# Final assignment

#### Prediction

We need to calculate P(survived | 2 positive nodes) =

## *P*(2 positive nodes| survived)P(survived)

 $\overline{P(2 \text{ positive nodes} | \text{ survived})P(\text{survived}) + P(2 \text{ positive nodes} | \text{ died})P(\text{died})}$ 

For the died distribution:

$$P(\text{died}) = \frac{81}{306} = 0.265$$

And according to geometric distribution 2nd version (with 0s):

Calculate p:

$$\frac{1-p}{p} = 7.5 \Rightarrow 7.5p = 1-p \Rightarrow 8.5p = 1 \Rightarrow p = 0.118$$

calculate P(2 positive nodes | died):

$$PMF(x = 2) = (1 - p)^2 p = (1 - 0.118)^2 \cdot 0.118 = 0.092$$

For the survived distribution:

P(survived) = 
$$\frac{225}{306}$$
 = 0.735

And according to geometric distribution 2nd version (with 0s):

Calculate p:

$$\frac{1-p}{p} = 2.8 \Rightarrow 2.8p = 1-p \Rightarrow 3.8p = 1 \Rightarrow p = 0.263$$

calculate P(2 positive nodes | survived):

$$PMF(x = 2) = (1 - p)^2 p = (1 - 0.263)^2 \cdot 0.263 = 0.143$$

So, the probability that the patient will survive within 5 years is:

$$P = \frac{0.143 \cdot 0.735}{0.143 \cdot 0.735 + 0.092 \cdot 0.265} = \frac{0.105}{0.105 + 0.024} = \frac{0.105}{0.129} = 0.814$$

### Likelihood

1. The likelihood function:

$$L(\theta|x_n) = \prod_{i=1}^n \frac{2x_i}{\theta} e^{\frac{-x_i^2}{\theta}} = \frac{2^n}{\theta^n} \cdot \prod_{i=1}^n x_i \cdot e^{\frac{-x_i^2}{\theta}}$$

Log likelihood:

$$\ln L(\theta|x_n) = \ln 2^n + \ln \theta^{-n} + \ln \prod_{i=1}^n x_i + \ln \prod_{i=1}^n e^{\frac{-x_i^2}{\theta}}$$

$$= n \ln 2 - n \ln \theta + \sum_{i=1}^n \ln (x_i) + \sum_{i=1}^n \ln \left( e^{\frac{-x_i^2}{\theta}} \right)$$

$$= n \ln 2 - n \ln \theta + \sum_{i=1}^n \ln (x_i) + \sum_{i=1}^n \frac{-x_i^2}{\theta} \ln(e)$$

$$= n \ln 2 - n \ln \theta + \sum_{i=1}^n \ln (x_i) + \sum_{i=1}^n \frac{-x_i^2}{\theta} \cdot 1$$

$$= n \ln 2 - n \ln \theta + \sum_{i=1}^n \ln (x_i) - \frac{1}{\theta} \sum_{i=1}^n x_i^2$$

The derivative of LL with respect to  $\theta$ :

$$\frac{\partial \ell}{\partial \theta} = -\frac{n}{\theta} + \frac{1}{\theta^2} \sum_{i=1}^{n} x_i^2$$

To find  $\theta$  that maximizes LL we will equal the derivative to zero:

$$\frac{\partial \ell}{\partial \theta} = 0 \Rightarrow -\frac{n}{\hat{\theta}} + \frac{1}{\hat{\theta}^2} \sum_{i=1}^{n} x_i^2 = 0$$

$$\frac{-n\hat{\theta} + \sum_{i=1}^{n} x_i^2}{\hat{\theta}^2} = 0$$

$$\theta > 0$$
, So:

$$-n\widehat{\theta} + \sum_{i=1}^{n} x_i^2 = 0$$

$$\widehat{\boldsymbol{\theta}} = \frac{\sum_{i=1}^{n} x_i^2}{n}$$

2. Estimate  $\theta$  from the sample (0.5,0.5,1):

$$\widehat{\theta} = \frac{\left(0.5^2 + 0.5^2 + 1^2\right)}{3} = 0.5$$

## Hypothesis

We need to test:

$$H_0$$
:  $p_x = p_y$  versus  $H_0$ :  $p_x > p_y$ 

$$\bar{X} \sim N\left(p_x, \frac{p_x(1-p_x)}{n_x}\right) \qquad \bar{Y} \sim N\left(p_y, \frac{p_y(1-p_y)}{n_y}\right)$$

$$\bar{X} - \bar{Y} \sim N \left( p_x - p_y, \frac{p_x(1-p_x)}{n_x} + \frac{p_y(1-p_y)}{n_y} \right)$$

Since sample sizes are "large enough" we can use normal approximation.

The test statistic:

$$z = \frac{\hat{p}_{x} - \hat{p}_{y}}{\sqrt{\frac{\hat{p}_{x}(1 - \hat{p}_{x})}{n_{x}} + \frac{\hat{p}_{y}(1 - \hat{p}_{y})}{n_{y}}}}$$

Where:

$$n_x = 100$$

$$n_y = 150$$

$$\hat{p}_x = \frac{60}{100} = 0.6$$

$$\hat{p}_y = \frac{70}{150} = 0.467$$

So:

$$z = \frac{0.6 - 0.467}{\sqrt{\frac{0.6 \cdot (1 - 0.6)}{100} + \frac{0.467 \cdot (1 - 0.467)}{150}}} = \frac{0.133}{\sqrt{\frac{0.24}{100} + \frac{0.249}{150}}} = 2.087$$

For 1% significance level we will reject  $H_0$  for z>2.326 (Upper-Tailed Test).

Since 2.087 < 2.326 we cannot reject  $H_0$  that  $p_x = p_y$ .