

Assignment 9

1. Let X be a continuous random variable with density f , expectation $\mathbb{E}[X] = \mu$ and variance $\text{Var}[X] = \sigma^2$. Define a new random variable $Y := aX + b$ for some $a, b \in \mathbb{R}$.

a) Calculate the standard deviation $\text{SD}[Y]$.

b) Express the moment generating function m_Y in terms of m_X .

$$\text{a) } \text{Var}[Y] = \text{Var}[aX + b] = a^2 \text{Var}[X] = a^2 \sigma^2$$

$$\text{SD}[Y] = \sqrt{\text{Var}[Y]} = \boxed{|a|\sigma}$$

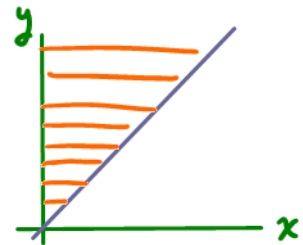
$$\text{b) } m_X(t) = E[e^{tX}]$$

$$m_Y(t) = E[e^{tY}] = E[e^{t(ax+b)}] = e^{tb} E[e^{t(ax)}] = \boxed{e^{tb} m_X(at)}$$

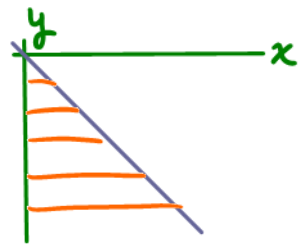
2. Prove that for an arbitrary continuous random variable X with density f we have

$$\mathbb{E}[X] = \int_0^\infty \mathbb{P}[X > x] dx - \int_0^\infty \mathbb{P}[X < -x] dx.$$

$$\begin{aligned} \int_0^\infty \mathbb{P}[X > x] dx &= \int_0^\infty \int_x^\infty \mathbb{P}[X=y] dy dx \\ &= \int_0^\infty \int_0^y \mathbb{P}[X=y] dx dy \\ &= \int_0^\infty y \mathbb{P}[X=y] dy \\ &= \int_0^\infty x \mathbb{P}[X=x] dx \end{aligned}$$



$$\begin{aligned} \int_0^\infty \mathbb{P}[X < -x] dx &= \int_0^\infty \int_{-\infty}^{-x} \mathbb{P}[X=y] dy dx \\ &= \int_{-\infty}^0 \int_0^{-y} \mathbb{P}[X=y] dx dy \\ &= \int_{-\infty}^0 -y \mathbb{P}[X=y] dy \\ &= - \int_{-\infty}^0 y \mathbb{P}[X=y] dy \\ &= - \int_{-\infty}^0 x \mathbb{P}[X=x] dx \end{aligned}$$



$$\begin{aligned} \int_0^\infty \mathbb{P}[X > x] dx - \int_0^\infty \mathbb{P}[X < -x] dx &= \int_0^\infty x \mathbb{P}[X=x] dx + \int_{-\infty}^0 x \mathbb{P}[X=x] dx \\ &= \int_{-\infty}^\infty x \mathbb{P}[X=x] dx \\ &= E[X] \end{aligned}$$

3. Assume that $U^{0,1}$ is a uniformly distributed random variable on the unit interval. Find a real-valued function $g : [0, 1] \rightarrow \mathbb{R}$ such that $Y := g(U^{0,1})$ is an exponentially distributed random variable with parameter $\lambda > 0$.

We want Y to have cdf $F_Y(y) = 1 - e^{-\lambda y}$. Solve $x = F_Y(y)$ for x to find $F_Y^{-1}(x) = y$.

$$1 - x = e^{-\lambda y}$$

$$y = F_Y^{-1}(x) = -\frac{1}{\lambda} \ln(1-x)$$

If we generate an x_0 from $X \sim U^{0,1}$ and compute $y_0 = -\frac{1}{\lambda} \ln(1-x_0)$, this y_0 has exponential distribution.

$$\therefore g(x) = -\frac{1}{\lambda} \ln(1-x)$$

4. The lifetime of an electrical device (in months) is given by the continuous random variable X with density

$$f(x) = \begin{cases} cxe^{-\frac{x}{2}} & \text{if } x > 0; \\ 0 & \text{if } x \leq 0. \end{cases}$$

- What is c ?
- What is the probability that the device functions more than 5 months?
- What is the expected lifetime of the device?

$$\begin{aligned} \text{a) } 1 &= \int_{-\infty}^{\infty} f(x) dx = c \int_0^{\infty} xe^{-x/2} dx = c \left[-2xe^{-x/2} \Big|_0^{\infty} + 2 \int_0^{\infty} e^{-x/2} dx \right] \\ &= c \left[0 - (4e^{-x/2}) \Big|_0^{\infty} \right] = -4c(0 - 1) = 4c \end{aligned}$$

$$\boxed{c = 1/4}$$

$$\begin{aligned} \text{b) } P[X > 5] &= \int_5^{\infty} f(x) dx = \frac{1}{4} \left[-2xe^{-x/2} - 4e^{-x/2} \right] \Big|_5^{\infty} \\ &= \frac{1}{4} \left[(0 - 0) - (-10e^{-5/2} - 4e^{-5/2}) \right] \\ &= \boxed{\frac{7}{2}e^{-5/2} \approx 0.287} \end{aligned}$$

$$\begin{aligned} \text{c) } E[X] &= \int_{-\infty}^{\infty} xf(x) dx = \frac{1}{4} \int_0^{\infty} x^2 e^{-x/2} dx \\ &= \frac{1}{4} \left[-2x^2 e^{-x/2} \Big|_0^{\infty} + 4 \int_0^{\infty} x e^{-x/2} dx \right] \\ &= \frac{1}{4} \left[0 + 4 \cdot 4 \right] \quad \left(\int_0^{\infty} x e^{-x/2} dx = 4 \text{ by part a) } \right) \\ &= \boxed{4} \end{aligned}$$

5. Assume that X is an exponentially distributed random variable with parameter $\lambda > 1$. Calculate

- a) $\mathbb{E}[X^3]$;
b) $\mathbb{E}[e^X]$.

Why did we impose the condition $\lambda > 1$ (instead of the "usual" one, $\lambda > 0$)?

$$a) M_X(t) = \mathbb{E}[e^{tx}] = \int_{-\infty}^{\infty} e^{tx} f(x) dx = \int_0^{\infty} \lambda e^{tx} e^{-\lambda x} dx = \frac{\lambda}{\lambda - t} \quad (t < \lambda)$$

$$M_X'''(t) = \frac{3! \lambda}{(\lambda - t)^4}$$

$$\mathbb{E}[X^3] = M_X'''(0) = \frac{6}{\lambda^3} \quad (\text{Note that } \lambda > 1 > 0 = t)$$

$$\begin{aligned} b) \mathbb{E}[e^X] &= \int_0^{\infty} \lambda e^x e^{-\lambda x} dx \\ &= \lambda \int_0^{\infty} e^{(1-\lambda)x} dx \\ &= \frac{\lambda}{1-\lambda} e^{(1-\lambda)x} \Big|_0^{\infty} \\ &= \boxed{\frac{\lambda}{\lambda-1} \quad \text{for } \lambda > 1} \end{aligned}$$

For $\lambda \leq 1$, $\int_0^{\infty} \lambda e^{(1-\lambda)x} dx$ diverges to $+\infty$.

6. Find the cumulative distribution function F such that it has hazard rate $\lambda(t) = \frac{1}{\sqrt{t}}$ (for $t > 0$). Can you express F in terms of an exponentially distributed random variable?

$$\text{Given the hazard rate } \lambda(t) = \frac{f(t)}{\bar{F}(t)} = -\frac{\bar{F}'(t)}{\bar{F}(t)} = -(\log \bar{F}(t))'$$

$$\text{we have } \log \bar{F}(t) = -\int_0^t \lambda(s) ds + c, \quad c \in \mathbb{R}$$

$$\log \bar{F}(t) = -\int_0^t \frac{1}{\sqrt{s}} ds + c = -2\sqrt{t} + c$$

$$\bar{F}(t) = d e^{-2\sqrt{t}}, \quad d \in \mathbb{R} \quad 1 = 1 - 0 = 1 - F(0) = \bar{F}(0) = d$$

$$F(t) = 1 - \bar{F}(t) = 1 - e^{-2\sqrt{t}}$$

$$f(t) = 2e^{-2\sqrt{t}}$$

Let $G(x)$ be the cdf of an exponential r.v. with $\lambda = 2$.

Since $G(x) = 1 - e^{-2x}$ for $x > 0$, $F(t) = G(\sqrt{t})$ for $t > 0$. $F(t)$ must be defined for all $t \in \mathbb{R}$ since F is a cdf. Set $F(t) = 0, t \leq 0$.

$$\therefore F(t) = \begin{cases} G(\sqrt{t}), & t > 0 \\ 0, & t \leq 0 \end{cases}$$