

weak stationary: E. Var constant.  $\exists \rho > 0$  + ACF is constant.

strict stationary: iid

$$\text{ACF: } |r_k| \leq r_0 \quad |p_k| \leq 1$$

$$r_k = r_{-k} \quad p_k = p_{-k}$$

$p_k=0 \Rightarrow$  white noise  $\Rightarrow$  can't predict

$r_k \neq 0 \Rightarrow$  auto-correlation  $\Rightarrow$  can used to predict  
dependence

ACF  $\neq 0$  不意味着 market inefficiency

conditional part is more important.

$$\text{ar: } E[\alpha_t | F_{t-1}] = 0 \Rightarrow E[\alpha_t] = 0 \Rightarrow r_k = 0.$$

$$\begin{aligned} &= E[E[\alpha_t | F_{t-1}]] \\ &= E[E[E[\alpha_t | \alpha_{t-i} | F_{t-1}]]] \\ &= 0 \end{aligned}$$

$$\sigma_t^2 = \text{Var}(r_t | F_{t-1})$$

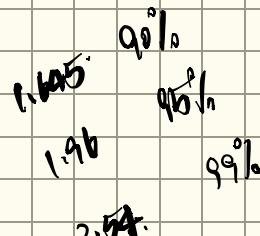
$$= E[(r_t - \mu_t)^2 | F_{t-1}] = E[\alpha_t^2 | F_{t-1}]$$

$$\begin{aligned} E[\varepsilon_t^2] &= E\left(\frac{\alpha_t^2}{\sigma_t^2}\right) = E\left[IE\left(\frac{\alpha_t^2}{\sigma_t^2} | F_{t-1}\right)\right] \\ &= IE\left[\frac{1}{\sigma_t^2} E[\alpha_t^2 | F_{t-1}]\right] \quad \text{在 } E[\alpha_t^2 | F_{t-1}] \text{ 是常数. } \sigma_t^2 \text{ is a constant} \\ &= IE\left[\frac{1}{\sigma_t^2} \sigma_t^2\right] = 1 \end{aligned}$$

$$\text{AR(1)} \quad r_t = \phi_1 r_{t-1} + \alpha_t = \alpha_t + \sum_{i=1}^{\infty} \phi_i^i \alpha_{t-i} \quad \text{ar iid } \sim N(0, \sigma^2)$$

how to know this model is correct?

$$\begin{aligned} E[r_t] &= \alpha \quad \text{Var}(r_t) = E[(r_t - \alpha)^2] \\ &= E\left[\left(\sum_{i=0}^{\infty} \phi_i^i \alpha_{t-i}\right)^2\right] \\ &= E\left[\sum_{i=0}^{\infty} (\phi_i^i \alpha_{t-i})^2 + \sum_{i \neq j} \phi_i^i \alpha_{t-i} \phi_j^j \alpha_{t-j}\right] \\ &= \sigma^2 \sum_{i=0}^{\infty} (1 + \phi_i^2 + \dots) \\ &= \sigma^2 \frac{1}{1 - \phi^2} = \frac{\sigma^2}{1 - \phi^2} \end{aligned}$$



Stationary:  $|\phi_1| < 1$

$$\begin{aligned} \text{ACF: } r_0 &= E[r_t r_{t-1}] = \frac{\sigma^2}{1 - \phi^2} \quad P_0 = 1 \\ r_1 &= E[r_t r_{t-1}] = E[(\alpha_t + \phi_1 \alpha_{t-1} + \phi_2 \alpha_{t-2} + \dots)] \\ &\quad (\alpha_{t-1} + \phi_1 \alpha_{t-2} + \phi_2 \alpha_{t-3} + \dots) \quad P_1 = \phi_1 \\ &= \phi_1 (\phi_0 + \phi_1^2 + \dots) \\ &= \frac{\phi_1 \sigma^2}{1 - \phi_1^2} \quad : \\ r_k &= \frac{\phi_1^k \sigma^2}{1 - \phi_1^2} \quad P_k = \phi_1^k \end{aligned}$$

mean reverting behavior

$$r_t = \phi_0 + \phi_1 r_{t-1} + \alpha_t \quad E[r_t] = [E[r_t | F_{t-1}]]$$

$$\begin{aligned} &= E[\phi_0 + \phi_1 r_{t-1}] \\ &= \phi_0 + \phi_1 E[r_{t-1}] \quad : E = \frac{\phi_0}{1 - \phi_1} \\ &\Leftarrow E[r_t] = E[r_{t-1}] \text{ stationary} \end{aligned}$$

$\text{Var}(r_t)$ :  $\text{Var}(\phi_1 r_{t-1} + \alpha_t)$  不变

$$r_t = \phi_0 + \phi_1 (\phi_0 + \phi_1 r_{t-2} + \alpha_{t-1}) + \alpha_t$$

ACF: 不变

$\Leftarrow$  成比例.

$$\begin{aligned} &= \phi_0 (1 + \phi_1) + \phi_1^2 r_{t-2} + \alpha_t + \phi_1 \alpha_{t-1} \\ &= \phi_0 (1 + \phi_1) + \phi_1^2 (\phi_0 + \phi_1 r_{t-3} + \alpha_{t-2}) + \alpha_t + \phi_1 \alpha_{t-1} \\ &= \phi_0 (1 + \phi_1 + \phi_1^2) + \phi_1^3 r_{t-3} + \alpha_t + \phi_1 \alpha_{t-2} + \phi_1^2 \alpha_{t-1} \\ &= \phi_0 \frac{1}{1 - \phi_1^2} + \sum_{i=0}^{\infty} \phi_1^i \alpha_{t-i} \end{aligned}$$

$$Y_{n+1} = \hat{Y}_{n+1} = E[Y_{n+1} | F_n] = \phi_0 + \phi_1 Y_n$$

$$\epsilon_{n+1} = Y_{n+1} - \hat{Y}_{n+1} = \phi_0 + \phi_1 Y_n + \alpha_{n+1} - \phi_0 - \phi_1 Y_n = \alpha_{n+1}$$

$$\therefore \text{Var}(\epsilon_{n+1}) = 6\sigma^2$$

$$\begin{aligned}\hat{r}_n(2) &= E[Y_{n+2} | F_n] = E[\phi_0 + \phi_1 Y_n + \alpha_{n+2} | F_n] \\ &= \phi_0 + \phi_1 E[Y_n | F_n] \\ &= \phi_0 + \phi_1 (\phi_0 + \phi_1 Y_n) = \phi_0 + \phi_1 \phi_0 + \phi_1^2 Y_n\end{aligned}$$

$$\begin{aligned}\epsilon_{n+2} &= \phi_0 + \phi_1 Y_{n+1} + \alpha_{n+2} - \hat{r}_n(2) \\ &= \underline{\phi_0 + \phi_1 (\phi_0 + \phi_1 Y_n + \alpha_{n+1})} + \alpha_{n+2} - \underline{\phi_0 + \phi_1 \phi_0 + \phi_1^2 Y_n} \\ &= \phi_1 \alpha_{n+1} + \alpha_{n+2}\end{aligned}$$

$$\text{Var}(\epsilon_{n+2}) = (\phi_1^2 + 1)6\sigma^2$$

$$\begin{aligned}Y_n(l) &= E[Y_{n+l} | F_n] = E\left[\frac{\phi_0}{1-\phi_1} + \alpha_{n+1} + \phi_1 \alpha_{n+2} + \phi_1^2 \alpha_{n+3} + \dots | F_n\right] \\ &= \frac{\phi_0}{1-\phi_1} + \phi_1^l \alpha_n + \phi_1^{l+1} \alpha_{n+1} + \dots\end{aligned}$$

$$\therefore \epsilon_n(l) = Y_{n+l} - \hat{Y}_n(l) = \alpha_{n+1} + \phi_1 \alpha_{n+2} + \dots + \phi_1^{l-1} \alpha_{n+l-(l-1)}$$

$$= \sum_{j=0}^{l-1} \phi_1^j \alpha_{n+l-j}$$

$$\therefore \text{Var}(\epsilon_n(l)) = \left[1 + (\phi_1^2 + (\phi_1^2)^2 + \dots + (\phi_1^{l-1})^2)\right] 6\sigma^2 = 6\sigma^2 \sum_{j=0}^{l-1} \phi_1^{2j}$$

forecast interval.  $[Y_n(l) - N \frac{\sigma}{\sqrt{6}} \sqrt{\sum_{j=0}^{l-1} \phi_1^{2j}}, Y_n(l) + N \frac{\sigma}{\sqrt{6}} \sqrt{\sum_{j=0}^{l-1} \phi_1^{2j}}]$

**AR(2)** :  $Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \alpha_t$

Stationary:

$$(1-\alpha_1 B)(1-\alpha_2 B)Y_t = \alpha_t.$$

$$U_t = (1-\alpha_2 B)Y_t = \alpha_t + \sum_{i=1}^{\infty} \alpha_2^i \alpha_{t-i}$$

$$Y_t = \alpha_2 Y_{t-1} + U_t = U_t + \sum_{i=1}^{\infty} \alpha_2^i U_{t-i}$$

$$\text{B.R.E. } |\alpha_1| < 1, |\alpha_2| < 1$$

$$|\alpha_1| = |\alpha_2| > 1, |\alpha_2| = \frac{1}{|\alpha_2|} > 1 \Rightarrow z_1, z_2 \text{ outside circle}$$

ACF:  $Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \alpha_t$

$$r_1 = E[Y_t Y_{t-1}]$$

$$= E[(\phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \alpha_t) Y_{t-1}]$$

$$r_1 = \phi_1 r_0 + \phi_2 r_1 \quad r_1 = \frac{\phi_1}{1-\phi_2} r_0$$

$$r_2 = E[Y_t Y_{t-2}]$$

$$= E[(\phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \alpha_t) Y_{t-2}]$$

$$= \phi_1 r_1 + \phi_2 r_0$$

$$r_3 = E[Y_t Y_{t-3}]$$

$$= E[(\phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \alpha_t) Y_{t-3}]$$

$$= \phi_1 r_2 + \phi_2 r_1$$

$$\rho_0 = 1$$

$$\rho_1 = \frac{\phi_1}{1-\phi_2}$$

$$\rho_2 = \phi_1 \rho_1 + \phi_2 = \phi_1 \rho_1 + \phi_2 \rho_0$$

$$\rho_k = \phi_1 \rho_{k-1} + \phi_2 \rho_{k-2}$$

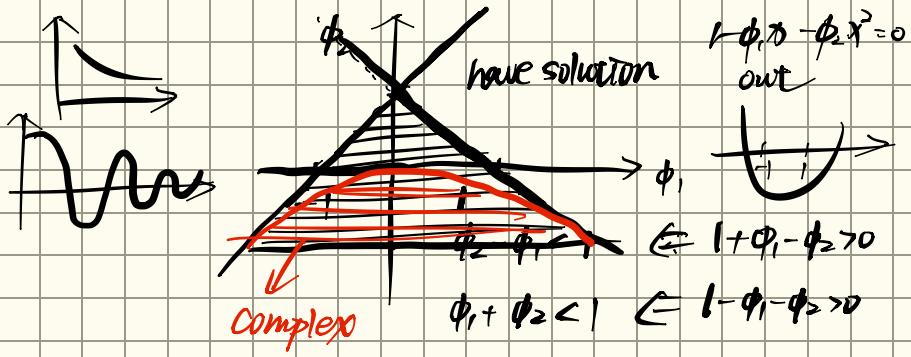
$$W_{1,2} = 1 - \phi_1 z - \phi_2 z^2 = 0$$

$$\rho_k = \frac{C_1}{w_1^k} + \frac{C_2}{w_2^k}$$

$$r_3 = \phi_1 r_2 + \phi_2 r_1 \Rightarrow r_k = \phi_1 r_{k-1} + \phi_2 r_{k-2}$$

If  $w_{1,2}$  all real

If  $w_{1,2}$  complex



Stochastic business cycle: 有 2 个周期

$$k = \frac{2\pi}{\cos(\phi_1/2\sqrt{\phi_2})}$$

$$k = \frac{2\pi}{\cos^2(a/\sqrt{a^2+b^2})}$$

$$\begin{aligned} x_1 + x_2 &= -\frac{\phi_1}{1} = -\phi_1 \\ x_1 x_2 &= \frac{-\phi_2}{1} = -\phi_2 \\ \frac{\phi_1}{2\sqrt{\phi_2}} &= \frac{-(x_1 + x_2)}{2\sqrt{x_1 x_2}} \\ &= \frac{2a}{2\sqrt{a^2+b^2}} = \frac{a}{\sqrt{a^2+b^2}} \end{aligned}$$

$$x_1, x_2 = a \pm bi$$

$$\text{forecast: } r_t = \phi_0 + \phi_1 r_{t-1} + \phi_2 r_{t-2} + a_t$$

$$\begin{aligned} \hat{r}_n(1) &= E[r_{n+1} | r_n] = E[\phi_0 + \phi_1 r_n + \phi_2 r_{n-1} + a_{n+1} | r_n] \\ &= \phi_0 + \phi_1 r_n + \phi_2 \hat{r}_{n-1} \end{aligned}$$

$$\text{Var}(\hat{r}_n(1)) = \sigma_a^2$$

**AR model**  $r_t = \phi_0 + \phi_1 r_{t-1} + \dots + \phi_p r_{t-p} + a_t$ . AR(p).

$$\text{PACF.} = \text{corr}(r_t, r_{t+k} | r_m, r_{m+1}, \dots, r_{m+k-1})$$

$$= \frac{\text{cov}(r_t - \hat{r}_t)(r_{t+k} - \hat{r}_{t+k})}{\sqrt{\text{Var}(r_t - \hat{r}_t) \text{Var}(r_{t+k} - \hat{r}_{t+k})}} = \phi_{kk}$$

有  $\hat{r}_t$ ,  $\hat{r}_{t+k}$   $\Rightarrow$  can't calculate.

$$\phi_{11} = p_1$$

$$\phi_{kk} = \frac{\begin{vmatrix} 1 & p_1 & p_2 & \dots & p_{k-2} & p_1 \\ p_1 & 1 & p_1 & \dots & p_{k-3} & p_2 \\ \vdots & \vdots & \vdots & \ddots & \vdots & \vdots \\ p_{k-1} & p_{k-2} & p_{k-3} & \dots & p_1 & p_k \end{vmatrix}}{\begin{vmatrix} 1 & p_1 & p_2 & \dots & p_{k-2} & p_{k-1} \\ p_1 & 1 & p_1 & \dots & p_{k-3} & p_{k-2} \\ p_2 & p_1 & 1 & \dots & p_{k-4} & p_{k-3} \\ \vdots & \vdots & \vdots & \ddots & \vdots & \vdots \\ p_{k-2} & p_{k-3} & p_{k-4} & \dots & 1 & \dots & p_1 \\ p_{k-1} & p_{k-2} & p_{k-3} & \dots & p_1 & 1 \end{vmatrix}}$$

$$\phi_{kk} =$$

$$\begin{vmatrix} 1 & p_1 & p_2 & \dots & p_{k-2} & p_{k-1} \\ p_1 & 1 & p_1 & \dots & p_{k-3} & p_{k-2} \\ p_2 & p_1 & 1 & \dots & p_{k-4} & p_{k-3} \\ \vdots & \vdots & \vdots & \ddots & \vdots & \vdots \\ p_{k-2} & p_{k-3} & p_{k-4} & \dots & 1 & \dots & p_1 \\ p_{k-1} & p_{k-2} & p_{k-3} & \dots & p_1 & 1 \end{vmatrix}$$

feature. AR(p).  $\phi_{11}, \phi_{22}, \dots, \phi_{pp} \neq 0$ . else = 0  $\phi_{kk} \sim N(0, \frac{1}{n})$

$\Rightarrow$  可用 Box test.

estimate parameters: least square or maximum likelihood.

$$\bar{J}_n(\hat{\phi} - \phi_0) \sim N(0, \hat{\sigma}^2)$$

$$\bar{J}_n(\hat{\phi}_i - \phi_{0i}) \sim N(0, \hat{\sigma}_{ii}^2) \quad \hat{\sigma}^2 = \left( \sum_{t=1}^n \tilde{z}_{t-1} \tilde{z}'_{t-1} \right)^{-1} \hat{\alpha}^2$$

$$\hat{\alpha}^2 = \frac{1}{n} \sum_{t=1}^n a_t^2$$

檢查 residual 是否是 white noise  $\Rightarrow$  acf of residual  $\Rightarrow$  Box test

是否有截距  $\frac{\bar{r}}{\sqrt{\text{Var}(r_t)}} \sim N(0, 1)$ . if  $E[r_t] = 0$

intercept  $\neq E[r_t]$  不見  $\phi_0$ . T<sub>t</sub> if  $E[r_t] \neq 0 \Leftrightarrow \phi_0 = 0$ .

MA(1)  $r_t = \mu + a_t - \theta a_{t-1}$

$$E[r_t] = \mu$$

$$\text{Var}(r_t) = (1 + \theta^2) \hat{\alpha}^2$$

$$\rho_0 = \text{Var}(r_t)$$

$$\begin{aligned} \rho_1 &= E[r_t r_{t-1}] = E[(\mu + a_t - \theta a_{t-1})(\mu + a_{t-1} - \theta a_{t-2})] \\ &= -\theta \hat{\alpha}^2 \end{aligned}$$

$$\rho_2 = E[r_t r_{t-2}] = E[(a_t - \theta a_{t-1})(a_{t-2} - \theta a_{t-3})] = 0$$

$$\rho_0 = 1$$

$$\rho_1 = -\frac{\theta}{1 + \theta^2}$$

$$\rho_k = 0, k \geq 2$$

forecast:  $\hat{r}_{n+1} = E[r_{n+1} | r_n] = E[\mu + a_{n+1} - \theta a_n | r_n] = \mu - \theta a_n$

$$\hat{e}_{n+1} = a_{n+1}$$

$$\text{Var} = \hat{\alpha}^2$$

$$\hat{r}_{n+2} = E[r_{n+2} | r_n] = E[\mu + a_{n+2} - \theta a_{n+1} | r_n] = \mu$$

$$\hat{e}_{n+2} = a_{n+2} - \theta a_{n+1}$$

$$\text{Var} = (1 + \theta^2) \hat{\alpha}^2$$

$$\hat{r}_{n+1} = \mu$$

$$\hat{e}_{n+1} = a_{n+1} - \theta a_{n+1}$$

$$\text{Var} = (1 + \theta^2) \hat{\alpha}^2$$

Invertibility: MA(1),  $a_t = (r_t - \mu) + \sum_{i=1}^{\infty} \theta^i (r_{t-i} - \mu)$

$$r_t = \frac{\mu}{1-\theta} + a_t - \sum_{i=1}^{\infty} \theta^i r_{t+i}$$

MA(2)  $r_t = \mu + a_t - \theta_1 a_{t-1} - \theta_2 a_{t-2}$

$$E r_t = \mu$$

$$\text{Var} = (1 + \theta_1^2 + \theta_2^2) 6a^2$$

$$\text{Invertibility: } (1 - \theta_1 z - \theta_2 z^2) = 0 \text{ outside.}$$

$$r_t = \mu + (1 - \alpha_1 B)(1 - \alpha_2 B)a_t$$

$$u_t = (1 - \alpha_2 B) a_t$$

$$\text{ACF: } \gamma_0 = \text{Var}(r_t) = (1 + \theta_1^2 + \theta_2^2) 6a^2$$

$$\begin{aligned} \gamma_1 &= E[(r_t - \mu)(r_{t+1} - \mu)] \\ &= E[(a_t - \theta_1 a_{t-1} - \theta_2 a_{t-2})(a_{t+1} - \theta_1 a_{t} - \theta_2 a_{t-1})] \\ &= (-\theta_1 + \theta_1 \theta_2) 6a^2 \end{aligned}$$

$$\begin{aligned} \gamma_2 &= E[(r_t - \mu)(r_{t+2} - \mu)] \\ &= E[(a_t - \theta_1 a_{t-1} - \theta_2 a_{t-2})(a_{t+2} - \theta_1 a_{t+1} - \theta_2 a_{t})] \\ &= -\theta_2 6a^2 \end{aligned}$$

MA(q)  $r_t = \mu + a_t - \theta_1 a_{t-1} - \theta_2 a_{t-2} \dots - \theta_q a_{t-q}$

$$E r_t = \mu$$

$$\text{Var} = (1 + \theta_1^2 + \theta_2^2 + \dots + \theta_q^2) 6a^2$$

$$\text{Invertibility: } (1 - \theta_1 z - \theta_2 z^2 - \dots - \theta_q z^q) = 0 \text{ outside.}$$

always stationary

only  $q \uparrow$  ACF  $\neq 0$ .  $\text{即为0} \Rightarrow \text{用ACF求参数}.$

$$\gamma_{q+1} = 0$$

① q : PACF

② estimate parameter

③ model checking

$$\sqrt{n} (\hat{\theta} - \theta_0) \sim N(0, \hat{\sigma}^2) \quad \hat{\sigma}^2 = \left( \sum_{t=1}^n \frac{\partial a_{t-1}(\hat{\theta})}{\partial \theta} \frac{\partial a_{t-1}(\hat{\theta})}{\partial \theta'} \right)^{-1}_{\hat{\theta}, \hat{\theta}'}$$

$$\sqrt{n} (\hat{\theta}_i - \theta_{0i}) \sim N(0, \hat{\sigma}_{ii}^2)$$

ARMA(1,1)

$$(1-\phi_1 B) r_t = \phi_0 + \underbrace{(1-\theta B)}_{w_t} a_t \cdot w_t$$

$$\begin{aligned}
r_t &= \phi_0 + \phi_1 r_{t-1} + w_t \\
&= \phi_0 + \phi_1 (\phi_0 + \phi_1 r_{t-2} + w_{t-1}) + w_t \\
&= \phi_0 + \phi_1 \phi_0 + \phi_1^2 r_{t-2} + \phi_1 w_{t-1} + w_t \\
&= \phi_0 + w_t + \phi_1 (\phi_0 + w_{t-1}) + \phi_1^2 r_{t-2} \\
&= \phi_0 (1 + \phi_1) + w_t + \phi_1 w_{t-1} + \phi_1^2 r_{t-2} \\
E r_t &= \phi_0 \frac{1}{1-\phi_1} = \frac{\phi_0}{1-\phi_1} \\
&= \frac{\phi_0}{1-\phi_1} + w_t + \sum_{i=1}^{\infty} \phi_1^i w_{t-i} \\
&= \frac{\phi_0}{1-\phi_1} + a_t - \theta a_{t-1} + \phi(a_{t-1} - \theta a_{t-2}) + \phi^2(a_{t-2} - \theta a_{t-3}) + \dots \\
&= \frac{\phi_0}{1-\phi_1} + a_t + (\phi - \theta) a_{t-1} + \phi(\phi - \theta) a_{t-2} \\
&= \frac{\phi_0}{1-\phi_1} + a_t + (\phi - \theta) \sum_{i=1}^{\infty} \phi^{i-1} a_{t-i} + \phi^2(\phi - \theta) a_{t-3} \dots \\
\text{Var} &= \left( 1 + (\phi - \theta)^2 [1^2 + \phi^2 + \phi^4 + \dots] \right) \sigma^2 \\
&= \left( 1 + (\phi - \theta)^2 \frac{1}{1-\phi^2} \right) 6\sigma^2
\end{aligned}$$

$$\begin{aligned}
\rho_1 &= E[(r_t - \mu)(r_{t-1} - \mu)] \\
&= E[(a_t + (\phi - \theta)a_{t-1} + (\phi - \theta)\phi a_{t-2} + (\phi - \theta)\phi^2 a_{t-3}, \\
&\quad (a_{t-1} + (\phi - \theta)a_{t-2} + (\phi - \theta)\phi a_{t-3} + \dots)) \\
&= [(1 - \phi) + (\phi - \theta)^2 \phi + (\phi - \theta)^2 \phi^3 + \dots] \sigma^2 \\
&= [(1 - \phi) + (\phi - \theta)^2 \frac{\phi}{1 - \phi^2}] \sigma^2 \\
&:= \rho_1 = \frac{(1 - \phi^2)(\phi - \theta) + (\phi - \theta)^2 \phi}{(1 - \phi^2) + (\phi - \theta)^2} \quad (\text{假设 } \mu = 0, \phi_0 = 0)
\end{aligned}$$

$$\begin{aligned}
\rho_2 &= E[(\phi_1 r_{t-1} + a_t - \theta a_{t-1})(r_{t-1})] \\
&= \phi_1 \rho_1 + 0 - \theta \sigma^2 \\
&= \underbrace{\phi_1 \rho_1 - \theta \sigma^2}_{\text{AR.}} \quad \rho_2 = \phi_1 \rho_1 \dots \quad \rho_K = \phi^{K-1} \rho_1
\end{aligned}$$

$\therefore$  ACF 和 PACF 均不等于 0.

$\downarrow$   
similar to AR.

$$\begin{aligned}
r_t &= \phi_0 + \phi_1 r_{t-1} + a_t - \theta a_{t-1} \\
\text{forecast.} \quad (1-\phi_1 B) r_t &= \phi_0 + \underbrace{(1-\theta B)}_{w_t} a_t \cdot w_t
\end{aligned}$$

$$\begin{aligned}
\hat{Y}_{n+1}(1) &= E[Y_{n+1} | Y_n] = E[\phi_0 + \phi_1 r_n + a_{n+1} - \theta a_n | F_n] \\
&= \phi_0 + \phi_1 r_n - \theta a_n
\end{aligned}$$

$$e_n(1) = a_{n+1}$$

$$\hat{r}_n(2) = E[r_{n+2}|r_n] = E[\phi_0 + \phi_1 r_{n+1} + \alpha_{n+2} - \theta \alpha_{n+1} | F_n]$$

$$= \phi_0 + \phi_1 \hat{r}_n(1)$$

$$\hat{r}_n(3) = E[\phi_0 + \phi_1 r_{n+2} + \alpha_{n+3} - \theta \alpha_{n+2} | F_n]$$

$$= \phi_0 + \phi_1 \hat{r}_n(2) \Rightarrow \text{AR model.}$$

**ARMA (p,q)**  $r_t = \phi_0 + \phi_1 r_{t-1} + \dots + \phi_p r_{t-p}$   
 $+ \alpha_t - \theta_1 \alpha_{t-1} - \dots - \theta_q \alpha_{t-q}$

$1 - \sum \phi_i z^i$  and  $1 - \sum \theta_i z^i$  all outside + no common solution.

representation:  $\phi_p(z) \Theta_q(z) = 1 - \sum_{i=1}^{\infty} \pi_i z^i$

$$\phi_p(z) r_t = \Theta_q(z) \alpha_t + \phi_0$$

$$\begin{aligned} \textcircled{1} \text{ AR: } \alpha_t &= \phi_p(z) \Theta_q^{-1}(z) r_t - \Theta_q^{-1}(z) \phi_0 \\ &= \left( 1 - \sum_{i=1}^{\infty} \pi_i z^i \right) r_t - \Theta_q^{-1}(z) \phi_0 \\ &= r_t - \pi_1 r_{t-1} - \pi_2 r_{t-2} \dots - \Theta_q^{-1}(z) \phi_0 \end{aligned}$$

$$r_t = \Theta_q^{-1}(z) \phi_0 + \alpha_t + \pi_1 r_{t-1} + \dots$$

$$\begin{aligned} \textcircled{2} \text{ MA: } \phi_p^{-1}(z) \Theta_q(z) &= 1 + \sum q_i z^i \\ r_t &= \phi_p^{-1}(z) \Theta_q(z) \alpha_t + \phi_p^{-1}(z) \phi_0 \\ &= \alpha_t + q_1 \alpha_{t-1} + q_2 \alpha_{t-2} + \dots + \phi_p^{-1}(z) \phi_0. \end{aligned}$$

forecast error  $\hat{r}_{n+1}$  MA representation.

$$\begin{aligned} \hat{r}_{n+1} &= E[\alpha_{n+1} + q_1 \alpha_{n+1-1} + \dots + q_{l-1} \alpha_{n+1-l} + \dots | F_n] \\ &= q_l \alpha_l + q_{l+1} \alpha_{l+1} \dots \end{aligned}$$

$$e_n = \alpha_{n+1} + q_1 \alpha_{n+1-1} + \dots + q_{l-1} \alpha_{n+1-l}$$

$$\text{Var} = l + q_1^2 + \dots + q_{l-1}^2 \quad (1 - \phi_1 B) r_t = \phi_0 + \underbrace{(1 - \phi_1 B)}_{\text{wt}} \alpha_t \text{ wt.}$$

$|B| = 1 \Rightarrow$  non-stationary

strong memory

unit-root test

Seasonal:  $w_t: (1 - \theta B^s) \alpha_t$  only  $r_s \neq 0$ .

$w_t: (1 - \theta B)(1 - \theta B^s) \alpha_t$  ~~if  $r_s \neq 0$~~

how to make sure if  $\sigma_t^2$  is a constant

of constant.  $AOF = 0$

ARCH:

$$GARCH. \sigma_t^2 = \alpha_0 + \sum_{i=1}^m \alpha_i \sigma_{t-i}^2 + \sum_{j=1}^s \beta_j \varepsilon_{t-j}^2$$

$$\hat{y}_t = \sigma_t^2 - \sigma_t^{*2}$$

$$\sigma_t^{*2} = \alpha_0 + \sum_{\max(m,s)}^{\infty} (\alpha_i + \beta_i) \sigma_{t-i}^2 - \frac{\sum_{j=1}^s \beta_j \hat{y}_{t-j} + \hat{y}_t}{\sigma_{t-1} \sigma_{t-2}}$$

$$GARCH(1,1) \quad \sigma_t^2 = \alpha_0 + \alpha_1 \sigma_{t-1}^2 + \beta_1 \varepsilon_{t-1}^2.$$

$$\textcircled{1} \quad E[\sigma_t^2] = 0 \quad \sigma_t^2 \text{ and } \varepsilon_t \text{ is independent.}$$

$$\textcircled{2} \quad \sigma_t^2 = \alpha_0 + \alpha_1 \varepsilon_{t-1}^2 \sigma_{t-1}^2 + \beta_1 \sigma_{t-1}^2$$

$$= \alpha_0 + (\alpha_1 \varepsilon_{t-1}^2 + \beta_1) \sigma_{t-1}^2$$

$$= \alpha_0 + (\alpha_1 \varepsilon_{t-1}^2 + \beta_1) [\alpha_0 + (\alpha_1 \varepsilon_{t-2}^2 + \beta_1) \sigma_{t-2}^2]$$

$$= \alpha_0 + \alpha_0 (\alpha_1 \varepsilon_{t-1}^2 + \beta_1) + (\alpha_1 \varepsilon_{t-2}^2 + \beta_1) (\alpha_1 \varepsilon_{t-1}^2 + \beta_1) \sigma_{t-2}^2 \dots$$

$$= \alpha_0 \left[ 1 + \sum_{i=1}^{\infty} \prod_{j=1}^i (\alpha_1 \varepsilon_{t-j}^2 + \beta_1) \right]$$

$\Leftarrow$  if  $E \ln(\alpha_1 \varepsilon_{t-1}^2 + \beta_1) < 0 \Rightarrow$  converge.

Cauchy  $S_{n+m} - S_n \rightarrow 0$

$$\textcircled{3} \quad E[\sigma_t^2] = E[\varepsilon_t^2] E[\sigma_{t-1}^2] = E[\sigma_{t-1}^2]$$

$$= \alpha_0 \left( 1 + \sum_{i=1}^{\infty} \prod_{j=1}^i (\alpha_1 \varepsilon_{t-j}^2 + \beta_1) \right)$$

$$= \alpha_0 \left( 1 + \sum_{i=1}^{\infty} \prod_{j=1}^i \alpha_1 \beta_1 \right)$$

$$= \alpha_0 \left( 1 + \sum_{i=1}^{\infty} (\alpha_1 + \beta_1)^i \right)$$

$$= \alpha_0 \frac{1}{1 - (\alpha_1 + \beta_1)} \quad \text{if } \alpha_1 + \beta_1 < 1$$

$$\text{for Garch } (m,s) \Rightarrow \sum_{i=1}^{\max(m,s)} (\alpha_i + \beta_i) < 1$$

$$\textcircled{4} \quad \text{if } 1 - 2\alpha_1 - (\alpha_1 + \beta_1)^2 > 0.$$

$$\frac{E[\sigma_t^2]}{[E[\sigma_t^2]]^2} > 3.$$

$\Rightarrow$  can catch fat tail feature

$\epsilon_t$  is strict stationary

forecast:  $\alpha_1 + \beta_1$  is important.

IGARCH:  $\alpha_1 + \beta_1 = 1$

GARCH-M:  $r_t = \mu + C\sigma_t^2 + \alpha_t$

$$\alpha_t = \sigma_t \varepsilon_t$$

$$\sigma_t^2 = \alpha_0 + \alpha_1 \alpha_{t-1}^2 + \beta_1 \sigma_{t-1}^2 \quad C \text{ is risk premium.}$$

EGARCH:  $g(\varepsilon_t) = \theta \varepsilon_t + |\varepsilon_t| - E|\varepsilon_t|$  asymmetry.

$$\alpha_t = \sigma_t \varepsilon_t \quad \ln \sigma_t^2 = \alpha_0 + \sum_{i=1}^m \alpha_i g(\varepsilon_{t-i}) + \sum_{i=1}^s \beta_i \ln(\sigma_{t-i}^2)$$

$\theta, \alpha$  is leverage.  $\theta_1 = \theta \alpha$ . (EGARCH(1,1))

CHARMA:  $\alpha_t = S_1 \alpha_{t-1} + S_2 \alpha_{t-2} + \dots + S_m \alpha_{t-m} + \gamma_t$

RCA:  $r_t = \phi_0 + \sum_{i=1}^p (\phi_i + S_{it}) r_{t-i} + \alpha_t$

when  $\phi_i$  stable: garch interval narrow

$\phi_i$  not stable: garch interval wide

Vector mean:  $\begin{bmatrix} \mu_{1t} \\ \mu_{2t} \\ \vdots \\ \mu_{Kt} \end{bmatrix} = \mu$

$\Sigma$ : variance =  $P(0) = \begin{bmatrix} r_{11}(0) & r_{12}(0) & \dots & r_{1K}(0) \\ \vdots & \ddots & \ddots & \vdots \\ r_{K1}(0) & \ddots & \ddots & r_{KK}(0) \end{bmatrix}$

$$r_{ii}(0) = E[(z_{it} - \mu_i)^2] \quad r_{ij}(0) = E[(z_{it} - \mu_i)(z_{jt} - \mu_j)]$$

$$P_{ij}(0) = \frac{r_{ij}(0)}{\sqrt{r_{ii}(0) r_{jj}(0)}}$$

$$P(0) = D^T P(0) D^{-1}$$

$$D = \begin{pmatrix} \sqrt{r_{11}(0)} & & \\ & \sqrt{r_{22}(0)} & \\ & & \ddots & \\ & & & \sqrt{r_{KK}(0)} \end{pmatrix}$$

$$T(1) = E[(z_{it} - \mu_i)(z_{it-1} - \mu_i)]$$

$$= \begin{pmatrix} r_{11}(1) & r_{12}(1) & \dots \\ \vdots & \vdots & \vdots \\ r_{K1}(1) & r_{K2}(1) & \dots \end{pmatrix}$$

$$r_{ij}(1) = E[(z_{it} - \mu_i)(z_{jt-1} - \mu_j)]$$

cointegration: CK. 1 independent  $\beta$ .

partial stationary: some root = 1. others outside the unit circle.

$$\Phi(B) = \Phi^*(B)(1 - B) + \Phi(1)B,$$

$$\Phi^*(B)(1 - B)y_t = \alpha\beta'y_{t-1} + \varepsilon_t$$

$$\Phi(B)y_t = \varepsilon_t$$

$$(\Phi^*(B)(1 - B) + \Phi(1)B)y_t = \varepsilon_t$$

$$\Phi^*(B)(1 - B)y_t + \Phi(1)y_{t-1} = \varepsilon_t$$

$$\Phi^*(B)(1 - B)y_t = -\Phi(1)y_{t-1} + \varepsilon_t$$

$$= \underbrace{\alpha\beta'y_{t-1}}_{\text{stationary.}} + \varepsilon_t$$

VECM  $\rightarrow z_t$

$\beta$ : long-run parameter.

$r = k - d$ .  $d$  # unit root #.

