log \frac{p_{\theta}(\mathbf{y}_{0:T}|\mathbf{x})}{q(\mathbf{y}_{1:T}|\mathbf{y}_{0})}\right]$$
(6)

The objective in (6) can be expanded in terms of $D_{\mathrm{KL}}(p(\mathbf{y}_{t-1}|\mathbf{y}_t)||q(\mathbf{y}_t|\mathbf{y}_{t-1})$ as detailed in (Ho 2020). We choose to adopt the simplified from of the variational bound, which emphasizes that the DDPM estimates the score $\nabla_{\mathbf{y}} \log p(\mathbf{y}|\mathbf{x})$ at each noise level (Song 2021)

$$\theta^* = \underset{\theta}{\operatorname{argmin}} \mathbb{E} \underset{(\hat{\mathbf{y}}, \mathbf{y}_0)(\epsilon, \gamma)}{\mathbb{E}} \left[s_{\theta} \left(x, \sqrt{\gamma} \mathbf{y}_0 + \sqrt{1 - \gamma} \epsilon \, \middle| \, \mathbf{y}_t, \gamma \right) - \epsilon \right], \tag{7}$$

After training, samples can be generated by

$$\mathbf{y}_{t-1} = \frac{1}{\sqrt{1-\beta_i}} \left(\mathbf{y}_i + \beta_i s_{\theta}(\mathbf{y}_t) \right) + \sqrt{\beta_i} \xi \tag{8}$$

The distribution of γ has a big impact on the quality of the model and the generated outputs. For our training noise schedule, we use a piecewise distribution for γ , $p(\gamma) = \frac{1}{T} \sum_{t=1}^{T} U(\gamma_{t-1}, \gamma_t)$ (Nanxin 2021). Specifically, during training, we first uniformly sample a time step $t \sim \{0, ..., T\}$ followed by sampling $\gamma \sim U(\gamma_{t-1}, \gamma_t)$. We set T=100 in all our experiments.

3.3 Optimization of the DeepSTORM architecture

A first pass at localization treats localization as a binary classification problem, such that 0 denotes 157 a vacant pixel and 1 denotes an occupied pixel containing an emitter. Direct learning of pixel-wise 158 classification with cross-entropy loss leads to an imbalance of occupied and unoccupied pixels in 159 dense localization problems (Nehme 2020). CE loss is usually either weighted [51], replaced with a 160 Focal loss [52], or applied to a "blobbed" version of the desired boolean volume e.g. by placing a disk 161 around each GT position [53–55]. Alternative methods take a soft version of the binary classification problem. That is, by placing a small Gaussian around each GT position (e.g. with std of 1 pixel), 164 and matching continuous heatmaps, backpropagation yields more meaningful gradients and eases the learning process convergence. 165

Localization heatmaps thus form a natural encoding for SMLM images, which can be input to our conditional diffusion model. Therefore, to encode raw data x into a more tractable representation, we train the DeepSTORM architecture (Nehme 2020). Raw coordinates θ are binned into an upsampled image z.

4 Experiments

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All training data was simulated under the likelihood and impulse reponse (2,10), drawing coordinates uniformly over a disc with a radius of 7 pixels.

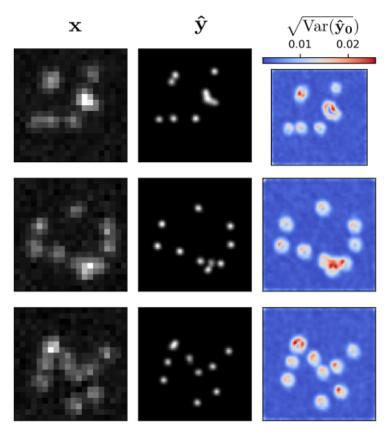


Figure 4: Kernel density estimates for various signal to noise ratios (SNR)

4.1 Model Precision on Simulated Ensembles

4.2 Model Uncertainty

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- We set T=100 for all experiments and treat forward process variances β_t as hyperparameters, with a linear schedule from $\beta_0=10^{-4}$ to $\beta_T=10^{-2}$. These constants were chosen to be small relative to data scaled to [-1,1], ensuring that reverse and forward processes have approximately
- the same functional form while keeping the signal-to-noise ratio at x_T as small as possible ($L_T = \frac{1}{2} \frac{1}{2$
- 179 $D_{KL}(q(x_T|x_0)||\mathcal{N}(0,I)) \approx 10^{-5}$ bits per dimension in our experiments).
- To represent the reverse process, we used the DDPM architecture based on a U-Net backbone (Ho
- 181 2020). Parameters are shared across time, which is specified to the network using the Transformer
- sinusoidal position embedding?. We use self-attention at the 16×16 feature map resolution??.
- Details are in Appendix A.
- and the channel multipliers at different resolutions (see Appendix A for details). To condition the
- model on the input x, we up-sample the low-resolution image to the target resolution using bicubic
- interpolation. The result is concatenated with y_t along the channel dimension. We experimented with
- more sophisticated methods of conditioning, such as using, but we found that the simple concatenation
- yielded similar generation quality.

189 5 Related Work

5.1 Diffusion Models

- 191 Prior work of diffusion models ?? require 1-2k diffusion steps during inference, making generation
- slow for large target resolution tasks. We adapt techniques from ? to enable more efficient inference.
- Our model conditions on γ directly (vs t as in ?), which allows us flexibility in choosing the number

- of diffusion steps, and the noise schedule during inference. This has been demonstrated to work
- well for speech synthesis?, but has not been explored for images. For efficient inference, we set the
- maximum inference budget to 100 diffusion steps, and hyper-parameter search over the inference
- noise schedule. This search is inexpensive as we only need to train the model once?. We use FID on
- held-out data to choose the best noise schedule, as we found PSNR did not correlate well with image
- 199 quality.

200 6 Conclusion

201 References

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A Appendix

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Standard SMLM localization algorithms based on maximum likelihood estimators or least squares optimization require tight control of activation and reactivation to maintain sparse emitters, presenting a tradeoff between imaging speed and labeling density. Recently, deep models have generalized SMLM to densely labeled structures by predicting high-resolution kernel density estimates (KDEs) from low resolution images with convolutional networks. However, estimated KDEs may contain irregularities due to finite sample sizes and limited model capacity.

The DeepSTORM CNN, initially proposed in (Nehme 2020) for 3D localization, can be viewed 247 as a deep kernel density estimator, reconstructing kernel density estimates y from low-resolution 248 inputs x. We utilize a simplified form of the original architecture for 2D localization, which we 249 denote ϕ hereafter, which consists of three main modules: a multi-scale context aggregation module, 250 an upsampling module, and a prediction module. For context aggregation, the architecture utilizes 251 dilated convolutions to increase the receptive field of each layer. The upsampling module is then 252 composed of two consecutive 2x resize-convolutions, computed by nearest-neighbor interpolation, 253 to increase the lateral resolution by a factor of 4. For a common sCMOS camera, each pixel has a 254 255 lateral size of approximately 108 nanometers, giving approximately 27 nanometer pixels in the KDE. The terminal prediction module contains three additional convolutional blocks for refinement of the 256 upsampled image, followed by an element-wise HardTanh. 257

Single molecule localization microscopy (SMLM) relies on the temporal resolution of fluorophores whose spatially overlapping point spread functions would otherwise render them unresolvable at the detector. Common strategies for the temporal separation of molecules involve molecular photoswitching from dark to fluorescent states, permitting resolution of fluorophores beyond the diffraction limit. Estimation of molecular coordinates is typically carried out by modeling the optical impulse response of the imaging system and fitting model functions to the data. However, such models are only well-suited to isolated molecules, reducing the number of molecules in the field of view and limiting temporal resolution in super resolution microscopy. This issue has incited a series of efforts to increase the density of fluorescent molecules imaged in a single frame while developing appropriate models for dense localization.

In fluorescence microscopy, each pixel is treated as a Poisson random variable (Smith 2010; Nehme 2020; Chao 2016), with expected value

$$\omega = i_0 \int O(u) du \int O(v) dv \tag{9}$$

where $i_0 = \eta N_0 \Delta$. The scalar parameters η , Δ are the photon detection probability of the sensor and the exposure time, respectively. Without loss of generality, we assume $\eta = \Delta = 1$. Most importantly, N_0 represents the signal amplitude, which we assume maintains a fixed value. The optical impulse response O(u,v) is often approximated as a 2D isotropic Gaussian with standard deviation σ (Zhang 2007). This approximation has the convenient property, that the effects of pixelation can be expressed in terms of error functions. For example, given a fluorescent emitter located at $\theta = (u_0, v_0)$, we have that

$$\int O(u)du = \frac{1}{2} \left(\operatorname{erf}\left(\frac{u_k + \frac{1}{2} - u_0}{\sqrt{2}\sigma}\right) - \operatorname{erf}\left(\frac{u_k - \frac{1}{2} - u_0}{\sqrt{2}\sigma}\right) \right)$$
 (10)

where we have used the common definition $\operatorname{erf}(z) = \frac{2}{\sqrt{\pi}} \int_0^t e^{-t^2} dt$. Our generative model also incorporates a normally distributed white noise per pixel ζ with offset o and variance σ^2 . Ultimately, we have a Poisson component of the signal, which scales with N_0 and a Gaussian component, which does not.

281 Consider,

$$\zeta_k - o_k + \sigma_k^2 \sim \mathcal{N}(\sigma_k^2, \sigma_k^2) \approx \text{Poisson}(\sigma_k^2)$$
 (11)

Since $\mathbf{x}_k = \mathbf{s}_k + \zeta_k$, we transform $\mathbf{x}_k' = \mathbf{x}_k - o_k + \sigma_k^2$, which is distributed according to

Consider the factorization $p(\hat{\mathbf{y}}|\mathbf{x},\mathbf{y})p(\mathbf{x}|\mathbf{y})p(\mathbf{y}) = p(\mathbf{x}|\mathbf{y},\hat{\mathbf{y}})p(\mathbf{y}|\hat{\mathbf{y}})p(\hat{\mathbf{y}})$. Given that \mathbf{x} is conditionally independent of $\hat{\mathbf{y}}$, we find

$$p_{\Psi}(\hat{\mathbf{y}}|\mathbf{x},\mathbf{y}) = p(\mathbf{y}|\hat{\mathbf{y}})$$