CARNEGIE MELLON UNIVERSITY APPLIED STOCHASTIC PROCESSES (COURSE 18-751) HOMEWORK 1

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Q.1 prove with Venn diagrams

(a) $A \cap B^c = A - B$

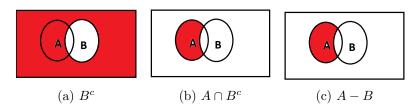


Figure 1

this implies $A \cap B^c = A - B$ is true

(b)
$$A \cup B^c = (A^c \cap B)^c$$

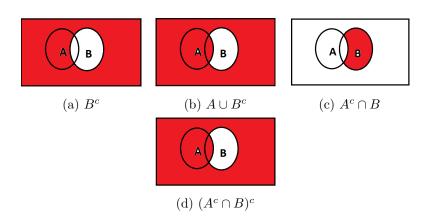


Figure 2

hence $A \cup B^c = (A^c \cap B)^c$ is true

(c)
$$B - A \neq A - B$$

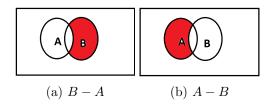


Figure 3

therefor $B-A \neq A-B$ is true

Q.2

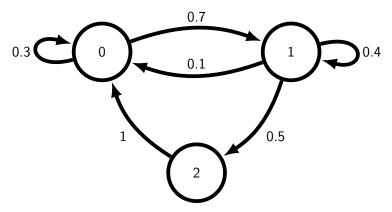


Figure 4: MC

(a) Irreduciblity and periodicity

Irreduciblity

A markov chain is said to be Irreducible if

$$\forall i, j \in S, \exists m < \infty : P(X_{n+m} = j | X_n = i) > 0 \tag{1}$$

i.e regardless the present state we can reach any other state in finit time.[1] On the Markov Chain in figure 4 we can go from any of the states to the other states in a finit number of steps hence it is irreducible.

Periodicity

for an irreducible Markov Chain the periodicity of state i is defined by

$$d(i) = g.c.d\{n \ge 1 | P^n(i,i) > 0\}$$
(2)

and d(i) has the same value d for all i. A Markov Chain is said to be aperiodic if d = 1 [2]

for the Markov Chain on figure 4 at n=1

$$P^1 = \begin{bmatrix} 0.3 & 0.7 & 0 \\ 0.1 & 0.4 & 0.5 \\ 1 & 0 & 0 \end{bmatrix}$$

already we can find i such that P(i,i) > 0. For example P(0,0) = 0.3 > 0 and the g.c.d of any set of integers that contains 1 is 1. hence figure 4 is

aperiodic.

Ans. Irreducible and Aperiodic.

(b) invariant distribution (π)

$$\pi_j = \sum_{k=1}^n \pi_k p_{kj} \tag{3}$$

$$\pi_0 = \pi_0 P_{00} + \pi_1 P_{10} + \pi_2 P_{20} \tag{4}$$

$$\pi_1 = \pi_0 P_{01} + \pi_1 P_{11} + \pi_2 P_{21} \tag{5}$$

$$\pi_2 = \pi_0 P_{02} + \pi_1 P_{12} + \pi_2 P_{22} \tag{6}$$

i.e $\pi = \pi P$

$$P = \begin{bmatrix} 0.3 & 0.7 & 0 \\ 0.1 & 0.4 & 0.5 \\ 1 & 0 & 0 \end{bmatrix}$$

$$\pi_0 = 0.3\pi_0 + 0.1\pi_1 + \pi_2 \tag{7}$$

$$\pi_1 = 0.7\pi_0 + 0.4\pi_1 \tag{8}$$

$$\pi_2 = 0.5\pi_1 \tag{9}$$

$$\pi_0 + \pi_1 + \pi_2 = 1 \tag{10}$$

lets write everything in terms of π_1

$$\pi_1 = \pi_1 \tag{11}$$

$$\pi_2 = 0.5\pi_1 \tag{12}$$

$$\pi_0 = 0.3\pi_0 + 0.1\pi_1 + 0.5\pi_1 \tag{13}$$

$$0.7\pi_0 = 0.6\pi_1 \tag{14}$$

$$\pi_0 = \frac{6}{7}\pi_1 \tag{15}$$

$$\pi_0 + \pi_1 + \pi_2 = \frac{6}{7}\pi_1 + \pi_1 + 0.5\pi_1 = 1 \tag{16}$$

$$\frac{33}{14}\pi_1 = 1\tag{17}$$

$$\pi_1 = \frac{14}{33} \tag{18}$$

$$\pi_0 = \frac{6}{7} \cdot \frac{14}{33} = \frac{4}{11} \tag{19}$$

$$\pi_{1} = \frac{14}{33}$$

$$\pi_{0} = \frac{6}{7} \cdot \frac{14}{33} = \frac{4}{11}$$

$$\pi_{2} = 0.5 \cdot \frac{14}{33} = \frac{7}{33}$$

$$(18)$$

$$(20)$$

Ans.

$$\pi = \begin{bmatrix} \frac{4}{11} & \frac{14}{33} & \frac{7}{33} \end{bmatrix}$$

$$\pi \approx \begin{bmatrix} 0.3636 & 0.4242 & 0.2121 \end{bmatrix}$$

(c) Expected Time from 0 to 2

We can calculate $\beta(0)$

i.e

 $\beta(0) = E[\text{average time to reach 2} \mid \text{current state is 0}]$

$$\beta(2) = 0 \tag{21}$$

$$\beta(0) = 1 + 0.7\beta(1) + 0.3\beta(0) \tag{22}$$

$$\beta(1) = 1 + 0.1\beta(1) + 0.4\beta(1) + 0.5\beta(2) \tag{23}$$

since $\beta(2) = 0$

$$\beta(1) = 1 + 0.1\beta(0) + 0.4\beta(1)\beta(0) = 1 + 0.3\beta(0) + 0.7\beta(1)$$
(24)

$$0.7\beta(0) - 0.7\beta(1) = 1\tag{25}$$

$$0.6\beta(1) - 0.1\beta(0) = 1 \tag{26}$$

solving the equations (25) and (26) we get $\beta(0) \approx 3.708$ and $\beta(1) \approx 2.286$

Ans. So the expected time from 0 to 2 is 3.708

(d) Probability that starting from 0, the MC has reached 2 after n-steps vs n

The n-step transition probability is defined by [3]

$$r_{ij}(n) = P(X_n = j | X_0 = i)$$
 (27)

 $r_{ij}(n)$ is the probability that the state after n time periods will be j, given that the current state is i.

We can get $r_{ij}(n)$ by using Chapman-Kolmogorov Equation [3]

$$r_{ij}(n) = \sum_{k=1}^{m} r_{ik}(n-1)p_{kj}$$
(28)

for n > 1, and all i, j, starting with

$$r_{ij}(1) = P_{ij}$$

[CODE] = Code Appendix 2. d

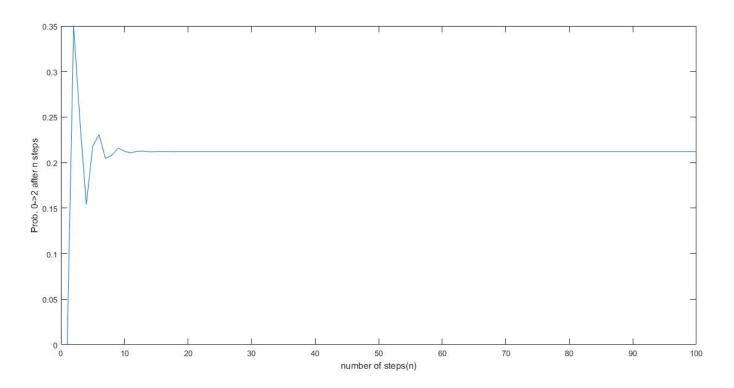


Figure 5: Probability that starting from 0, the MC has reached 2 after n-steps vs n

(e) Fraction of Time

[CODE] = Code Appendix 2. e

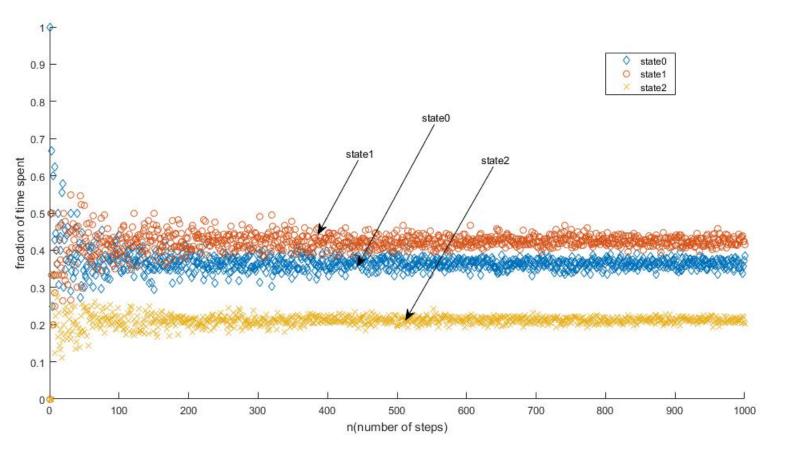


Figure 6: Fraction of time it spends in each of the states vs n

Remark:

On every run we get slightly different outcome just because of the inherent randomness of the Markov chain. But for large n all of the states will converge to the values in the invariant distribution.

(f) π_n vs n [CODE] = Code Appendix 2. f

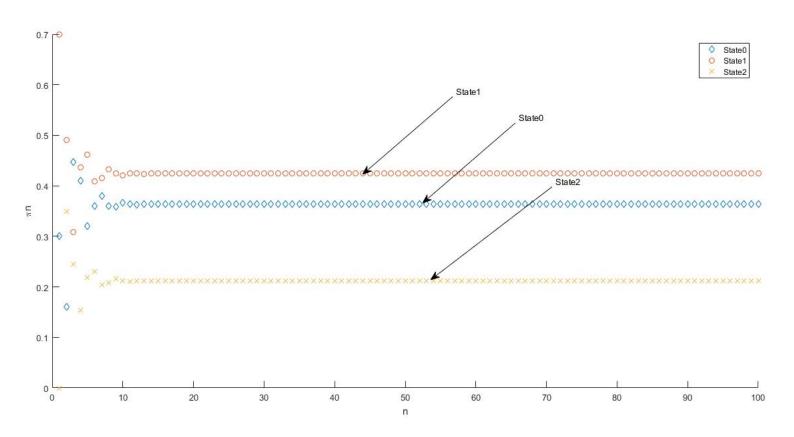


Figure 7: π_n vs n

$\mathbf{Q3}$

(a) Page Rank

We can rank the states by their value in the invariant distribution (π)

- 1. state(page) 1
- 2. state(page) 0
- 3. state(page) 2

(b) Removing self connections and Normalizing

After removing the self connections, there are an infinit number of ways of normalizing the Markov Chain. One way is

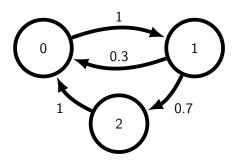


Figure 8: Updated MC

(c) Adding States to Trick the MC

After adding 1a and 0a we get a Markov Chain that looks like figure 9

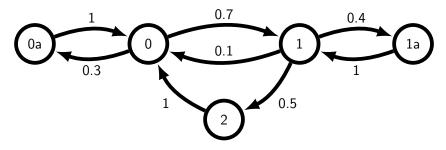


Figure 9: Updated MC

$$P = \begin{bmatrix} 0 & 1 & 0 & 0 & 0 \\ 0.3 & 0 & 0.7 & 0 & 0 \\ 0 & 0.1 & 0 & 0.5 & 0.4 \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \end{bmatrix}$$

$$\pi_{0a} = 0.3\pi_0 \tag{29}$$

$$\pi_0 = \pi_{0a} + 0.1\pi_1 + \pi_2 \tag{30}$$

$$\pi_1 = 0.7\pi_0 + \pi_{1a} \tag{31}$$

$$\pi_2 = 0.5\pi_1 \tag{32}$$

$$\pi_{1a} = 0.4\pi_1 \tag{33}$$

$$\pi_{0a} + \pi_0 + \pi_1 + \pi_2 + \pi_{1a} = 1 \tag{34}$$

lets write everything in terms of π_1

$$\pi_1 = \pi_1 \tag{35}$$

$$\pi_2 = 0.5\pi_1 \tag{36}$$

$$\pi_{1a} = 0.4\pi_1 \tag{37}$$

$$\pi_0 = 0.3\pi_0 + 0.1\pi_1 + 0.5\pi_1 \tag{38}$$

$$0.7\pi_0 = 0.6\pi_1 \tag{39}$$

$$\pi_0 = \frac{6}{7}\pi_1 \tag{40}$$

$$\pi_{0a} = 0.3. \frac{6}{7} \pi_1 = \frac{18}{70} \pi_1 \tag{41}$$

$$\frac{18}{70}\pi_1 + \frac{6}{7}\pi_1 + \pi_1 + \frac{1}{2}\pi_1 + \frac{4}{10}\pi_1 = 1 \tag{42}$$

$$\pi_1 = \frac{70}{211} \approx 0.331 \tag{43}$$

(44)

similarly $\pi_{oa}=\frac{18}{70}\pi_1\approx 0.085,\ \pi_0=\frac{6}{7}\pi_1\approx 0.284,\ \pi_2=0.5\pi_1\approx 0.165,\ \pi_{1a}=0.4\pi_1\approx 0.132$

$$\pi = \begin{bmatrix} 0.085 & 0.284 & 0.331 & 0.165 & 0.132 \end{bmatrix}$$

We can rank the states (0, 1 and 2) by their value in the invariant distribution (π) form highest to lowest

- 1. state(page) 1
- 2. state(page) 0
- 3. state(page) 2

for the MC in (b)

$$\pi_0 = 0.3\pi_1 + \pi_2 \tag{45}$$

$$\pi_1 = \pi_0 \tag{46}$$

$$\pi_2 = 0.7\pi_1 \tag{47}$$

$$\pi_0 + \pi_1 + \pi_2 = 1 \tag{48}$$

solving the above equations we get

$$\pi = \begin{bmatrix} \frac{10}{27} & \frac{10}{27} & \frac{7}{27} \end{bmatrix}$$

So the ranking for (b) would be

1. state(page) 1 = state(page) 0 i.e page1 and page 2 have the same rank 2. state(page) 2

Remark:

Compared to (a) adding the new states 0a and 1a do not change the ranking. On the other hand in (b), after removing and normalizing it, Pages 0 and 1 will have the same ranking because the exit probabilty from 0 to 1 is 1(i.e if we are at Page 0 the next Page we will look at is guaranteed to be Page 1). In all cases page 2 is ranked last.

Code Appendix

2. d

```
1 function [R, chap_kol] = chapmanKolmogorov(P, N, start, ending)
_{2} \operatorname{len}_{p} = \operatorname{size}(P);
R = zeros(len_p(1), len_p(1), N);
^{4} R(:,:,1) = P;
5 \text{ for } n=2:N
        for i = 1: len_p(1)
         for j=1:len_p(1)
              accum = 0;
              for k=1:len_p(1)
9
              accum = accum + R(i, k, n-1) * P(k, j);
10
11
              R(\,i\,\,,j\,\,,n\,)\,\,=\,\,accum\,;
12
         end
13
        end
14
15 end
chap_kol = R(start, ending,:);
17 end
1 N = 100;
^{2} P = \begin{bmatrix} 0.3 & 0.7 & 0; 0.1 & 0.4 & 0.5; 1 & 0 & 0 \end{bmatrix};
3 \text{ start} = 1;\% \text{state } 0
4 \text{ ending} = 3;\% \text{state } 2
[R, ch] = chapmanKolmogorov(P, N, start, ending);
6 \operatorname{ch} = \operatorname{ch}(:);
7 plot (ch);
8 xlabel('number of steps(n)');
9 ylabel ('Prob. 0->2 after n steps')
```

2. e [2] Appendix C.3

```
function T = discrete(P)
Pnorm = [0 P]/sum(P);
Pcum = cumsum(Pnorm);
R = rand(1);
[~,T] = histc(R,Pcum);
end

function X = simMC(M,A,P)
X = zeros(1,M);
X(1) = A;
for m=1:M-1
X(m+1) = discrete(P(X(m),:));
end
end
```

```
{\tt 1 \ function \ frac\_dist = get\_frac\_dist (state\_tans \ , number\_of\_states)}
_{2} frac_{dist} = zeros(1, number_{of\_states});
3 N = size(state\_tans);
4 N = N(2);
6 for i=1:number_of_states
  frac_dist(i) = sum(state_tans=i);
9 frac_dist = frac_dist/N;
10 end
1 N = 1000;
_{2} Number_of_states = 3;
start_state = 1;
^{4} P = [0.3 \ 0.7 \ 0; 0.1 \ 0.4 \ 0.5; 1 \ 0 \ 0];
5 Y = zeros(N, Number_of_states);
6 simMC(10, start_state, P)
7 \text{ for } i=1:N
       Y(i,:)=get_frac_dist(simMC(i,start_state,P),Number_of_states
9 end
10
11 x = (1:N);
scatter(x,Y(:,1),'d');
13 hold on;
14 scatter (x,Y(:,2),'o');
scatter (x, Y(:,3), 'x');
hold off;
17 legend('state0', 'state1', 'state2');
18 xlabel('n(number of steps)');
19 ylabel ('fraction of time spent')
```

2. f

```
function pin = PI(pi0,P,n)
 pin = pi0*(P^n);
 з end
function mat = PIN(pi0,P,n)
 _{2} len_{pi} = size(pi0);
 _{4} \text{ mat} = \operatorname{zeros}(n, \operatorname{len-pi}(2));
5 for i=1:n
 6
              mat(i,:) = PI(pi0,P,i);
 7
         end
 8 end
 1 N = 100;
\begin{array}{l} P = \begin{bmatrix} 0.3 & 0.7 & 0; 0.1 & 0.4 & 0.5; 1 & 0 & 0 \end{bmatrix}; \\ pi0 = \begin{bmatrix} 1 & 0 & 0 \end{bmatrix}; \end{array}

\begin{array}{ll}
\text{pin} &= \text{PIN}(\text{pio}, P, N); \\
\text{s} &= [1:N];
\end{array}

6 scatter (x, pin (:,1), 'd');
 7 hold on;
8 scatter(x, pin(:,2), 'o');
9 scatter(x, pin(:,3), 'x');
10 hold off;
legend('State0', 'State1', 'State2');
xlabel('n');
13 ylabel('\pin')
```

Bibliography

- [1] https://pages.dataiku.com/hubfs/Dataiku
- [2] Jean Walrand. Probability in Electrical Engineering and Computer science. Jean Walrand, 2014.
- [3] Dimitri P. Bertsekas and John N. Tsitsiklis. *Introduction to Probability*. LECTURE NOTES Course 6.041-6.431 M.I.T. FALL 2000