# Adaptive Shrinkage and False Discovery Rates by Laplace Approximation

Matthew Stephens

2013/5/13

# Outline

Prelude

## Outline

- Prelude
- Allegro (ma non troppo)

## Outline

- Prelude
- Allegro (ma non troppo)
- Coda

• Consider testing the null hypothesis  $H_0: \beta = 0$ , vs the alternative  $H_1: \beta \neq 0$  in the logistic regression model:

$$\log \frac{p(Y_i = 1 | X_i = x)}{p(Y_i = 0 | X_i = x)} = \mu + x\beta$$

• Consider testing the null hypothesis  $H_0: \beta = 0$ , vs the alternative  $H_1: \beta \neq 0$  in the logistic regression model:

$$\log \frac{p(Y_i = 1 | X_i = x)}{p(Y_i = 0 | X_i = x)} = \mu + x\beta$$

Specifically, consider computing the Bayes Factor

$$BF:=\frac{p(Y|X,H_1)}{p(Y|X,H_0)}.$$

• Consider testing the null hypothesis  $H_0$ :  $\beta = 0$ , vs the alternative  $H_1$ :  $\beta \neq 0$  in the logistic regression model:

$$\log \frac{p(Y_i = 1 | X_i = x)}{p(Y_i = 0 | X_i = x)} = \mu + x\beta$$

Specifically, consider computing the Bayes Factor

$$BF:=\frac{p(Y|X,H_1)}{p(Y|X,H_0)}.$$

 In genome-wide association studies, we may wish to do this for millions of different genetic variants (X).

$$BF = \frac{\int p(Y|\mu, \beta, X)p_1(\mu, \beta|X) d\mu d\beta}{\int p(Y|\mu, \beta = 0, X)p_0(\mu|X) d\mu},$$

where  $p_0$  and  $p_1$  denote priors under  $H_0$  and  $H_1$ .

 These integrals generally don't have closed forms, but being low-dimensional they are simple to approximate.

$$BF = \frac{\int p(Y|\mu, \beta, X)p_1(\mu, \beta|X) d\mu d\beta}{\int p(Y|\mu, \beta = 0, X)p_0(\mu|X) d\mu},$$

where  $p_0$  and  $p_1$  denote priors under  $H_0$  and  $H_1$ .

- These integrals generally don't have closed forms, but being low-dimensional they are simple to approximate.
- For  $p_1: \beta \sim N(0,\phi^2)$ , Wakefield, 2009 (see also Johnson, 2008) suggested a particularly simple *Approximate Bayes Factor* (ABF) based on the maximum likelihood estimate,  $\hat{\beta}$ , and its (estimated) standard error s.

$$ABF = \sqrt{1 - k} \exp(0.5kT^2)$$

where  $k := \phi^2/(s^2 + \phi^2)$  and  $T := \hat{\beta}/s$ .

• ABF arises if we assume  $\hat{\beta}|s, \beta \sim N(\beta, s^2)$  and treat  $\hat{\beta}$  as the observed "data".

$$ABF = \sqrt{1 - k} \exp(0.5kT^2)$$

where  $k := \phi^2/(s^2 + \phi^2)$  and  $T := \hat{\beta}/s$ .

- ABF arises if we assume  $\hat{\beta}|s,\beta\sim N(\beta,s^2)$  and treat  $\hat{\beta}$  as the observed "data".
- Equivalently ABF can be derived as a "Laplace approximation", approximating the likelihood  $L(\beta)$  as Normal, centered on  $\hat{\beta}$ , with variance  $s^2$ :

$$L(\beta) \propto \exp[-0.5(\beta - \hat{\beta})^2/s^2].$$



 This is not the moxt accurate Laplace approximation one might consider.

- This is not the moxt accurate Laplace approximation one might consider.
- However, it has some nice features.

- This is not the moxt accurate Laplace approximation one might consider.
- However, it has some nice features.
  - The approximation is independent of prior.

- This is not the moxt accurate Laplace approximation one might consider.
- However, it has some nice features.
  - The approximation is independent of prior.
  - ullet Applicable to any regression where  $\hat{eta}$  and s are available.

- This is not the moxt accurate Laplace approximation one might consider.
- However, it has some nice features.
  - The approximation is independent of prior.
  - ullet Applicable to any regression where  $\hat{eta}$  and s are available.
  - Easily computed using results of standard software or published analyses (e.g. CI).

- This is not the moxt accurate Laplace approximation one might consider.
- However, it has some nice features.
  - The approximation is independent of prior.
  - ullet Applicable to any regression where  $\hat{eta}$  and s are available.
  - Easily computed using results of standard software or published analyses (e.g. CI).
- A simple transformation of T can improve accuracy for small samples (analogous to t test vs Z test); Wen and Stephens, Arxiv.

#### Extensions of ABF

 Similar ideas can be used to compute ABFs in slightly more complex settings.

#### Extensions of ABF

- Similar ideas can be used to compute ABFs in slightly more complex settings.
- Eg In Wen and Stephens, we consider S subgroups, and approximate the BF for  $H_0: \beta_s = 0$  for all s, vs a general alternative  $H_0: \beta_s \neq 0$ .

# Allegro (ma non troppo)

• The problem: you have imperfect measurements of many "similar" things, and wish to estimate their values.

# Allegro (ma non troppo)

- The problem: you have imperfect measurements of many "similar" things, and wish to estimate their values.
- Particularly common in genomics. For example, a very common goal is to compare the mean expression (activity) level of many genes in two conditions.

## Example: Mouse Heart Data

 Data on 150 mouse hearts, dissected into left and right ventricle (courtesy Scott Schmemo, Marcelo Nobrega)

```
##
        gene
              lv1 lv2 rv1
                                 rv2 genelength
       Itm2a 2236 2174 9484 10883
                                            1626
## 1
##
   2
      Sergef
                97
                     90
                          341
                                 408
                                            1449
    Fam109a 383
                         1864
                    314
                                2384
                                            2331
##
        Dhx9 2688 2631
                        18501
                               20879
                                            4585
               762
                    674
## 5
       Ssu72
                         2806
                                3435
                                            1446
              736
                    762
## 8
      Eif2b2
                         3081
                                3601
                                            1565
```

• Standard practice: analyses use False Discovery Rates

- Standard practice: analyses use False Discovery Rates
  - e.g. Benjamini and Hochberg, 1995; Storey and Tibshirani, 2003, which have roughly 18k and 4k citations respectively!

- Standard practice: analyses use False Discovery Rates
  - e.g. Benjamini and Hochberg, 1995; Storey and Tibshirani, 2003, which have roughly 18k and 4k citations respectively!
- Typical analysis proceeds roughly as follows:

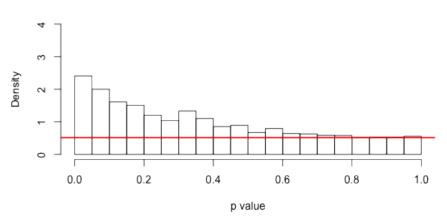
- Standard practice: analyses use False Discovery Rates
  - e.g. Benjamini and Hochberg, 1995; Storey and Tibshirani, 2003, which have roughly 18k and 4k citations respectively!
- Typical analysis proceeds roughly as follows:
  - Estimate an effect size  $\beta_j$  and standard error  $s_j$  for each gene.

- Standard practice: analyses use False Discovery Rates
  - e.g. Benjamini and Hochberg, 1995; Storey and Tibshirani, 2003, which have roughly 18k and 4k citations respectively!
- Typical analysis proceeds roughly as follows:
  - Estimate an effect size  $\beta_j$  and standard error  $s_j$  for each gene.
  - Convert this to a p value for each gene, e.g. by a t test on  $\beta_j/s_j$ .

- Standard practice: analyses use False Discovery Rates
  - e.g. Benjamini and Hochberg, 1995; Storey and Tibshirani, 2003, which have roughly 18k and 4k citations respectively!
- Typical analysis proceeds roughly as follows:
  - Estimate an effect size  $\beta_j$  and standard error  $s_j$  for each gene.
  - Convert this to a p value for each gene, e.g. by a t test on  $\beta_j/s_j$ .
  - Use the distribution of p values to estimate the false discovery rate (FDR) at a given threshold.

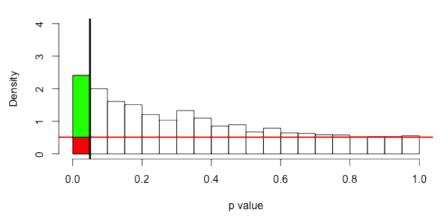
# False Discovery Rates

#### p value distribution, all genes



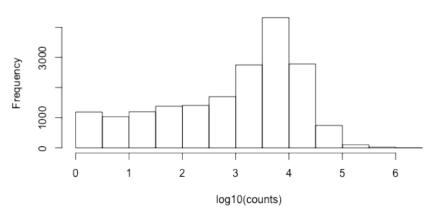
# False Discovery Rates

#### p value distribution, all genes



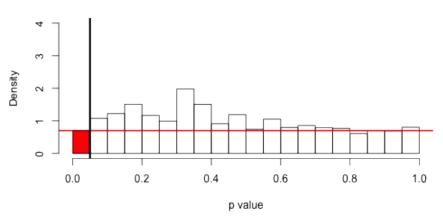
# FDR problem: different genes have different precision/power

#### Counts vary considerably across genes



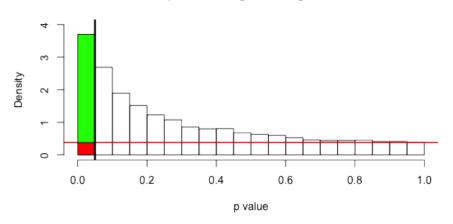
# FDR problem: lower count genes, less power, add noise





# FDR problem: higher count genes, more power

#### p values, high count genes



# Adaptive Shrinkage

• Fundamental idea: use hierarchical modelling so measurements of  $\beta_j$  for each gene improve inference for  $\beta$  at other genes.

# Adaptive Shrinkage

- Fundamental idea: use hierarchical modelling so measurements of  $\beta_j$  for each gene improve inference for  $\beta$  at other genes.
- Despite a long-standing literature on these types of methods e.g. Greenland and Robins 1991, Efron and Tibshirani 2002, Gelman
  et al 2012 they are much less widely used (in genomics at least).

# Adaptive Shrinkage

- Fundamental idea: use hierarchical modelling so measurements of  $\beta_j$  for each gene improve inference for  $\beta$  at other genes.
- Despite a long-standing literature on these types of methods e.g. Greenland and Robins 1991, Efron and Tibshirani 2002, Gelman
  et al 2012 they are much less widely used (in genomics at least).
- Possibly this is due, in part, to the lack of a simple, flexible, and generic implementation?

• Summarize data on each gene by two numbers,  $\hat{\beta}_j$  and its standard error  $s_i$ . (a la Wakefield; Greenland and Robins 1991)

- Summarize data on each gene by two numbers,  $\hat{\beta}_j$  and its standard error  $s_j$ . (a la Wakefield; Greenland and Robins 1991)
- Approximate likelihood for  $\beta_j$  by

$$L(\beta_j) \propto \exp(-0.5(\beta_j - \hat{\beta}_j)^2/s_j^2).$$

("Laplace Approximation")

- Summarize data on each gene by two numbers,  $\hat{\beta}_j$  and its standard error  $s_i$ . (a la Wakefield; Greenland and Robins 1991)
- Approximate likelihood for  $\beta_j$  by

$$L(\beta_j) \propto \exp(-0.5(\beta_j - \hat{\beta}_j)^2/s_j^2).$$

("Laplace Approximation")

• Borrow information by assuming  $\beta_j$  are iid  $\sim g(\cdot; \pi)$ , where  $\pi$  are hyperparameters to be estimated.

- Summarize data on each gene by two numbers,  $\hat{\beta}_j$  and its standard error  $s_j$ . (a la Wakefield; Greenland and Robins 1991)
- Approximate likelihood for  $\beta_j$  by

$$L(\beta_j) \propto \exp(-0.5(\beta_j - \hat{\beta}_j)^2/s_j^2).$$

("Laplace Approximation")

- Borrow information by assuming  $\beta_j$  are iid  $\sim g(\cdot; \pi)$ , where  $\pi$  are hyperparameters to be estimated.
- Letting  $g(\cdot; \pi)$  be a mixture of normal distributions provides both flexibility, and analytic calculations.

- Summarize data on each gene by two numbers,  $\hat{\beta}_j$  and its standard error  $s_i$ . (a la Wakefield; Greenland and Robins 1991)
- Approximate likelihood for  $\beta_j$  by

$$L(\beta_j) \propto \exp(-0.5(\beta_j - \hat{\beta}_j)^2/s_j^2).$$

("Laplace Approximation")

- Borrow information by assuming  $\beta_j$  are iid  $\sim g(\cdot; \pi)$ , where  $\pi$  are hyperparameters to be estimated.
- Letting  $g(\cdot; \pi)$  be a mixture of normal distributions provides both flexibility, and analytic calculations.
  - very small variances can capture effects that are "effectively" zero.



• Focus on the special case where  $g(\cdot; \pi)$  can be assumed unimodal and symmetric about zero.

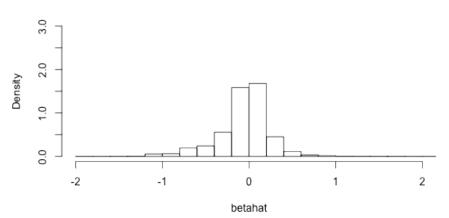
- Focus on the special case where  $g(\cdot; \pi)$  can be assumed unimodal and symmetric about zero.
- Then the posterior mean,  $E(\beta_j|\hat{\beta}, s, \hat{\pi})$  is a "shrinkage" estimate of  $\beta_j$ .

- Focus on the special case where  $g(\cdot; \pi)$  can be assumed unimodal and symmetric about zero.
- Then the posterior mean,  $E(\beta_j|\hat{\beta}, s, \hat{\pi})$  is a "shrinkage" estimate of  $\beta_j$ .
- And  $p(\beta_j > 0 | \hat{\beta}, s, \hat{\pi})$  can be used to identify j for which the sign of  $\beta_j$  can be confidently determined (analogous to test of  $\beta_j = 0$ ; Gelman et al, 2012).

- Focus on the special case where  $g(\cdot; \pi)$  can be assumed unimodal and symmetric about zero.
- Then the posterior mean,  $E(\beta_j|\hat{\beta}, s, \hat{\pi})$  is a "shrinkage" estimate of  $\beta_j$ .
- And  $p(\beta_j > 0 | \hat{\beta}, s, \hat{\pi})$  can be used to identify j for which the sign of  $\beta_j$  can be confidently determined (analogous to test of  $\beta_j = 0$ ; Gelman et al, 2012).
- Because  $\pi$  is estimated from the data, the amount of shrinkage is adaptive to the data. And because of the role of  $s_j$ , the amount of shrinkage adapts to the information on each gene.

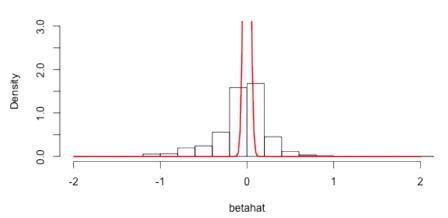
## Example: ASH applied to mouse data

#### Raw effect size estimates



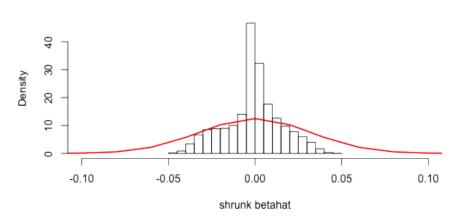
# Example: ASH applied to mouse data

#### Raw effect size estimates

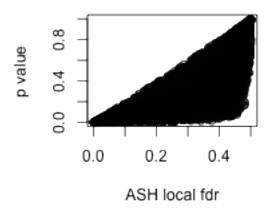


## Example: ASH applied to mouse data

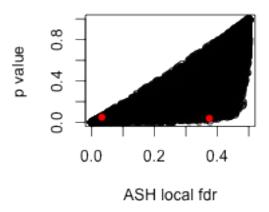
#### Shrunken estimates



### Shrinkage is adaptive to information



### Shrinkage is adaptive to information



#### Shrinkage is adaptive to information

```
## gene lv1 lv2 rv1 rv2 pval zdat.ash$localfdr
## 19422 Mgat5b 7 10 320 452 0.03795 0.37448
## 20432 Sec63 1042 1034 5496 6649 0.04908 0.03251
```

• Both provide a rational approach to identifying "significant" findings.

- Both provide a rational approach to identifying "significant" findings.
- Both are generic and modular: once you have the summary data, you can forget where they came from.

- Both provide a rational approach to identifying "significant" findings.
- Both are generic and modular: once you have the summary data, you can forget where they came from.
- But by using two numbers  $(\hat{\beta}, s)$  instead of one (p values) precision of different measurements can be better accounted for.

- Both provide a rational approach to identifying "significant" findings.
- Both are generic and modular: once you have the summary data, you can forget where they came from.
- But by using two numbers  $(\hat{\beta}, s)$  instead of one (p values) precision of different measurements can be better accounted for.
- ASH borrows information for estimation, as well as testing.

#### Guarantees?

• "I think you have some nice ideas. How will you convince people to use them?" (C Morris)

#### Next steps?

• Extend to allow  $g(\cdot; \pi)$  to depend on covariates X.

#### Next steps?

- Extend to allow  $g(\cdot; \pi)$  to depend on covariates X.
- Extend to allow for correlations in the measured  $\hat{\beta}_j$ .

• Bayesian variable selection for large-scale linear regression.

- Bayesian variable selection for large-scale linear regression.
- BSLMM:

$$Y = X\beta + \epsilon$$

, with 
$$\beta_j \sim \pi N(0, \sigma_b^2) + (1 - \pi)N(0, \sigma_a^2 + \sigma_b^2)$$
.

- Bayesian variable selection for large-scale linear regression.
- BSLMM:

$$Y = X\beta + \epsilon$$

, with 
$$\beta_j \sim \pi N(0, \sigma_b^2) + (1 - \pi)N(0, \sigma_a^2 + \sigma_b^2)$$
.

• Particular focus on prior specification (reparameterize in terms of regression  $R^2$ ).

- Bayesian variable selection for large-scale linear regression.
- BSLMM:

$$Y = X\beta + \epsilon$$

, with 
$$\beta_j \sim \pi N(0, \sigma_b^2) + (1 - \pi)N(0, \sigma_a^2 + \sigma_b^2)$$
.

- Particular focus on prior specification (reparameterize in terms of regression  $R^2$ ).
- BSLMM software, runs with thousands of individuals, hundreds of thousands of variables. (Zhou et al, 2013)

- Bayesian variable selection for large-scale linear regression.
- BSLMM:

$$Y = X\beta + \epsilon$$

, with 
$$\beta_j \sim \pi N(0, \sigma_b^2) + (1 - \pi)N(0, \sigma_a^2 + \sigma_b^2)$$
.

- Particular focus on prior specification (reparameterize in terms of regression  $R^2$ ).
- BSLMM software, runs with thousands of individuals, hundreds of thousands of variables. (Zhou et al, 2013)
- Also variational approximations (Carbonetto and Stephens, Bayesian Analysis)

#### **Thanks**

• to the several postdoctoral researchers and students who have worked with me on related topics.

#### **Thanks**

- to the several postdoctoral researchers and students who have worked with me on related topics.
- Especially: William Wen, Timothee Flutre, Scott Powers, Heejung Shim, Zhengrong Xing, and Ester Pantaleo.

#### **Thanks**

- to the several postdoctoral researchers and students who have worked with me on related topics.
- Especially: William Wen, Timothee Flutre, Scott Powers, Heejung Shim, Zhengrong Xing, and Ester Pantaleo.
- And to the NIH for funding, and i-like for inviting me.

### Reproducible research

• This document is produced with **knitr**, **Rstudio** and **Pandoc**.

#### Reproducible research

- This document is produced with **knitr**, **Rstudio** and **Pandoc**.
- For more details see my stephens999/ash repository at http://www.github.com/stephens999/ash

#### Reproducible research

- This document is produced with knitr, Rstudio and Pandoc.
- For more details see my stephens999/ash repository at http://www.github.com/stephens999/ash
- Website: http://stephenslab.uchicago.edu

#### Pandoc Command used

Matthew Stephens

```
pandoc -s -S -i -template=my.beamer -t beamer -V theme:CambridgeUS

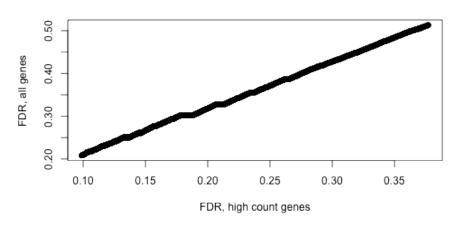
    V colortheme:beaver ilike-slides.md -o ilike-slides.pdf

Here is my session info:
print(sessionInfo(), locale = FALSE)
## R version 2.15.1 (2012-06-22)
## Platform: x86_64-apple-darwin9.8.0/x86_64 (64-bit)
##
## attached base packages:
## [1] stats
                 graphics grDevices utils
                                                  datasets
                                                            meth
##
## other attached packages:
## [1] qvalue_1.30.0 knitr_1.1
##
## loaded via a namespace (and not attached):
      codetable 0 2-8 digast 0 6
```

Adaptive Shrinkage and False Discovery Rates

2013/5/13

# FDRs for higher count genes affected by lower count genes



## Some odd things in the data

