

## Bayesian Spike and Slab with Random Intercept

This work builds upon Naqvi et al. Fast Laplace Approximation for Sparse Bayesian Spike and Slab Models, by adding in a random effect to produce more robust posterior inclusion probabilities.

When the main goal is variable selection, we treat the variance components (fixed effects ( $\sigma^2$ ), random effects ( $\delta^2$ )) as hyperparameters and estimate them by empirical bayes.

The model implemented can be described as follows:

$$Y|\beta, \mu, \sigma^2 \sim N(X\beta + Z\mu, \sigma^2)$$

$$\beta|\tau_1^2, \tau_0^2 \sim \pi N(0, \tau_1^2) + (1 - \pi)N(0, \tau_0^2)$$

$$\mu|\delta^2 \sim N(0, \delta^2)$$

$$P(\beta|Y, \mu, \sigma^2, \tau_1^2, \tau_0^2, \delta^2) = \frac{P(Y|\beta, \mu, \sigma^2)P(\beta|\tau_1^2, \tau_0^2)P(\mu|\delta^2)}{P(Y|\mu, \sigma^2, \tau_1^2, \tau_0^2)P(\mu|\delta^2)}$$

Where  $Y \in \mathbb{R}^{N \times 1}$  is a response vector;  $X \in \mathbb{R}^{N \times p}$  is a matrix of predictors;  $Z \in \mathbb{R}^{N \times m}$  is a one-hot-encoded group ID matrix;  $\mu$  the random effect vector;  $\beta$  the (sparse) regression weights;  $\sigma^2$  the fixed effects variance;  $\tau_0^2, \tau_1^2$  the spike and slab variances, respectively;  $\pi$  the spike and slab mixture proportion;  $\delta^2$  the random effects variance component.

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### Step 1 – Marginalize the random effects

First, we integrate out the random effects,  $\mu$

$$\begin{aligned} P(Y, \beta) &= \int P(Y|\beta, \mu, \sigma^2)P(\beta|\tau_1^2, \tau_0^2)P(\mu|\delta^2)d\mu \\ &= P(\beta|\tau_1^2, \tau_0^2) \int P(Y|\beta, \mu, \sigma^2)P(\mu|\delta^2)d\mu \end{aligned}$$

This is a well-known integral (Gaussian convolution):

$$= P(\beta|\tau_1^2, \tau_0^2)N(Y|X\beta, \Sigma); \quad \Sigma = \delta^2 Z Z^T + \sigma^2 I$$

So the joint probability of  $Y$  and  $\beta$  is:

$$\begin{aligned}\log P(Y, \beta) = & -\frac{N}{2} \log 2\pi - \frac{1}{2} \log |\Sigma| - \frac{1}{2} (Y - X\beta)^T \Sigma^{-1} (Y - X\beta) \\ & + \sum_{j=1}^p \log \left( \pi N(\beta_j | 0, \tau_1^2) + (1 - \pi) N(\beta_j | 0, \tau_0^2) \right)\end{aligned}$$


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## Step 2 – Maximum *a posteriori* estimation

We set  $\sigma^2, \delta^2$  to a constant. Maximize the posterior with respect to  $\beta$ . To do this, we need the gradients of the posterior with respect to  $\beta$ :

Gradient of the likelihood:

$$\nabla_{\beta} \log P(Y | \beta, \sigma^2) = X^T \Sigma^{-1} Y - X^T \Sigma^{-1} X \beta$$

Gradient of the prior:

$$\nabla_{\beta_j} \log P(\beta_j | \tau_1^2, \tau_0^2) = -\beta_j \frac{\left( \frac{\pi}{\tau_1^2} N(\beta_j | 0, \tau_1^2) + \frac{(1 - \pi)}{\tau_0^2} N(\beta_j | 0, \tau_0^2) \right)}{\pi N(\beta_j | 0, \tau_1^2) + (1 - \pi) N(\beta_j | 0, \tau_0^2)}$$


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## Step 3 – Hessian calculation

Compute the Hessian at  $\hat{\beta}$

Hessian of the likelihood:

$$\nabla_{\beta}^2 \log P(Y | \beta, \sigma^2) = -X^T \Sigma^{-1} X$$

Hessian of the prior:

$$\nabla_{\beta}^2 \log P(\beta | \tau_1^2, \tau_0^2) = \text{diag} \left( \frac{\nabla_{\beta}^2 P(\beta | \tau_1^2, \tau_0^2)}{P(\beta | \tau_1^2, \tau_0^2)} - \left[ \frac{\nabla_{\beta} P(\beta)}{P(\beta)} \right]^2 \right)$$

Where we have:

$$\nabla_{\beta_j} P(\beta) = -\beta_j \left( \frac{\pi}{\tau_1^2} N(\beta_j | 0, \tau_1^2) + \frac{(1-\pi)}{\tau_0^2} N(\beta_j | 0, \tau_0^2) \right)$$

$$\nabla_{\beta_j}^2 P(\beta | \tau_1^2, \tau_0^2) = \pi \left( \frac{\beta_j^2}{\tau_1^4} - \frac{1}{\tau_1^2} \right) N_1 + (1-\pi) \left( \frac{\beta_j^2}{\tau_0^4} - \frac{1}{\tau_0^2} \right) N_0$$

$$H(\beta) = \nabla_{\beta}^2 \log P(Y | \beta, \sigma^2) + \nabla_{\beta}^2 \log P(\beta | \tau_1^2, \tau_0^2)$$

#### Step 4 – Estimation of variance components (Type II likelihood)

Use the Laplace approximation to integrate  $\beta$  out. Maximize the marginal posterior with respect to  $\sigma^2, \delta^2$ .

The Laplace approximation is found by Taylor expansion of the log exponent of the posterior, leading to:

$$P(\beta, Y | \sigma^2, \tau_1^2, \tau_0^2, \delta^2) \approx P(\hat{\beta}, Y | \sigma^2, \tau_1^2, \tau_0^2, \delta^2) \cdot N(\beta | \hat{\beta}, -H^{-1}(\hat{\beta}))$$

$$\therefore P(Y | \sigma^2, \tau_1^2, \tau_0^2, \delta^2) = P(\hat{\beta}, Y | \sigma^2, \tau_1^2, \tau_0^2, \delta^2) (2\pi)^{\frac{p}{2}} \det(-H(\hat{\beta}))^{-\frac{1}{2}}$$

Where  $\hat{\beta}$  denotes the MAP estimate.

We can use this to integrate out  $\beta$ , leading to the log marginal posterior:

$$\log P(Y | \sigma^2, \tau_1^2, \tau_0^2, \delta^2) = -\frac{1}{2} \log |\Sigma| - \frac{1}{2} (Y - X\beta)^T \Sigma^{-1} (Y - X\beta) - \frac{1}{2} \log |-H(\hat{\beta})| + \text{const}$$

Where const denotes terms not dependent on  $\sigma^2, \delta^2$

To take the derivative, we use the following matrix identities:

$$\frac{\partial}{\partial \theta} \log |\Sigma| = \text{tr}(\Sigma^{-1} \partial_{\theta} \Sigma)$$

$$\frac{\partial}{\partial \theta} (Y - X\beta)^T \Sigma^{-1} (Y - X\beta) = -(Y - X\beta)^T \Sigma^{-1} (\partial_{\theta} \Sigma) \Sigma^{-1} (Y - X\beta)$$

$$\frac{\partial}{\partial \theta} \log |-H(\beta)| = \text{tr}(-H^{-1}(\beta) (-\partial_{\theta} H(\beta)))$$

$$\partial_{\theta_i} \Sigma = \begin{matrix} e^{\theta_1} I; & i = 1 \\ e^{\theta_2} Z Z^T; & i = 2 \end{matrix}$$

$$\partial_{\theta}(-H(\beta)) = -X^T \Sigma^{-1} (\partial_{\theta} \Sigma) \Sigma^{-1} X$$

So we have:

$$\begin{aligned} \nabla_{\theta} \log P(Y|\sigma^2, \tau_1^2, \tau_0^2, \delta^2) = & -\frac{1}{2} \text{tr}(\Sigma^{-1} \partial_{\theta} \Sigma) + \frac{1}{2} (Y - X\beta)^T \Sigma^{-1} (\partial_{\theta} \Sigma) \Sigma^{-1} (Y - X\beta) \\ & -\frac{1}{2} \text{tr}(-H^{-1}(\beta)(-\partial_{\theta} H(\beta))) \end{aligned}$$

And we maximize with respect to  $\theta = (\sigma^2, \delta^2)$ , parametrized as  $(e^{\theta_1}, e^{\theta_2})$  so we can do unconstrained optimization.

### Step 5 – Model iteration

Repeat step 2-4 until convergence. We use the following convergence criterion:

$$\log P(\beta^{(t)}|X, Y, \theta^{(t)}) - \log P(\beta^{(t-1)}|X, Y, \theta^{(t-1)}) \leq 0.0001$$

### Step 6 – Estimation of posterior inclusion probabilities (PIPs)

Once we have estimated the MAP for  $\beta, \sigma^2, \delta^2$ , we can find the PIP.

$$PIP_{\beta_j} = \int \frac{\pi N(\beta|0, \tau_1^2)}{\pi N(0, \tau_1^2) + (1 - \pi) N(0, \tau_0^2)} N(\beta|\hat{\beta}, -H^{-1}(\hat{\beta}))_{jj} d\beta_j$$

While this integral has no closed form, it is a simple 1 dimensional integral that can be efficiently computed using Monte Carlo integration or Gaussian quadrature, as in Naqvi et al.

### Appendix – derivations

Gradient of the likelihood:

$$\begin{aligned}
& \nabla_{\beta} \left( -\frac{N}{2} \log 2\pi - \frac{1}{2} \log |\Sigma| - \frac{1}{2} (Y - X\beta)^T \Sigma^{-1} (Y - X\beta) \right) \\
&= \nabla_{\beta} \left( -\frac{1}{2} (Y - X\beta)^T \Sigma^{-1} (Y - X\beta) \right) \\
&= \nabla_{\beta} \left( -\frac{1}{2} (Y^T - \beta^T X^T) \Sigma^{-1} (Y - X\beta) \right) \\
&= \nabla_{\beta} \left( -\frac{1}{2} (Y^T \Sigma^{-1} - \beta^T X^T \Sigma^{-1}) (Y - X\beta) \right) \\
&= \nabla_{\beta} \left( -\frac{1}{2} (Y^T \Sigma^{-1} Y - Y^T \Sigma^{-1} X\beta - \beta^T X^T \Sigma^{-1} Y + \beta^T X^T \Sigma^{-1} X\beta) \right) \\
&= \nabla_{\beta} \left( \beta^T X^T \Sigma^{-1} Y - \frac{1}{2} \beta^T X^T \Sigma^{-1} X\beta \right) \\
&= X^T \Sigma^{-1} Y - X^T \Sigma^{-1} X \beta
\end{aligned}$$

Gradient of the prior

$$\begin{aligned}
& \nabla_{\beta_j} \sum_{j=1}^p \log[\pi N(\beta_j|0, \tau_1^2) + (1 - \pi)N(\beta_j|0, \tau_0^2)] \\
& \nabla_{\beta_j} \log[\pi N(\beta_j|0, \tau_1^2) + (1 - \pi)N(\beta_j|0, \tau_0^2)] \\
&= \frac{-\beta_j \left( \frac{\pi}{\tau_1^2} N_1 + \frac{(1 - \pi)}{\tau_0^2} N_0 \right)}{\pi N_1 + (1 - \pi) N_0}; \quad N_k = N(\beta_j|0, \tau_k^2)
\end{aligned}$$

Hessian of the likelihood

$$\begin{aligned}
\nabla_{\beta}^2 \log P(Y, \beta | \sigma^2, \tau_1^2, \tau_0^2, \delta^2) &= \nabla_{\beta} (X^T \Sigma^{-1} Y - X^T \Sigma^{-1} X \beta) \\
&= -X^T \Sigma^{-1} X
\end{aligned}$$

Hessian of the prior

$$\nabla_{\beta}^2 \log P(\beta | \tau_1^2, \tau_0^2) = \nabla_{\beta_j} \frac{-\beta_j \left( \frac{\pi}{\tau_1^2} N_1 + \frac{(1 - \pi)}{\tau_0^2} N_0 \right)}{\pi N_1 + (1 - \pi) N_0} = \nabla_{\beta_j} \left( \frac{\nabla_{\beta_j} P(\beta)}{P(\beta)} \right)$$

$$\left[ \frac{f(x)}{g(x)} \right]' = \frac{g(x)f'(x) - f(x)g'(x)}{g^2(x)} = \frac{P(\beta)\nabla_{\beta}^2 P(\beta)}{[P(\beta)]^2} - \frac{\nabla_{\beta} P(\beta)\nabla_{\beta} P(\beta)}{[P(\beta)]^2}$$

$$= \frac{\nabla_{\beta_j}^2 P(\beta|\tau_1^2, \tau_0^2)}{P(\beta|\tau_1^2, \tau_0^2)} - \left[ \frac{\nabla_{\beta_j} P(\beta|\tau_1^2, \tau_0^2)}{P(\beta|\tau_1^2, \tau_0^2)} \right]^2$$

$$\begin{aligned} \nabla_{\beta_j}^2 P(\beta) &= \nabla_{\beta_j} \left[ \frac{-\beta_j \pi}{\tau_1^2} N_1 - \frac{\beta_j (1 - \pi)}{\tau_0^2} N_0 \right] \\ &= \frac{\partial}{\partial \beta_j} \left( \frac{-\beta_j \pi}{\tau_1^2} \right) N_1 + \left( \frac{-\beta_j \pi}{\tau_1^2} \right) \frac{\partial}{\partial \beta_j} N_1 + \frac{\partial}{\partial \beta_j} \left( \frac{-\beta_j (1 - \pi)}{\tau_0^2} \right) N_0 + \left( \frac{-\beta_j (1 - \pi)}{\tau_0^2} \right) \frac{\partial}{\partial \beta_j} N_0 \end{aligned}$$