

STAT 215A Fall 2023

Week 8

Chengzhong Ye

Announcements

- Midterm: 10/26 in class at the standard time. **Starts at 12:40pm SHARP!**
 - Will released a practice midterm today
 - Solutions will be posted next week
 - Review lecture Tuesday 10/24

Outline for today

- EM algorithm

EM Algorithm

Dempster et al., 1977



Various resources

Intuition:

- Quick: <https://stackoverflow.com/questions/11808074/what-is-an-intuitive-explanation-of-the-expectation-maximization-technique#answer-43561339>

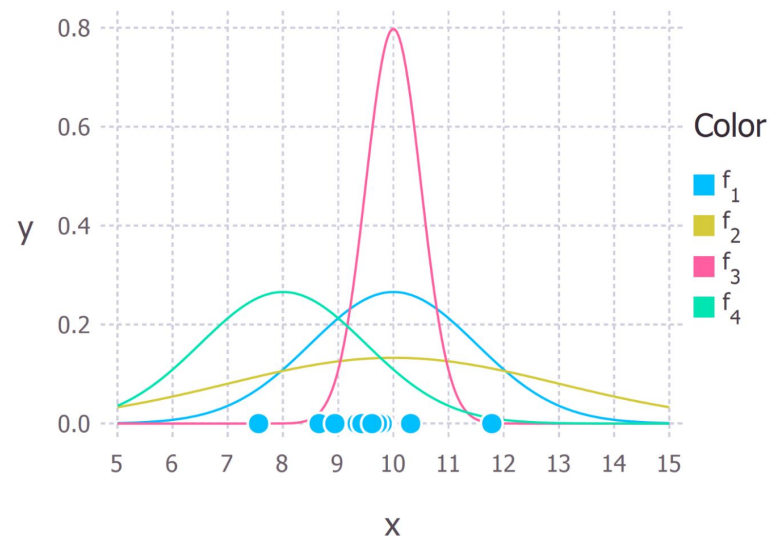
Math:

- Quick overview: http://www.seanborman.com/publications/EM_algorithm.pdf
 - Included in the week8 folder
- More in-depth: <https://arxiv.org/pdf/1105.1476.pdf>

Maximum Likelihood Estimation Review

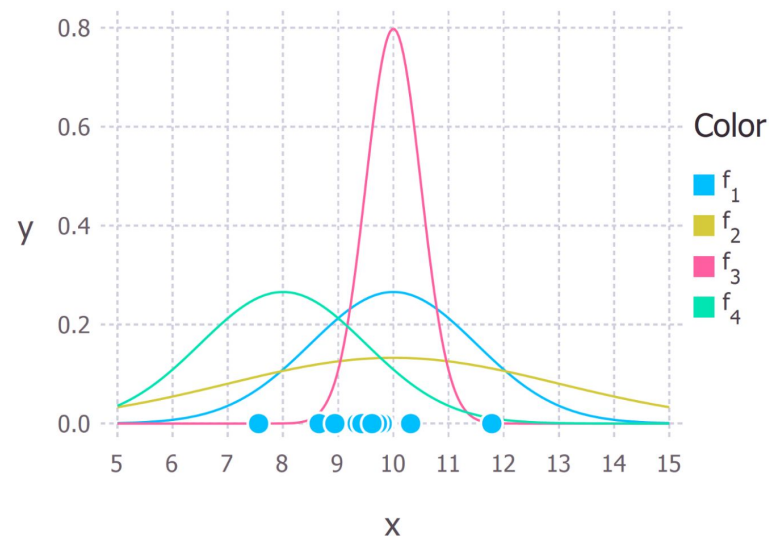
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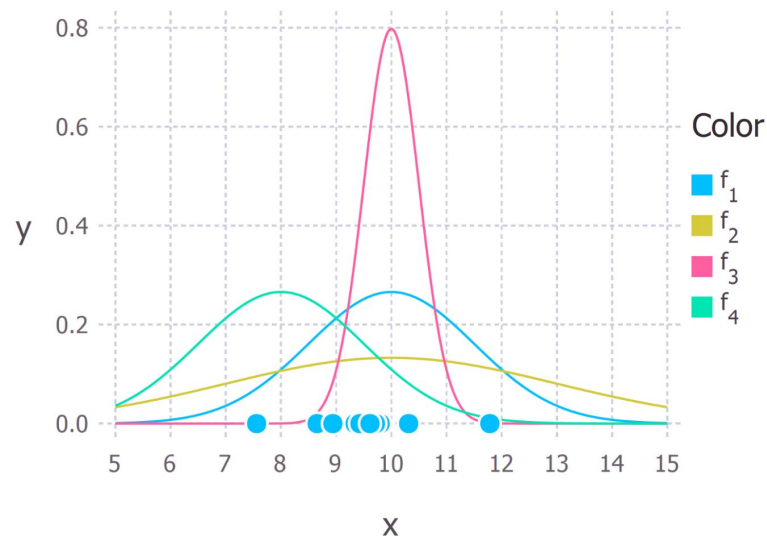
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- Likelihood: $\mathcal{L}(\theta) = \prod_{i=1}^n p_\theta(x_i)$
- Log-likelihood: $\ell(\theta) = \log \mathcal{L}(\theta)$
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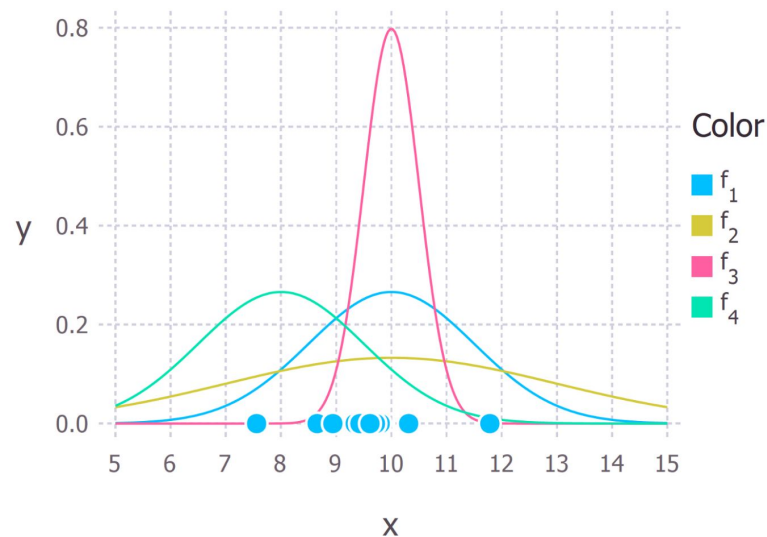
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- **Intuition:** Find the value of θ under which we would be least surprised to see a sample like the observed one.
- **Optimization problem:** take derivative, set equal to zero, solve for parameter.

EM Overview

Motivation 1: “Hard” Maximum Likelihood Estimation Problems

Say we have the following mixture of Gaussians problem:

$$X_1, \dots, X_n \stackrel{i.i.d.}{\sim} \pi_1 N(\mu_1, \sigma_1^2) + \pi_2 N(\mu_2, \sigma_2^2), \pi_1 + \pi_2 = 1$$

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- Because of the sum of normals, the log-likelihood isn't as helpful as before and taking derivatives w.r.t $\theta = (\pi_1, \mu_1, \sigma_1, \mu_2, \sigma_2)$ and setting equal to zero, etc., won't be particularly successful.
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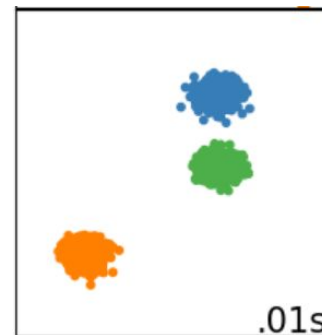
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- Instead, we can introduce **latent variables** which tell us which of the two Gaussians each observation comes from: $Z_i \stackrel{i.i.d.}{\sim} \text{Bernoulli}(\pi_1)$
- If we know the latent allocations then the problem simply becomes two easy MLE exercises.

$$\begin{aligned} X_i | Z_i = 1 &\stackrel{i.i.d.}{\sim} N(\mu_1, \sigma_1^2) \\ X_i | Z_i = 2 &\stackrel{i.i.d.}{\sim} N(\mu_2, \sigma_2^2) \end{aligned}$$

EM Overview

Motivation 2: Clustering

Since EM helps with finding solutions to mixture problems, it lends itself naturally to clustering.



Recall:

- K-means performs well when clusters are homogeneous.
-

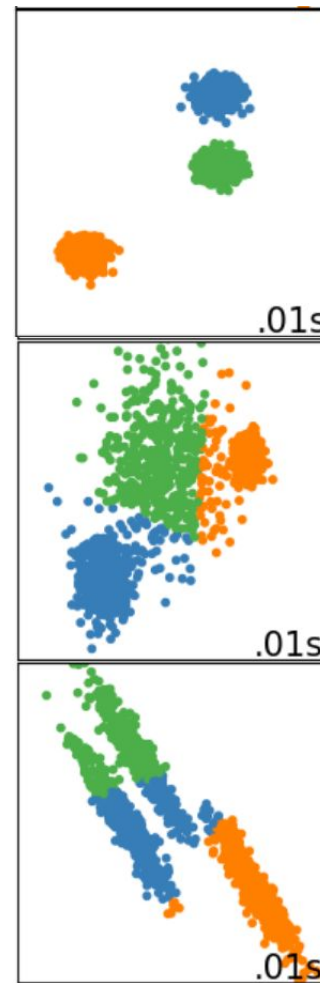
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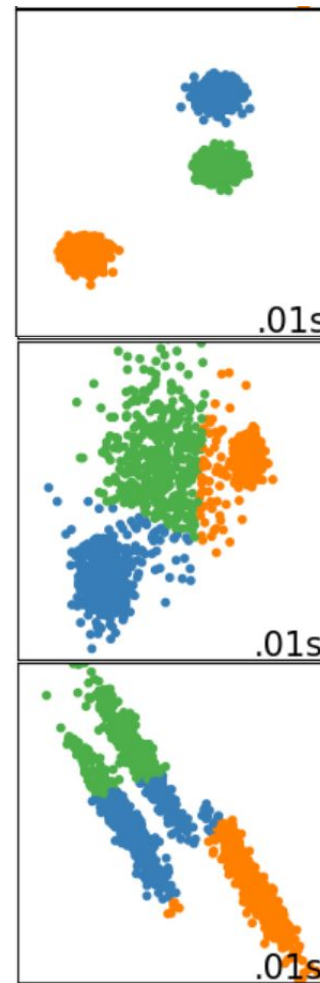
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The EM generalizes K-means:

- Still performs great where K-means does.



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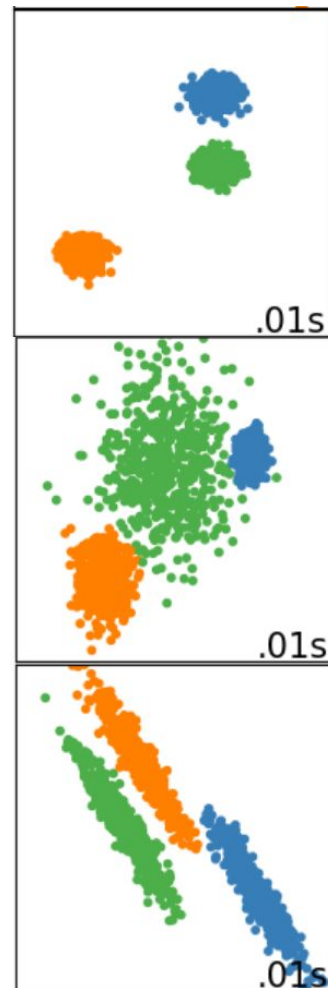
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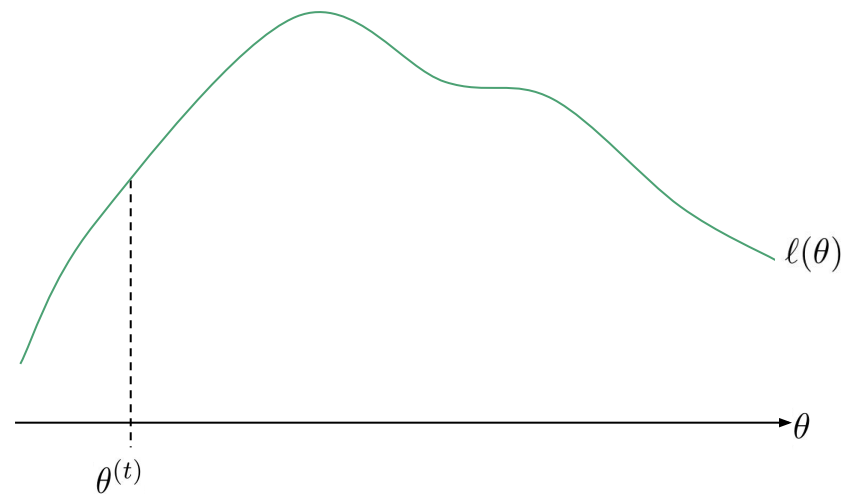
The EM generalizes K-means:

- Still performs great where K-means does.
- But it can also handle differences in spread and symmetry.
- EM is a “soft” clustering algorithm.



EM algorithm intuition

Say we make a guess $\theta^{(t)}$

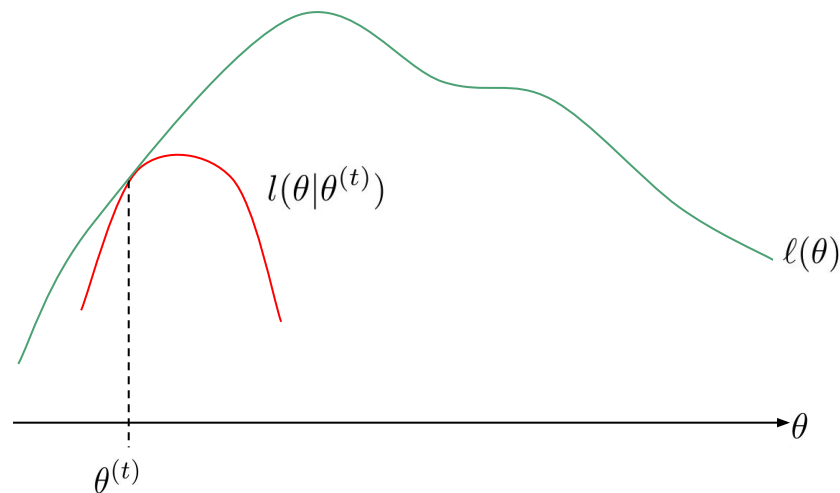


EM algorithm intuition

Say we make a guess $\theta^{(t)}$

The insight of the EM algorithm is that we can find a function $l(\theta|\theta^{(t)})$ such that

- $l(\theta|\theta^{(t)}) \leq \ell(\theta)$
- $l(\theta^{(t)}|\theta^{(t)}) = \ell(\theta^{(t)})$



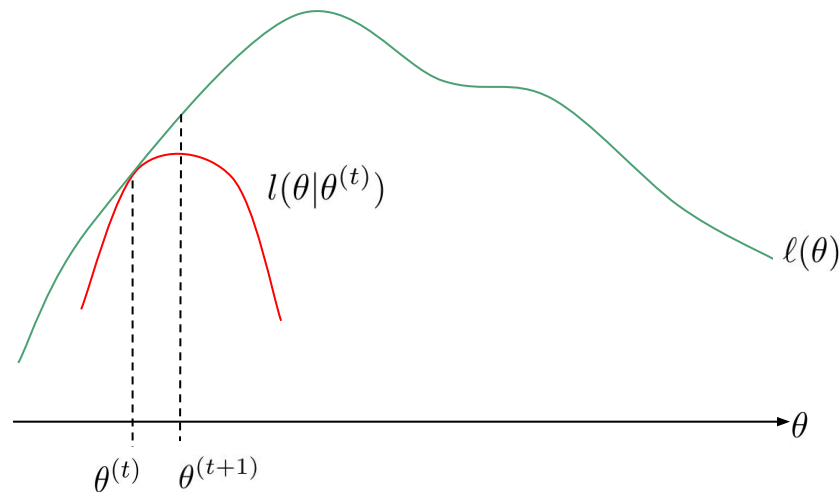
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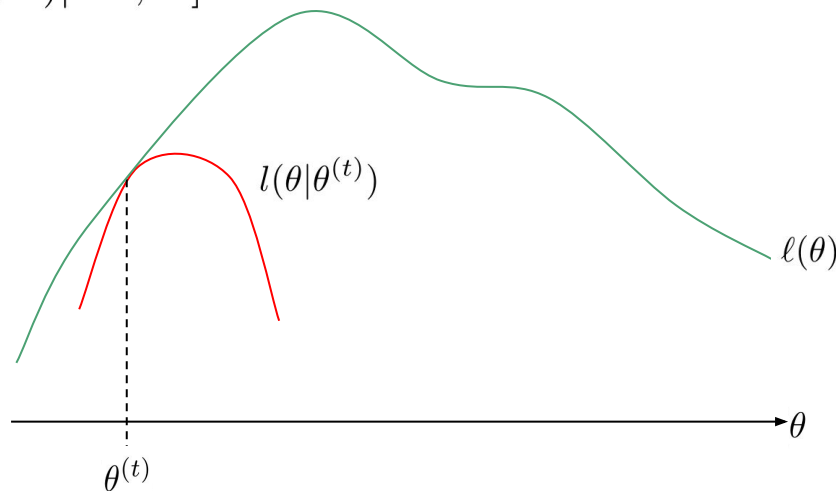
- $l(\theta|\theta^{(t)}) \leq \ell(\theta)$
- $l(\theta^{(t)}|\theta^{(t)}) = \ell(\theta^{(t)})$

So any θ that increases $l(\theta|\theta^{(t)})$ also increases $\ell(\theta)$.



EM algorithm steps

It turns out that maximizing $l(\theta|\theta^{(t)})$ is equivalent* to maximizing the expectation $Q(\theta|\theta^{(t)}) = \mathbb{E}[\ell(\theta; X, Z)|\theta^{(t)}, X]$



* see the Borman tutorial for the derivation

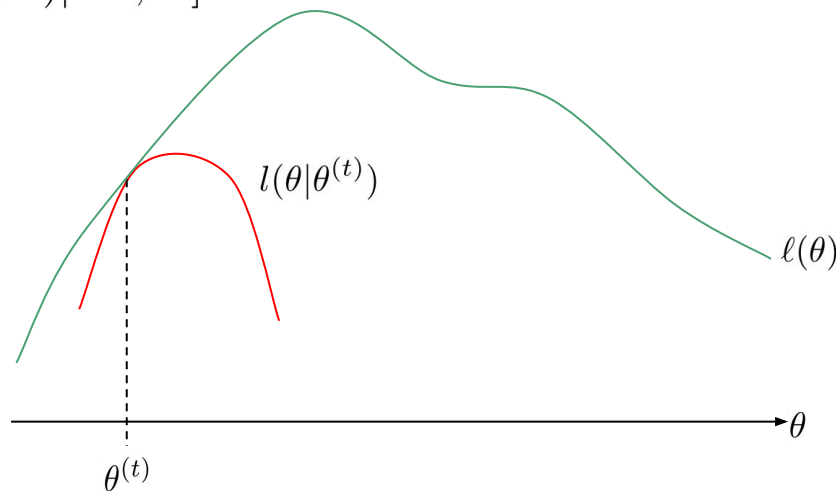
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The EM algorithm involves two steps that are repeated until convergence:

1. **E:** Calculate the expectation

$$Q(\theta|\theta^{(t)}) = \mathbb{E}[\ell(\theta; X, Z)|\theta^{(t)}, X]$$



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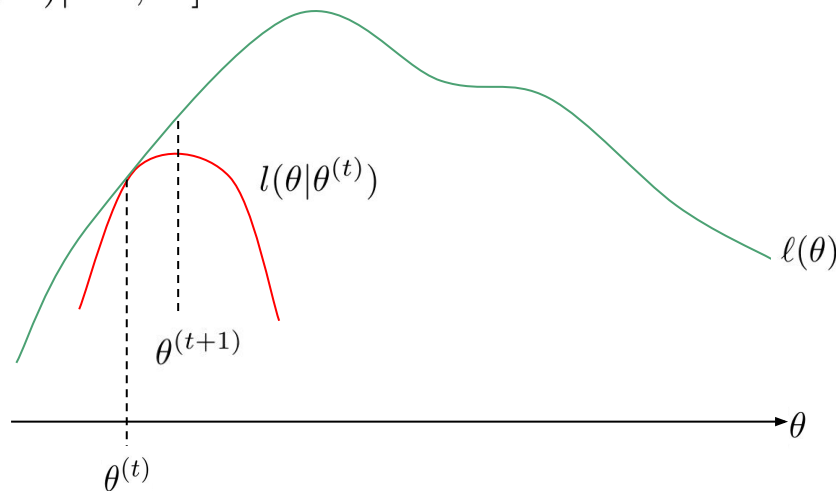
The EM algorithm involves two steps that are repeated until convergence:

1. **E:** Calculate the expectation

$$Q(\theta|\theta^{(t)}) = \mathbb{E}[\ell(\theta; X, Z)|\theta^{(t)}, X]$$

2. **M:** Maximize $Q(\theta|\theta^{(t)})$ w.r.t. θ

Note: we can initialize with a random guess $\theta^{(0)}$



EM: Gaussian Mixture Example

Same setup from before:

$$X_1, \dots, X_n \stackrel{i.i.d.}{\sim} \pi_1 N(\mu_1, \sigma_1^2) + \pi_2 N(\mu_2, \sigma_2^2), \pi_1 + \pi_2 = 1$$

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$$Z_i \stackrel{i.i.d.}{\sim} \text{Bernoulli}(\pi_1) + 1$$

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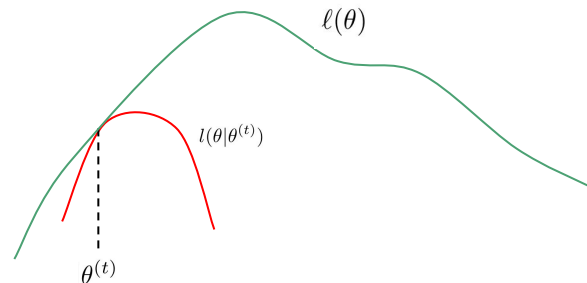
$$Z_i \stackrel{i.i.d.}{\sim} \text{Bernoulli}(\pi_1)$$

Likelihood: $p_\theta(x_i, z_i) = p_\theta(x_i | z_i) p_\theta(z_i) = \begin{cases} \frac{1}{\sqrt{2\pi\sigma_1^2}} \exp\left\{-\frac{(x_i - \mu_1)^2}{2\sigma_1^2}\right\} \pi_1, & \text{if } Z_i = 1 \\ \frac{1}{\sqrt{2\pi\sigma_2^2}} \exp\left\{-\frac{(x_i - \mu_2)^2}{2\sigma_2^2}\right\} \pi_2, & \text{if } Z_i = 2 \end{cases}$

$$\log p_\theta(x_i, z_i) = \begin{cases} -\frac{1}{2} \log 2\pi - \log \sigma_1 - \frac{(x_i - \mu_1)^2}{2\sigma_1^2} + \log \pi_1, & \text{if } Z_i = 1 \\ -\frac{1}{2} \log 2\pi - \log \sigma_2 - \frac{(x_i - \mu_2)^2}{2\sigma_2^2} + \log \pi_2, & \text{if } Z_i = 2 \end{cases}$$

E-Step: Compute $Q(\theta|\theta^{(t)}) = \mathbb{E}[\ell(\theta; X, Z)|\theta^{(t)}, X]$

$$\begin{aligned} Q(\theta|\theta^{(t)}) &= \sum_{i=1}^n \mathbb{E}[\ell(\theta; X_i, Z_i)|\theta^{(t)}, X] \\ &= \sum_{i=1}^n \left\{ \mathbb{E}[\ell(\theta; X_i, Z_i)|\theta^{(t)}, X, Z_i = 1] \mathbb{P}(Z_i = 1|\theta^{(t)}, X) + \right. \\ &\quad \left. \mathbb{E}[\ell(\theta; X_i, Z_i)|\theta^{(t)}, X, Z_i = 2] \mathbb{P}(Z_i = 2|\theta^{(t)}, X) \right\} \\ &\quad \text{(law of total expectation)} \end{aligned}$$



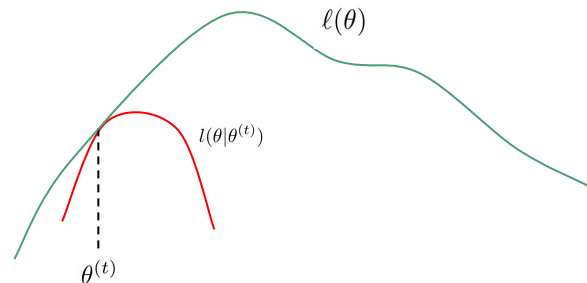
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(law of total expectation)

$$\begin{aligned} &= \sum_{i=1}^n \left\{ \log \pi_1 - \frac{1}{2} \log 2\pi - \log \sigma_1 - \frac{(x_i - \mu_1)^2}{2\sigma_1^2} Z_{i,1}^{(t)} + \right. \\ &\quad \left. \log \pi_2 - \frac{1}{2} \log 2\pi - \log \sigma_2 - \frac{(x_i - \mu_2)^2}{2\sigma_2^2} Z_{i,2}^{(t)} \right\} \end{aligned}$$

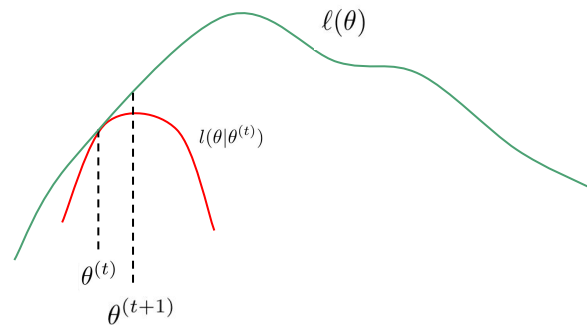


$$\begin{aligned} Z_{i,j}^{(t)} &= \mathbb{P}(Z_i = j|\theta^{(t)}, X) \\ &= \frac{\pi_j^{(t)} \phi(\frac{x_i - \mu_j}{\sigma_j})/\sigma_j}{\sum_{k=1}^2 \pi_k^{(t)} \phi(\frac{x_i - \mu_k}{\sigma_k})/\sigma_k} \end{aligned}$$

Standard
normal pdf

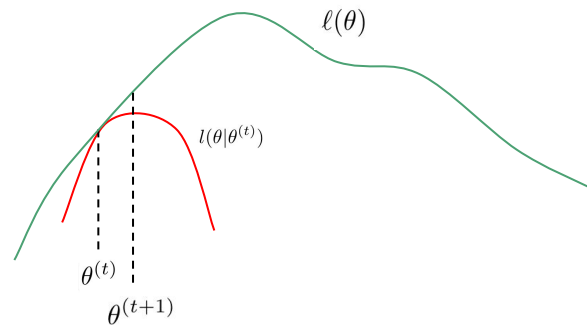
M-Step: Maximize $Q(\theta|\theta^{(t)})$ w.r.t. $\theta = (\pi_1, \mu_1, \sigma_1, \mu_2, \sigma_2)$

- π_1 :
$$\frac{\partial Q}{\partial \pi_1} = \frac{\partial}{\partial \pi_1} \sum_{i=1}^n \left(\log \pi_1 Z_{i,1}^{(t)} + \log(1 - \pi_1) Z_{i,2}^{(t)} \right)$$



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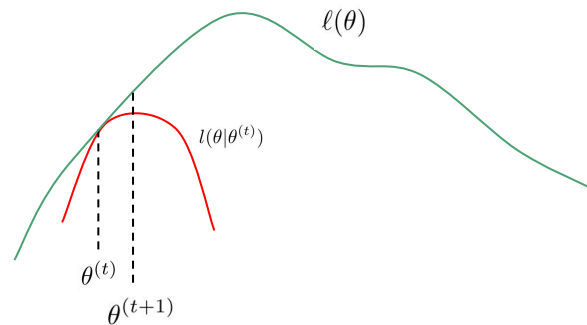


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$$= \sum_{i=1}^n \frac{Z_{i,1}^{(t)}}{\pi_1} + \sum_{i=1}^n \frac{Z_{i,2}^{(t)}}{1 - \pi_1}$$

$$\stackrel{\text{set}}{=} 0 \implies \boxed{\pi_1^{(t+1)} = \frac{\sum_i Z_{i,1}^{(t)}}{\sum_i Z_{i,1}^{(t)} + Z_{i,2}^{(t)}} = \frac{1}{n} \sum_i Z_{i,1}^{(t)}}$$

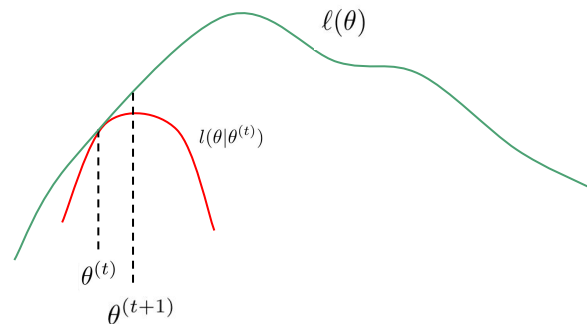


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- μ_1 :
$$\frac{\partial Q}{\partial \mu_1} = \frac{\partial}{\partial \mu_1} \sum_{i=1}^n \left(-\frac{(x_i - \mu_1)^2}{2\sigma_1^2} \right) Z_{i,1}^{(t)} \implies \mu_1^{(t+1)} = \frac{\sum_i Z_{i,1}^{(t)} X_i}{\sum_i Z_{i,1}^{(t)}}$$

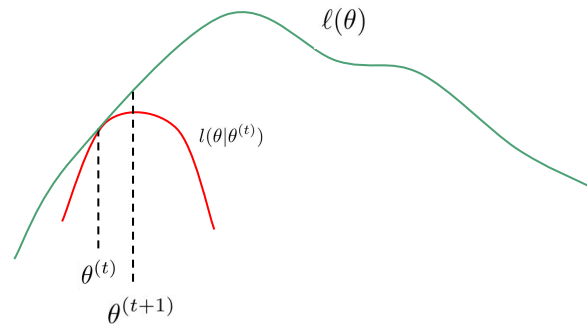
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- $$\begin{aligned} \sigma_1: \frac{\partial Q}{\partial \sigma_1} &= \frac{\partial}{\partial \sigma_1} \sum_{i=1}^n \left(-\log \sigma_1 - \frac{(x_i - \mu_1)^2}{2\sigma_1^2} \right) Z_{i,1}^{(t)} \\ &= \sum_{i=1}^n \left(-\frac{1}{\sigma_1} + \frac{(x_i - \mu_1)^2}{\sigma_1^3} \right) Z_{i,1}^{(t)} \stackrel{set}{=} 0 \end{aligned}$$

$$(\sigma_1^2)^{(t+1)} = \frac{\sum_i Z_{i,1}^{(t)} (X_i - \mu_1^{(t+1)})^2}{\sum_i Z_{i,1}^{(t)}}$$



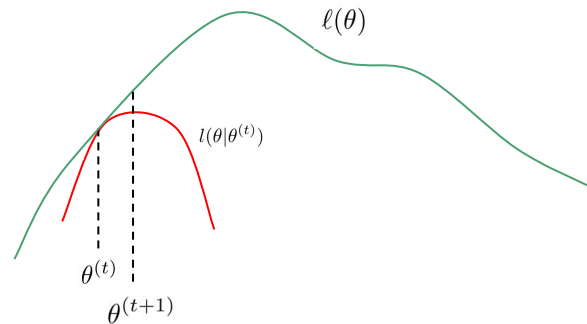
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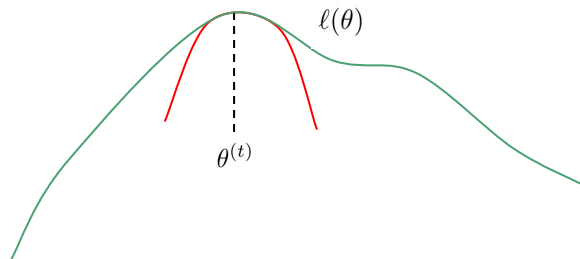
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- Similar process for μ_2 & σ_2



Repeat this process until we find a (local) maximum



Go to `week8/em_lab.R`
