# Chapter 2. Discrete Distributions

Math 3215 Spring 2024

Georgia Institute of Technology

Section 1.
Random Variables of the Discrete
Type

## Random variables

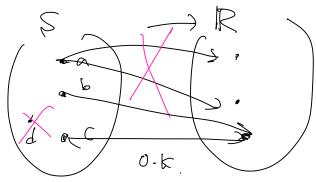
#### **Definition**

the set of all outcomes

Given a random experiment with a sample space S, a function X that assigns one and only one real number X(s) = r to each elements in S is called a random variable.

The space of X is the set of real numbers  $\{x: X(s) = x, s \in S\}$  and denoted by S(X) = the support of X

Male 
$$\longrightarrow -1$$



## Random variables

## Example

A rat is selected at random from a cage and its sex is determined.

The set of possible outcomes is female and male. Thus, the sample space is  $S = \{\text{female}, \, \text{male}\}.$ 

## Random variables

#### **Example**

Consider a random experiment in which we roll a six-sided die.

The sample space associated with this experiment is  $S = \{1, 2, 3, 4, 5, 6\}$ .

Let X(s) = s. Compute  $\mathbb{P}(2 \le X \le 4)$ .

$$X : S^{\downarrow} \to \mathbb{R}$$

ex & = [0,1]: uncontable outcomes

#### **Definition**

Let X be a random variable defined on a sample space S.

If S consists of finite outcomes or countable outcomes, then X is called a discrete  $f: S(x) \xrightarrow{\mathbb{R}} \mathbb{R}$ 

random variable.

$$f(x) = P(X = x)$$

The probability mass function (pmf) of X is only for disorde RV

The cumulative distribution function (cdf) of X is  $\vdash (x) = \mathcal{P} (X \leq x)$ 



$$f(x) = P(X = x)$$

#### **Properties of PMF**

The pmf f(x) of a discrete random variable X is a function that satisfies the following properties:

- $f(x) \ge 0$  for all x,  $= \mathbb{P}(S^f)$   $\sum_{x \in S(X)} f(x) = 1$ , and
- $\mathbb{P}(X \in A) = \sum_{x \in A} f(x)$ .

 $5 = \{1, 2, 3, 4, 5, 6\}$ come. 1 = 2Roll a die, let X be the outcome.

Find the pmf and the cdf of X.

PMF: 
$$J(x) = P(X = x) = \frac{1}{6}$$
 for  $x = 1, ---.6$ 

CDF:  $F(x) = P(X \le x) = 0$  of  $X < 1$ 

$$F(-1) = P(X \le -1) = 0$$

$$F(0) = P(X \le 0) = 0$$

$$\frac{2}{6}$$
 for  $1 \le x < 2$ 

$$\frac{2}{6}$$
 for  $3 \le x < 4$ 

$$\frac{2}{6}$$
 for  $1 \le x < 6$ 

## Example

Roll a fair four-sided die twice.

Let X equal the larger of the two outcomes if they are different and the common value if they are the same.

 $5 = \begin{cases} (1,1) & (1,2), --- \end{cases}$  (16 octomes

Find the pmf and the cdf of X. 
$$S(x) = \{1, 2, 3, 4\}$$

$$PMF = F(x) = P(x = x) = \begin{cases} 1/16 & x = 1 \\ 3/16 & x = 2 \\ 5/16 & x = 3 \end{cases}$$

$$P(x = 3) = P(x = 3) =$$

RV: 
$$X: S^1 \rightarrow \mathbb{R}$$

Souple.

$$S^1(X) = \{S: X = S \mid P(X = X)\}$$

PMF
$$F(X) = F_X(X) = P(X \leq X)$$

CDF
$$F(X) = F_X(X) = P(X \leq X)$$

# Bar graph, Probability histogram, relative frequency histogram

$$5 = \{ (1,1) (1,2) (1,3), ---- \}$$
 $2$ 
 $3$ 
 $4$ 

#### **Example**

A fair four-sided die with outcomes 1, 2, 3, and 4 is rolled twice.

Let X equal the sum of the two outcomes.

$$S(x) = \begin{cases} 2, 3, - \cdots, 8 \end{cases}$$

$$PMF \quad f(x) = \begin{cases} 1/16 & x = 2 \\ 2/16 & x = 3 \\ 3/16 & x = 4 \end{cases}$$

$$4/16 & x = 5 \\ 7/16 & x = 6 \end{cases}$$

$$2/16 & x = 8$$

$$0 \quad \text{otherwise}$$

# Bar graph, Probability histogram, relative frequency histogram

## Example

Two fair four-sided dice are rolled. Write down the sum of the two outcomes. Repeat this 1000 times.

| The sum of two outcomes | 2          | 3   | 4          | 5   | 6   | 7   | 8    |            |
|-------------------------|------------|-----|------------|-----|-----|-----|------|------------|
| Number of Observations  | 71         | 124 | 194        | 258 | 177 | 122 | 54   | relative   |
| '                       | <b>→</b> 1 | ام  | /r         |     |     |     | F1 8 | frequency. |
|                         | (000       | 12  | 4 <u> </u> |     |     |     | 1000 | , 0 1      |



Section 2. Mathematical Expectation

## **Definition of Expectation**

$$\mathbb{P}(u(x)) = \mathbb{P}(A) \cdot 1 + \mathbb{P}(B) \cdot 2 + \mathbb{P}(c) \cdot 3$$

Zu(x) P(X=x)

#### **Example**

Consider the following game. A player roll a fair die.

If the event  $A = \{1, 2, 3\}$  occurs, he receives one dollar.

If  $B = \{4, 5\}$  occurs, he receives two dollars.

If  $C = \{6\}$  occurs, he receives three dollars.

If the game is repeated a large number of times, what is the average payment?

$$= \frac{\sum_{x \in S(x)} x \cdot P(x = x)}{x \cdot f(x)} = E[x]$$

$$= \frac{\sum_{x \in S(x)} x \cdot f(x)}{\sum_{x \in S(x)} \sum_{x \in S(x)} x \cdot f(x)} = \frac{E[x]}{\sum_{x \in S(x)} x \cdot f(x)}$$

$$= \frac{\sum_{x \in S(x)} x \cdot P(x = x)}{\sum_{x \in S(x)} x \cdot f(x)} = \frac{E[x]}{\sum_{x \in S(x)} x \cdot f(x)}$$

$$= \frac{\sum_{x \in S(x)} x \cdot P(x = x)}{\sum_{x \in S(x)} x \cdot f(x)} = \frac{E[x]}{\sum_{x \in S(x)} x \cdot f(x)}$$

$$= \frac{E[x]}{\sum_{x \in S(x)} x \cdot f(x)} = \frac{E[x]}{\sum_{x \in S(x)} x \cdot f(x)}$$

## **Definition of Expectation**

#### **Definition**

If f(x) is the pmf of a discrete random variable X with the space S(X), and if the summation

$$\sum_{x \in S(X)} u(x) f(x)$$

exists, then the sum is called the mathematical expectation or the expected value of u(X), and denoted by  $\mathbb{E}[u(X)]$ .

$$\frac{EX}{E[X]} = x_1 \cdot P(X = x_1) + x_2 P(X = x_2) + \cdots - E[X^2] = x_2^2 \cdot P(X = x_1) + x_2^2 P(X = x_2) + \cdots - \cdots$$

## **Definition of Expectation**

#### **Example**

Let the random variable X have the pmf  $f(x) = \frac{1}{3}$  for  $x \in \{-1, 0, 1\} = S(X)$ .

Let  $Y = u(X) = X^2$ .

Find the pmf of Y and  $\mathbb{E}[Y] = \mathbb{E}[X^2]$ .

$$f_{Y}(y) = P(Y = y) = \int \frac{1}{3} \qquad y = 0$$

$$\int \frac{1}{3} \qquad y = 1$$
o otherwise

$$F[Y] = \sum_{y \in S(Y)} y \cdot f_{Y}(y) = 0 \cdot f_{Y}(0) + 1 - f_{Y}(1) = \frac{2}{3}$$
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$$E[x^{2}] = \sum_{x \in S(x)}^{2} x^{2} \cdot f_{X}(x)$$

$$= (-1)^{2} \cdot f_{X}(-1) + 0^{2} \cdot f_{X}(0) + 1^{2} \cdot f_{X}(1) = 3$$

## **Properties of Expectation**

#### Theorem

- 1. If c is a constant, then  $\mathbb{E}[c] = c$ .
- 2. If c is a constant and u is a function, then  $\mathbb{E}[cu(X)] = c\mathbb{E}[u(X)]$ .
- 3. If  $c_1$  and  $c_2$  are constants and  $u_1$  and  $u_2$  are functions. then

$$\mathbb{E}[c_1u_1(X) + c_2u_2(X)] = c_1\mathbb{E}[u_1(X)] + c_2\mathbb{E}[u_2(X)].$$

$$E[X(X-2)] = Z[X(X-2)] + (X-2) = E[X^2] - E[2X]$$

$$= E[X^2] - 2 \cdot E[X]$$

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## **Properties of Expectation**

Example

Let 
$$X$$
 have the pmf  $f(x) = \frac{x}{10}$  for  $x = 1, 2, 3, 4$ .

Find  $\mathbb{E}[X]$ ,  $\mathbb{E}[X^2]$  and  $\mathbb{E}[X(5-X)]$ .

$$\begin{cases}
\frac{3}{\sqrt{5}} & \chi = 2 \\
\frac{3}{\sqrt{5}} & \chi = 3 \\
\frac{4}{\sqrt{5}} & \chi = 4
\end{cases}$$

$$\mathbb{E}[X] = 1 \cdot \frac{1}{6} + 2 \cdot \frac{2}{6} + 3 \cdot \frac{3}{6} + 4 \cdot \frac{4}{6}$$

$$= \frac{1}{6} \cdot (1^{2} + 2^{2} + 3^{2} + 4^{2}) = 3$$

$$\mathbb{E}[X^{2}] = 1^{2} \cdot \frac{1}{6} + 2^{2} \cdot \frac{2}{6} + 3^{2} \cdot \frac{3}{6} + 4^{2} \cdot \frac{4}{6}$$

$$= \frac{1}{6} \cdot (1^{3} + 2^{2} + 3^{3} + 4^{3}) = 10$$

$$= \frac{1}{6} \cdot (1^{3} + 2^{2} + 3^{3} + 4^{3}) = 10$$

$$\mathbb{E}[x(z-x)] = \mathbb{E}[z-x] = z \cdot \mathbb{E}[x] - \mathbb{E}[x^2]$$

Note
$$E(x^2) \neq (E(x))^2$$

$$E(u(x)) \neq u(E(x))$$

$$E[X] = \sum_{x \in S(X)} \frac{x - f(x)}{x - f(x)}$$

$$E[u(X)] = \sum_{x \in S(X)} u(x) f(x)$$
New RV

# **Properties of Expectation**

$$S(x) = \{1, 2, 3, ---- \}$$

#### **Example**

An experiment has probability of success p(0,1) and probability of failure q=1-p.

This experiment is repeated independently until the first success occurs.

Let X be the number of trials. Find  $\mathbb{E}[X]$ .

$$E[X] = b$$

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$$X=1:H$$
  $f(1)=P(X=1)=p$   
 $X=2:TH$   $f(2)=P(X=2)=(I-p)p$   
 $X=3:TH$   $f(3)=(I-p)p$   
 $X=3:TH$   $f(4)=(I-p)p$ 

$$(A) = E[x] = 1 \cdot P + (2)(1-p)P + 3 \cdot (1-p)P + 4(1-p)P + \cdots$$

$$(4)(1-p)P + 2(1-p)P + 3(1-p)P + 3(1-p)P + \cdots$$

$$A - (I-p)A = 1 \cdot p + 1 \cdot (I-p) \cdot p + 1 \cdot (I-p)^{3} p + 1 \cdot (I-p)$$

If trial 
$$A$$
 H,  $E[\# \text{ of frians}] = 1$   
 $A$  = 1 + A  
 $A$  =  $A$ 

Section 3.

Special Mathematical Expectations

The expectation or mean of a random variable X is

$$\mu = \mathbb{E}[X'] = \sum x f(x).$$

This is also called the first moment about the origin. The first moment about the mean 
$$\mu$$
 is  $\mathbb{E}[X-\mu] = \mathbb{E}[X] - \mathbb{E}[X] = \mathbb{E}[X] - \mu = 0$  If moment of  $X$  about  $b = \mathbb{E}[X-b]$ 

The second moment of X about b is  $\mathbb{E}[(X-b)^2]$ .

If  $b = \mu$ , it is also called the variance of X and denoted by  $Var(X) = \sigma^2$ .

Its positive square root is **the standard deviation** of X and denoted by  $Std(X) = \sigma$ .

$$M = MX = \mathbb{E}(X) \qquad \qquad \qquad \qquad = \mathbb{E}[(X - M)^2] = \text{Vor}(X)$$

$$= \mathbb{E}[(X - \mathbb{E}(X))^2]$$

Std (X) = 
$$\sigma_{x} = \sigma = \sqrt{\mathbb{E}[(x-\mu)^{2}]}$$

with pubability

with probability

$$X = \begin{cases} 1 & \omega.p. \frac{1}{2} \\ -1 & \omega.p. \end{cases} \qquad \begin{cases} f_{\chi}(x) = \frac{1}{2} & \text{for } x=1,-1 \\ 0 & \text{for } x=1,-1 \end{cases}$$

$$A = \begin{pmatrix} 1 & 1 & 2 \\ 1 & 1 & 2 \end{pmatrix}$$

$$\mathbb{E}[X] = M_{X} = 1 \cdot \frac{1}{L} + (-1) \cdot \frac{1}{L} = 0 \quad \text{i} \quad \mathbb{E}[Y] = (0 \cdot \frac{1}{L} + (-10) \cdot \frac{1}{L} = 0$$

$$\text{Var}(X) = \mathbb{E}[(X - \mu_{X})^{2}] = \mathbb{E}[X^{2}] \quad \text{var}(Y) = \mathbb{E}[(Y - \mu_{Y})^{2}] = \mathbb{E}[X^{2}]$$

$$Y = \begin{cases} 10 & \text{w.p.} & \frac{1}{2} \\ -10 & \text{w.p.} & \frac{1}{2} \end{cases}$$
 $(x,y) = \frac{1}{2} (x,y) = \frac{1}{2} ($ 

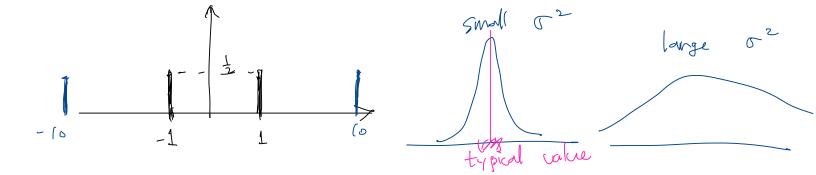
$$[Y] = [0 \cdot \frac{1}{2} + (-10) \cdot \frac{1}{2} = 0]$$

$$[Var(Y) = \mathbb{E}[(Y-\mu_Y)^2] = \mathbb{E}[Y]$$

$$Var(X) = \mathbb{E}[(X - \mu_X^2) = \mathbb{E}[X^2]$$

$$= 1^2 \cdot \frac{1}{2} + (-1)^2 \cdot \frac{1}{2} = 1$$

$$= 10 \cdot \frac{1}{2} + (-10)^2 \cdot \frac{1}{2} = 100$$



#### **Example**

Roll a fair die and let X be the outcome.

Find  $\mathbb{E}[X]$  and Var(X).

$$E[X] = 1 \cdot \frac{1}{6} + 2 \cdot \frac{1}{6} + 3 \cdot \frac{1}{6} + 4 \cdot \frac{1}{6} + 5 \cdot \frac{1}{6} + 6 \cdot \frac{1}{6}$$

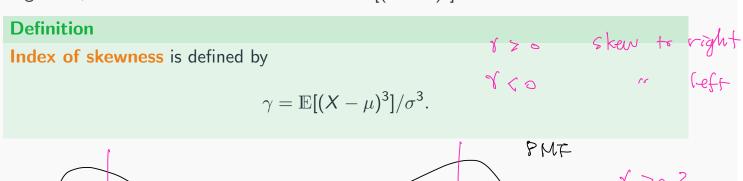
$$= \frac{1}{6} \cdot (1 + 2 + - - + 6) = \frac{21}{6} = \frac{7}{2} = M$$

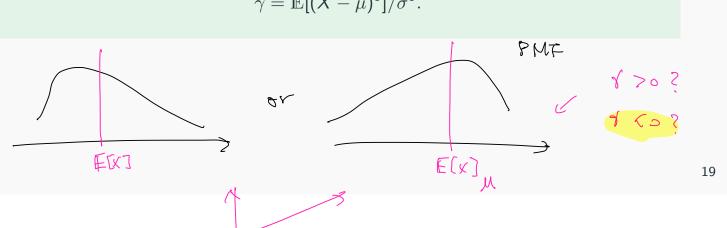
$$Var(X) = E[(X - M)^2] = E[(X - \frac{7}{2})^2]$$

$$= (1 - \frac{7}{2})^{2} - \frac{1}{6} + (2 - \frac{7}{2})^{2} - \frac{1}{6} + \cdots + (6 - \frac{7}{2})^{2} + (\frac{3}{2})^{2} + (\frac{3}{$$

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In general, the r-th moment of X about b is  $\mathbb{E}[(X-b)^r]$ .





$$\int = \underbrace{\mathbb{E}\left(\left(X - W\right)^{3}\right)}_{C}$$

#### **Example**

Let  $f(x) = \frac{4-x}{6}$  for x = 1, 2, 3 be the pmf of X. Compute the index of skewness.

$$E[X] = \frac{5}{3} = 1 \cdot \frac{(4-1)}{6} + 2 \cdot \frac{(4-2)}{6} + 3 \cdot \frac{(4-3)}{6}$$

$$= \frac{1}{6} \cdot (3 + 2 \cdot 2 + 3 \cdot 1) = \frac{10}{6} = \frac{5}{3}.$$

$$E^{2} = Vov(X) = E[(X - \frac{5}{3})^{2} \cdot \frac{3}{6} + (2 - \frac{5}{3})^{2} \cdot \frac{2}{6} + (3 - \frac{5}{3})^{2} \cdot \frac{1}{6}^{20}$$

$$= \frac{4}{9} \cdot \frac{1}{2} + \frac{1}{9} \cdot \frac{1}{3} + \frac{16}{9} \cdot \frac{1}{6}$$

$$= \frac{1}{9} \cdot \frac{1}{3} \cdot (2 + 2 + 16) = \frac{5}{3} \cdot \frac{3}{6} \cdot \frac{2}{6} + (3 - \frac{5}{3})^{2} \cdot \frac{2}{6} + (3 - \frac{5}{3})^{2} \cdot \frac{2}{6}$$

$$E[(X - M)^{3}] = (1 - \frac{5}{3})^{3} \cdot \frac{3}{6} + (2 - \frac{5}{3})^{3} \cdot \frac{2}{6} + (3 - \frac{5}{3})^{2} \cdot \frac{2}{6}$$

 $=\frac{1}{6} \cdot \left[ -\frac{8}{27} \cdot 3 + \frac{1}{27} \cdot 2 + \frac{64}{27} \cdot 1 \right]$ 

$$C^2 = \sum_{n=0}^{\infty} (x-n)^2 f(x)$$

$$0 \leqslant \sigma^2 = \mathbb{E}[(X - \mu)^2] = \mathbb{E}[X^2] - \mu^2 \qquad = \quad \mathbb{E}(X^2) - \left(\mathbb{E}[X]\right)$$

$$= \mathbb{E}\left[\chi^2 - 2\sqrt{\chi_{-M}} + M^2\right]$$

$$= \mathbb{E}(X_5) - 5N_5 + N_7 = \mathbb{E}(X_5) - N_5$$

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$$E[X^2] > (E[X])$$

## Moment generating functions

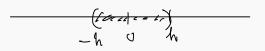
For 
$$u(x) = e^{+x}$$
,  $E[e^{+X}] = E[u(x)]$ 

#### **Definition**

Let X be a discrete random variable and assume that there exists h > 0 such that

$$\sum_{x \in A} \frac{\mathbb{E}[e^{tX}]}{x^{x}} = \sum_{x \in A} e^{tx} f(x)$$

 $\lim_{t\to\infty} \int_{\text{overly}} \mathbb{E}[e^{tX}] = \sum_{t\to\infty} e^{tx} f(x)$  is finite for all  $t\in(-h,h)$ . Then,  $M(t)=\mathbb{E}[e^{tX}]$  is called **the moment generating** function (mgf).



## Moment generating functions

#### **Properties**

- 1. M(0) = 1
- 2.  $M'(0) = \mathbb{E}[X]$
- 3.  $M''(0) = \mathbb{E}[X^2]$
- 4. In general,  $M^{(r)}(0) = \mathbb{E}[X^r]$ .

$$M(o) = \mathbb{E}[e^{o \cdot X}] = \mathbb{E}[1] = 1$$

$$M(o) = \frac{d}{dt} \mathbb{E}[e^{t \cdot X}] = \mathbb{E}\left[\frac{d}{dt}(e^{t \cdot X})\right]|_{t=0}$$

$$= \mathbb{E}(X)$$

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## Moment generating functions

## **Example**

Let  $f(x) = q^{x-1}p$  where  $p \in (0,1)$  and q = 1 - p.

Compute M(t).

$$\mathbb{E}(x) = ?$$

$$M(t) = E(e^{tX}) = e^{t-1} \cdot g^{0} \cdot p + e^{t-2} g^{-1} \cdot p + e^{t-3} \cdot g^{-2} \cdot p + e^{t$$

Section 4.
The Binomial Distribution

## Bernoulli random variables

A Bernoulli experiment, more commonly called a Bernoulli trial, is a random experiment with two outcomes.

Say  $S = \{\text{success, failure}\}$  and  $\mathbb{P}(\text{success}) = p$  for some  $p \in (0,1)$ . Then  $\mathbb{P}(\text{failure}) = q = 1 - p$ .

A random variable X is a **Bernoulli random variable** with success probability p is X = 1 if success and 0 otherwise.

# Bernoulli random variables

## Theorem

Let X be a Bernoulli random variable with success probability p.

$$\mathbb{E}[X] =$$

$$Var[X] =$$

Consider a sequence of independent Bernoulli experiments with success probability p.

Let X be the number of success trials in the first n experiments.

This is called a **Binomial random variable** with the number of trials n and success probability p.

We use the notation  $X \sim b(n, p) = \text{Bin}(n, p)$ .

#### **Theorem**

Let X a binomial random variable with the number of trials n and success probability p.

The pmf of X is

$$\mathbb{E}[X] =$$

$$Var[X] =$$

## **Example**

Out of millions of instant lottery tickets, suppose that 20% are winners. If eight such tickets are purchased, what is the probability of purchasing two winning ticket?

#### **Example**

H5N1 is a type of influenza virus that causes a severe respiratory disease in birds called avian influenza (or "bird flu").

Although human cases are rare, they are deadly; according to the World Health Organization the mortality rate among humans is 60%.

Let X equal the number of people, among the next 25 reported cases, who survive the disease.

Assuming independence, the distribution of X is b(25, 0.4). What is the probability that ten or fewer of the cases survive?

#### **Theorem**

The  $\operatorname{mgf}$  of a binomial random variable X is

$$M(t) =$$

### Binomial random variables

#### **Exercise**

It is believed that approximately 75% of American youth now have insurance due to the health care law.

Suppose this is true, and let X equal the number of American youth in a random sample of n=15 with private health insurance.

How is X distributed? Find the probability that X is at least 10. Find the mean, variance, and standard deviation of X.

Section 5.
The Hypergeometric Distribution

There is a collection of  $N_1$  red balls and  $N_2$  blue balls.

Sample *n* balls at random without replacement  $(n \le N_1 + N_2)$ .

Let X be the number of red balls chosen.

Then, X is called a **hypergeometric random variable** with parameters  $N_1$ ,  $N_2$ , n, and denoted by  $HG(N_1, N_2, n)$ .

### **Example**

In a small pond there are 50 fish, ten of which have been tagged.

If a fisherman's catch consists of seven fish selected at random and without replacement, and X denotes the number of tagged fish,

what is the probability that exactly two tagged fish are caught?

### **Theorem**

$$\mathbb{P}(X=k)=$$

$$\mathbb{E}[X] = n \frac{N_1}{N_1 + N_2}$$

$$Var[X] = n \frac{N_1}{N_1 + N_2} \frac{N_2}{N_1 + N_2}$$

### Exercise

In a lot (collection) of 100 light bulbs, there are five bad bulbs.

An inspector inspects ten bulbs selected at random.

Find the probability of finding at least one defective bulb.

Section 6.
The Negative Binomial Distribution

## Geometric random variables

Consider a sequence of independent Bernoulli trials with success probability

Let *X* be the number of trials until the first success.

This random variable is called a **geometric random variable**.

# Geometric random variables

### Theorem

The pmf of X is

$$\mathbb{E}[X] = \frac{1}{p}$$

$$Var[X] = \frac{q}{p^2}$$

$$M(t) = \frac{pe^t}{1-(1-p)e^t}$$

### Geometric random variables

### Example

Some biology students were checking eye color in a large number of fruit flies.

For the individual fly, suppose that the probability of white eyes is 1/4 and the probability of red eyes is 3/4, and that we may treat these observations as independent Bernoulli trials.

What is the probability that at least four flies have to be checked for eye color to observe a white-eyed fly?

Consider a sequence of independent Bernoulli trials with success probability

Let *X* be the number of trials until the *r*-th success.

This random variable is called a negative binomial random variable.

#### **Theorem**

The pmf of X is

$$f(k) = \binom{k-1}{r-1} p^r (1-p)^{k-r}$$

for  $k = r, r + 1, \cdots$  and otherwise zero.

$$\mathbb{E}[X] = \frac{r}{p}$$

$$Var[X] = \frac{rq}{p^2}$$

$$M(t) = \left(rac{pe^t}{1-(1-p)e^t}
ight)^r$$

A negative binomial random variable can be written as a sum of independent geometric random variables.

### **Example**

Suppose that during practice a basketball player can make a free throw 80% of the time.

Furthermore, assume that a sequence of free-throw shooting can be thought of as independent Bernoulli trials.

Let X equal the minimum number of free throws that this player must attempt to make a total of ten shots.

Find the mean of X.

#### **Exercise**

One of four different prizes was randomly put into each box of a cereal.

If a family decided to buy this cereal until it obtained at least one of each of the four different prizes, what is the expected number of boxes of cereal that must be purchased?

Section 7.
The Poisson Distribution

Some experiments result in counting the number of times particular events occur at given times or with given physical objects.

#### **Example**

- the number of cell phone calls passing through a relay tower between 9 and 10am.
- the number of flaws in 100 feet of wire
- the number of customers that arrive at a ticket window between noon and 2pm.
- the number of defects in a 100-foot roll of aluminum screen that is 2 feet wide.

Counting such events can be looked upon as observations of a random variable associated with an **approximate Poisson process**, provided that the conditions in the following definition are satisfied.

Let the number of occurrences of some event in a given continuous interval be counted. Then we have an approximate Poisson process with parameter  $\lambda > 0$  if

- The numbers of occurrences in nonoverlapping subintervals are independent.
- The probability of exactly one occurrence in a sufficiently short subinterval of length h is approximately  $\lambda h$ .
- The probability of two or more occurrences in a sufficiently short subinterval is essentially zero.

Under these assumption, consider the number of occurrences in a time interval [0,1].

Split [0,1] into n subintervales  $[0,\frac{1}{n}],[\frac{1}{n},\frac{2}{n}],\cdots,[\frac{n-1}{n},1].$ 

In each subinterval, at most one event occurs with probability  $\frac{\lambda}{n}$ .

Thus, the number of occurrences is a binomial random variable with n nad  $\frac{\lambda}{n}$ .

As  $n \to \infty$ , the random variable gets close to some random variable X.

We say X is a Poisson random variable with parameter  $\lambda$  if its pmf is

$$\mathbb{P}(X=k) = \frac{e^{-\lambda}\lambda^k}{k!}$$

for  $k = 0, 1, 2, \cdots$ .

### Theorem

 $\mathbb{E}[X] =$ 

Var[X] =

M(t) =

### **Example**

In a large city, telephone calls to 911 come on the average of two every 3 minutes.

If one assumes an approximate Poisson distribution, what is the probability of five or more calls arriving in a 9 minute period?

# Poisson Approximation to Binomial

Supose X is a binomial random variable b(n, p), n is large, and p is small but np converges to some constant, say  $\lambda$ .

In this case, X can be approximated by a Poisson random variable with parameter  $\lambda$ .

This approximation is quite accurate if  $n \ge 20$ ,  $p \le 0.05$  or  $n \ge 100$ ,  $p \le 0.1$ .

## Poisson Approximation to Binomial

#### **Example**

A manufacturer of Christmas tree light bulbs knows that 2% of its bulbs are defective.

Assuming independence, the number of defective bulbs in a box of 100 bulbs has a binomial distribution with parameters n=100 and p=0.02.

Find the probability that a box of 100 of these bulbs contains at most three defective bulbs.

# Poisson Approximation to Binomial

#### Exercise

Suppose that the probability of suffering a side effect from a certain flu vaccine is 0.005. If 1000 persons are vaccinated, approximate the probability that (a) At most one person suffers. (b) Four, five, or six persons suffer.