# Difference strokes for proportions, folks

HYPOTHESIS TESTING IN PYTHON



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#### Chapter 1 recap

• Is a claim about an unknown population proportion feasible?

- 1. Standard error of sample statistic from bootstrap distribution
- 2. Compute a standardized test statistic
- 3. Calculate a p-value
- 4. Decide which hypothesis made most sense

• Now, calculate the test statistic without using the bootstrap distribution

#### Standardized test statistic for proportions

p: population proportion (unknown population parameter)

 $\hat{p}$ : sample proportion (sample statistic)

 $p_0$ : hypothesized population proportion

$$z = rac{\hat{p} - ext{mean}(\hat{p})}{ ext{SE}(\hat{p})} = rac{\hat{p} - p}{ ext{SE}(\hat{p})}$$

Assuming  $H_0$  is true,  $p=p_0$ , so

$$z = rac{\hat{p} - p_0}{ ext{SE}(\hat{p})}$$

#### Simplifying the standard error calculations

$$SE_{\hat{p}} = \sqrt{rac{p_0*(1-p_0)}{n}} o$$
 Under  $H_0$ ,  $SE_{\hat{p}}$  depends on hypothesized  $p_0$  and sample size  $n$ 

Assuming  $H_0$  is true,

$$z=rac{\hat{p}-p_0}{\sqrt{rac{p_0*(1-p_0)}{n}}}$$

ullet Only uses sample information ( $\hat{p}$  and n) and the hypothesized parameter  $(p_0)$ 

#### Why z instead of t?

$$t = rac{\left(ar{x}_{
m child} - ar{x}_{
m adult}
ight)}{\sqrt{rac{s_{
m child}^2}{n_{
m child}} + rac{s_{
m adult}^2}{n_{
m adult}}}}$$

- ullet s is calculated from  $ar{x}$ 
  - $\circ \;\; ar{x}$  estimates the population mean
  - $\circ$   $\;s$  estimates the population standard deviation
  - † uncertainty in our estimate of the parameter
- t-distribution fatter tails than a normal distribution
- $oldsymbol{\hat{p}}$  only appears in the numerator, so z-scores are fine

#### Stack Overflow age categories

 $H_0$ : Proportion of Stack Overflow users under thirty =0.5

 $H_A$ : Proportion of Stack Overflow users under thirty eq 0.5

```
alpha = 0.01
```

```
stack_overflow['age_cat'].value_counts(normalize=True)
```

```
Under 30 0.535604
```

At least 30 0.464396

Name: age\_cat, dtype: float64



#### Variables for z

```
p_hat = (stack_overflow['age_cat'] == 'Under 30').mean()
```

#### 0.5356037151702786

```
p_0 = 0.50
```

n = len(stack\_overflow)

2261

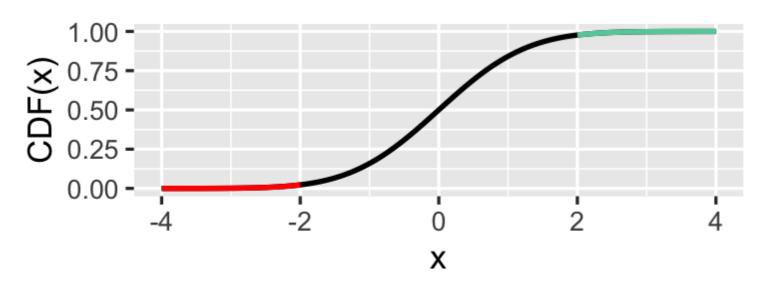
#### Calculating the z-score

$$z=rac{\hat{p}-p_0}{\sqrt{rac{p_0*(1-p_0)}{n}}}$$

```
import numpy as np
numerator = p_hat - p_0
denominator = np.sqrt(p_0 * (1 - p_0) / n)
z_score = numerator / denominator
```

3.385911440783663

#### Calculating the p-value



Left-tailed ("less than"):

```
from scipy.stats import norm
p_value = norm.cdf(z_score)
```

Right-tailed ("greater than"):

```
p_value = 1 - norm.cdf(z_score)
```

Two-tailed ("not equal"):

```
p_value = norm.cdf(-z_score) +
   1 - norm.cdf(z_score)
```

```
p_value = 2 * (1 - norm.cdf(z_score))
```

0.0007094227368100725

```
p_value <= alpha</pre>
```

True

## Let's practice!

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# A sense of proportion

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#### Comparing two proportions

 $H_0$ : Proportion of hobbyist users is the same for those under thirty as those at least thirty

$$H_0$$
:  $p_{\geq 30} - p_{< 30} = 0$ 

 $H_A$ : Proportion of hobbyist users is different for those under thirty to those at least thirty

$$H_A$$
:  $p_{\geq 30} - p_{<30} 
eq 0$ 

alpha = 0.05

#### Calculating the z-score

• z-score equation for a *proportion test*:

$$z = rac{(\hat{p}_{\geq 30} - \hat{p}_{<30}) - 0}{ ext{SE}(\hat{p}_{>30} - \hat{p}_{<30})}$$

• Standard error equation:

$$ext{SE}(\hat{p}_{\geq 30} - \hat{p}_{< 30}) = \sqrt{rac{\hat{p} imes (1 - \hat{p})}{n_{\geq 30}}} + rac{\hat{p} imes (1 - \hat{p})}{n_{< 30}}$$

•  $\hat{p}$  ightarrow weighted mean of  $\hat{p}_{\geq 30}$  and  $\hat{p}_{<30}$ 

$$\hat{p} = rac{n_{\geq 30} imes \hat{p}_{\geq 30} + n_{<30} imes \hat{p}_{<30}}{n_{>30} + n_{<30}}$$

• Only require  $\hat{p}_{>30}$ ,  $\hat{p}_{<30}$ ,  $n_{>30}$ ,  $n_{<30}$  from the sample to calculate the z-score

```
p_hats = stack_overflow.groupby("age_cat")['hobbyist'].value_counts(normalize=True)
```

```
n = stack_overflow.groupby("age_cat")['hobbyist'].count()
```

```
age_cat
At least 30 1050
Under 30 1211
Name: hobbyist, dtype: int64
```



```
p_hats = stack_overflow.groupby("age_cat")['hobbyist'].value_counts(normalize=True)
```

```
p_hat_at_least_30 = p_hats[("At least 30", "Yes")]
p_hat_under_30 = p_hats[("Under 30", "Yes")]
print(p_hat_at_least_30, p_hat_under_30)
```

```
0.773333 0.843105
```



```
n = stack_overflow.groupby("age_cat")['hobbyist'].count()
```

```
age_cat
At least 30 1050
Under 30 1211
Name: hobbyist, dtype: int64
```

```
n_at_least_30 = n["At least 30"]
n_under_30 = n["Under 30"]
print(n_at_least_30, n_under_30)
```

1050 1211



-4.223718652693034

#### Proportion tests using proportions\_ztest()

```
stack_overflow.groupby("age_cat")['hobbyist'].value_counts()
```

```
age_cat hobbyist

At least 30 Yes 812

No 238

Under 30 Yes 1021

No 190

Name: hobbyist, dtype: int64
```

```
(-4.223691463320559, 2.403330142685068e-05)
```



## Let's practice!

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# Declaration of independence

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#### Revisiting the proportion test

```
age_by_hobbyist = stack_overflow.groupby("age_cat")['hobbyist'].value_counts()
```

```
age_cat hobbyist

At least 30 Yes 812

No 238

Under 30 Yes 1021

No 190

Name: hobbyist, dtype: int64
```

```
(-4.223691463320559, 2.403330142685068e-05)
```



#### Independence of variables

Previous hypothesis test result: evidence that hobbyist and age\_cat are associated

**Statistical independence** - proportion of successes in the response variable is the same across all categories of the explanatory variable



#### Test for independence of variables

```
chi2 dof
                        lambda
                test
                                                   pval
                                                           cramer
                                                                      power
             pearson 1.000000 17.839570 1.0 0.000024
                                                         0.088826
                                                                   0.988205
0
                                               0.000024
        cressie-read 0.666667 17.818114 1.0
                                                         0.088773
                                                                   0.988126
      log-likelihood 0.000000 17.802653
                                               0.000025
2
                                          1.0
                                                         0.088734
                                                                   0.988069
3
       freeman-tukey -0.500000 17.815060
                                          1.0
                                               0.000024
                                                         0.088765
                                                                   0.988115
  mod-log-likelihood -1.000000 17.848099 1.0
                                               0.000024
                                                         0.088848
                                                                   0.988236
5
              neyman -2.000000 17.976656
                                          1.0
                                               0.000022
                                                         0.089167
                                                                   0.988694
```

$$\chi^2$$
 statistic = 17.839570 =  $(-4.223691463320559)^2$  =  $(z$ -score) $^2$ 

#### Job satisfaction and age category

```
stack_overflow['age_cat'].value_counts()
```

stack\_overflow['job\_sat'].value\_counts()

Under 30 1211
At least 30 1050
Name: age\_cat, dtype: int64

```
Very satisfied 879
Slightly satisfied 680
Slightly dissatisfied 342
Neither 201
Very dissatisfied 159
Name: job_sat, dtype: int64
```

#### Declaring the hypotheses

 $H_0$ : Age categories are independent of job satisfaction levels

 $H_A$ : Age categories are not independent of job satisfaction levels

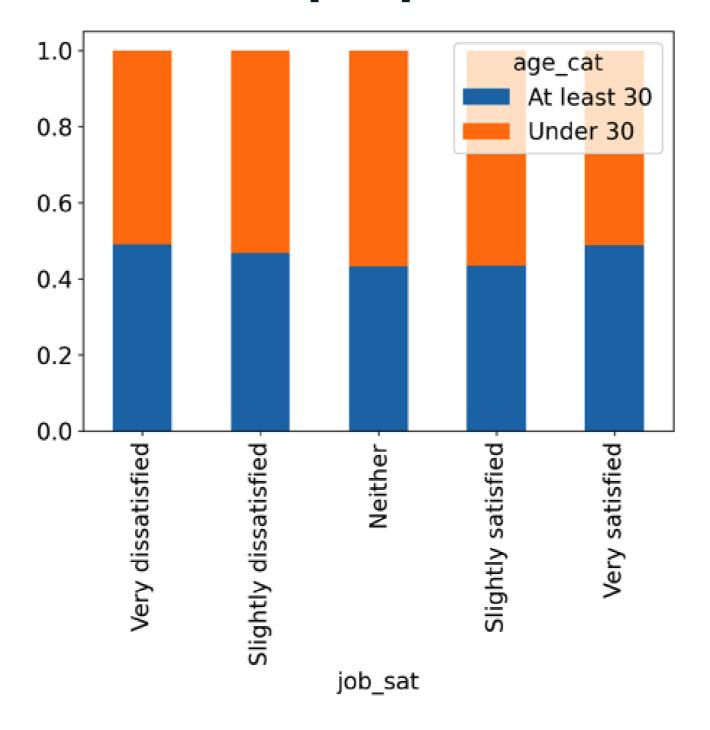
```
alpha = 0.1
```

- Test statistic denoted  $\chi^2$
- Assuming independence, how far away are the observed results from the expected values?

#### Exploratory visualization: proportional stacked bar plot

```
props = stack_overflow.groupby('job_sat')['age_cat'].value_counts(normalize=True)
wide_props = props.unstack()
wide_props.plot(kind="bar", stacked=True)
```

#### Exploratory visualization: proportional stacked bar plot



#### Chi-square independence test

```
import pingouin
expected, observed, stats = pingouin.chi2_independence(data=stack_overflow, x="job_sat", y="age_cat")
print(stats)
```

```
test
                       lambda
                                  chi2 dof
                                                 pval
                                                        cramer
                                                                   power
             pearson 1.000000 5.552373 4.0 0.235164 0.049555
                                                                0.437417
                                        4.0 0.235014
        cressie-read 0.666667
                               5.554106
                                                      0.049563
                                                                0.437545
                                       4.0 0.234632 0.049583
      log-likelihood 0.000000
                              5.558529
                                                               0.437871
       freeman-tukey -0.500000 5.562688 4.0 0.234274 0.049601
3
                                                                0.438178
  mod-log-likelihood -1.000000 5.567570 4.0 0.233854
                                                      0.049623
                                                               0.438538
              neyman -2.000000 5.579519 4.0 0.232828
                                                      0.049676 0.439419
5
```

#### Degrees of freedom:

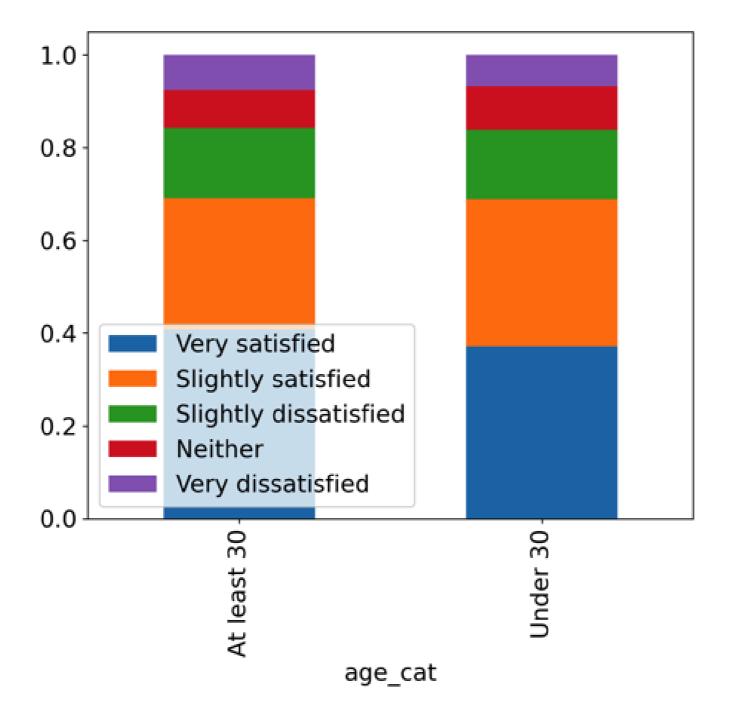
(No. of response categories -1)  $\times$  (No. of explanatory categories -1)

$$(2-1)*(5-1)=4$$

#### Swapping the variables?

```
props = stack_overflow.groupby('age_cat')['job_sat'].value_counts(normalize=True)
wide_props = props.unstack()
wide_props.plot(kind="bar", stacked=True)
```

#### Swapping the variables?





#### chi-square both ways

```
expected, observed, stats = pingouin.chi2_independence(data=stack_overflow, x="age_cat", y="job_sat")
print(stats[stats['test'] == 'pearson'])
```

```
test lambda chi2 dof pval cramer power
O pearson 1.0 5.552373 4.0 0.235164 0.049555 0.437417
```

Ask: Are the variables X and Y independent?

Not: Is variable X independent from variable Y?

#### What about direction and tails?

- Observed and expected counts squared must be non-negative
- ullet chi-square tests are almost always right-tailed  $^1$

<sup>&</sup>lt;sup>1</sup> Left-tailed chi-square tests are used in statistical forensics to detect if a fit is suspiciously good because the data was fabricated. Chi-square tests of variance can be two-tailed. These are niche uses, though.



## Let's practice!

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# Does this dress make my fit look good?

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#### Purple links

You search for a coding solution online and the first result link is purple because you already visited it. How do you feel?

```
purple_link_counts = stack_overflow['purple_link'].value_counts()

purple_link_counts = purple_link_counts.rename_axis('purple_link').reset_index(name='n')
```

```
purple_link n
0 Hello, old friend 1225
1 Indifferent 405
2 Amused 368
3 Annoyed 263
```

#### Declaring the hypotheses

```
hypothesized = pd.DataFrame({
    'purple_link': ['Hello, old friend', 'Amused', 'Indifferent', 'Annoyed'],
    'prop': [1/2, 1/6, 1/6, 1/6]})
```

```
purple_link prop
0 Hello, old friend 0.500000
1 Amused 0.166667
2 Indifferent 0.166667
3 Annoyed 0.166667
```

 $H_0$ : The sample matches with the hypothesized distribution

 $H_A$ : The sample does not match with the hypothesized distribution

 $\chi^2$  measures how far observed results are from expectations in each group

```
alpha = 0.01
```

#### Hypothesized counts by category

```
n_total = len(stack_overflow)
hypothesized["n"] = hypothesized["prop"] * n_total
```

```
      purple_link
      prop
      n

      0 Hello, old friend
      0.500000
      1130.500000

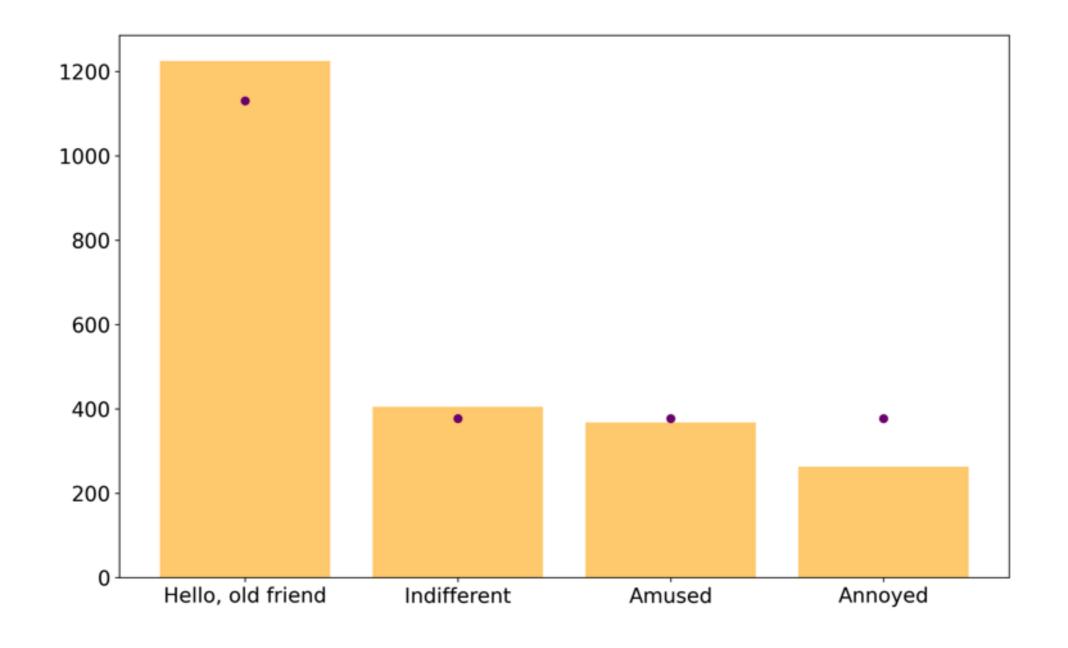
      1 Amused
      0.166667
      376.833333

      2 Indifferent
      0.166667
      376.833333

      3 Annoyed
      0.166667
      376.833333
```

#### Visualizing counts

#### Visualizing counts





#### chi-square goodness of fit test

print(hypothesized)

```
purple_link prop n

0 Hello, old friend 0.500000 1130.500000

1 Amused 0.166667 376.833333

2 Indifferent 0.166667 376.833333

3 Annoyed 0.166667 376.833333
```

```
from scipy.stats import chisquare
chisquare(f_obs=purple_link_counts['n'], f_exp=hypothesized['n'])
```

Power\_divergenceResult(statistic=44.59840778416629, pvalue=1.1261810719413759e-09)

## Let's practice!

HYPOTHESIS TESTING IN PYTHON

