```
title: "jess daly spatial hw"
output: html document
```{r}
library(vegan)
data(BCI)
UTM Coordinates (in metres)
BCI xy = data.frame(x = rep(seq(625754, 626654, by=100), each=5),
 y = rep(seq(1011569, 1011969, by=100), len=50))
First check out quantities for all species present in the dataset.
```{r}
BCIcol<-colSums (BCI, na.rm = FALSE, dims = 1)
```{r}
hist(BCIcol)
Choosing Alseis blackiana as common Andira inermis as rare species (check out
total number of trees to get an idea of what common and rare mean here).
```{r}
sum(BCI [,11])
sum(BCI [,14])
```{r}
com<-BCI [,11]
rare<-BCI [,14]
First conduct spatial analysis of common species
```{r}
plot(BCI xy, cex = com/max(com))
```{r}
col brks = hist(com, plot=F)$breaks
col indices = as.numeric(cut(com, col brks))
cols = rev(terrain.colors(length(col brks)))
plot(BCI xy, cex=2, pch=19, col=cols[col indices])
```

Most of the dots indicate that most patches have few trees, but towards the top right there is a cluster of high density, indicating this species may display spatial dependence.

```
```{r}
BCI dist = dist(com)
xy dist = dist(BCI xy)
max dist = max(BCI xy) / 2
```{r}
plot(xy dist, BCI dist)
abline(lm(BCI dist ~ xy dist), lwd=3, col='red')
lines(lowess(xy dist, BCI dist), lwd=3, col='pink')
abline(v = max dist, col='red', lwd=3, lty=2)1
```{r}
obs_cor = cor(xy_dist, BCI dist)
obs cor
```{r}
nperm = 1000
null cor = obs cor
for (i in 2:nperm) {
 tmp xy = BCI xy[sample(nrow(BCI xy)),]
 null cor[i] = cor(dist(tmp xy), BCI dist)
sum(null cor >= obs cor) / nperm
My p-value indicates there is no significant correlation here. Taking a look
at the plot, there doesn't appear to be much there in terms of trends in the
data.
BCI mantel = mantel(BCI, BCI dist)
BCI mantel
The mantel test returns a value very close to 0, indicating a lack of
correlation. Overall there seems to be a lack of evidence for spatial
dependance in Alseias blackiana. This makes sense because this is a commonly
found species, so we would expect to find it everywhere and not in specific
patches.
I will perform my analysis again for my rare species.
```{r}
```

plot(BCI xy, cex = rare/max(rare))

```
```{r}
col brks1 = hist(rare, plot=F)$breaks
col indices1 = as.numeric(cut(rare, col brks1))
cols1 = rev(terrain.colors(length(col brks1)))
plot(BCI xy, cex=2, pch=19, col=cols1[col indices1])
There looks like there might be some pattern in dispersion since some patches
have no trees at all, but it is hard to tell.
```{r}
BCI dist1 = dist(rare)
xy dist = dist(BCI xy)
max dist = max(BCI xy) / 2
```{r}
plot(xy dist, BCI dist1)
abline(lm(BCI dist1 ~ xy dist), lwd=3, col='red')
lines(lowess(xy dist, BCI dist1), lwd=3, col='pink')
abline(v = max dist, col='red', lwd=3, lty=2)1
```{r}
obs cor1 = cor(xy dist, BCI dist1)
obs cor1
```{r}
nperm = 1000
null cor = obs cor1
for (i in 2:nperm) {
 tmp xy = BCI xy[sample(nrow(BCI xy)),]
 null cor[i] = cor(dist(tmp xy), BCI dist1)
sum(null cor >= obs cor) / nperm
The p-value indicates no significant correlation.
```{r}
BCI mantel1 = mantel(BCI, BCI dist1)
BCI mantel1
The models returned no evidence of spatial dependence in either species. This
makes less sense for the rare species, since you would expect that if it's
rare it may only be able to live in very specific condtions and thus certain
areas would have higer concentrations. However, the rarer the tree, the harder
it is to detect patterns in spatial dependence, which may be the issue here.
```

```
I am creating two linear models to predict the abundance of Drypetes
standleyi. In the first one I am using a single species, Cordia lasiocalyx
(randomly selected) as my predictor.
```{r}
sp ids = c("Cordia.lasiocalyx", "Hirtella.triandra",
 "Picramnia.latifolia", "Quassia.amara",
 "Tabernaemontana.arborea", "Trattinnickia.aspera",
 "Xylopia.macrantha")
allIDS<-cbind("Cordia.lasiocalyx", "Hirtella.triandra",
 "Picramnia.latifolia", "Quassia.amara",
 "Tabernaemontana.arborea", "Trattinnickia.aspera",
 "Xylopia.macrantha", "Drypetes.standleyi")
```{r}
allpre <- apply(allIDS, 1, function(x) sum(x > 0))
pre dat<-data.frame(allpre, BCI, BCI xy)</pre>
sub dat<-subset(pre dat, select=allIDS)</pre>
```{r}
spe<-BCI$Drypetes.standleyi</pre>
pre<-BCI$Cordia.lasiocalyx</pre>
Variograms show variability between points as a function of distance. I will
run a series of variograms to determine which best fits my data.
```{r}
BCI dat = data.frame(spe, BCI xy)
BCI lm = gls(spe ~ pre, data=BCI dat)
plot(Variogram(BCI lm, form= ~ x + y))
Points seem to fit this model quite well.
```{r}
res = residuals(BCI lm)
plot(dist(BCI dat[, c('x', 'y')]), dist(res))
lines(lowess(dist(BCI dat[, c('x', 'y')]), dist(res)), col='red', lwd=2)
abline(v = max dist, col='red', lwd=3, lty=2)
```{r}
BCI exp = update(BCI lm, corr=corExp(form=~x + y))
```{r}
```

```
plot(Variogram(BCI exp, maxDist = max dist))
This one not so good. The model includes a nugget, or non-zero y-intercept.
Next examine residuals for patterns.
```{r}
plot(Variogram(BCI exp, resType='normalized', maxDist = max dist))
The residuals show a positive trend. Ideally they should be randomly
disributed and show no pattern, so this model may not be a good fit.
```{r}
BCI exp nug = update(BCI exp, corr=corExp(form=~x + y, nugget=T))
plot(Variogram(BCI exp nug, maxDist = max dist))
The same model with a nugget looks to fit this data very well.
```{r}
plot(Variogram(BCI exp nug, resType='n', maxDist = max dist))
However, residuals still show a positive trend.
```{r}
BCI rat nug = update(BCI lm, corr=corRatio(form=~x + y, nugget=T))
plot(Variogram(BCI rat nug, maxDist = max dist))
The rational quadratic model is not a good fit.
```{r}
plot(Variogram(BCI rat nug, resType='n', maxDist = max dist))
```{r}
anova(BCI lm, BCI exp, BCI exp nug, BCI rat nug, test=F)
Interestingly according to the ANOVA, all of the models are nearly equal in
terms of fitting the data, though lm is a little worse than the others.
```{r}
summary(BCI exp nug)
```{r}
col brks2 = hist(residuals(BCI exp nug), plot=F)$breaks
col indices2 = as.numeric(cut(residuals(BCI exp nug), col brks2))
cols2 = rev(terrain.colors(length(col brks2)))
```

```
plot(BCI xy, cex=2, pch=19, col=cols2[col indices2])
Judging from the plot it certainly looks like there is a spatial pattern here,
with a high concentration in the bottom right corner.
Now I will rerun my variograms using all 7 given species together as a
predictor.
```{r}
x<-BCI xy$x
y<-BCI xy$y
```{r}
BCI lm1 = gls(Drypetes.standleyi ~ Cordia.lasiocalyx + Hirtella.triandra +
Picramnia.latifolia + Quassia.amara + Tabernaemontana.arborea +
Trattinnickia.aspera + Xylopia.macrantha, data=sub dat)
plot(Variogram(BCI lm1, form= ~ x + y))
```{r}
res1 = residuals(BCI lm1)
plot(dist(BCI dat[, c('x', 'y')]), dist(res1))
lines(lowess(dist(BCI dat[, c('x', 'y')]), dist(res1)), col='red', lwd=2)
abline(v = max dist, col='red', lwd=3, lty=2)
```{r}
BCI exp1 = update(BCI lm1, corr=corExp(form=~x + y))
```{r}
plot(Variogram(BCI exp1, maxDist = max dist))
```{r}
plot(Variogram(BCI exp1, resType='normalized', maxDist = max dist))
```{r}
BCI exp nug1 = update(BCI exp1, corr=corExp(form=~x + y, nugget=T))
plot(Variogram(BCI exp nug1, maxDist = max dist))
```{r}
plot(Variogram(BCI exp nug1, resType='n', maxDist = max dist))
```{r}
BCI rat nug1 = update(BCI lm1, corr=corRatio(form=~x + y, nugget=T))
plot(Variogram(BCI rat nug1, maxDist = max dist))
```

```
```{r}
plot(Variogram(BCI rat nug1, resType='n', maxDist = max dist))
Similarly to when I used only a single species, the rational quadratic error
model does not appear to be as useful as the others; however, it is the only
model in which the residuals have no obvious pattern.
```{r}
anova (BCI lm1, BCI exp1, BCI exp nug1, BCI rat nug1, test=F)
Once again my ANOVA indicates all model are about even when it comes to
representing the data at hand.
```{r}
summary(BCI exp nug1)
```{r}
col brks3 = hist(residuals(BCI exp nug1), plot=F)$breaks
col indices3 = as.numeric(cut(residuals(BCI exp nug1), col brks3))
cols3 = rev(terrain.colors(length(col brks3)))
plot(BCI xy, cex=2, pch=19, col=cols3[col indices3])
```

Once again there appears to be evidence of a spatial relationship in this plot, with higher concentrations on the far right side.

Using a single vs. multiple species overall produced similar results. Including the nugget (the spatial error term) in analysis did not change the effectivness of the models in either case. If you look at the variograms for the first 2 models, the scale of the y-axis is very small. This may cause the nugget effect to essentally become non-existant, since the value is so close to 0.