Stochastic Optimisation - Problem Sheet 4

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Answer 1) i)

$$\mathbb{P}^{\pi}(X_{0:t} = s_{0:t}, Y_{0:t} = a_{0:t})$$

$$= \mathbb{P}(X_0 = s_0) \cdot \frac{\mathbb{P}(X_0 = s_0, Y_0 = a_0)}{\mathbb{P}(X_0 = s_0)} \prod_{k=0}^{t-1} \frac{\mathbb{P}^{\pi}(X_{0:k+1} = s_{0:k+1}, Y_{0:k} = a_{0:k})}{\mathbb{P}^{\pi}(X_{0:k} = s_{0:k}, Y_{0:k} = a_{0:k})} \cdot \frac{\mathbb{P}^{\pi}(X_{0:k+1} = s_{0:k}, Y_{0:k} = a_{0:k})}{\mathbb{P}^{\pi}(X_{0:k} = s_{0:k}, Y_{0:k+1} = a_{0:k})}$$

$$= p_{X_0}(s_0)\mathbb{P}(Y_0 = a_0|X_0 = s_0) \prod_{k=0}^{t-1} \mathbb{P}^{\pi}(X_{k+1} = s_{k+1}|X_{0:k} = s_{0:k}, Y_{0:k} = a_{0:k}) \cdot \mathbb{P}^{\pi}(Y_k = a_k|X_{0:k} = s_{0:k}, Y_{0:k-1} = a_{0:k})$$

$$= p_{X_0}(s_0)q_0(a_0|s_0) \prod_{k=0}^{t-1} \mathbb{P}^{\pi}(X_{k+1} = s_{k+1}|X_k = s_k, Y_k = a_k) \cdot \mathbb{P}^{\pi}(Y = a_k|X_k = s_k, Y_{k-1} = a_{k-1}) \text{ by Markov Proposition of the proposition of t$$

Answer 1) ii)

$$\mathbb{P}^{\pi}(X_{0:t} = s_{0:k})$$
= $\mathbb{P}(X_0 = s_0) \prod_{k=1}^{t} \mathbb{P}^{\pi}(X_k = s_k | X_{0:k-1} = s_{0:k-1})$ by Bayes Rule

= $p_{X_0}(s_0) \prod_{k=1}^{t} \mathbb{P}^{\pi}(X_k = s_k | X_{k-1} = s_{k-1})$ by Markov Property

= $p_{X_0}(s_0) \prod_{k=1}^{t} \left(\sum_{a \in A(s_k)} \mathbb{P}^{\pi}(X_k = s_k, Y_{k-1} = a | X_{k-1} = s_{k-1}) \right)$ by Marginalisation

= $p_{X_0}(s_0) \prod_{k=1}^{t} \left(\sum_{a \in A(s_k)} \mathbb{P}^{\pi}(X_k = s_k | Y_{k-1} = a, X_{k-1} = s_{k-1}) \mathbb{P}(Y_{k-1}a | X_{k-1} = s_{k-1}) \right)$ by Bayes Rule

= $p_{X_0}(s_0) \prod_{k=1}^{t} \left(\sum_{a \in A(s_k)} \mathbb{P}^{\pi}(X_k = s_k | Y_{k-1} = a, X_{k-1} = s_{k-1}) \mathbb{P}(Y_{k-1}a | X_{k-1} = s_{k-1}) \right)$ by Bayes Rule

Answer 1) iii)

$$\begin{split} & = \sum_{s_{0:t} \in S^{t+1}}^{\pi} \mathbb{P}^{\pi}(X_{0:t} = a_{0:t}) \\ & = \sum_{s_{0:t} \in S^{t+1}}^{\pi} \mathbb{P}^{\pi}(X_{0:t} = s_{0:t}, Y_{0:t} = a_{0:t}) \text{ by Marginalisation} \\ & = \sum_{s_{0:t} \in S^{t+1}} \left[p_{X_0}(s_0) q_0(a_0|s_0) \prod_{k=1}^t p_k(s_{k+1}|s_k, a_k) q_k(a_k|s_{0:k}, a_{0:k-1}) \right] \text{ by 1) i)} \end{split}$$

Answer 2) a.

To show that $\{X_t\}_{t\in T}$ is a markov chain it is sufficient to show that

$$\mathbb{P}(X_{t+1} = s_{t+1} | X_{0:t} = s_{0:t}) = \mathbb{P}(X_{t+1} = s_{t+1} | X_t = s_t)$$

By the definition of conditional probabilities we have

$$\mathbb{P}(X_{t+1} = s_{t+1} | X_{0:t} = s_{0:t}) = \frac{\mathbb{P}(X_{0:t+1} = s_{0:t+1})}{\mathbb{P}(X_{0:t} = s_{0:t})} \quad [1]$$

By marginalising we can re-express the numerator as

$$\mathbb{P}(X_{0:t+1} = s_{0:t+1}) = \sum_{a_{0:t} \in A^{t+1}} \mathbb{P}(X_{0:t+1} = s_{0:t+1}, Y_{0:t} = a_{0:t}) \quad [2]$$

Further, we can expand each term of this summation

$$\begin{split} & & \mathbb{P}(X_{0:t+1} = s_{0:t+1}, Y_{0:t} = a_{0:t}) \\ & = & \frac{\mathbb{P}(X_{0:t+1} = s_{0:t+1}, Y_{0:t} = a_{0:t})}{\mathbb{P}(X_{0:t} = s_{0:t}, Y_{0:t} = a_{0:t})} \cdot \frac{\mathbb{P}(X_{0:t} = s_{0:t} = Y_{0:t} = a_{0:t})}{\mathbb{P}(X_{0:t} = s_{0:t}, Y_{0:t-1} = a_{0:t-1})} \cdot \mathbb{P}(X_{0:t} = s_{0:t}, Y_{0:t-1} = a_{0:t-1}) \\ & = & \mathbb{P}(X_{t+1} = s_{t+1} | X_{0:t} = s_{0:t}, Y_{0:t} = a_{0:t}) \mathbb{P}(Y_t = a_t | X_{0:t} = s_{0:t}, Y_{0:t-1} = a_{0:t-1}) \mathbb{P}(X_{0:t} = s_{0:t}, Y_{0:t-1} = a_{0:t-1}) \\ & = & \mathbb{P}(X_{t+1} = s_{t+1} | X_t = s_t, Y_t = a_t) \mathbb{P}(Y_t = a_t | X_t = s_t, Y_{t-1} = a_{t-1}) \mathbb{P}(X_{0:t} = s_{0:t}, Y_{0:t-1} = a_{0:t-1}) \\ & = & p_t(s_{t+1} | s_t, a_t) q_t(a_t | s_t) \mathbb{P}(X_{0:t} = s_{0:t}, Y_{0:t-1} = a_{0:t-1}) \end{split}$$

where the penultimate step is due to the use of Markovian randomised decision rules. Substituting this expression back into [2] gives the following expression

$$\mathbb{P}(X_{0:t+1} = s_{0:t+1}) = \sum_{a_{0:t} \in A^{t+1}} p_t(s_{t+1}|s_t, a_t) q_t(a_t|s_t) \mathbb{P}(X_{0:t} = s_{0:t}, Y_{0:t-1} = a_{0:t-1})$$

This expression can be restated as a recursive expression

$$\begin{split} & \mathbb{P}(X_{0:t+1} = s_{0:t+1}) \\ & = \sum_{a_{0:t} \in A^{t+1}} p_t(s_{t+1}|s_t, a_t) q_t(a_t|s_t) \mathbb{P}(X_{0:t} = s_{0:t}, Y_{0:t-1} = a_{0:t-1}) \\ & = \sum_{a_t \in A} \sum_{a_{0:t-1} \in A^t} p_t(s_{t+1}|s_t, a_t) q_t(a_t|s_t) \mathbb{P}(X_{0:t} = s_{0:t}, Y_{0:t-1} = a_{0:t-1}) \\ & = \left(\sum_{a_t \in A} p_t(s_{t+1}|s_t, a_t) q_t(a_t|s_t)\right) \left(\sum_{a_{0:t-1} \in A^t} \mathbb{P}(X_{0:t} = s_{0:t}, Y_{0:t-1} = a_{0:t-1})\right) \\ & = \left(\sum_{a_t \in A} p_t(s_{t+1}|s_t, a_t) q_t(a_t|s_t)\right) \mathbb{P}(X_{0:t} = s_{0:t}) \end{split}$$

Substituting this expression into [1] we get

$$\mathbb{P}(X_{t+1} = s_{t+1} | X_{0:t} = s_{0:t}) = \frac{\mathbb{P}(X_{0:t+1} = s_{0:t+1})}{\mathbb{P}(X_{0:t} = s_{0:t})} \\
= \frac{\sum_{a_t \in A} p_t(s_{t+1} | s_t, a_t) q_t(a_t | s_t)}{\mathbb{P}(X_{0:t} = s_{0:t})} \\
= \sum_{a_t \in A} p_t(s_{t+1} | s_t, a_t) q_t(a_t | s_t)$$

Since the RHS of this final expression is independent of $s_{0:t-1}$, $\{X_t\}_{t\in T}$ is a Markov chain.

Answer 2) b.

From 2) a. we deduced that the transition kernel of our process is

$$\mathbb{P}(X_{t+1} = s' | X_t = s) = \sum_{a \in A} p_t(s' | s, a) q_t(a | s)$$

Since the transition and decision probabilities are stationary we have that $p(\cdot|\cdot,\cdot) = p_t$ and $q(\cdot|\cdot) = q_t(\cdot|\cdot)$ for all $t \in T$. Meaning the transition kernel can be expressed homogeneously as

$$\mathbb{P}(X_{t+1} = s' | X_t = s) = \sum_{a \in A} p(s' | s, a) q(a | s)$$

Answer 3)

For ease of notation let $p_{Y_t}(a) := \mathbb{P}(Y_t = a | X_{0:t} = s_{0:t}, Y_{0:t-1} = a_{0:t-1}).$

In order to show that when the agent uses the procedure defined in the question they do indeed implement the decision rule with decision function $q_t(a|s_{0:t}, a_{0:t-1})$, I shall show that

$$p_{Y_t}(a) = q_t(a|s_{0:t}, a_{0:t-1})$$

By the definition of Y_t , we have

$$p_{Y_t}(a) = \mathbb{P}(\Phi_t(U_t|s_{0:t}, a_{0:t-1}) = a|X_{0:t} = s_{0:t}, Y_{0:t-1} = a_{0:t-1})$$

By the definition of $\Phi_t(\cdot)$, we have three separate cases

$$p_{Y_t}(a) = \begin{cases} \mathbb{P}(U_t \leq q_t(a(1)|s_{0:t}, a_{0:t-1})) & \text{if } a = a(1) \\ \mathbb{P}\left(U_t \in \left(\sum_{j=1}^{k-1} q_t(a(j)|s_{0:t}, a_{0:t-1}), \sum_{j=1}^k q_t(a(j)|s_{0:t}, a_{0:t-1})\right]\right) & \text{if } a = a(k) \text{ with } k \in [2, M-1] \\ \mathbb{P}\left(U_t > \sum_{j=1}^{M-1} q_t(a(j)|s_{0:t}, a_{0:t-1})\right) & \text{if } a = a(M) \end{cases}$$

Since U_t is a uniform distribution on [0,1] and $q_t(\cdot)$ is a probability distribution, we can restate these cases as

$$p_{Y_t}(a) = \begin{cases} q_t(a(1)|s_{0:t}, a_{0:t-1}) & \text{if } a = a(1) \\ \sum_{j=1}^k q_t(a(j)|s_{0:t}, a_{0:t-1}) - \sum_{j=1}^{k-1} q_t(a(j)|s_{0:t}, a_{0:t-1}) & \text{if } a = a(k) \text{ with } k \in [2, M-1] \\ 1 - \sum_{j=1}^{M-1} q_t(a(j)|s_{0:t}, a_{0:t-1})) & \text{if } a = a(M) \end{cases}$$

By the fact that $q_t(\cdot)$ is a probability distribution, and the definition of summation, the cases can be further simplified to

$$p_{Y_t}(a) = \begin{cases} q_t(a(1)|s_{0:t}, a_{0:t-1}) & \text{if } a = a(1) \\ q_t(a(k)|s_{0:t}, a_{0:t-1}) & \text{if } a = a(k) \text{ with } k \in [2, M-1] \\ q_t(a(M)|s_{0:t}, a_{0:t-1}) & \text{if } a = a(M) \end{cases}$$

This final set of cases are just a partition of the decision function, and thus can be simplified to

$$p_{Y_t}(a) = q_t(a|s_{0:t}, a_{0:t-1}) \ \forall \ a \in A$$