

ST518 - Mixed effects models

Mixed Effects Models

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Outline

Topic: Mixed effects models

- One-way random effects model to study *variances*
- Mixed effects models
- Subsampling
- Expected mean squares for mixed models

One-way random effects model

Example:

- Genetics study w/ beef animals. Measure birthweight Y (lbs).
- $t = 5$ sires, each mated to a separate group of $n = 8$ dams.
- $N = 40$, completely randomized.



| Sire # | Level | Birthweights | | | | | | | | $\bar{y}_{i.}$ | s_i |
|--------|-------|--------------|-----|-----|-----|-----|-----|-----|-----|----------------|-------|
| | | Sample | | | | | | | | | |
| 177 | 1 | 61 | 100 | 56 | 113 | 99 | 103 | 75 | 62 | 83.6 | 22.6 |
| 200 | 2 | 75 | 102 | 95 | 103 | 98 | 115 | 98 | 94 | 97.5 | 11.2 |
| 201 | 3 | 58 | 60 | 60 | 57 | 57 | 59 | 54 | 100 | 63.1 | 15.0 |
| 202 | 4 | 57 | 56 | 67 | 59 | 58 | 121 | 101 | 101 | 77.5 | 25.9 |
| 203 | 5 | 59 | 46 | 120 | 115 | 115 | 93 | 105 | 75 | 91.0 | 28.0 |



Q: Statistical model for these data? $Y_{ij} = \mu + \tau_i + E_{ij}$

Random effects model

The one-way random effects model:

$$Y_{ij} = \underbrace{\mu}_{\text{fixed}} + \underbrace{T_i}_{\text{random}} + \underbrace{E_{ij}}_{\text{random}} \quad \text{for } i = 1, 2, \dots, t \text{ and } j = 1, \dots, n$$

with

- $T_1, T_2, \dots, T_t \stackrel{iid}{\sim} N(0, \sigma_T^2)$ 
- $E_{11}, \dots, E_{tn} \stackrel{iid}{\sim} N(0, \sigma^2)$ 
- T_1, T_2, \dots, T_t independent of E_{11}, \dots, E_{tn}

Features

- T_1, T_2, \dots denote *random* effects, drawn from some population of interest. That is, T_1, T_2, \dots is a random sample!
- σ_T^2 and σ^2 are called variance components
- conceptually different from one-way fixed effects model

Beef animal genetic study, continued

With $t = 5$ and $n = 8$, the random effects T_1, T_2, \dots, T_5 reflect sire-to-sire variability. (ie genetics)

No particular interest in $\tau_1, \tau_2, \dots, \tau_5$ from the (misspecified) fixed effects model:

$$Y_{ij} = \underbrace{\mu}_{\text{fixed}} + \underbrace{\tau_i}_{\text{fixed}} + \underbrace{E_{ij}}_{\text{random}} \quad \text{for } i = 1, 2, \dots, t \text{ and } j = 1, \dots, n$$

with

fixed

- $\tau_1, \tau_2, \dots, \tau_t$ unknown model parameters
- $E_{11}, \dots, E_{tn} \stackrel{iid}{\sim} N(0, \sigma^2)$

We're not trying to *estimate* linear combos of fixed effects such as $\mu + \tau_1$. Instead, we care about the population from which T_1 was sampled, which is $N(0, \sigma_T^2)$.

One-way random effects model continued

Exercise: Using the random effects model, specify

$$\begin{aligned}
 E(Y_{ij}) &= E(\mu + T_i + \bar{E}_{ij}) = E(\mu) + E(T_i) + E(\bar{E}_{ij}) = \mu + 0 + 0 = \mu \\
 V(Y_{ij}) &= V(\mu + T_i + \bar{E}_{ij}) = 0 + V(T_i) + V(\bar{E}_{ij}) = \sigma_T^2 + \frac{\sigma^2}{5} = V(Y_{ij}) \\
 V(\bar{Y}_{..}) &= V\left(\frac{1}{40} \sum \sum Y_{ij}\right) = V\left(\frac{1}{40} \sum \sum (\mu + T_i + \bar{E}_{ij})\right) = V\left(\mu + \bar{T}_{..} + \bar{\bar{E}}_{..}\right) \\
 &= 0 + \frac{\sigma_T^2}{5} + \frac{\sigma^2}{40}
 \end{aligned}$$

- Two *components* to variability in data: σ^2, σ_T^2
- T_1, T_2, T_3, T_4, T_5 a random sample of sire effects
- Sire effects is a population in its own right.

Contrast this situation with the binding fractions. Why not model antibiotic effects as random? Why fixed?

$$E(\bar{Y}_{..}) = E(\mu + \bar{T}_{..} + \bar{\bar{E}}_{..}) = \mu$$

Model parameters: $\sigma^2, \sigma_T^2, \mu$

Sums of squares, mean squares - same as in fixed effects ANOVA:

$$SS[T] = \sum_i \sum_j (\bar{y}_{i.} - \bar{y}_{..})^2$$

$$SS[E] = \sum_i \sum_j (y_{ij} - \bar{y}_{i.})^2$$

$$SS[Tot] = \sum_i \sum_j (y_{ij} - \bar{y}_{..})^2$$

The ANOVA table is almost the same, it just has a different expected mean squares column:

| Source | SS | df | MS | Expected MS |
|-----------|-----------|---------|-----------|--------------------------|
| Treatment | $SS[T]$ | $t - 1$ | $MS[Trt]$ | $\sigma^2 + n\sigma_T^2$ |
| Error | $SS[E]$ | $N - t$ | $MS[E]$ | σ^2 |
| Total | $SS[Tot]$ | $N - 1$ | | |

BTW, if $H_0 : \sigma_T^2 = 0$, what is $E(MS(Trt))/E(MS(E))$ 

Estimating parameters of one-way random effects model

(Solve a linear system of *estimating equations* obtained by equating statistics to their expected values and solving for unknown parameters:)

$$\begin{aligned}
 E(\bar{\bar{y}}) &= E(\hat{\mu}) = \bar{y} \\
 E(MS(T)) &= \sigma^2 + n\sigma_T^2 = MS(T) \\
 E(MS(E)) &= \sigma^2 = MS(E)
 \end{aligned}$$

set \uparrow & solve

leading to the solution

$$\begin{aligned}
 \hat{\mu} &= \bar{\bar{y}} \\
 \hat{\sigma}^2 &= MS(E) \\
 \hat{\sigma}_T^2 &= \frac{MS(T) - MS(E)}{n}
 \end{aligned}$$

We've derived these estimators:

$$\begin{aligned}\hat{\mu} &= \bar{y}_{..} \\ \hat{\sigma}^2 &= MS[E] \\ \hat{\sigma}_T^2 &= \frac{MS[T] - MS[E]}{n}\end{aligned}$$

For sires data, we observed $\bar{y}_{..} = 82.6$ and

| Source | SS | df | MS | Expected MS |
|--------|-------|----|------|--------------------------|
| Sire | 5591 | 4 | 1398 | $\sigma^2 + 8\sigma_T^2$ |
| Error | 16233 | 35 | 464 | σ^2 |
| Total | 21824 | 39 | | |

leading to the observed estimates

$$\begin{aligned}\hat{\mu} &= \frac{82.6}{1} (lbs) \\ \hat{\sigma}^2 &= \frac{464}{1} (lbs^2) \\ \hat{\sigma}_T^2 &= (1398 - 464) / 8 \\ &= \frac{1168}{8} (lbs^2)\end{aligned}$$

Questions pertaining to this type of study:

Consider the birthweight of a randomly sampled calf

- ① What is the estimated variance of such a calf? $V(\hat{y}_i) = \hat{\sigma}^2 + \hat{\sigma}_\tau^2 = 464 + 116.8 = 581$
- ② Estimate how much of this variation is due to the sire effect. $= 581$
- ③ Estimate how much of this variation is not due to the sire effect.

General questions: (2) $116.8/581 = 20\%$ (3) $464/581 = 80\%$

- ① Is it possible for an estimated variance component to be negative? *yes*
- ② How? *if $MS(T) < MS(E)$*
- ③ What do you do in that case? *Set $\hat{\sigma}_\tau^2 = 0$*

not necessarily dropping the random effect from the model

$$\begin{aligned}
 \text{Var}(Y_{ij}) &= \sigma^2 + \sigma_T^2 \\
 \widehat{\text{Var}}(Y_{ij}) &= \text{MS}(\text{E}) + \frac{\text{MS}(\text{T}) - \text{MS}(\text{E})}{n} \\
 \text{Var}(T_i) / \text{Var}(Y_{ij}) &= \sigma_T^2 / (\sigma^2 + \sigma_T^2) \\
 \text{Var}(E_{ij}) / \text{Var}(Y_{ij}) &= \sigma^2 / (\sigma^2 + \sigma_T^2)
 \end{aligned}$$

- ① Yes, it is possible for $\hat{\sigma}_T^2 < 0$ even though $\sigma_T^2 \geq 0$.
- ② $\hat{\sigma}_T^2 < 0 \Leftrightarrow$ _____?
- ③ Inference concerning σ_T^2 ? _____

Other parameters of interest in random effects models

Coefficient of variation (CV):

$$CV(Y_{ij}) = \frac{\sqrt{\text{Var}(Y_{ij})}}{|E(Y_{ij})|} = ? \frac{\sqrt{\text{Var}(Y_{ij})}}{|E(\mu)|} = \frac{\sqrt{\sigma^2 + \sigma_T^2}}{|\mu|}$$

Note: this is *not* estimated by Coeff Var in PROC GLM output.

Intraclass correlation coefficient

$$\rho_I = \frac{\text{Cov}(Y_{ij}, Y_{ik})}{\sqrt{\text{Var}(Y_{ij})\text{Var}(Y_{ik})}} = \frac{\text{Cov}(\mu + T_i + \bar{E}_{ij}, \mu + T_i + \bar{E}_{ik})}{\sqrt{(\sigma^2 + \sigma_T^2)(\sigma^2 + \sigma_T^2)}} = \frac{\text{Cov}(T_i, T_i)}{\sigma^2 + \sigma_T^2}$$

Handwritten notes:
 $\text{Cov}(T_i, \bar{E}_{ik}) = 0$
 $\frac{\sigma_T^2}{\sigma^2 + \sigma_T^2} \times 100\%$

- Interpretation: the correlation between two responses receiving the same level of the random factor.
- Bigger values of ρ_I correspond to (bigger)/smaller(?) random treatment effects.
- Answers questions like: How much of this variation is due to the sire effect?

For sires,

$$\begin{aligned} \widehat{CV} &= \frac{\sqrt{\hat{\sigma}^2 + \hat{\sigma}_\tau^2}}{\bar{y}} = \frac{\sqrt{464 + 116.8}}{82.6} = 0.29 \text{ or } 29\% \\ \hat{\rho}_I &= \frac{\hat{\sigma}_\tau^2}{\hat{\sigma}^2 + \hat{\sigma}_\tau^2} = \frac{116.8}{464 + 116.8} = 0.20 \end{aligned}$$

Interpretations:

- The estimated standard deviation of a birthweight, 24.1 is 29% of the estimated mean birthweight, 82.6.
- The estimated correlation between any two calves with the same sire for a male parent, or the estimated *intrasire* correlation coefficient, is 0.20