
SUPPLEMENTARY MATERIAL

Joint estimation of relative risk of dengue and Zika infections

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Let us assume the observed counts O_{ij} of dengue or Zika virus disease (ZVD) are Poisson distributed with mean parameter (μ_{ij}) where i is the aggregation area ($i = 1, \dots, n$, and $n = 87$ municipalities at departmental level; or $n = 293$ census section for the municipal level), and j is the disease ($j = 1, \dots, p$, and $p = 1$ for ZVD, or $p = 2$ for dengue).

$$O_{ij} \sim \text{Poisson}(\mu_{ij})$$

$$\mu_{ij} = E_{ij} \times r_{ij}$$

$$r_{ij} = \exp(\lambda_{ij})$$

Then, the mean parameter μ_{ij} is equal to the product of the expected values E_{ij} and the relative risk r_{ij} , with linear predictor λ_{ij} . The relative risk is the additive effect of spatially structured and unstructured random effects, and covariates. Spatially structured random effects are unobserved variables recovering a clustered risk pattern, or the fact that the risk in one area is highly associated with the neighboring areas. The lack of spatial association is accounted by the spatially unstructured random effects [1].

Model 1

Model 1 contains independent and identically distributed (IID) Normal spatially unstructured random effects for every disease (dengue and ZVD).

$$O_{ij} \sim \text{Poisson}(\mu_{ij})$$

$$\log(\mu_{ij}) = \log(E_{ij}) + \alpha_j + \phi_{ij}$$

$$\boldsymbol{\phi}_j \sim \text{Normal}(\mathbf{0}, \sigma_{\phi_j}^2 \mathbf{I})$$

$$\alpha_j \sim \text{Normal}(0, 1000)$$

$$1/\sigma_{\phi_j}^2 \sim \text{Gamma}(0.01, 0.01)$$

where ϕ_j are spatially unstructured random effects, α_j are intercepts, $\sigma_{\phi_j}^2$ are variance parameters of the ϕ_j , and \mathbf{I} is $n \times n$ identity matrix.

Model 2

Model 2 presents IID Normal spatially unstructured random effects linearly correlated for both diseases.

$$\begin{aligned}O_{ij} &\sim \text{Poisson}(\mu_{ij}) \\ \log(\mu_{ij}) &= \log(E_{ij}) + \alpha_j + \phi_{ij} \\ \begin{bmatrix} \phi_1 \\ \phi_2 \end{bmatrix} &\sim \text{Normal}(\mathbf{0}, \mathbf{\Sigma} \otimes \mathbf{I}) \\ \mathbf{\Sigma} &= \begin{bmatrix} \sigma_1^2 & \rho\sigma_1\sigma_2 \\ \rho\sigma_1\sigma_2 & \sigma_2^2 \end{bmatrix} \\ \mathbf{\Sigma}^{-1} &\sim \text{Wishart}(\mathbf{R}_{2 \times 2}, 2) \\ \mathbf{R} &= \begin{bmatrix} 1/5 & 0 \\ 0 & 1/5 \end{bmatrix} \\ \alpha_j &\sim \text{Normal}(0, 1000)\end{aligned}$$

$\mathbf{\Sigma}$ is an $n \times n$ variance-covariance matrix accounting for the association of the spatially unstructured random effects ϕ_j , \otimes corresponds to the Kronecker product, and the rest of the parameters similar to model 1.

Model 3

Model 3 accommodates conditionally autoregressive (CAR) [4] Normal spatially structured random effects for every disease.

$$\begin{aligned}O_{ij} &\sim \text{Poisson}(\mu_{ij}) \\ \log(\mu_{ij}) &= \log(E_{ij}) + \alpha_j + \phi_{ij} \\ \boldsymbol{\phi}_j &\sim \text{Normal}(\mathbf{0}, \sigma_{\phi_j}^2 (\mathbf{D} - \mathbf{W})^{-1}) \\ \alpha_j &\sim \text{Normal}(0, 1000) \\ 1/\sigma_{\phi_j}^2 &\sim \text{Gamma}(0.01, 0.01) \\ \mathbf{W}_{n \times n} &= \begin{cases} w_{ij} = 1 & \text{if } i \sim j \\ w_{ij} = 0 & \text{if } i = j \\ w_{ij} = 0 & \text{otherwise} \end{cases} \\ w_{i+} &= \sum_j w_{ij} \\ \mathbf{D}_{n \times n} &= \text{diagonal}(w_{1+}, \dots, w_{n+})\end{aligned}$$

where $\boldsymbol{\phi}_j$ are spatially structured random effects, with structure given by the proximity matrix \mathbf{W} , and variance parameters $\sigma_{\phi_j}^2$

Model 4

Model 4 contains CAR Normal spatially structured random effects linearly correlated for both diseases.

$$\begin{aligned}O_{ij} &\sim \text{Poisson}(\mu_{ij}) \\ \log(\mu_{ij}) &= \log(E_{ij}) + \alpha_j + \phi_{ij} \\ \begin{bmatrix} \phi_1 \\ \phi_2 \end{bmatrix} &\sim \text{Normal}(\mathbf{0}, \mathbf{\Sigma} \otimes (\mathbf{D} - \mathbf{W})^{-1}) \\ \mathbf{\Sigma} &= \begin{bmatrix} \sigma_1^2 & \rho\sigma_1\sigma_2 \\ \rho\sigma_1\sigma_2 & \sigma_2^2 \end{bmatrix} \\ \mathbf{\Sigma}^{-1} &\sim \text{Wishart}(\mathbf{R}_{2 \times 2}, 2) \\ \mathbf{R} &= \begin{bmatrix} 1/5 & 0 \\ 0 & 1/5 \end{bmatrix} \\ \alpha_j &\sim \text{Normal}(0, 1000) \\ \mathbf{W}_{n \times n} &= \begin{cases} w_{ij} = 1 & \text{if } i \sim j \\ w_{ij} = 0 & \text{if } i = j \\ w_{ij} = 0 & \text{otherwise} \end{cases} \\ w_{i+} &= \sum_j w_{ij} \\ \mathbf{D}_{n \times n} &= \text{diagonal}(w_{1+}, \dots, w_{n+})\end{aligned}$$

where ϕ_j are spatially structured random effects, $\mathbf{\Sigma}$ is the variance-covariance matrix accounting the spatial association of the ϕ_j , and proximity matrix \mathbf{W} .

Model 5

Model 5 includes IID Normal spatially unstructured random effects for every disease with a CAR shared-parameter.

$$\begin{aligned}O_{ij} &\sim \text{Poisson}(\mu_{ij}) \\ \log(\mu_{i,1}) &= \log(E_{i,1}) + \alpha_1 + \psi_i \times \gamma + \phi_{i,1} \\ \log(\mu_{i,2}) &= \log(E_{i,2}) + \alpha_2 + \psi_i/\gamma + \phi_{i,2} \\ \psi &\sim \text{Normal}(\mathbf{0}, \sigma_\psi^2 (\mathbf{D} - \mathbf{W})^{-1}) \\ \phi_j &\sim \text{Normal}(\mathbf{0}, \sigma_{\phi_j}^2 \mathbf{I}) \\ \alpha_j &\sim \text{Normal}(0, 1000) \\ \sigma_\psi &\sim \text{Uniform}(0, 10) \\ \sigma_{\phi_j} &\sim \text{Uniform}(0, 10) \\ \gamma &\sim \text{Normal}(0, 100) \\ \mathbf{W}_{n \times n} &= \begin{cases} w_{ij} = 1 & \text{if } i \sim j \\ w_{ij} = 0 & \text{if } i = j \\ w_{ij} = 0 & \text{otherwise} \end{cases} \\ w_{i+} &= \sum_j w_{ij} \\ \mathbf{D}_{n \times n} &= \text{diagonal}(w_{1+}, \dots, w_{n+})\end{aligned}$$

ψ are the spatially structured shared-parameter components, γ is a scaling parameter, ϕ_j are spatially unstructured random effects, \mathbf{W} is the proximity matrix, and $\sigma_{\phi_j}^2$, σ_ψ^2 are variance parameters for ϕ_j and ψ .

Model 6

Model 6 includes CAR Normal spatially structured random effects for every disease with a CAR shared-parameter.

$$\begin{aligned}
O_{ij} &\sim \text{Poisson}(\mu_{ij}) \\
\log(\mu_{i,1}) &= \log(E_{i,1}) + \alpha_1 + \psi_i \times \gamma + \phi_{i,1} \\
\log(\mu_{i,2}) &= \log(E_{i,2}) + \alpha_2 + \psi_i / \gamma + \phi_{i,2} \\
\psi &\sim \text{Normal}(\mathbf{0}, \sigma_\psi^2 (\mathbf{D} - \mathbf{W})^{-1}) \\
\phi_j &\sim \text{Normal}(\mathbf{0}, \sigma_{\phi_j}^2 (\mathbf{D} - \mathbf{W})^{-1}) \\
\alpha_j &\sim \text{Normal}(0, 1000) \\
\sigma_\psi &\sim \text{Uniform}(0, 10) \\
\sigma_{\phi_j} &\sim \text{Uniform}(0, 10) \\
\gamma &\sim \text{Normal}(0, 100) \\
\mathbf{W}_{n \times n} &= \begin{cases} w_{ij} = 1 & \text{if } i \sim j \\ w_{ij} = 0 & \text{if } i = j \\ w_{ij} = 0 & \text{otherwise} \end{cases} \\
w_{i+} &= \sum_j w_{ij} \\
\mathbf{D}_{n \times n} &= \text{diagonal}(w_{1+}, \dots, w_{n+})
\end{aligned}$$

ψ are the spatially structured shared-parameter components, γ is a scaling parameter, ϕ_j are spatially structured random effects, \mathbf{W} is the proximity matrix, and $\sigma_{\phi_j}^2$, σ_ψ^2 are variance parameters for ϕ_j and ψ .

Models 7 and 8

Models 7 accommodates the Generalized Multivariate CAR model [3], where the CAR Normal spatially structured random effects of ZVD by small area are conditioned by the CAR Normal spatially structured random effects of dengue. Model 8 presents the Generalized Multivariate CAR model [3], where the CAR Normal spatially structured random effects of dengue per area are conditioned by the CAR Normal spatially structured random effects of ZVD.

$$\begin{aligned}
O_{ij} &\sim \text{Poisson}(\mu_{ij}) \\
\log(\mu_{ij}) &= \log(E_{ij}) + \phi_{ij} \\
\phi_1 | \phi_2 &\sim \text{Normal}(\boldsymbol{\delta}_1, \boldsymbol{\Xi}_1) \\
\boldsymbol{\delta}_1 &= \beta_1 \mathbf{1} + (\eta_0 \mathbf{I} + \eta_1 \mathbf{W})(\phi_2 - \beta_2 \mathbf{1}) \\
\boldsymbol{\Xi}_1 &= \sigma_1^2 (\mathbf{D} - \kappa_1 \mathbf{W})^{-1} \\
\phi_2 &\sim \text{Normal}(\boldsymbol{\delta}_2, \boldsymbol{\Xi}_2) \\
\boldsymbol{\delta}_2 &= \beta_2 \mathbf{1} \\
\boldsymbol{\Xi}_2 &= \sigma_1^2 (\mathbf{D} - \kappa_2 \mathbf{W})^{-1} \\
\beta_j &\sim \text{Normal}(0, 100) \\
\eta_0, \eta_1 &\sim \text{Normal}(0, 10) \\
\sigma_j &\sim \text{Uniform}(0, 10) \\
\kappa_j &\sim \text{Uniform}(0, 0.99) \\
\mathbf{W}_{n \times n} &= \begin{cases} w_{ij} = 1 & \text{if } i \sim j \\ w_{ij} = 0 & \text{if } i = j \\ w_{ij} = 0 & \text{otherwise} \end{cases} \\
w_{i+} &= \sum_j w_{ij} \\
\mathbf{D}_{n \times n} &= \text{diagonal}(w_{1+}, \dots, w_{n+})
\end{aligned}$$

where ϕ_j are spatially structured random effects, \mathbf{W} is the proximity matrix, $\phi_1 | \phi_2$ is the conditional distribution of ϕ_1 , and ϕ_2 is a marginal distribution.

REFERENCES

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