

# Problem Sheet 7

Statistics 2

*Dom Hutchinson*

## Question 1

a)

By assuming that the colour each bulb comes up is independent we can define  $X \sim \text{Binomial}(12, p)$  to represent the number of bulbs which come up red.

Define simple hypotheses  $H_0 : p = 0.25$  (*i.e.*  $ii$ ) is true) and  $H_1 : p = 0.6$  (*i.e.*  $i$ ) is true)

Define test statistic  $T(x) = x$ .

Since the farmer decides to accept  $i$  if 8 or more bulbs come up red, we have critical value  $c = 8$  and critical region  $R = [8, 12]$ .

b)

Power Function

$$\pi(\theta; T, c) = \pi(p; X, 8) = \mathbb{P}(X \geq 8; p) = \sum_{i=8}^{12} \binom{12}{i} p^i (1-p)^{12-i}$$

Significance Level

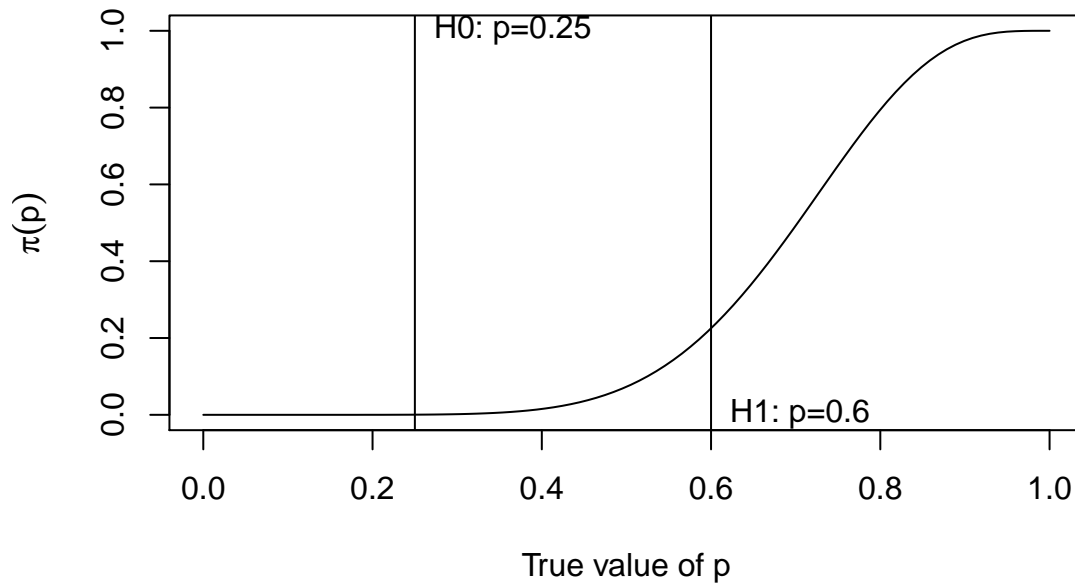
$$\alpha = \mathbb{P}(\text{Type I Error}) = \pi(0.25) = 0.0027815$$

Type II Error Level

$$\beta = \mathbb{P}(\text{Type II Error}) = 1 - \pi(0.6) = 0.7746627$$

c)

```
x<-seq(0,1,length.out=100)
y<-sapply(x,function(p) 1-pbinom(8,12,p))
plot(x,y,type="l",xlab="True value of p",ylab=expression(pi(p)),main="Power Function for n=12 & c=8")
abline(v=.25)
text(0.25,1,labels="H0: p=0.25",pos=4)
abline(v=.6)
text(0.6,0,labels="H1: p=0.6",pos=4)
```

**Power Function for n=12 & c=8**

d)

Define  $T_{NP}(x)$  to be the Neyman-Pearson test. Then

$$\begin{aligned}
 T_{NP}(x) &= \frac{L(p = 0.6; x)}{L(p = 0.25; x)} \\
 &= \frac{f_X(x; p = 0.6)}{f_X(x; p = 0.25)} \\
 &= \frac{\binom{12}{x} (0.6)^x (0.4)^{12-x}}{\binom{12}{x} (0.25)^x (0.75)^{12-x}} \\
 &= \frac{\left(\frac{12}{5}\right)^x \left(\frac{8}{15}\right)^{12-x}}{12^x 8^{12-x}} \\
 &= \frac{5^x 15^{12-x}}{2^{2x} 3^x 2^{3(12-x)}} \\
 &= \frac{5^x 3^{12-x} 5^{12-x}}{2^{36-x} 3^{2x-12}} \\
 &= \frac{5^{12}}{2^{36-x} 3^{2x-12}} \\
 &\propto 2^{36-x} 3^{2x-12}
 \end{aligned}$$

Thus  $T_{NP}(x)$  is an increasing function with  $x$ .

In **a)** we defined  $T(x) = x$  meaning  $T(x)$  is an equivalent test statistic to  $T_{NP}(X)$ . Meaning the farmer's test statistic is optimal in the Neyman-Pearson sense.

e)

Consider the power function of  $(T, c)$  with  $n$  &  $c$  not fixed.

$$\pi_n(p; X, c) := \sum_{i=c}^n \binom{n}{i} p^i (1-p)^{n-i}$$

We want to find a test  $(T, c)$  with significance level  $\alpha = 0.05$  and rate of type II error  $\beta = 0.1$ . Thus we want to find a combination of  $n$  &  $c$  which statisfies both the following equalities

$$\pi_n(0.25; X, c) = 0.05 \quad \text{and} \quad 1 - \pi_n(0.6; X, c) = 0.1$$

We note that

$$\pi_n(0.25; X, c) = \sum_{i=c}^n \binom{n}{i} \left(\frac{1}{4}\right)^i \left(\frac{3}{4}\right)^{n-i} \quad \text{and} \quad \pi_n(0.6; X, c) = \sum_{i=c}^n \binom{n}{i} \left(\frac{3}{5}\right)^i \left(\frac{2}{5}\right)^{n-i}$$

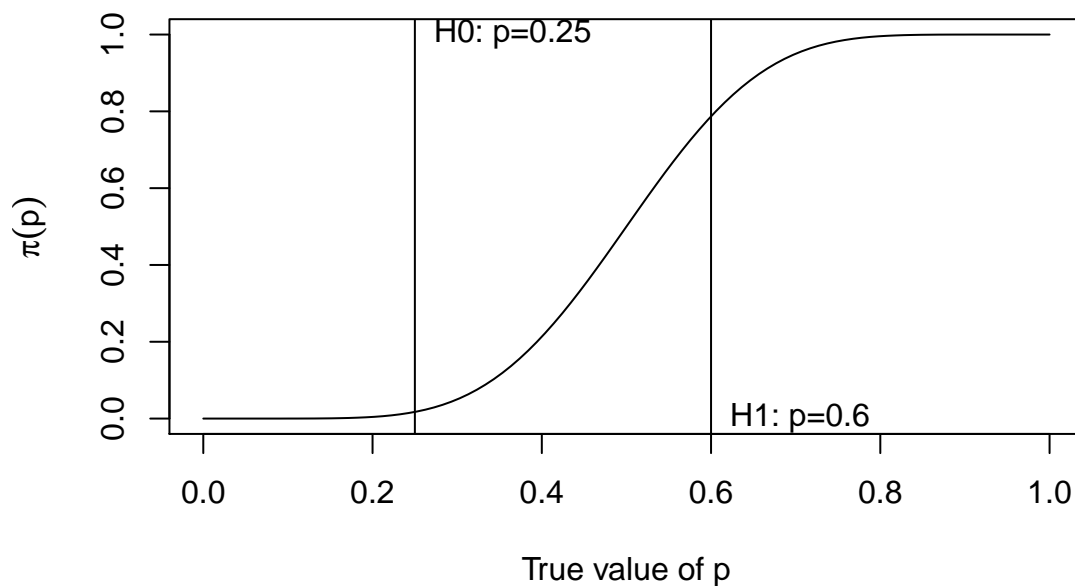
Using trial-and-error

$n$	$c$	$\pi_n(0.25)$	$1 - \pi_n(0.6)$
15	4	0.002	0.539
15	5	0.009	0.314
15	6	0.034	0.148
15	7	0.057	0.095

For  $n = 15$  &  $c = 7$  we have significance level  $\alpha = 0.057$  & rate of type-II-error  $\beta = 0.095$  which are both within 1 percentage point of our targets of 0.05 & 0.1 respectively.

```
x<-seq(0,1,length.out=100)
y<-sapply(x,function(p) 1-pbinom(7,15,p))
plot(x,y,type="l",xlab="True value of p",ylab=expression(pi(p)),main="Power Function for n=15 & c=7")
abline(v=.25)
text(0.25,1,labels="H0: p=0.25",pos=4)
abline(v=.6)
text(0.6,0,labels="H1: p=0.6",pos=4)
```

### Power Function for n=15 & c=7



## Question 2

Consider a test between two simple hypotheses. For each of the following statistical models, derive the Neyman-Pearson optimal test statistic, and try to find the simplest equivalent representation.

a)

Let  $\mathbf{X} \stackrel{\text{iid}}{\sim} \text{Poisson}(\lambda)$  and consider the tests  $H_0 : \lambda = \lambda_0$  &  $H_1 : \lambda = \lambda_1$  with  $0 < \lambda_1 < \lambda_0$ . Then

$$\begin{aligned} T_{NP}(\mathbf{x}) &= \prod_{i=1}^n \frac{f_X(x_i; \lambda_1)}{f_X(x_i; \lambda_0)} \\ &= \prod_{i=1}^n \frac{\frac{e^{-\lambda_1} \lambda_1^{x_i}}{x_i!}}{\frac{e^{-\lambda_0} \lambda_0^{x_i}}{x_i!}} \\ &= \prod_{i=1}^n e^{\lambda_0 - \lambda_1} \left( \frac{\lambda_1}{\lambda_0} \right)^{x_i} \\ &= e^{n(\lambda_0 - \lambda_1)} \left( \frac{\lambda_1}{\lambda_0} \right)^{\sum x_i} \end{aligned}$$

Since  $\lambda_1 < \lambda_0$  then  $\frac{\lambda_1}{\lambda_0} < 1$  then  $T_{NP}(\mathbf{x})$  is increasing with  $-S_n(\mathbf{x}) := -\sum x_i$ . Meaning  $T(\mathbf{x}) := -S_n(\mathbf{x})$  is an equivalent statistic to  $T_{NP}$ .

b)

Let  $\mathbf{X} \stackrel{\text{iid}}{\sim} \text{Exponential}(\lambda)$  and consider the tests  $H_0 : \lambda = \lambda_0$  &  $H_1 : \lambda = \lambda_1$  with  $0 < \lambda_0 < \lambda_1$ . Then

$$\begin{aligned} T_{NP}(\mathbf{x}) &= \prod_{i=1}^n \frac{\mathbb{1}\{x \geq 0\} \lambda_1 e^{-\lambda_1 x_i}}{\mathbb{1}\{x \geq 0\} \lambda_0 e^{-\lambda_0 x_i}} \\ &= \frac{\mathbb{1}\{x \geq 0\}}{\mathbb{1}\{x \geq 0\}} \prod_{i=1}^n \frac{\lambda_1 e^{-\lambda_1 x_i}}{\lambda_0 e^{-\lambda_0 x_i}} \\ &= \frac{\mathbb{1}\{x \geq 0\}}{\mathbb{1}\{x \geq 0\}} \left( \frac{\lambda_1}{\lambda_0} \right)^{\sum x_i} \prod_{i=1}^n e^{x_i(\lambda_0 - \lambda_1)} \\ &= \frac{\mathbb{1}\{x \geq 0\}}{\mathbb{1}\{x \geq 0\}} \left( \frac{\lambda_1}{\lambda_0} \right)^{\sum x_i} e^{(\lambda_0 - \lambda_1) \sum x_i} \end{aligned}$$

Note that  $T_{NP}$  is undefined if  $\exists x_i \in \mathbf{x}$  st  $x_i < 0$ .

Otherwise, since  $\lambda_0 < \lambda_1 \implies \lambda_0 - \lambda_1 < 0$  meaning  $T_{NP}$  is increasing with  $-S_n(\mathbf{x}) := \sum x_i$ .

Thus,  $T(\mathbf{x}) := -S_n(\mathbf{x})$  is an equivalent statistic to  $T_{NP}(\mathbf{x})$ .

c)

Let  $\mathbf{X} \stackrel{\text{iid}}{\sim} \text{Normal}(\mu, \sigma^2)$  with  $\sigma^2$  known and consider the tests  $H_0 : \mu = \mu_0$  &  $H_1 : \mu = \mu_1$  with  $\mu_0 < \mu_1$ . Then

$$\begin{aligned} T_{NP}(\mathbf{x}) &= \prod_{i=1}^n \frac{e^{-\frac{(x_i - \mu_1)^2}{2\sigma^2}}}{e^{-\frac{(x_i - \mu_0)^2}{2\sigma^2}}} \\ &= \prod_{i=1}^n e^{-\frac{1}{2\sigma^2} [(x_i - \mu_0)^2 - (x_i - \mu_1)^2]} \\ &= e^{-\frac{1}{2\sigma^2} \sum [(x_i - \mu_0)^2 - (x_i - \mu_1)^2]} \\ &= e^{-\frac{1}{2\sigma^2} \sum [2x_i(\mu_1 - \mu_0) + \mu_0^2 - \mu_1^2]} \\ &= e^{-\frac{1}{2\sigma^2} [n\mu_0^2 - n\mu_1^2 + 2(\mu_1 - \mu_0) \sum x_i]} \end{aligned}$$

By the constraints we know that  $\mu_1 - \mu_0 > 0$  meaning  $T_{NP}$  is increasing with  $-S_n(\mathbf{x}) := \sum x_i$ . Thus,  $T(\mathbf{x}) := -S_n(\mathbf{x})$  is an equivalent statistic to  $T_{NP}(\mathbf{x})$ .

d)

Let  $\mathbf{X} \stackrel{\text{iid}}{\sim} \text{Normal}(\mu, \sigma^2)$  with  $\mu$  known and consider the tests  $H_0 : \sigma^2 = \sigma_0^2$  &  $H_1 : \sigma^2 = \sigma_1^2$  with  $0 < \sigma_0^2 < \sigma_1^2$ . Then

$$\begin{aligned} T_{NP}(\mathbf{x}) &= \prod_{i=1}^n \frac{e^{-\frac{(x_i - \mu)^2}{2\sigma_1^2}}}{e^{-\frac{(x_i - \mu)^2}{2\sigma_0^2}}} \\ &= \prod_{i=1}^n e^{-(x_i - \mu)^2 \left[ \frac{1}{2\sigma_0^2} - \frac{1}{2\sigma_1^2} \right]} \\ &= e^{-\left[ \frac{1}{2\sigma_0^2} - \frac{1}{2\sigma_1^2} \right] \sum (x_i - \mu)^2} \end{aligned}$$

Since  $\sigma_0^2 < \sigma_1^2 \implies 0 < \frac{1}{2\sigma_0^2} - \frac{1}{2\sigma_1^2}$  Thus  $T_{NP}$  is increasing with  $-\sum (x_i - \mu)^2$  meaning  $T(\mathbf{x}) = -\sum (x_i - \mu)^2$  is an equivalent statistic to  $T_{NP}$ .