

Probability 2 - Notes

Dom Hutchinson

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1 General

Definition 1.1 - σ -Field, \mathcal{F}

A σ -Field is a collection of subsets of the *Sample Space* which can be used to establish a formal definition of the probability distribution of the *Sample Space*. \mathcal{F} is a σ -Field if

- i) $\emptyset \in \mathcal{F}$;
- ii) $\forall \{A_1, A_n\} \subseteq \mathcal{F}, \bigcup_{i=1}^{\infty} A_i \in \mathcal{F}$; and,
- iii) $\forall A \in \mathcal{F}, A^c \in \mathcal{F}$.

The events, $A_i \in \mathcal{F}$, are said to be \mathcal{F} -Measurable. $\emptyset \in \mathcal{F}$ is known as the *Impossible Event* & $\Omega \in \mathcal{F}$ is known as the *Certain Event*. For a collection of events, \mathcal{C} , $\sigma(\mathcal{C})$ is the smallest σ -Field that contains \mathcal{C} .

Theorem 1.1 - Properties of σ -Fields

- i) $\forall \sigma$ -fields $\mathcal{F}_1, \mathcal{F}_2, \mathcal{F}_1 \cap \mathcal{F}_2$ is a σ -field.
- ii) 2^Ω is a σ -field.

Definition 1.2 - Probability Measure, \mathbb{P}

A *Probability Measure* is a function $\mathbb{P} : \mathcal{F} \rightarrow [0, 1]$ which satisfies

- i) $\mathbb{P}(\emptyset) = 0$ & $\mathbb{P}(\Omega) = 1$; and,
- ii) If $A_1, \dots, A_n \in \mathcal{F}$ are pair-wise disjoint then $\mathbb{P}\left(\bigcup_{i=1}^n A_i\right) = \sum_{i=1}^n \mathbb{P}(A_i)$. (σ -Additivity)

Definition 1.3 - Probability Space

A *Probability Space* is a triple formed of a *Sample Space*, Ω ; a σ -Field, \mathcal{F} , on Ω ; and, a *Probability Measure*, \mathbb{P} , on \mathcal{F} . Denoted $(\Omega, \mathcal{F}, \mathbb{P})$.

Definition 1.4 - Random Variable

A *Random Variable*, X on a *Probability Space* $(\Omega, \mathcal{F}, \mathbb{P})$ is a function $X : \Omega \rightarrow \mathbb{R}$.

Definition 1.5 - Filtration, \mathcal{F}_t

A *Filtration* is a family of σ -Fields, $\mathcal{F}_t = \{\mathcal{F}_t : t \geq 0\}$ st $\mathcal{F}_{t_1} \subset \mathcal{F}_{t_2} \forall 0 \leq t_1 \leq t_2$.

Definition 1.6 - Stochastic Process

A *Stochastic Process* is a collection of random variables which represent the state of a system at different times (e.g. $X = \{X_t\}_{t \in \Delta}$ with $\Delta \subseteq \mathbb{R}$ where X_t is the state of the system at time t). The *Stochastic Process* has an associated *Filtration*, \mathcal{F}_t , st X_t is \mathcal{F}_t -Measurable (X is *Adapted* to \mathcal{F}_t). The *State Space* of a *Stochastic Process* is the set of all possible values at a specific time. N.B. - *Stochastic Processes* can be in discrete or continuous time.

Definition 1.7 - Time-Homogenous

A *Stochastic Process*, X , is *Time-Homogenous* if $\mathbb{P}(X_{n+1} = j | X_n = i)$ depends on i & j , but not on n .

Definition 1.8 - Sequences

Let $X = A_1, A_2, \dots$ be a sequence of events. If $A_n \subseteq A_{n+1} \forall n \in \mathbb{N}$ then X (i.e. A_n occurs $\Rightarrow A_{n+1}$ occurs) is an *Increasing Sequence*. If $A_{n+1} \subseteq A_n \forall n \in \mathbb{N}$ then X is a *Decreasing Sequence*.

Theorem 1.2 - Continuity of Probability

Let A_1, A_2, \dots be an *Increasing Sequence* of events. Define $A := \bigcup_{n=1}^{\infty} A_n$. Then

$$\mathbb{P}(A) = \lim_{n \rightarrow \infty} \mathbb{P}(A_n)$$

Let B_1, B_2, \dots be a *Decreasing Sequence* of events. Define $B := \bigcap_{n=1}^{\infty} B_n$. Then

$$\mathbb{P}(B) = \lim_{n \rightarrow \infty} \mathbb{P}(B_n)$$

2 Markov Chains

Definition 2.1 - Markov Property

A *Stochastic Process* $X := \{X_n\}_{n \in \mathbb{N}}$ or $X := \{X_t\}_{t \in \mathbb{R}}$ has the *Markov Property* if values only depend on the value immediately before them (in time) and nothing earlier. Equivalently

$$\begin{aligned} X_{n+1} &= f(X_n) \\ \mathbb{P}(X_{n+1} = x_{n+1} | \mathcal{F}_n) &= \mathbb{P}(X_{n+1} = x_{n+1} | X_n = x_n) \\ \mathbb{P}(X_{n+1} = x_{n+1} | X_n = x_n, \dots, X_0 = x_0) &= \mathbb{P}(X_{n+1} = x_{n+1} | X_n = x_n) \\ \mathbb{P}(X_t = j | \mathcal{F}_s) &= \mathbb{P}(X_t = j | X_s) \end{aligned}$$

Definition 2.2 - Markov Chain

A *Markov Chain* is a state space process with the *Markov Property*.

Definition 2.3 - Transition Matrix

A *Transition Matrix* are a *Stochastic Matrix* which describes the probability of transitioning between each state. For a *Continuous Time* process we define different *Transition Matrices* for each time t .

$$\begin{aligned} \text{Discrete} \quad [P]_{ij} &:= p_{ij} \\ \text{Continuous} \quad [P_t]_{ij} &:= p_{ij}(t) = \mathbb{P}(X_t = j | X_0 = i). \end{aligned}$$

Definition 2.4 - Stationary Distributions

Let $X := \{X_n\}_{n \in \mathbb{N}}$ (or continuous) be a *Markov Chain*, on state space S , with transition matrix P . Let π be a horizontal vector. π is a *Stationary Distribution* of X if

- i) $\pi_j \geq 0 \forall j \in S$;
- ii) $\sum_{j \in S} \pi_j = 1$;
- iii) $\pi = \pi P_t$ or $\pi = \pi P \Leftrightarrow \pi_j = \sum_{i \in S} \pi_i p_{ij} \forall j \in S$

An *Irreducible Markov Chain* on S has a *Stationary Distribution* iff all the states in S are positive recurrent. In this case $\pi_i = \frac{1}{m_{ii}}$ (For continuous time this is equivalent to $p_{ij}(t) \xrightarrow{t \rightarrow \infty} \pi_j$).

N.B. - π is a stationary distribution for G iff $\nu H = \nu$ where H is transition matrix for jump chain & $\nu_i = -\pi_i g_{ii}$.

2.1 Discrete Time

Definition 2.5 - Communication

A state $i \in S$ *Communicates* with state $j \in S$ ($i \rightarrow j$) if $\exists n \in \mathbb{N}$ st $p_{ij}(n) > 0$. State $i, j \in S$ *Intercommunicate* if $i \rightarrow j$ & $j \rightarrow i$. If two states *Intercommunicate* and one is *Recurrent*, then both are *Recurrent*. *Intercommunication* is an *Equivalence Relation*. The state space can be partitioned into *Communicating Classes* $\{E_1, \dots, E_n\}$ st $\forall i, j \in E_k$ $i \leftrightarrow j$.

Definition 2.6 - Closed & Irreducible

Let $C \subseteq S$ be a set of states. C is *Closed* if $p_{ij} = 0 \forall i \in C, j \notin C$. C is *Irreducible* if

$i \leftrightarrow k \forall i, j \in C$. A *Closed, Singleton Set* $\{i\}$ is called an *Absorbing State*. If C is a *Communicating Class* and is not closed then all states in C are *Transient*.

Remark 2.1 - Partitioning State Space

The state space can be uniquely partitioned $S = T \cup C_1 \cup \dots \cup C_n$ where T is a set of *Transient* states & each C_i is an *Irreducible-Closed* set of *Recurrent* States.

Definition 2.7 - Period

Let $j \in S$ be a state st $p_{jj}(n) > 0$, $n \in \mathbb{N}$. Define $\mathcal{N}_j := \{n \geq 1 : p_{jj}(n) > 0\}$. The *period* of j is defined as $d_j := \gcd(\mathcal{N}_j)$, if $d_j = 1$ then j is said to be *Aperiodic*. For a *Communicating Class* $C \subseteq S$, $d_i = d_j \forall i, j \in C$.

Theorem 2.1 - Stationary Distribution & Transition Probabilities

Let $X := \{X_n\}_{n \in \mathbb{N}}$ be an irreducible aperiodic *Markov Chain*, with *Stationary Distribution* π . Then

$$p_{ij}(n) \xrightarrow{n \rightarrow \infty} \pi_j \forall i, j \in S$$

Theorem 2.2 - Chapman-Kolmogorov Equation

$$p_{ij}(n) = \sum_{k \in S} p_{ik}(r) p_{kj}(n-r) \forall i, j \in S, n \in \mathbb{N} \text{ \& } r \in [0, n]$$

2.2 Continuous Time

Definition 2.8 - Generator Matrix

A *Generator Matrix* is an alternative way of displaying continuous-time *Markov Chains*. A *Generator Matrix* G is defined st

- i) $g_{ij} \geq 0 \forall i \neq j$;
- ii) $g_{ii} = -\sum_{j \neq i} g_{ij}$.

Definition 2.9 - Jump Chain

A *Jump Chain* $\{Y_n\}_{n \in \mathbb{N}}$, for a *Continuous-Time Markov Chain* $\{X_t\}_{t \in \mathbb{R}^{\geq 0}}$, is a *Discrete-Time Markov Chain* which describes the different states X_t moves to. $Y_n := X_{S_n}$ where S_n is an arrival time.

Definition 2.10 - Recurrence & Transience - Continuous Time

A state $i \in S$ is *Recurrent* if $\mathbb{P}(\{t \geq 0 : X_t = i\} \text{ is unbounded} | X_0 = i) = 1$. A state is only recurrent in continuous time iff it is recurrent for the *Jump Chain* Y_n . A state $i \in S$ is *Transient* if $\mathbb{P}(\{t \geq 0 : X_t = i\} \text{ is unbounded} | X_0 = i) = 0$.

Definition 2.11 - Poisson Process, N_t

Let N_t be the number of events to have occurred by time t . Then $\{N_t\}_{t \in \mathbb{R}^{\geq 0}}$ is a *Poisson Process* with rate λ if

- i) $N_0 = 0$;
- ii) $N_t \in \mathbb{N} \forall t \in \mathbb{R}^{\geq 0}$;
- iii) $N_{t+s} - N_t$ depends on s only (Stationary Increments);
- iv) $N_{t_2} - N_{t_1}, \dots, N_{t_n} - N_{t_{n-1}}$ are all independent. (Independent Increments);

v) $\forall t, h > 0$ we have

$$\begin{aligned}\mathbb{P}(N_{t+h} - N_t < 0) &= 0 \\ \mathbb{P}(N_{t+h} - N_t = 0) &= 1 - \lambda h + o(h) \\ \mathbb{P}(N_{t+h} - N_t = 1) &= \lambda h + o(h) \\ \mathbb{P}(N_{t+h} - N_t > 1) &= o(h)\end{aligned}$$

For $t \geq 0$, $N_t \sim \text{Po}(\lambda t) \implies \mathbb{E}(N_t) = \text{Var}(N_t) = \lambda t$. The filtration \mathcal{F}_t of a *Poisson process* is generated by the process itself.

Definition 2.12 - Arrival Times

For a *Poisson Process* $\{N_t\}_{t \in \mathbb{R}^{\geq 0}}$ we define $S_i := \inf\{t \geq 0 : N_t = i\}$ to be the *Arrival Time* of the occurrence of the i^{th} event. Since $N_t \sim \text{Po}(\lambda t)$ then $S_1 \sim \text{Exp}(\lambda)$. Since $S_n = \sum_{j=1}^n T_j$ then $S_i \sim \Gamma(i, \lambda)$.

Definition 2.13 - Inter-Arrival Times

For a *Poisson Process* $\{N_t\}_{t \in \mathbb{R}^{\geq 0}}$ we define $T_1 = S_1$, $T_i = S_i - S_{i-1} \forall i \in \mathbb{N}^{\geq 2}$ to be the *Inter-Arrival Time* of events. *Inter-Arrival Times* are independent of each other. $T_i \sim \text{Exp}(\lambda)$.

Proposition 2.1 - Waiting Times

Let T be a random variable that models how long customers have to wait. We want T to have the following property $\mathbb{P}(T \in (t, t+h] | T > t) = \lambda h + o(h)$ where $\lim_{h \rightarrow 0} \frac{o(h)}{h} = 0$ & λ is the average rate that customers are served at. Thus T is modelled by $T \sim \text{Exp}(\lambda)$. If we have multiple *servers* who serve customers at different rates we can model them as independent random variables $T_i \sim \text{Exp}(\lambda_i)$ where λ_i is the rate of the i^{th} server.

Proposition 2.2 - First Expected

Suppose we have n servers whose serving times $T_i \sim (\lambda_i)$. Then

$$\begin{aligned}\min(T_1, \dots, T_n) &\sim \text{Exp}(\sum_{i=1}^n \lambda_i) \\ \mathbb{P}(T_i = \min(T_1, \dots, T_n)) &= \frac{\lambda_i}{\sum_{i=1}^n \lambda_i}\end{aligned}$$

Definition 2.14 - Birth & Death Processes

Birth Death Processes are continuous time processes where we consider a population of independent individuals where population members either: give birth (increase population by 1); or, die (decrease population by 1). We define

- i) N_t to be the population size at time t (with $N_0 = 1$);
- ii) $p_n(t) := \mathbb{P}(N_t = n)$;
- iii) $S_i := \inf\{t \geq 0 : N_t = i\}$; and,
- iv) $T_i := S_1 - S_{i-1}$.

Definition 2.15 - Generalised Birth & Death Process

A *Continuous Time Stochastic Process*, $\{N_t\}_{t \in \mathbb{R}^{\geq 0}}$, is a *Generalised Birth & Death Process* if

- i) $N_t \in \mathbb{N} \forall t \in \mathbb{R}^{\geq 0}$; and,
- ii) For $\lambda_n, \mu_n \geq 0 \forall n$ we have

$$\begin{aligned}\mathbb{P}(N_{t+h} - N_t = 1 | N_t = n) &= \lambda_n h + o(h) \\ \mathbb{P}(N_{t+h} - N_t = -1 | N_t = n) &= \mu_n h + o(h) \\ \mathbb{P}(N_{t+h} - N_t = 0 | N_t = n) &= 1 - (\lambda_n + \mu_n)h + o(h)\end{aligned}$$

iii)

$$\begin{aligned} p'_n(t) &= \lambda_{n-1}p_{n-1}(t) - (\lambda_n + \mu_n)p_n(t) + \mu_{n+1}p_{n+1}(t) \quad n \geq 1 \\ p'_0(t) &= -\lambda_0p_0(t) + \mu_1p_1(t) \end{aligned}$$

Proposition 2.3 - Pure Immigration

For a *Birth & Death Process* with a $\lambda_n = \lambda$, $\mu_n = 0 \quad \forall n \in \mathbb{N}$, for $\lambda \in [0, 1]$, we have that $\{N_t\}_{t \in \mathbb{R}^s}$ is a *Poisson Process*.

Proposition 2.4 - Linear Birth Process (Yule Process)

The *Yule Process* is the *Linear Birth Process*. This is a *Birth & Death Process* where $\lambda_n = n\lambda$, $\mu_n = 0 \quad \forall n \in \mathbb{N}$ for $\lambda, \mu \in [0, 1]$. Then

- i) $N_t \sim \text{Geo}(e^{-\lambda t})$;
- ii) $p_n(t) = e^{-\lambda t}(1 - e^{-\lambda t})^{n-1}$;
- iii) $\mathbb{E}(S_i) \simeq \frac{1}{\lambda} \ln n$;
- iv) $T_i \sim \text{Exp}(i\lambda)$;

$$\text{v) } g_{ij} = \begin{cases} \mu & , j = i - 1 \\ \lambda & , j = i + 1 \\ -(\lambda + \mu)i & , j = i \\ 0 & \text{otherwise} \end{cases}$$

Proposition 2.5 - Linear Birth & Death Process

The *Linear Birth & Death Process* is a *Birth & Death Process* where $\lambda_n = n\lambda$, $\mu_n = n\mu \quad \forall n \in \mathbb{N}$ for $\lambda, \mu \in [0, 1]$. Then

$$\text{i) } g_{ij} = \begin{cases} 0 & j < i - 1 \text{ or } j > i + 1 \\ \mu i & h = i - 1 \\ \lambda i & j = i + 1 \\ -(\lambda + \mu)i & j = i \end{cases}$$

Proposition 2.6 - Holding Times

Consider a *Continuous-Time Markov Chain* on state space S . Suppose the chain enters state $i \in S$ at time t , define $U := \inf\{s \geq - : X_{t+s} \neq i\} \sim \text{Exp}(\lambda_i)$ to be *Holding Time* of the chain in state i . We define H to be the matrix st $h_{ij} = \mathbb{P}(X_{t+U} = j | X_t = i)$ (probability of transitioning to state i from state j). Then $h_{ij} = -\frac{g_{ij}}{g_{ii}} = \frac{g_{ij}}{\lambda_i}$.

Definition 2.16 - Kolmogorov Equations

Forwards Equations $P'_t = P_t G$ since $P_0 = I \implies P_t = e^{tG}$

Backwards Equations $P'_t = G P_t$.

Further, we derive that

$$\left. \frac{d^k}{dt^k} P_t \right|_{t=0} = G^k \text{ for } k \geq 0 \implies g_{ij} = p'_{ij}(0)$$

3 Random Walks

Definition 3.1 - Random Walk

A *Random Walk* is a discrete-time stochastic process where the value of the process increases or decreases by 1, only (it always changes value). *Random Walks* have the *Markov Property*.

Random Walks can be defined to include *Absorbing Barriers* which are values that if the process ever achieves, the process stops.

Definition 3.2 - Probability of Visiting

Consider a time-homogenous stochastic process X on state space S . Let $i, j \in S$ be a state. We define

$$f_{ij} := \mathbb{P}(\text{We will visit } j | X_0 = i) \quad \& \quad f_{ij}(n) := \mathbb{P}(X_n \text{ is first visit to } j | X_0 = i)$$

This can be equated with *n-Step Transition Probabilities* as

$$p_{ij}(n) = \sum_{m=1}^n p_{jj}(n-m) f_{ij}(m)$$

If we are guaranteed to return to a state $s \in S$ (i.e. $f_{ss} = 1$) then s is *Recurrent*, else (i.e. if $f_{ss} < 1$) s is *Transient*. A state j is *Recurrent* iff $P_{ij}(1) = \infty$ for all i . If a state j is *Transient* then $p_{ij}(n) \xrightarrow{n \rightarrow \infty} 0$.

Proposition 3.1 - Generating Functions

For *n-Step Transition Probabilities*, $p_{ij}(n)$, & *First Passage Probabilities*, $f_{ij}(n)$, we define the following generating functions

$$\begin{aligned} P_{ij}(s) &= \sum_{n=1}^{\infty} p_{ij}(n) s^n & F_{ij}(s) &= \sum_{n=1}^{\infty} f_{ij}(n) s^n \\ \implies P_{ij}(s) &= \mathbb{1}_{i=j} + F_{ij}(s) P_{jj}(s) \end{aligned}$$

These can be used to relate $f_{ij}(n)$ with $p_{ij}(n)$.

Proposition 3.2 - Mean Time of First Passage

Consider a time-homogenous stochastic process X on state space S . We can calculate the *Mean Time of First Passage* from state $i \in S$ to state $j \in S$ as

$$m_{ij} = \sum_{n=1}^{\infty} n f_{ij}(n)$$

N.B. $-m_{ii}$ is called the *Mean Time of Return*.

Proposition 3.3 - Gambler's Ruin

The *Gambler's Ruin* is a common scenario used to demonstrate a *Random Walk*. A gambler starts with $\mathcal{L}k$ & stops gambling if they reach $\mathcal{L}0$ or $\mathcal{L}N$ (i.e. There are absorbing barriers at 0 & N). Each time $\mathcal{L}1$ is gambled & either it is lost (with probability 1) or $\mathcal{L}1$ is won (with the bet returned), with probability $q := 1 - p$.

$$p_k := \mathbb{P}(\text{Gambler ever ruined} | X_0 = k) \implies p_k = p_{k+1}p + p_{k-1}q = \frac{\left(\frac{q}{p}\right)^k - \left(\frac{q}{p}\right)^N}{1 - \left(\frac{q}{p}\right)^N}$$

Theorem 3.1 - Probability of Ruin against Upper Barrier

Let p_k be the probability of ruin for a gambler starting at $\mathcal{L}k$ & no upper barrier. Let $p_k^{(N)}$ be the probability of ruin for a gambler with an upper barrier at N . Then

$$\lim_{N \rightarrow \infty} p_k^{(N)} = p_k$$

Definition 3.3 - Stopping Time

Let $X = \{X_n\}_{n \in \mathbb{N}}$ be a *Stochastic Process*. Non-negative integer-valued random variable T is a

Stopping Time Process for X if $\forall n \in \mathbb{N}$ the event $\{T \leq n\}$ is determined by X_0, \dots, X_n only. Often a *Stopping Time* is an event where a first passage occurs.

Theorem 3.2 - Wald's Lemma

Let Z_1, Z_2, \dots be a sequence of iid rv with $\mathbb{E}(|Z_n|) < \infty$ & $X_n := \sum_{i=1}^n Z_i$. If T is a *Stopping Time Process* if $X = \{X_n\}_{n \in \mathbb{N}}$ with $\mathbb{E}(T) < \infty$. Then

$$\mathbb{E}(X_T) = \mathbb{E}(Z_1)\mathbb{E}(T)$$

4 Brownian Motion / Wiener Process

Definition 4.1 - Simple Symmetric Random Walk

A *Simple Symmetric Random Walk* is a process S_1, S_2, \dots where $S_i = \sum_{j=1}^i Y_j$ with Y_1, Y_2, \dots being independent rvs which take values $\{-1, 1\}$ with probability $\frac{1}{2}$.

Definition 4.2 - Brownian Motion

Brownian Motion is a *continuous-time stochastic process* which models random motion in continuous space $\mathbb{R}^n, n \in \mathbb{N}$.

Let \mathcal{F}_t be a filtration. An *Adapted Stochastic Process* $W := \{W_t\}_{t \geq 0}$ is a *Brownian Motion* if

- i) $W_0 = x, x \in \mathbb{R}$;
- ii) $W_{t+u} - W_t \sim \mathcal{N}(0, u)$ is independent of $\mathcal{F}_t \forall t, u \geq 0$ ($\implies W_t \sim \mathcal{N}(0, t)$); and,
- iii) $t \mapsto W_t(\omega)$ is a continuous function of $t \forall \omega \in \Omega$.

N.B. $-W$ is a *Standard Brownian Motion* if $x = 0$.

Proposition 4.1 - Constructing Brownian Motion from a Simple Symmetric Random Walk

By defining $X_t^n = \frac{1}{\sqrt{n}} S_{nt}$ for $t = \frac{j}{n} \in [0, 1]$ (and linearly interpreting other values) we can show that as $n \rightarrow \infty$ X_t^n satisfies the definition for *Brownian Motion*.

Theorem 4.1 - Transition Density of Brownian Motion

$$p(t, x, y) = \frac{1}{\sqrt{2\pi t}} e^{-\frac{(x-y)^2}{2t}}$$

Proposition 4.2 - Properties of a Standard Brownian Motion

Let W be a *Standard Brownian Motion*, then

- i) $\forall W_t \sim \mathcal{N}(0, t) \implies \mathbb{E}(W_t) = 0$ & $\text{Var}(W_t) = t$.
- ii) $\forall 0 \leq s \leq t$ $\text{Cov}(W_s, W_t) = 0$.
- iii) $\mathbb{P}(\sup_{t \geq 0} W_t = \infty, \inf_{t \geq 0} W_t = -\infty) = 1$.
- iv) $-W_t$ is a *Standard Brownian Motion*.
- v) $X = \{X_t\}_{t \geq 0}$ where $X_t := W_{t+s} - W_s$ is a *Standard Brownian Motion* for fixed $s > 0$.
- vi) $X = \{X_t\}_{t \geq 0}$ where $X_t := \frac{1}{\sqrt{\alpha}} W_{\alpha t}$ is a *Standard Brownian Motion* for fixed $\alpha > 0$.

Definition 4.3 - The Reflection Principle

Let W_t be a *Standard Brownian Motion*. Let $a \in \mathbb{R}^{\geq 0}$ and $\tau_a := \inf\{t > 0 : W_t = a\}$. Define

$$\widetilde{W}_t = \begin{cases} W_t & \text{for } t \leq \tau_a \\ a - (W_t - a) & \text{for } t > \tau_a \end{cases}. \text{ Then } \widetilde{W}_t \text{ is a Standard Brownian Motion.}$$

Proposition 4.3 - Properties of First Passage Time

Let W_t be a *Standard Brownian Motion*. Let $a \in \mathbb{R}^{\geq 0}$ and $\tau_a := \inf\{t > 0 : W_t = a\}$. Then

- i) $\mathbb{P}(\tau_a \leq T) = 2\mathbb{P}(W_t > a)$.
- ii) $f_{\tau_a}(t) = \frac{|a|}{\sqrt{2\pi t^3}} e^{-\frac{a^2}{2t}}$ for $t > 0$.
- iii) $\mathbb{E}(\tau_a) = \infty$ for fixed $a \neq 0$.

Definition 4.4 - Gaussian Process

A *Gaussian Process* is a *continuous-time stochastic process* with continuous sample paths (Set of possible values for a stochastic process). and finite dimensional distributions that are multivariate normal. A *Gaussian Process* is completely determined by its mean function $\mu_t = \mathbb{E}(X_t)$ & auto-covariance function $\mathbb{E}((X_s - \mu_s)(X_t - \mu_t))$.

5 Martingales

Definition 5.1 - Discrete-Time Martingales

A *Discrete-Time Martingale*, wrt a filtration \mathcal{F}_n (often generated by the process itself), is a *Stochastic Process* $Y := \{Y_n\}_{n \in \mathbb{N}}$ where $\forall n \in \mathbb{N}$

- i) $\mathbb{E}(|Y_n|) < \infty$;
- ii) $\mathbb{E}(Y_{n+1} | \mathcal{F}_n) = Y_n$.

Definition 5.2 - Continuous-Time Martingales

A *Continuous-Time Martingale*, wrt a filtration \mathcal{F}_t (often generated by the process itself), is a *Stochastic Process* $Y := \{Y_n\}_{n \in \mathbb{N}}$ where $\forall 0 \leq s \leq t$

- i) $\mathbb{E}(|Y_t|) < \infty$;
- ii) $\mathbb{E}(Y_t | \mathcal{F}_s) = Y_s$.

Definition 5.3 - Supermartingale & Submartingale

A *Supermartingale* is a process where $\mathbb{E}(Y_t | \mathcal{F}_s) \leq Y_s$.

A *Submartingale* is a process where $\mathbb{E}(Y_t | \mathcal{F}_s) \geq Y_s$.

Proposition 5.1 - Properties of Martingales

Let $Y := \{Y_n\}_{n \in \mathbb{N}}$ be a *Discrete-Time Martingale*. Note the following hold, with appropriate alterations, for *Continuous-Time Martingales*, submartingales & supermartingales.

- i) $\mathbb{E}(Y_{n+1}) = \mathbb{E}(\mathbb{E}(Y_{n+1} | \mathcal{F}_n)) = \mathbb{E}(Y_n)$

Remark 5.1 - Stopped Martingales are still Martingales**Theorem 5.1 - Stopped Supermartingales**

Let $Y := \{Y_n\}_{n \in \mathbb{N}}$ be a *Supermartingale* wrt \mathcal{F}_n . Let T be a stopping time of Y . Then $Z := \{Z_n\}_{n \in \mathbb{N}}$ with $Z_n := Y_{T \wedge n}$ is a *Supermartingale*.

N.B. - A similar results holds for *Continuous-Time Supermartingales*

Theorem 5.2 - Optional Stopping Theorem - Discrete Time

Let $Y := \{Y_n\}_{n \in \mathbb{N}}$ be a *Discrete-Time Martingale* wrt \mathcal{F}_n , let T be a stopping time of \mathcal{F}_n .

If any of the following conditions holds then $\mathbb{E}(Y_T) = \mathbb{E}(Y_0)$

- i) T is bounded (i.e. $\exists K \in \mathbb{R}^{\geq 0}$ st $\text{prob}(T < K) = 1$).
- ii) T is finite and $\exists K > 0$ st $|Y_{T \wedge n}| < K \forall n \in \mathbb{N}$.
- iii) $\mathbb{E}(T) < \infty$ and $\exists K > 0$ st $|Y_n - Y_{n-1}| \leq K \forall n < T$.

iv) T is finite and $\mathbb{E}(|Y_T|) < \infty$ and $\mathbb{E}(Y_n \mathbf{1}_{T > n}) \xrightarrow{n \rightarrow \infty} 0$.

Theorem 5.3 - Martingale Convergence Theorem - Discrete Time

Let $Y := \{Y_n\}_{n \in \mathbb{N}}$ be a *Supermartingale* wrt \mathcal{F}_n . Suppose there exists $A > 0$ st $\mathbb{E}(|Y_n|) \leq A \forall n \in \mathbb{N}$. Then \exists rv Y_∞ st $\mathbb{P}(\lim_{n \rightarrow \infty} Y_n = Y_\infty) = 1$.

Proposition 5.2 - Gambler's Ruin - Application

Let X_1, X_2, \dots be iid rv st $\mathbb{P}(X_i = 1) = p$ & $\mathbb{P}(X_i = -1) = q := 1 - p$. Define $S_0 = k$ & $S_n = S_{n-1} + X_n$, then $\{S_n\}_{n \in \mathbb{N}}$ is an unrestricted random walk. Let $T := \min\{n : S_n = 0 \text{ or } S_n = N\}$, then $\{Y_n\}_{n \in \mathbb{N}}$ is a random walk with absorbing barriers at 0 & N .

$p = q$ $Y := \{Y_n\}_{n \in \mathbb{N}}$ is a martingale wrt $\mathcal{F}_n := \sigma(X_1, \dots, X_n)$.

The following properties hold

- i) $\mathbb{P}(T < \infty) = 1$.
- ii) $\mathbb{P}(Y_T = N) = \frac{Y_0}{N}$.
- iii) $\mathbb{E}(T) = k(N - k)$.

$p \neq q$ Define $B_n = \left(\frac{q}{p}\right)^{Y_n}$. Then $\{V_n\}_{n \in \mathbb{N}}$ is a martingale wrt $\mathcal{F}_n := \sigma(X_1, \dots, X_n)$.

The following properties hold

- i) $\mathbb{P}(T < \infty) = 1$.
- ii) $\mathbb{P}(Y_T = N) = \frac{1 - \left(\frac{q}{p}\right)^k}{1 - \left(\frac{q}{p}\right)^N}$.

Proposition 5.3 - Brownian Motion - Application

Let $W := \{W_t\}_{t \geq 0}$ be a *Standard Brownian Motion* wrt $\sigma_t := \sigma(W_1, \dots, W_t)$. Then

- i) W_t is a martingale.
- ii) $W_t^2 - t$ is a martingale.
- iii) $X_t := at + \sigma W_t$ is a martingale iff $a = 0$.
- iv) $X_t := e^{at + \sigma W_t}$ is a martingale iff $a = -\frac{1}{2}\sigma^2$.

Let $a, b > 0$ and define $\tau := \min\{t \geq 0 : W_t \in \{a, -b\}\}$ then $\mathbb{P}(W_\tau = a) = \frac{b}{a+b}$ and $\mathbb{E}(\tau) = ab$. Let $X_t := \mu t + \sigma W_t$ with $\mu < 0$ & define $M := \max_{t \geq 0} X_t$. For $a, b > 0$ then

$$\mathbb{P}(\tau_a < \tau_{-b}) = \frac{1 - e^{-\alpha b}}{e^{\alpha a} - e^{-\alpha b}} \text{ with } \alpha := -\frac{2\mu}{\sigma^2}. \text{ Thus } \mathbb{P}(M \geq a) = e^{-\alpha a}$$

0 Reference

0.1 Definitions

Definition 0.1 - Equivalence Relation

A process is an *Equivalence Relation* if it is: Reflexive ($a = a$); symmetric ($a = b \implies b = a$); and, Transitive ($a = b, b = c \implies a = c$).

Definition 0.2 - Lack of Memory Property

A random variable X is said to have the *Lack of Memory Property* if

$$\mathbb{P}(X > t + h | T > h) = \mathbb{P}(T > h)$$

Definition 0.3 - Positive Definite Matrix

A *Positive Definite Matrix* is a matrix with strictly positive eigenvalues. A positive-definite matrix A can be defined as $A = P^{-1}DP$ where D is a diagonal matrix & P is a *unitary matrix*.

Definition 0.4 - Stochastic Matrix

A matrix M is a *Stochastic Matrix* if $[M]_{ij} \geq 0 \forall i, j \in [1, n]$ & $\sum_{j=1}^n M_{ij} = 1 \forall i \in [1, n]$.

0.2 Notation

Notation 0.1 - Transition Probabilities

For a *Time-Homogenous Stochastic Process* we define *Transition Probabilities* as

$$\begin{aligned} p_{ij} &:= \mathbb{P}(X_1 = j | X_0 = i) = \mathbb{P}(X_{n+1} = j | X_n = i) \\ p_{ij}(n) &:= \mathbb{P}(X_n = j | X_0 = i) \end{aligned}$$

Notation 0.2 - Stopped

Let $Y := \{Y_n\}_{n \in \mathbb{N}}$ be a stochastic process. Let T be a stopping time of Y . Then

$$Y_{T \wedge n} := \begin{cases} Y_n & \text{for } n \leq t \\ Y_T & \text{for } n > t \end{cases}$$

0.3 Theorems

Theorem 0.1 - Cauchy-Schwarz Inequality

Let X & Y be random variables with finite variance. Then

$$\mathbb{E}(|XY|) \leq \sqrt{\mathbb{E}(X^2)\mathbb{E}(Y^2)}$$

Theorem 0.2 - Central Limit Theorem

Let X_1, X_2, \dots be iid rb with $\mu = \mathbb{E}(X)$ & $\sigma^2 = \text{Var}(X) \neq 0$. Then

$$Z_n = \frac{1}{\sigma\sqrt{n}} \sum_{i=1}^n (X_i - \mu) \sim \mathcal{N}(0, 1)$$

Theorem 0.3 - Equating Exponential & Binomial Distribution

For an event that occurs at time $T \sim \text{Exp}(\lambda)$ then $\mathbb{P}(T < t) = 1 - e^{-\lambda t} \simeq \lambda t$ (for small λt). For n of these events (with the same distribution) the number of events that occur before time t is modelled by $N \sim \text{Bi}(n, \lambda t)$.

Theorem 0.4 - Jensen's Inequalities

Let g be a *Convex Function*. Then

$$\mathbb{E}(g(X)) \geq g(\mathbb{E}(X))$$

Let g be a *Concave Function*. Then

$$\mathbb{E}(g(X)) \leq g(\mathbb{E}(X))$$

Theorem 0.5 - Markov's Inequality

Let X be a non-negative random variable. Then

$$\mathbb{P}(X > c) \leq \frac{\mathbb{E}(X)}{c} \quad \forall c > 0$$

Theorem 0.6 - One-Step Conditioning Argument

Let A be an event that is dependent on the events X & Y . Then

$$\mathbb{P}(A) = \mathbb{P}(A|X)\mathbb{P}(X) + \mathbb{P}(A|Y)\mathbb{P}(Y)$$

0.4 Probability Distributions

Definition 0.5 - Binomial Distribution

Let X be a discrete random variable modelled by a *Binomial Distribution* with n events and rate of success p .

$$\begin{aligned} p_X(k) &= \binom{n}{k} p^k (1-p)^{n-k} \\ \mathbb{E}(X) &= np \quad \& \quad \text{Var}(X) = np(1-p) \end{aligned}$$

Definition 0.6 - Gamma Distribution

Let T be a continuous random variable modelled by a *Gamma Distribution* with shape parameter α & scale parameter λ . Then

$$\begin{aligned} f_T(x) &= \frac{\lambda^\alpha x^{\alpha-1} e^{-\lambda x}}{\Gamma(\alpha)} \quad \text{for } x > 0 \\ \mathbb{E}(T) &= \frac{\alpha}{\lambda} \quad \& \quad \text{Var}(T) = \frac{\alpha}{\lambda^2} \end{aligned}$$

N.B. - $\alpha, \lambda > 0$.

Definition 0.7 - Exponential Distribution

Let T be a continuous random variable modelled by a *Exponential Distribution* with parameter λ . Then

$$\begin{aligned} f_T(t) &= \lambda e^{-\lambda t} \quad \text{for } t > 0 \\ F_T(t) &= 1 - e^{-\lambda t} \quad \text{for } t > 0 \\ \mathbb{E}(X) &= \frac{1}{\lambda} \quad \& \quad \text{Var}(X) = \frac{1}{\lambda^2} \end{aligned}$$

N.B. - Exponential Distribution is used to model the wait time between decays of a radioactive source.

Definition 0.8 - Normal Distribution

Let X be a continuous random variable modelled by a *Normal Distribution* with mean μ & variance σ^2 .

Then

$$\begin{aligned} f_X(x) &= \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \\ F_X(x) &= \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^x e^{-\frac{(y-\mu)^2}{2\sigma^2}} dy \\ M_X(\theta) &= e^{\mu\theta + \sigma^2\theta^2(1/2)} \\ \mathbb{E}(X) &= \mu \quad \& \quad \text{Var}(X) = \sigma^2 \end{aligned}$$

Definition 0.9 - Poisson Distribution

Let X be a discrete random variable modelled by a *Poisson Distribution* with parameter λ . Then

$$\begin{aligned} p_X(k) &= \frac{e^{-\lambda} \lambda^k}{k!} \quad \text{For } k \in \mathbb{N}_0 \\ \mathbb{E}(X) &= \lambda \quad \& \quad \text{Var}(X) = \lambda \end{aligned}$$

N.B. - Poisson Distribution is used to model the number of radioactive decays in a time period.